

02450: Introduction to Machine Learning and Data Mining

Performance evaluation, Bayes, and Naive Bayes

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Department of Applied Mathematics and Computer Science



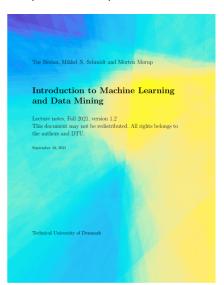


Feedback Groups of the day:

Christa Skytte Jensen, Mathias Fynbo Jensen, Magnus Nikolai Nyholm Jensen. Emma Kerstine Cens Jensen, Caroline Amalie Bastholm Jensen, Michael Rene Lund Jensen, Rikke Schjeldrup Jessen, Sindri Jonsson, Jonas Dalsberg Jørgensen, Martin Koch Jørgensen, Kasper Jørgensen, Mia Ann Jørgensen, Marie-Louise Birch Jørgensen, Teakosheen Joulak, Fredrik Junghus, Martin de Fries Justinussen, Maria Kalimantzali Liana. Kristofer Rhea Rebien Kandle. Baris Kara, Dimosthenis Karafylias, Spyridon Karamolegkos, Athanasios Karanatsios, Gustaw Karol Chojecki, Karthikraja Karunakaran, Sylvia Katharina Kessler, Per Kiil, Anna Kikidi, Christian Valentin Kiær, Povl Árni Klarlund, Anton Faber Klausen, Gabriel Franz Koenig, Casper Rosenstjerne Kølle, Tsz Hong Ivan Kong, Kristoffer Ioannis Tang Kordatos, Augusta Frost Korsgaard, Christos Koumparakis, Nicolai Blumensaat Kragh, Miranda Niemann Kristensen, Andreas Borg Kristensen, Karolina Zofia Krzesinska, Matias Kühnau, Lukas Kvzlik

Reading material:

Chapter 11, Chapter 13



Schedule



Introduction

1 February: C1

Data: Feature extraction, and visualization

2 Data, feature extraction and PCA 8 February: C2, C3

Measures of similarity, summary statistics and probabilities 15 February: C4. C5

Probability densities and data visualization

22 February: C6, C7

Supervised learning: Classification and regression

5 Decision trees and linear regression

6 Overfitting, cross-validation and Nearest Neighbor

8 March: C10, C12 (Project 1 due before 13:00)

Performance evaluation, Bayes, and Naive Bayes

15 March: C11, C13

3 Artificial Neural Networks and Bias/Variance
22 March: C14, C15

AUC and ensemble methods
 Auch: C16, C17

Unsupervised learning: Clustering and density estimation

K-means and hierarchical clustering 5 April: C18

Mixture models and density estimation 19 April: C19, C20 (Project 2 due before 13:00)

Association mining

Recap

Recap and discussion of the exam

3 May: C1-C21

Lecture stream: See DTU Learn Q&A during lectures: Zoom chat or Piazza Lecture recordings: https://video.dtu.dk Online assistance (exercises): Microsoft Teams Online assistance (24/7): Piazza forum



Evaluation, interpretation, and visualization Data modelling Evaluation Data preparation Classification Anomaly detection Feature extraction · Similarity measures Regression Decision making Result Daritai Summary statistics Clustering Result visualization Data visualization. Density estimation Dissemination Domain knowledge

Learning Objectives

- Understand the two different evaluation setups
- Apply appropriate statistical tests to evaluate and compare models
- Account for the assumptions made in Naïve Bayes
- Apply Bayes Theorem to obtain the class posterior likelihood



- A social media company wish to know if a new ad-placement method increases the click-through rate
- How many customers are likely click adds next month?
- How well can a neural network model learn to distinguish between diseased/non-diseased X-rays?
- Should I recommend my neural network model over a competing method?

All involve induction beyond the dataset



Tests can provide:

- An objective way to choose between methods
- A quantification of model performance which takes uncertainty into account



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Tests do not provide

- Certain conclusions (Model A is better than B)
- A black-box recipe



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- An objective way to choose between methods
- A quantification of model performance which takes uncertainty into account

Tests **do not** provide

- Certain conclusions (Model A is better than B)
- A black-box recipe

Use statistical tests to aid your interpretation of your results **not** as an argument in itself

Outline



- What is our overall **objective**? What conclusions do we want?
- What tools do we have available?
- What specific test should I use? (classification, regression, etc.)





- Models are compared based on how well they generalize to future data
- ullet Suppose we have data $\mathcal{D}=(oldsymbol{X},oldsymbol{y})$ and two models \mathcal{M}_A , \mathcal{M}_B
- \bullet Training on $\mathcal{D},$ we obtain prediction rules

$$f_{\mathcal{D},A}: \boldsymbol{x} o y, \quad \text{ and } \quad f_{\mathcal{D},B}: \boldsymbol{x} o y.$$





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• If $z_{\mathcal{D}} < 0$, it means that \mathcal{M}_A is better than \mathcal{M}_B ...





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Setup I Statistical tests of performance considering the **specific** training set \mathcal{D}



$$z_{\mathcal{D}} = E_{\mathcal{D},A}^{\text{gen}} - E_{\mathcal{D},B}^{\text{gen}}$$

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- Therefore, the conclusion is not independently reproducible



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- ullet To overcome this, test if \mathcal{M}_A is better than \mathcal{M}_B when averaging over dataset

$$\begin{split} z &= \mathbb{E}_{\mathcal{D}}[z_{\mathcal{D}}] < 0 \\ E^{\text{gen}} &= \int \left[\int L(f_{\mathcal{D}}(\boldsymbol{x}), y) p(\boldsymbol{x}, \boldsymbol{y}) d\boldsymbol{x} dy \right] p(\mathcal{D}) d\mathcal{D} \end{split}$$



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• If z < 0, it means \mathcal{M}_A is better than \mathcal{M}_B ... using a typical training set Setup II Statistical tests of performance considering a dataset of size N

Choices, choices



Setup I Statistical tests of performance considering the **specific** training set \mathcal{D} ?

Setup II Statistical tests of performance considering **a dataset** of size N

Which to choose fundamentally depends on what you want to conclude

Choices, choices



Setup I Statistical tests of performance considering the **specific** training set \mathcal{D} ?

Setup II Statistical tests of performance considering **a dataset** of size N

Which to choose fundamentally depends on what you want to conclude

- Setup II is a more general (impressive) conclusion
- Setup II is probably what we want in science
- Setup II requires (a lot of) cross-validation
- If you have a single train/test split, use setup I

We will consider **setup I** here



Let z be a quantity of interest (for instance $z=E_{\mathcal{A}}^{\mathsf{gen}}-E_{\mathcal{B}}^{\mathsf{gen}})$



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Estimation Determine a likely value $z\approx \hat{z}$ and an interval $[z_L,z_U]$ that likely contains z



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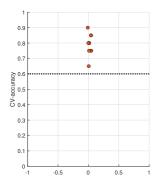
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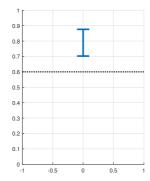
- Evidence against H_0 is measured by a p-value (low p is evidence for an effect $z \neq 0$)
- Estimation of $[z_L, z_U]$ done using an α -confidence interval (lower α means a more conservative, wider, interval)

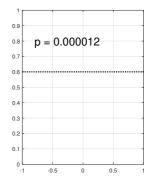
Choosing the right tool



- Consider binary classification using N=200 samples
- We estimate test error using K=10-fold CV (10 test-error estimates)
- Question: Is accuracy $E_A^{\text{gen}} pprox \frac{1}{K} \sum_{k=1}^K E_i^{\text{test}}$ greater than baseline θ_0 ?
- (Baseline classify everything as maximum class, accuracy 60%)



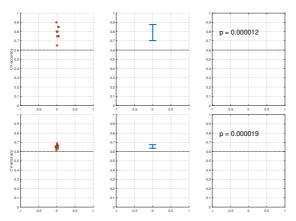




Which tool to use



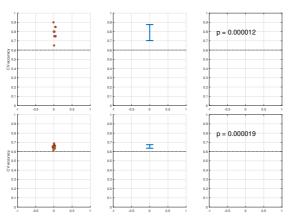
- Top: N = 200 sample example
- \bullet Bottom: Harder problem using $N=2000~\mathrm{samples}$



Which tool to use



- Top: N = 200 sample example
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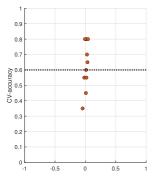


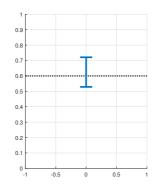
- p-value primarily measure of sample size (not effect size!)
- Which do you think are more informative?

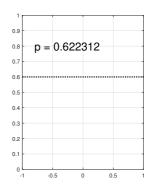
Variability



ullet New problem using N=200 samples. Is there an effect?



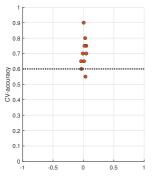


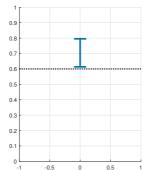


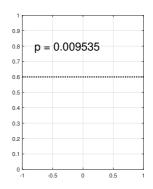
Variability



ullet Another problem using N=200 samples. Is there an effect?

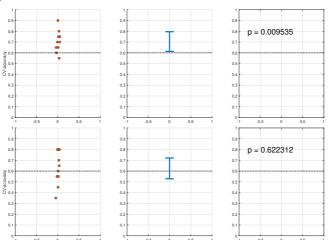






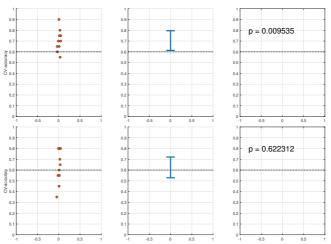
The nasty bit





The nasty bit





- Only difference is random variability in dataset
- Low *p*-value does **not necessarily** mean reproducible
 - Training many models will lead to false positives
 - Statistics will not fix unclear results; probably just lead to false positives

Connecting objective to numbers



• We want to draw conclusions about the difference in performance:

$$\begin{split} z_{\mathcal{D}} &= E_{\mathcal{D},A}^{\text{gen}} - E_{\mathcal{D},B}^{\text{gen}} \\ E_{\mathcal{D},A}^{\text{gen}} &= \int p(\boldsymbol{x},y) L(f_{\mathcal{D},A}(\boldsymbol{x}),y) d\boldsymbol{x} dy, \quad E_{\mathcal{D},B}^{\text{gen}} = \int p(\boldsymbol{x},y) L(f_{\mathcal{D},B}(\boldsymbol{x}),y) d\boldsymbol{x} dy. \end{split}$$

This can be estimated as

$$egin{aligned} \hat{z}_{\mathcal{D}} &= rac{1}{N^{ ext{test}}} \sum_{i=1}^{N^{ ext{test}}} \left[L(f_{\mathcal{D},A}(oldsymbol{x}_i), y_i) - L(f_{\mathcal{D},B}(oldsymbol{x}_i), y_i)
ight] \ &= rac{1}{N^{ ext{test}}} \sum_{i=1}^{N^{ ext{test}}} z_i, \quad ext{where:} \quad z_i = L(f_{\mathcal{D},A}(oldsymbol{x}_i), y_i) - L(f_{\mathcal{D},B}(oldsymbol{x}_i), y_i). \end{aligned}$$

Abstracting to a statistical question



Consider data as the n numbers

$$D=(z_1,\ldots,z_n). (1)$$

General form of the problem: Draw conclusions about

$$\theta = E_{A,\mathcal{D}}^{\mathsf{gen}} - E_{B,\mathcal{D}}^{\mathsf{gen}}$$

Based on the estimate:

$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^{n} z_i. \tag{2}$$

Statistical tools: Parameter



- ullet Assume z_i is a realization of a random variable Z_i
- It has density

$$p(Z_i = z_i | \theta) = p_{\theta}(z_i)$$

• Density of all dataset

$$p_{\theta}(D) = \prod_{i=1}^{n} p_{\theta}(z_i). \tag{3}$$

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- Returning to our goals:
 - ullet estimating plausible ranges of heta
 - \bullet hypothesis testing such as whether θ takes a particular value
- Let's look at the statistical tools to accomplish this

Statistical tools: Statistic and estimator



Statistic A statistic is a function of the data D and will be denoted t. For instance, the mean and variance are both statistics:

$$t_0(D) = \frac{1}{n} \sum_{i=1}^n Z_i$$
, or $t_1(D) = \frac{1}{n} \sum_{i=1}^n (Z_i - t_0(D))^2$.

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Estimator An estimator is a statistic t of D such that t(D) is close to θ . In the examples we will consider the mean

$$t_0(D) = \frac{1}{n} \sum_{i=1}^n Z_i$$

Statistical tools: Confidence interval



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Statistical tools: Confidence interval



- A confidence interval (CI) is an interval $[\theta_L, \theta_U]$ which likely contains θ
- The CI is a function of the data D. θ_L and θ_U are two statistics and for a concrete dataset the interval is computed to be

$$[\theta_L(D), \theta_U(D)]. \tag{4}$$

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$$[\theta_L(D), \theta_U(D)]. \tag{4}$$

• With probability $1-\alpha$, the true value θ should fall within the confidence interval $[\theta_L(D),\theta_U(D)]$ as we randomize over different datasets

$$P_{\theta}(\theta \in [\theta_L, \theta_U]) = 1 - \alpha. \tag{5}$$





• Determining whether a **null hypothesis** H_0 about the parameters is true or false

$$H_0: \theta = 0$$
 vs. $H_1: \theta \neq 0$

- ullet Intuitively, if H_0 is true, the data should behave in a certain way
 - ullet We test if the data is implausible assuming H_0





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$$t(D) = \frac{1}{n} \sum_{i=1}^{n} Z_i$$

On our dataset it has a particular value $t_0 = \frac{1}{n} \sum_{i=1}^n z_i$

• We can compute the density t(D) takes a particular value given H_0 is true:

$$p(t(D) = t|H_0) = p_{\theta = \theta_0}(t(D) = t)$$





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ullet p-value is the chance t(D) is at least as extreme as what we actually observed:

$$p$$
-value: $p = P(t(D) > |t_0| | H_0) = P_{\theta = \theta_0}(t(D) \ge |t_0|).$ (6)

Setup I: Fixed training set



Suppose we carry out cross-validation to obtain:

$$(\mathcal{D}_1^{\mathsf{train}}, \mathcal{D}_1^{\mathsf{test}}), \dots, (\mathcal{D}_K^{\mathsf{train}}, \mathcal{D}_K^{\mathsf{test}}). \tag{7}$$

We collect these into (paired) vectors of predictions and true values:

$$\hat{m{y}} = egin{bmatrix} \hat{m{y}}_1 \\ \hat{m{y}}_2 \\ \vdots \\ \hat{m{y}}_K \end{bmatrix}, \quad m{y} = egin{bmatrix} m{y}_1^{ ext{test}} \\ m{y}_2^{ ext{test}} \\ \vdots \\ m{y}_K^{ ext{test}} \end{bmatrix}.$$
 (8)

Evaluation of a single classifier



• Define:

$$c_i = \begin{cases} 1 & \text{if } \hat{y}_i = y_i \\ 0 & \text{if otherwise.} \end{cases}$$

• Number of accurate guesses:

$$m = \sum_{i=1}^{n} c_i.$$

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• Let the chance the classifier is correct be θ . Then, from Lecture 4, we know

$$p(\theta|m,n) = \text{Beta}(\theta|a,b), \quad a = m + \frac{1}{2}, \text{ and } b = n - m + \frac{1}{2}.$$
 (9)

Evaluating a single classifier (Jeffreys interval)



 \bullet If m is the number of accurate guesses, then

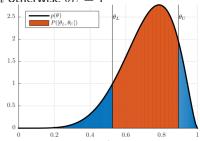
$$p(\theta|m,n) = \mathrm{Beta}(\theta|a,b), \quad a = m + \frac{1}{2}, \text{ and } b = n - m + \frac{1}{2}.$$

• The $1-\alpha$ confidence interval is given as $[\theta_L, \theta_U]$:

$$heta_L = \mathrm{cdf}_B^{-1}\left(rac{lpha}{2}|a,b
ight)$$
 if $m>0$ otherwise $heta_L=0$

$$heta_U = \mathrm{cdf}_B^{-1} \left(1 - rac{lpha}{2} | a, b
ight) \, \, ext{if} \, \, m < n \, \, ext{otherwise} \, \, heta_{^{II}} = 1$$

$$\hat{\theta} = \mathbb{E}[\theta] = \frac{a}{a+b}$$



Comparing two classifiers



• Assume we have predictions from both classifiers:

$$\hat{\boldsymbol{y}}^{A} = \hat{y}_{1}^{A}, \dots, \hat{y}_{n}^{A}, \quad \hat{\boldsymbol{y}}^{B} = \hat{y}_{1}^{B}, \dots, \hat{y}_{n}^{B}.$$

• As before, we want to know if the classifiers are correct or not:

$$c_i^A = \begin{cases} 1 & \text{if } \hat{y}_i^A = y_i \\ 0 & \text{if otherwise.} \end{cases} \quad \text{and} \quad c_i^B = \begin{cases} 1 & \text{if } \hat{y}_i^B = y_i \\ 0 & \text{if otherwise.} \end{cases}$$





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$$c_i^A = \begin{cases} 1 & \text{if } \hat{y}_i^A = y_i \\ 0 & \text{if otherwise.} \end{cases} \quad \text{and} \quad c_i^B = \begin{cases} 1 & \text{if } \hat{y}_i^B = y_i \\ 0 & \text{if otherwise.} \end{cases}$$

• The relevant information is the contingency table:

$$\begin{split} n_{11} &= \sum_{i=1}^n c_i^A c_i^B &= \{ \text{Both classifiers are correct} \} \\ n_{12} &= \sum_{k=1}^n c_i^A (1 - c_i^B) &= \{ A \text{ is correct, } B \text{ is wrong} \} \\ n_{21} &= \sum_{k=1}^n (1 - c_i^A) c_i^B &= \{ A \text{ is wrong, } B \text{ is correct} \} \\ n_{22} &= \sum_{k=1}^n (1 - c_i^A) (1 - c_i^B) = \{ \text{Both classifiers are wrong} \} \end{split}$$





- ullet We want to compare the accuracy difference: $heta= heta_A- heta_B$
- It is possible to show (approximately)

$$p(\theta|\mathbf{n}) = \frac{1}{2} \operatorname{Beta} \left(\frac{\theta+1}{2} \mid a=f, b=g \right),$$

$$f = \frac{E_{\theta}+1}{2} (Q-1) \quad g = \frac{1-E_{\theta}}{2} (Q-1)$$

$$E_{\theta} = \frac{n_{12}-n_{21}}{n}, \quad Q = \frac{n^{2}(n+1)(E_{\theta}+1)(1-E_{\theta})}{n(n_{12}+n_{21})-(n_{12}-n_{21})^{2}}.$$

$$\theta_{L} = 2\operatorname{cdf}_{B}^{-1} \left(\frac{\alpha}{2} \mid a=f, b=g \right) - 1, \quad \theta_{U} = 2\operatorname{cdf}_{B}^{-1} \left(1 - \frac{\alpha}{2} \mid a=f, b=g \right) - 1.$$

$$(10)$$

- For a p-value, note that A is better than B if $n_{12} > n_{21}$
- A p-value can be obtained as:

$$p = 2\text{cdf}_{\text{binom}}\left(m = \min\{n_{12}, n_{21}\} \mid \theta = \frac{1}{2}, N = n_{12} + n_{21}\right)$$





• Use cross-validation to obtain predictions \hat{y}_i and true values y_i . Select loss

$$z_i = |\hat{y}_i - y_i|$$
 or $z_i = (\hat{y}_i - y_i)^2$ (11)

- Estimated error is: $\hat{z} = \frac{1}{n} \sum_{i=1}^{n} z_i$.
- Assume each error is normally distributed (warning!)

$$p(D|u, \sigma^2) = \prod_{i=1}^n \mathcal{N}(z_i|u, \sigma^2)$$

ullet It is possible to show u follows a generalized Student's t-distribution:

$$p(u|D) = p_{\mathcal{T}}(u|\nu = n - 1, \mu = \hat{z}, \sigma = \tilde{\sigma})$$

with parameters
$$\hat{z}=\frac{1}{n}\sum_{i=1}^n z_i$$
 and $\tilde{\sigma}=\sqrt{\sum_{i=1}^n \frac{(z_i-\hat{z})^2}{n(n-1)}}.$

The Student's t-distribution has density

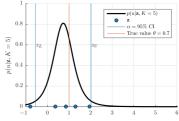
Student t-distribution
$$p_{\mathcal{T}}(x|\nu,\mu,\sigma) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)\sqrt{\pi\nu\sigma^2}} \left(1 + \frac{1}{\nu}\left[\frac{x-\mu}{\sigma}\right]^2\right)^{-\frac{\nu+1}{2}}$$
.

Confidence interval for a regression model



ullet Step back: Assuming $z_i = L(y_i, \hat{y}_i)$ and

$$z_i \sim \mathcal{N}(z_i|\mu=u,\sigma^2)$$



• In this case u is the average error rate. Since we have shown:

$$p(u|D) = p_{\mathcal{T}}(u|\nu = n - 1, \mu = \hat{z}, \sigma = \tilde{\sigma})$$

• An approximate $1-\alpha$ confidence interval is:

$$z_L = \operatorname{cdf}_{\mathcal{T}}^{-1} \left(\frac{\alpha}{2} \mid \nu, \hat{z}, \tilde{\sigma} \right), \ z_U = \operatorname{cdf}_{\mathcal{T}}^{-1} \left(1 - \frac{\alpha}{2} \mid \nu, \hat{z}, \tilde{\sigma} \right).$$
 (12)

Comparing two regression models



ullet Use cross-validation to obtain (paired) predictions along with true values y_i

$$\hat{y}_1^A,\dots,\hat{y}_n^A,\quad\text{ and }\quad \hat{y}_1^B,\dots,\hat{y}_n^B. \tag{13}$$

Select a loss-function to compute the per-observation losses as in

$$z_1^A, \ldots, z_n^A,$$
 and $z_1^B, \ldots, z_n^B.$

Note that

$$\begin{split} z &= E_{A,\mathcal{D}}^{\text{gen}} - E_{B,\mathcal{D}}^{\text{gen}} \approx \hat{z} = \left(\frac{1}{n}\sum_{i=1}^n z_i^A\right) - \left(\frac{1}{n}\sum_{i=1}^n z_i^B\right) \\ &= \frac{1}{n}\sum_{i=1}^n z_i, \quad \text{ where } z_i = z_i^A - z_i^B \end{split}$$

- Assume $z_i \sim \mathcal{N}(z_i|\mu=u,\sigma^2)$
- ullet Compute a 1-lpha CI using methods on previous slide

Comparing two regression models: p-values



$$z=E_A^{\mathsf{gen}}-E_B^{\mathsf{gen}}pprox\hat{z}=rac{1}{n}\sum_{i=1}^n z_i, \quad ext{ where } z_i=z_i^A-z_i^B$$

Assuming

$$z_i \sim \mathcal{N}(z_i|\mu=u,\sigma^2)$$

where u is the true difference in error function we have shown:

$$p(u|D) = p_{\mathcal{T}}(u|\nu = n - 1, \mu = \hat{z}, \sigma = \tilde{\sigma})$$

• Therefore, we can test the hypothesis

$$H_0$$
: Model \mathcal{M}_A and \mathcal{M}_B have the same performance, $u=0$ (14)

$$H_1$$
: Model \mathcal{M}_A and \mathcal{M}_B have different performance, $u \neq 0$. (15)

A p-value can be computed as

$$p = 2\operatorname{cdf}_{\mathcal{T}}(-|\hat{z}| \mid \nu = n - 1, \mu = 0, \sigma = \tilde{\sigma}).$$
 (16)



- When using **setup I** choose K as large as feasible (leave-one-out)
- Hold-out has the benefit the training/test data is fixed



- When using **setup I** choose *K* as large as feasible (leave-one-out)
- Hold-out has the benefit the training/test data is fixed
- Results will be significant with enough data



- When using **setup I** choose *K* as large as feasible (leave-one-out)
- Hold-out has the benefit the training/test data is fixed
- Results will be significant with enough data
 - Focus on estimated an effect size



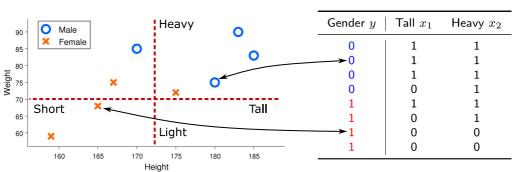
- When using **setup I** choose K as large as feasible (leave-one-out)
- Hold-out has the benefit the training/test data is fixed
- Results will be significant with enough data
 - Focus on estimated an effect size
 - Multiple-comparison problem



- When using **setup I** choose K as large as feasible (leave-one-out)
- Hold-out has the benefit the training/test data is fixed
- Results will be significant with enough data
 - Focus on estimated an effect size
 - Multiple-comparison problem
 - Transparency, availability of datasets/code, breadth of testing, self-criticism guarantees reprehensibility, not a sophisticated test
- In setup II, correlation of training data is taken into account and K-fold is optimal
 - Your setup I results do not generalize beyond your training data

Bayes and Naive-Bayes

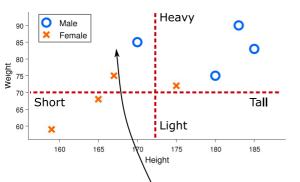




$$p(y|x_1, x_2) = \frac{p(x_1, x_2|y)p(y)}{\sum_{k=0}^{1} p(x_1, x_2|y=k)p(y=k)}$$

Example 1: Normal Bayes

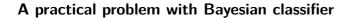




$Gender\ y$	Tall x_1	Heavy x_2		
0	1	1		
0	1	1		
0	1	1		
0	0	1		
1	1	1		
1	0	1		
1	0	0		
1	0	0		

Probability a short, heavy person is male:

$$P(y=0|x_1=0,x_2=1) = \frac{p(x_1=0,x_2=1|y=0)p(y=0)}{\sum_{k=0}^{1} p(x_1=0,x_2=1|y=k)p(y=k)}$$

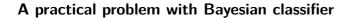




• In general:

$$p(y|x_1,x_2,\ldots,x_M) = \frac{p(x_1,x_2,\ldots,x_M|y)p(y)}{\sum_{k=0}^{K-1}p(x_1,x_2,\ldots,x_M|y=k)p(y=k)}$$

$$p(x_1,\ldots,x_M|y=k) = \frac{\mathsf{Nr. obs where }y=k \text{ and we measure }x_1,\ldots,x_M}{\mathsf{Observations where }y=k}$$





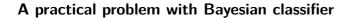
• In general:

$$p(y|x_1,x_2,\ldots,x_M) = \frac{p(x_1,x_2,\ldots,x_M|y)p(y)}{\sum_{k=0}^{K-1}p(x_1,x_2,\ldots,x_M|y=k)p(y=k)}$$

$$p(x_1,\ldots,x_M|y=k) = \frac{\text{Nr. obs where }y=k \text{ and we measure }x_1,\ldots,x_M}{\text{Observations where }y=k}$$

Naive Bayes assumption

$$p(x_1, x_2, \dots, x_M | y) = p(x_1 | y) p(x_2 | y) \times \dots \times p(x_M | y)$$





• In general:

$$p(y|x_1,x_2,\ldots,x_M) = \frac{p(x_1,x_2,\ldots,x_M|y)p(y)}{\sum_{k=0}^{K-1}p(x_1,x_2,\ldots,x_M|y=k)p(y=k)}$$

$$p(x_1,\ldots,x_M|y=k) = \frac{\text{Nr. obs where }y=k \text{ and we measure }x_1,\ldots,x_M}{\text{Observations where }y=k}$$

Naive Bayes assumption

$$p(x_1, x_2, \dots, x_M | y) = p(x_1 | y) p(x_2 | y) \times \dots \times p(x_M | y)$$

Naive Bayes classifier

$$p(y|x_1, x_2, ..., x_M) = \frac{p(x_1, x_2, ..., x_M|y)p(y)}{\sum_{k=0}^{1} p(x_1, x_2, ..., x_M|y = k)p(y = k)}$$

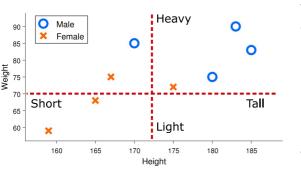
$$= \frac{p(x_1|y)p(x_2|y) \times \cdots \times p(x_M|y)p(y)}{\sum_{k=0}^{1} p(x_1|y = k)p(x_2|y = k) \times \cdots \times p(x_M|y = k)p(y = k)}$$

Example 2:

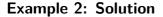


 Naive Bayes classifier (Probability someone is a female given they are heavy and tall)

$$p(y=1|x_1=1,x_2=1) = \frac{p(x_1|y)p(x_2|y)p(y)}{\sum_{k=0}^{1} p(x_1|y=k)p(x_2|y=k)p(y=k)}$$



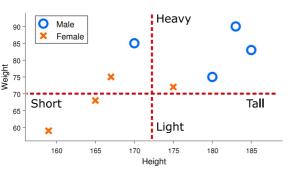
$Gender\ y$	Tall x_1	Heavy x_2		
0	1	1		
0	1	1		
0	1	1		
0	0	1		
1	1	1		
1	0	1		
1	0	0		
1	0	0		





 Naive Bayes classifier (Probability someone is a female given they are heavy and tall)

$$p(y=1|x_1=1, x_2=1) = \frac{p(x_1|y)p(x_2|y)p(y)}{\sum_{k=0}^{1} p(x_1|y=k)p(x_2|y=k)p(y=k)}$$
$$= \frac{\frac{1}{4}\frac{2}{4}\frac{1}{2}}{\frac{1}{4}\frac{2}{4}\frac{1}{2} + \frac{3}{4}\frac{4}{4}\frac{1}{2}} = \frac{2}{2+12} = \frac{1}{7}$$



Gender y	Tall x_1	Heavy x_2		
0	1	1		
0	1	1		
0	1	1		
0	0	1		
1	1	1		
1	0	1		
1	0	0		
1	0	0		

Quiz 1, Naive-Bayes (Spring 2012)

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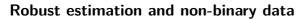
	S1	S2	S3	S4	S5	S6	S7	S8	S9	S10
		0								
P2	1	0	1	0	0	1	1	1	0	0
		1								
		1								
P5	1	0	0	1	1	0	0	1	0	1
P6	1	0	1	1	1	1	1	0	1	0

Table 1: Table indicating whether 10 songs denoted S1–S10 are downloaded to 6 different phones denoted P1–P6. P1 and P2 given in red are phones that belong to females whereas P3, P4, P5, and P6 given in blue belong to males.

The phones P1 and P2 are owned by females whereas P3, P4, P5 and P6 are owned by males (this is indicated in red and blue respectively in Table 1). We would like to predict whether a phone is owned by a male based on whether or not the songs S1, S2 and S3 have been downloaded. We will therefore classify whether the phone belongs to a male or female considering only the attributes S1, S2 and S3 and the data in Table 1. We will apply a Naïve Bayes classifier that assumes independence between these attributes Given that a phone has installed songs 1, 2 and 3 (i.e., S1=1, S2=1 and S3=1) What is the probability that the phone is owned by a male according to the Naïve Bayes classifier?

. Don't know.

$$p(y|x_1, x_2, ..., x_M) = \frac{p(x_1|y) \times ... \times p(x_M|y)p(y)}{\sum_{k=0}^{1} p(x_1|y=k) \times ... \times p(x_M|y=k)p(y=k)}$$





Assume

$$p(x_1, ..., x_M | y) = \prod_{k=1}^{M} p(x_k | y)$$

Defining $n_{x_i=k|y=c} = \sum_{i=1}^N \delta_{X_{ij},k} \delta_{y,c}$ we have more generally:

Binary case:
$$p(x_j = 1|y = c) = \frac{n_{x_j=1|y=c} + \alpha}{N_c + 2\alpha}$$
.

Categorical case:
$$p(x_j = k|y = c) = \frac{n_{x_j = k|y = c} + \alpha}{N_c + K\alpha}$$
.

Continious case:
$$p(x_j = x | y = c) = \mathcal{N}(x | \mu = \mu_{j|c}, \sigma^2 = (\sigma_{j|c} + \alpha)^2)$$

$$\mu_{j|c} = \mathbb{E}_{y=c}[x_j] = \frac{1}{N_c} \sum_{i=1}^{N} \delta_{y_i,c} X_{ij},$$

$$\sigma_{j|c} = \hat{\text{std}}_{y=c}[x_j] = \sqrt{\frac{1}{N_c - 1} \sum_{i=1}^N \delta_{y_i,c} (X_{ij} - \mu_c)^2}$$

Select these parameters using cross-validation.



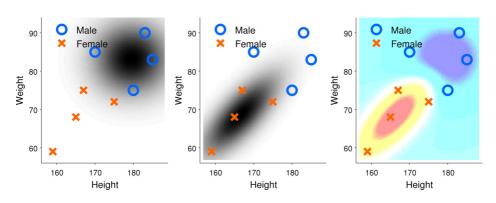
Bayesian classification by the multivariate normal distribution

Continuous density estimation

$$P(\boldsymbol{x}|y=c) = \frac{1}{(2\pi)^{M/2} det(\boldsymbol{\Sigma}_c)^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu}_c)^{\top} \boldsymbol{\Sigma}_c^{-1} (\boldsymbol{x} - \boldsymbol{\mu}_c)\right)$$

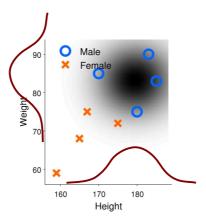
- Fit a Normal distribution to each class
 - Compute class mean and covariance
- Classify using Bayes rule as before

$$P(y = c|\boldsymbol{x}) = \frac{P(\boldsymbol{x}|y = c)P(y = c)}{\sum_{c'} P(\boldsymbol{x}|y = c')P(y = c')}$$





• What does the Naive Bayes assumption of independence of the attributes correspond to in terms of the parameters of the multivariate normal distribution?



Midterm practice test



Look at the test on DTU Learn. Note the test is not part of your evaluation.



In the analysis of house prices the following attributes were collected for a house: The year the house was built (denoted YEAR), the size of the house given in square meters (denoted SIZE) the county in which the house is located (denoted LOCATION). Which statement about the three attributes is correct?

- A. YEAR is ratio, SIZE is interval and LOCATION is nominal
- B. YEAR is interval, SIZE is ratio and LOCATION is nominal
- C. YEAR is interval, SIZE is ratio and LOCATION is ordinal
- $\mbox{D. YEAR is interval, SIZE is ratio and LOCATION } \\ \mbox{is interval}$
- E. Don't know.



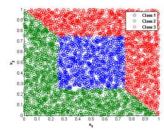
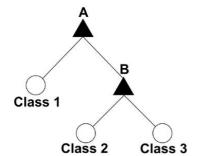


Figure 1



Consider the classification problem given in figure 1 and the Decision Tree shown below it with two decision nodes denoted A and B. We will let $\boldsymbol{x}_n = (x,y)$ denote a 2-dimensional observation such that $\boldsymbol{x}_n - 0.5 \cdot 1$ denotes the subtraction of 0.5 from each of the two coordinates of \boldsymbol{x}_n . Which one of the following classification rules would lead to a correct classification of the data?

A. A:
$$\|\boldsymbol{x}_n - 0.5 \cdot \boldsymbol{1}\|_1 \le 0.25$$
, B: $\|\boldsymbol{x}_n\|_{\infty} \le 1$

$$\text{B.} \quad \text{A:} \ \left\|\boldsymbol{x}_{n}\right\|_{1} \leq 1, \ \text{B:} \ \left\|\boldsymbol{x}_{n}-0.5 \cdot \boldsymbol{1}\right\|_{2} \leq \infty$$

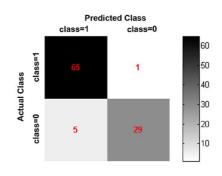
C. A:
$$\|\boldsymbol{x}_n - 0.5 \cdot \boldsymbol{1}\|_2 \le 0.25$$
, B: $\|\boldsymbol{x}_n\|_{\infty} \le 1$

D. A:
$$\|\boldsymbol{x}_n - 0.5 \cdot \boldsymbol{1}\|_{\infty} \le 0.25$$
, B: $\|\boldsymbol{x}_n\|_1 \le 1$

E. Don't know.



A classifier has the confusion matrix given in the figure below. Which statement about the classifier is correct?



- A. The Accuracy is 94% and the Error rate is 6%
- B. The Accuracy is 6% and the Error rate is 94%
- $\mathsf{C}.$ The Accuracy is 65% and the Error rate is 35%
- $\mbox{D. There is insufficient information in the confusion} \\ \mbox{matrix to determine the Accuracy and Error rate}.$
- E. Don't know.



Which statement about crossvalidation is wrong?

- A. Cross-validation can be used to estimate the generalization error.
- B. Leave one out cross-validation is more computationally expensive than 10 fold crossvalidation.
- C. Holding out one third of the data for validation is faster but less accurate than performing 10 fold cross-validation.
- D. The same test set can be used for model selection as well as evaluation of the generalization performance of the model.
- E. Don't know.



Consider a data set of four features: A, B, C, and D that are applied in a classification algorithm. The table below shows the cross-validated Error rate when using different combinations of the features.

Feature(s)	Error rate			
A	0.40			
В	0.45			
$^{\mathrm{C}}$	0.33			
D	0.42			
A and B	0.20			
A and C	0.25			
A and D	0.34			
B and C	0.29			
B and D	0.42			
C and D	0.40			
A and B and C	0.13			
A and B and D	0.17			
B and C and D	0.10			
A and C and D	0.15			
A and B and C and D	0.28			

We will apply a forward feature selection algorithm. Which feature set will the selection algorithm choose?

- A. C
- $\mathsf{B.}\ B \ \mathrm{and}\ C \ \mathrm{and}\ D$
- $\mathsf{C.}\ A \text{ and } B$
- $\mathsf{D.}\ A \ \mathrm{and}\ B \ \mathrm{and}\ C$
- E. Don't know.



When training a decision tree we will use the classification error as impurity measure I(t) given by $I(t) = 1 - \max_i [p(i|t)]$ where p(i|t) denotes the fraction of data objects belonging to class i at a given node t. We will use $\operatorname{Hunt\^a} \mathfrak{C}^{\mathbb{T}^{\mathsf{N}}}$ s algorithm to grow the tree and recall that the purity gain is given by:

$$\Delta = I(\text{ Parent }) - \sum_{j=1}^{k} \frac{N(v_j)}{N} I(v_j)$$

where N is the total number of data objects at the parent node, k is the number of child nodes and $N(v_j)$ is the number of data objects associated with the child node, v_j . We will consider classification of Iris flowers into Iris-Setosa, Iris-Virginica and Iris-Versicolor. At a potential split we have:

 Before the split: 5 Iris-Setosa, 10 Iris-Virginica and 10 Iris Versicolor.

After the split

- 0 Iris-Setosa, 8 Iris-Virginica and 2 Iris-Versicolor in the left node.
- 5 Iris-Setosa, 2 Iris-Viriginica and 8 Iris-Versicolor in the right node.

Which statement is correct?

- A. The purity gain is $\Delta = \frac{3}{5}$
- B. The purity gain is $\Delta = \frac{3}{15}$
- C. The purity gain is $\Delta = \frac{6}{25}$
- D. The purity gain is $\Delta = \frac{7}{15}$
- E. Don't know.



When people are well rested and take an exam their chance of passing the exam is 90%, however, when people are not well rested there chance of passing the exam is only 40%. On any given day 80% of people are well-rested. What is the chance that a person passing

the test is well rested?

- A. $\frac{4}{10}$
- B. $\frac{8}{10}$
- C. $\frac{9}{10}$
- D. $\frac{10}{11}$
- E. Don't know.



When carrying out a principal component analysis of a dataset with four attributes we obtain the following singular values $\sigma_1 = 4$, $\sigma_2 = 2$, $\sigma_3 = 1$, and $\sigma_4 = 0$.

Which one of the following statements is wrong?

- A. The first principal component accounts for more than 60% of the variation in the data.
- B. The third principal component accounts for less than 5% of the variation in the data.
- C. The second principal component accounts for more than 20% of the variation in the data.
- D. The data can be perfectly represented in a three dimensional sub-space.
- E. Don't know.



Consider the following sequence of numbers

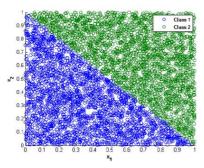
$$x = \begin{bmatrix} 0 & 1 & 1 & 1 & 2 & 3 & 4 & 4 & 5 & 14 \end{bmatrix}.$$

What is the sum of the mean, the median and the mode of these numbers, i.e. what is the value: y =

mean(x) + median(x) + mode(x)?

- A. y = 1
- B. y = 6
- C. y = 7
- D. y = 11
- E. Don't know.





Consider the classification problem given in the figure below where x_1 and x_2 are used as features for a logistic regression classifier and a decision tree. The considered logistic regression models all include the constant term w0. Which one of the following

statements is wrong?

- A. The two classes can be perfectly separated by a logistic regression model using x_1 and x_2 as features.
- B. A decision tree with less than five nodes, all of the usual axis-aligned form $x_1 > a$ or $x_2 > b$ for different values of a, b, can perfectly separate the classes using only x_1 and x_2 as features.
- C. A logistic regression model can perfectly separate the two classes using only the feature z given by $z=x_1+x_2.$
- D. In logistic regression the probability that each observation belong to the two classes can be derived from the logistic function.
- E. Don't know.

Resources



https://www.youtube.com Video explaining Naive Bayes

(https://www.youtube.com/watch?v=8yvBqhm92xA)

https://machinelearningmastery.com Statistical comparison of the cross-validation estimate of the generalization error is not a solved problem. This reference provides an overview of various issues and proposed solutions. Note no simple solution exists.

(https://machinelearningmastery.com/

statistical-significance-tests-for-comparing-machine-learning-algorithms/)