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1 Introduction

Objectives

- Understand the concept of Jointpoint regression
- Learn how to perform Jointpoint regression with R
- Peform data analyses where the scientific question is to determine changes in temporal trends (incidence or mortality rates)

2 Package installation

Package can be installed from CRAN

```
install.packages("ljr")
```

After that it is loaded as usual

```
library("ljr")
```

It contains the following functions:

```
ls(2)
## [1] "ljr0" "ljr01" "ljr1" "ljr11" "ljrb" "ljrf" "ljrjk" "ljrk" "ljrkk"
```

- kcm: Kentucky yearly cancer mortality from 1999-2005
- *ljr0*: MLE with 0 joinpoints
- *ljr01*: Perform test of 0 vs 1 joinpoints}
- *ljr1*: MLE with 1 joinpoint}
- *ljr11*: Test coefficients conditioned on K=1 joinpoint}

- *Ijrb*: Perform backward joinpoint selection algorithm with upper bound K}
- Ijrf: Perform forward joinpoint selection algorithm with unlimited upper bound
- *ljrjk*: Perform test of j vs k joinpoints
- *ljrk*: MLE with k joinpoints}
- *ljrkk*: Test coefficients conditioned on K=k joinpoint}

3 Data analysis

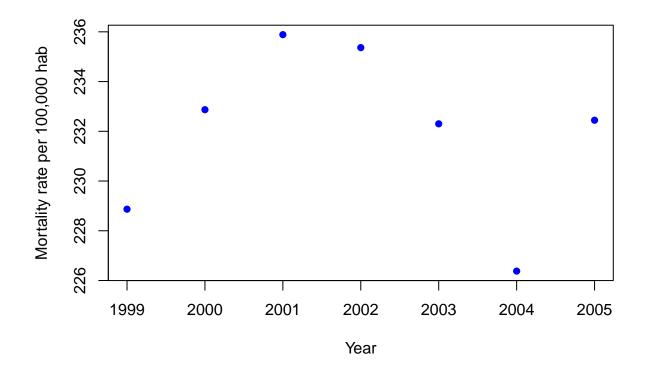
Let us use data available at the package

```
data(kcm)
head(kcm)
## Year Count Population
## 1 1999 9196 4018053
## 2 2000 9412 4041769
## 3 2001 9595 4067643
## 4 2002 9624 4088977
## 5 2003 9558 4114489
## 6 2004 9373 4140427
```

3.1 Data visualization

As in any other statistical data analysis, first, let us have a look at the data

```
kcm$rate <- (kcm$Count/kcm$Population)*100000
plot(kcm$Year, kcm$rate, xlab="Year",
    ylab="Mortality rate per 100,000 hab",
    type="n")
points(kcm$Year, kcm$rate, pch=16, col="blue")</pre>
```



3.2 Simple trend analysis

One may be interested in estimating a Poisson model to determine whether there is any change in global tren. Here the null hypothesis is that slope is 0.

```
modPoisson <- glm(Count~Year+offset(log(Population)),</pre>
               family=poisson, data=kcm)
modPoisson
##
## Call: glm(formula = Count ~ Year + offset(log(Population)), family = poisson,
##
       data = kcm)
##
## Coefficients:
## (Intercept)
                        Year
##
   -4.1972066
                 -0.0009335
##
## Degrees of Freedom: 6 Total (i.e. Null); 5 Residual
## Null Deviance:
                         12.2
## Residual Deviance: 11.96
                                 AIC: 92.94
```

Hence, the annual percentage change would be:

```
round((exp(modPoisson$coef[2])-1)*100, 2)
## Year
## -0.09
```

with confidence interval

```
round((exp(confint(modPoisson)[2,])-1)*100, 2)
### 2.5 % 97.5 %
### -0.47 0.29
```

In some ocassions there is overdispersion and negative binomial (NB) distribution must be used instead. Dispersion can be estimated by using residual deviance (null deviance / df residual). When this coefficient is >1 Poisson distribution is not adequate and NB has to be used. Here you can see an approximate test. Other better tests can be found in the library pscl.

 H_0 : There is no overdispersion

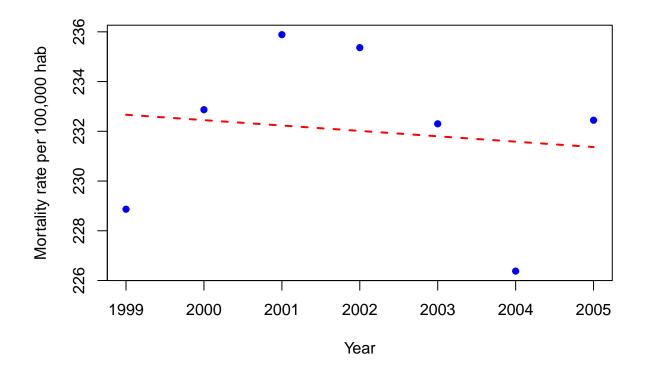
The p-value corresponding to this test can be obtained by

```
1 - pchisq(modPoisson$deviance, modPoisson$df.res)
## [1] 0.03527432
modPoisson
##
## Call: glm(formula = Count ~ Year + offset(log(Population)), family = poisson,
##
      data = kcm)
##
## Coefficients:
## (Intercept)
                      Year
## -4.1972066 -0.0009335
##
## Degrees of Freedom: 6 Total (i.e. Null); 5 Residual
## Null Deviance:
                        12.2
## Residual Deviance: 11.96
                                AIC: 92.94
```

When rejecting the null hypothesis, negative binomial model can be fitted as

```
library(MASS)
modNB <- glm.nb(Count~Year+ offset(log(Population)),</pre>
                data=kcm)
modNB
##
## Call: glm.nb(formula = Count ~ Year + offset(log(Population)), data = kcm,
       init.theta = 13398.06067, link = log)
##
##
## Coefficients:
## (Intercept)
                       Year
## -4.2218596 -0.0009212
##
## Degrees of Freedom: 6 Total (i.e. Null); 5 Residual
## Null Deviance:
                       7.141
## Residual Deviance: 7.01 AIC: 93.73
```

However, as illustrated in the next plot, linear trend is not valid



4 Joinpoint analysis

4.1 Model for a given number of joinpoints

In general, one may be interested in estimating the best model for a given number of joinpoints (obtained by visualy inspecting the overal tren). For example, for 1 joinpoint the model is estimated with the function ljrk

```
ljrk(1, kcm$Count, kcm$Population, kcm$Year+.5)
## Model:
## y~Binom(n,p) where p=invlogit(eta)
## eta=b0+g0*t+g1*max(t-tau1,0)
##
##
          Variables
                             Coef
## b0
          Intercept -40.81272431
## g0
                  t
                      0.01737196
## g1 max(t-tau1,0)
                     -0.02418284
##
## Joinpoints:
##
## 1 tau1= 2001.273
## $Coef
##
                              t max(t-tau1,0)
       Intercept
##
   -40.81272431
                    0.01737196
                                  -0.02418284
##
```

```
## $Joinpoints
## tau1=
## 2001.273
##
## $wlik
## [1] -0.112523
```

Notice that we have used the Year variable +0.5 for interpreting purposes. The model for 2 joinpoints in estimated by

```
ljrk(2, kcm$Count, kcm$Population, kcm$Year+.5)
## Model:
## y~Binom(n,p) where p=invlogit(eta)
## eta=b0+g0*t+g1*max(t-tau1,0)+g2*max(t-tau2,0)
##
         Variables
                           Coef
## b0
         Intercept -36.31521597
             t
                    0.01512302
## g0
## g1 max(t-tau1,0) -0.03460547
## g2 max(t-tau2,0) 0.04383173
##
## Joinpoints:
##
## 1 tau1= 2004.500
## 2 tau2= 2002.039
## $Coef
##
      Intercept
                           t max(t-tau1,0) max(t-tau2,0)
##
  -36.31521597 0.01512302 -0.03460547
                                               0.04383173
##
## $Joinpoints
##
     tau1= tau2=
## 2004.500 2002.039
##
## $wlik
## [1] -0.1125223
```

4.2 Model selection using LRT

When two different models may fit the data, we can select the best one by using a likelihood ratio test (LRT) where the p-value is computed by using Montecarlo method. The function for that purpose is ljrjk and can be fitted in the case of comparing 1 an 2 joinpoints by executing

```
ljrjk(1, 2, kcm$Count, kcm$Population, kcm$Year+.5,
            R=1000, alpha=.05)
## Testing HO: 1 joinpoint(s) vs. H1: 2 joinpoints
## p-value= 0.05
## Null hypothesis is rejected
##
## Model:
## y~Binom(n,p) where p=invlogit(eta)
## eta=b0+g0*t+g1*max(t-tau1,0)+g2*max(t-tau2,0)
##
##
          Variables
## b0
          Intercept -36.31521597
## g0
                 t
                    0.01512302
```

```
## g1 max(t-tau1,0)
                     -0.03460547
## g2 max(t-tau2,0)
                      0.04383173
##
## Joinpoints:
##
## 1 tau1= 2004.500
## 2 tau2= 2002.039
## $Coef
##
                             t \max(t-tau1,0) \max(t-tau2,0)
       Intercept
   -36.31521597
                    0.01512302 -0.03460547
                                                 0.04383173
##
##
## $Joinpoint
##
     tau1=
               tau2=
## 2004.500 2002.039
##
## $wlik
## [1] -0.1125223
## $pval
## [1] 0.05
```

Here we can conclude that

Once the model is selected, we can compute the anual percentage of change of each segment by

```
mod <- ljrk(2, kcm$Count, kcm$Population, kcm$Year+.5)</pre>
## Model:
## y~Binom(n,p) where p=invlogit(eta)
## eta=b0+g0*t+g1*max(t-tau1,0)+g2*max(t-tau2,0)
##
##
          Variables
                            Coef
## b0
          Intercept -36.31521597
## g0
                 t
                     0.01512302
## g1 max(t-tau1,0) -0.03460547
## g2 max(t-tau2,0)
                     0.04383173
##
## Joinpoints:
##
## 1 tau1= 2004.500
## 2 tau2= 2002.039
cbind(year=c(1999, mod$Joinpoints),
      APC=round((exp(mod\$Coef[-1])-1)*100,2))
##
             year APC
         1999.000 1.52
##
## tau1= 2004.500 -3.40
## tau2= 2002.039 4.48
```

4.3 Automatic method of model selection

In some ocassion the user is interested in selecting the best modelby using an automatic method by using backward or forward approaches Si queremos escoger el modelo de forma automática, utilizamos un método de selección 'backward' o 'forward'. This can be performed by using the function ljrb or ljrf. Note that in 'barckward' method the argument K is required. This indicates the maximum number of joinpoints to be tested. The 'forward' method starts from the null model (e.g. linear trend or no joinpoints) and tests whether the inclusion of a new joinpoint is statistically significant or not.

```
ljrb(K=3, kcm$Count, kcm$Population, kcm$Year+.5)
## Backward algorithm for determining the number of joinpoints:
## Step 1 : Test HO: O joinpoint(s) vs H1: 3 joinpoint(s)
## p-value= 0.023
## Step 2 : Test HO: 0 joinpoint(s) vs H1: 2 joinpoint(s)
## p-value= 0.011
## Step 3 : Test HO: 1 joinpoint(s) vs H1: 2 joinpoint(s)
## p-value= 0.042
##
## Estimated number of joinpoints= 1
##
## Model:
## y~Binom(n,p) where p=invlogit(eta)
## eta=ofst+b0+g0*t+g1*max(t-tau1,0)
##
         Variables
##
## b0 Intercept -40.81272431
## g0
         t 0.01737196
## g1 max(t-tau1,0) -0.02418284
##
## Joinpoints:
##
## 1 tau1= 2001.273
## $Coef
                    t max(t-tau1,0)
##
      Intercept
## -40.81272431 0.01737196 -0.02418284
##
## $Joinpoints
## tau1=
## 2001.273
##
## $wlik
## [1] -0.112523
##
## $pvals
## [1] 0.023 0.011 0.042
ljrf(kcm$Count, kcm$Population, kcm$Year+.5)
## Forward algorithm for determining the number of joinpoints:
## Step 1 : Test HO: 0 joinpoint(s) vs H1: 1 joinpoint(s)
## p-value= 0.05
## Estimated number of joinpoints= 0
##
## Model:
## y~Binom(n,p) where p=invlogit(eta)
## eta=b0+g0*t
##
     Variables
## b0 Intercept -4.190069631
## g0
      t -0.000935695
## $Coef
## Intercept
## -4.190069631 -0.000935695
```

```
##
## $wlik
## [1] -0.1125237
##
## $pvals
## [1] 0.05 0.00 0.00 0.00
```

We observe that

5 References

■ The pscl package

6 Session information

```
## R version 3.3.2 (2016-10-31)
## Platform: x86_64-w64-mingw32/x64 (64-bit)
## Running under: Windows 10 x64 (build 14393)
##
## locale:
## [1] LC_COLLATE=Spanish_Spain.1252 LC_CTYPE=Spanish_Spain.1252
## [3] LC_MONETARY=Spanish_Spain.1252 LC_NUMERIC=C
## [5] LC_TIME=Spanish_Spain.1252
## attached base packages:
## [1] stats
            graphics grDevices utils datasets methods
                                                               base
##
## other attached packages:
## [1] MASS_7.3-45
                  ljr_1.4-0
                                    knitr_1.15.1
                                                    BiocStyle_2.2.1
##
## loaded via a namespace (and not attached):
## [1] backports_1.0.5 magrittr_1.5 rprojroot_1.2 tools_3.3.2
## [5] htmltools_0.3.5 yaml_2.1.14 Rcpp_0.12.9
                                                     stringi_1.1.2
## [9] rmarkdown_1.3 stringr_1.1.0 digest_0.6.11
                                                     evaluate_0.10
```