

Relationships between Probability Distributions

Note: The gamma distribution mentioned in this notes uses parameters shape and rate: i.e. $\text{Gamma}(\alpha, \beta)$ where α is shape, β is rate. There is a different parametrisation using shape and scale

DISCRETE DISTRIBUTIONS							
	range \mathbb{X}	parameters	pmf f_X	cdf F_X	$E[X]$	$\text{Var}[X]$	mgf M_X
$\text{Bernoulli}(\theta)$	$\{0, 1\}$	$\theta \in (0, 1)$	$\theta^x (1 - \theta)^{1-x}$		θ	$\theta(1 - \theta)$	$1 - \theta + \theta e^t$
$\text{Binomial}(n, \theta)$	$\{0, 1, \dots, n\}$	$n \in \mathbb{Z}^+, \theta \in (0, 1)$	$\binom{n}{x} \theta^x (1 - \theta)^{n-x}$		$n\theta$	$n\theta(1 - \theta)$	$(1 - \theta + \theta e^t)^n$
$\text{Poisson}(\lambda)$	$\{0, 1, 2, \dots\}$	$\lambda \in \mathbb{R}^+$	$\frac{e^{-\lambda} \lambda^x}{x!}$		λ	λ	$\exp\{\lambda(e^t - 1)\}$
$\text{Geometric}(\theta)$	$\{1, 2, \dots\}$	$\theta \in (0, 1)$	$(1 - \theta)^{x-1} \theta$	$1 - (1 - \theta)^x$	$\frac{1}{\theta}$	$\frac{(1 - \theta)}{\theta^2}$	$\frac{\theta e^t}{1 - e^t(1 - \theta)}$
$\text{NegBinomial}(n, \theta)$	$\{n, n + 1, \dots\}$	$n \in \mathbb{Z}^+, \theta \in (0, 1)$	$\binom{x-1}{n-1} \theta^n (1 - \theta)^{x-n}$		$\frac{n}{\theta}$	$\frac{n(1 - \theta)}{\theta^2}$	$\left(\frac{\theta e^t}{1 - e^t(1 - \theta)}\right)^n$
or	$\{0, 1, 2, \dots\}$	$n \in \mathbb{Z}^+, \theta \in (0, 1)$	$\binom{n+x-1}{x} \theta^n (1 - \theta)^x$		$\frac{n(1 - \theta)}{\theta}$	$\frac{n(1 - \theta)}{\theta^2}$	$\left(\frac{\theta}{1 - e^t(1 - \theta)}\right)^n$

CONTINUOUS DISTRIBUTIONS							
		parameters	pdf	cdf	$E[X]$	$\text{Var}[X]$	mgf
$\text{Uniform}(\alpha, \beta)$ (stand. model $\alpha = 0, \beta = 1$)	(α, β)	$\alpha < \beta \in \mathbb{R}$	$\frac{1}{\beta - \alpha}$	$\frac{x - \alpha}{\beta - \alpha}$	$\frac{(\alpha + \beta)}{2}$	$\frac{(\beta - \alpha)^2}{12}$	$\frac{e^{\beta t} - e^{\alpha t}}{t(\beta - \alpha)}$
$\text{Exponential}(\lambda)$ (stand. model $\lambda = 1$)	\mathbb{R}^+	$\lambda \in \mathbb{R}^+$	$\lambda e^{-\lambda x}$	$1 - e^{-\lambda x}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\left(\frac{\lambda}{\lambda - t}\right)$
$\text{Gamma}(\alpha, \beta)$ (stand. model $\beta = 1$)	\mathbb{R}^+	$\alpha, \beta \in \mathbb{R}^+$	$\frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}$		$\frac{\alpha}{\beta}$	$\frac{\alpha}{\beta^2}$	$\left(\frac{\beta}{\beta - t}\right)^\alpha$
$\text{Weibull}(\alpha, \beta)$ (stand. model $\beta = 1$)	\mathbb{R}^+	$\alpha, \beta \in \mathbb{R}^+$	$\alpha \beta x^{\alpha-1} e^{-\beta x^\alpha}$	$1 - e^{-\beta x^\alpha}$	$\frac{\Gamma(1 + 1/\alpha)}{\beta^{1/\alpha}}$	$\frac{\Gamma(1 + \frac{2}{\alpha}) - \Gamma(1 + \frac{1}{\alpha})^2}{\beta^{2/\alpha}}$	
$\text{Normal}(\mu, \sigma^2)$ (stand. model $\mu = 0, \sigma = 1$)	\mathbb{R}	$\mu \in \mathbb{R}, \sigma \in \mathbb{R}^+$	$\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x - \mu)^2}{2\sigma^2}\right\}$		μ	σ^2	$e^{\{\mu t + \sigma^2 t^2/2\}}$
$\text{Student}(\nu)$	\mathbb{R}	$\nu \in \mathbb{R}^+$	$\frac{(\pi\nu)^{-\frac{1}{2}} \Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right) \left\{1 + \frac{x^2}{\nu}\right\}^{(\nu+1)/2}}$		0 (if $\nu > 1$)	$\frac{\nu}{\nu - 2}$ (if $\nu > 2$)	
$\text{Pareto}(\theta, \alpha)$	\mathbb{R}^+	$\theta, \alpha \in \mathbb{R}^+$	$\frac{\alpha \theta^\alpha}{(\theta + x)^{\alpha+1}}$	$1 - \left(\frac{\theta}{\theta + x}\right)^\alpha$	$\frac{\theta}{\alpha - 1}$ (if $\alpha > 1$)	$\frac{\alpha \theta^2}{(\alpha - 1)^2(\alpha - 2)}$ (if $\alpha > 2$)	
$\text{Beta}(\alpha, \beta)$	$(0, 1)$	$\alpha, \beta \in \mathbb{R}^+$	$\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1 - x)^{\beta-1}$		$\frac{\alpha}{\alpha + \beta}$	$\frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$	

Binomial distribution is sum of iid Bernoulli

$$X \sim \text{Binomial}(n, \theta) \implies X = \sum_{i=1}^n X_i \text{ where } X_i \sim \text{Bernoulli}(\theta)$$

Poisson distribution can be cut into n iid small fragments

$$X \sim \text{Poisson}(\lambda) \implies X = \sum_{i=1}^n X_i \text{ where } X_i \sim \text{Poisson}(\lambda/n), \text{ so that } \mu = E(X_i) = \lambda/n \text{ and } \sigma^2 = \text{Var}(X_i) = \lambda/n$$

Negative Binomial is sum of iid geometric distributions

$$X \sim \text{Negative Binomial}(n, \theta) \implies X = \sum_{i=1}^n X_i \text{ where } X_i \sim \text{Geometric}(\theta), \text{ so that } \mu = E(X_i) = 1/\theta \text{ and } \sigma^2 = \text{Var}(X_i) = (1 - \theta) / \theta^2, \text{ and hence}$$

Linearity of Normal distribution

if $X \sim N(\mu_1, \sigma_1^2)$, $Y \sim N(\mu_2, \sigma_2^2)$, then

$$X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$X - Y \sim N(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2)$$

if $X \sim N(\mu, \sigma^2)$, then $kX \sim N(k\mu, k^2\sigma^2)$

Student t-distribution

If $X_1, X_2, \dots, X_n \sim N(0, \sigma^2)$ where σ^2 is unknown, then

$$(\bar{x}_n - \mu) / (S/\sqrt{n}) \sim t_{n-1} \text{ where } \bar{x}_n \text{ is mean of } X_1, X_2, \dots, X_n$$

Chi-squared •

$$\chi_n^2 = \sum Z_i^2 \text{ where } Z_i \text{ are iid } N(0, 1)$$

Therefore, for integer k, sum of k iid χ_n^2 is χ_{kn}^2

Exp and gamma

$$\text{Exp}(\lambda) = \text{Gamma}(1, \lambda)$$

sum of n iid $\text{Gamma}(1, \lambda) \sim \text{Gamma}(n, \lambda)$, sum of k iid $\text{Gamma}(\lambda, 1) \sim \text{Gamma}(\lambda, k)$

so sum of n iid $\text{Exp}(\lambda)$ is $\text{Gamma}(n, \lambda)$

$$\chi_n^2 \sim \text{Gamma}(n/2, 1/2)$$

$$\text{so } \text{Exp}(1/2) = \chi_2^2$$

$$\text{Gamma}(a, \lambda) = \text{Gamma}(a, 1) / \lambda$$

so $2\lambda \text{Gamma}(n, \lambda) = \text{Gamma}(n, 1/2) = \chi^2_{2n}$

Gamma and F-distribution

If $U \sim \text{Gamma}(\alpha_1, \beta_1)$, $V \sim \text{Gamma}(\alpha_2, \beta_2)$, then

$$\alpha_2 \beta_1 U / \alpha_1 \beta_2 V \sim F_{2\alpha_1, 2\alpha_2}$$

Dirichlet Distribution :

A general result is that if for $i = 1, 2, \dots, k$, $U_i \sim \Gamma(\alpha_i, \beta)$ independent, then

$$\frac{1}{\sum_{i=1}^k U_i} (U_1, U_2, \dots, U_k) \sim \text{DIRICHLET}(\alpha_1, \alpha_2, \dots, \alpha_k),$$