#### **Model Error**

**Empirical Risk**  $\hat{R}_D(f) = \frac{1}{n} \sum \ell(y, f(x))$ **Population Risk**  $R(f) = \mathbb{E}_{x,y \sim p}[\ell(y, f(x))]$ 

It holds that  $\mathbb{E}_D[\hat{R}_D(\hat{f})] \leq R(\hat{f})$ . We call  $R(\hat{f})$ the generalization error.

#### **Bias Variance Tradeoff:**

Pred. error = 
$$\operatorname{Bias}^2$$
 + Variance + Noise  

$$\mathbb{E}_D[R(\hat{f})] = \mathbb{E}_x[(f^*(x) - \mathbb{E}_D[\hat{f}(x)])^2] + \mathbb{E}_x[(\hat{f}(x) - \mathbb{E}_D[\hat{f}(x)])^2] + \sigma$$

**Bias**: how close  $\hat{f}$  can get to  $f^*$ 

**Variance**: how much  $\hat{f}$  changes with D

# Regression

**Squared loss** (convex)

$$\frac{1}{n}\sum(y_i - f(x_i))^2 = \frac{1}{n}||y - Xw||_2^2$$
$$\nabla_w L(w) = 2X^{\top}(Xw - y)$$

Solution:  $\hat{w} = (X^{\top}X)^{-1}X^{\top}v$ 

## Regularization

**Lasso Regression** (sparse)

$$\underset{w \in \mathbb{R}^d}{\operatorname{argmin}} ||y - \Phi w||_2^2 + \lambda ||w||_1$$

# **Ridge Regression**

$$\underset{w \in \mathbb{R}^d}{\operatorname{argmin}} ||y - \Phi w||_2^2 + \lambda ||w||_2^2$$

$$\nabla_w L(w) = 2X^{\top} (Xw - y) + 2\lambda w$$
Solution:  $\hat{w} = (X^{\top} X + \lambda I)^{-1} X^{\top} y$ 

#### **Cross-Validation**

- For all folds i = 1, ..., k:
  - Train  $\hat{f}_i$  on  $D' D'_i$
  - Val. error  $R_i = \frac{1}{|D'|} \sum \ell(\hat{f}_i(x), y)$
- Compute CV error  $\frac{1}{k} \sum_{i=1}^{k} R_i$
- Pick model with lowest CV error

## **Gradient Descent**

Converges only for convex case.

$$w^{t+1} = w^t - \eta_t \cdot \nabla \ell(w^t)$$

For linear regression:

||
$$w^t - w^*$$
|| $_2 \le ||I - \eta X^\top X||_{op}^t ||w^0 - w^*||_2$   
 $\rho = ||I - \eta X^\top X||_{op}^t \text{ con. speed for const. } \eta.$   
Opt. fixed  $\eta = \frac{2}{\lambda_{\min} + \lambda_{\max}}$  and max.  $\eta \le \frac{2}{\lambda_{\max}}$ .  
Momentum:  $w^{t+1} = w^t + \gamma \Delta w^{t-1} - \eta_t \nabla \ell(w^t)$ 

## Classification

**Zero-One loss** not convex or continuous  $\ell_{0-1}(\hat{f}(x), y) = \mathbb{I}_{y \neq \operatorname{sgn}\hat{f}(x)}$ 

**Logistic loss** 
$$\log(1 + e^{y\hat{f}(x)})$$
  $\nabla \ell(\hat{f}(x), y) = \frac{y_i x_i}{1 + e^{y_i \hat{f}(x)}}$ 

**Hinge loss**  $\max(0, 1 - y\hat{f}(x))$ 

## **Softmax Transformation**

$$\mathbb{P}[y = +1|x] = \frac{1}{1 + e^{-\hat{f}(x)}}$$

Multi-Class:  $\hat{p}_k = \frac{e^{\hat{f}_k(x)}}{\sum_{i=1}^K e^{\hat{f}_j(x)}}$ 

#### Linear Classifiers

 $f(x) = w^{\top}x$ , the decision boundary f(x) = 0.

# **Maximum-Margin Solution:**

 $w_{\text{MM}} = \operatorname{argmax} \operatorname{margin}(w) \text{ with } ||w||_2 = 1$ Where margin(w) =  $y_i w^{\top} x_i$ .

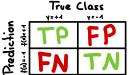
# Support Vector Machines

#### Hard SVM

 $\hat{w} = \min_{w} ||w||_2 \text{ s.t. } \forall i \ y_i w^{\top} x_i \ge 1$ **Soft SVM** allow "slack" in the constraints  $\hat{w} = \min_{w,\xi} \frac{1}{2} ||w||_2^2 + \lambda \sum_{i=1}^n \max(0, 1 - y_i w^\top x_i)$ 

#### Metrics

Choose +1 as the more important class.



error<sub>2</sub>/FNR :  $\frac{FN}{TP + FN}$ Precision TPR / Recall :  $\frac{1P}{TP + FN}$ 

AUROC: Plot TPR vs. FPR and compare different ROC's with area under the curve. **F1-Score**:  $\frac{2TP}{2TP + FP + FN}$ 

## Kernels

Parameterize:  $w = \Phi^{\top} \alpha$ ,  $K = \Phi \Phi^{\top}$ A kernel is **valid** if *K* is sym.: k(x,z) = k(z,x)and psd:  $z^{\top}Kz > 0$ 

**lin.**: 
$$k(x,z) = x^{\top}z$$
, **poly**.:  $k(x,z) = (x^{\top}z+1)^m$   
**rbf**:  $k(x,z) = \exp(-\frac{||x-z||\alpha}{z})$ 

 $\alpha = 1 \Rightarrow \text{laplacian kernel}$ 

 $\alpha = 2 \Rightarrow$  gaussian kernel

# **Kernel composition rules**

 $\forall f \text{ convex. } k = f(k_1)$ 

Mercers Theorem: Valid kernels can be decomposed into a lin. comb. of inner products. Initialize with k-Means++:

**Kern. Ridge Reg.**  $\frac{1}{n}||y-K\alpha||_2^2 + \lambda \alpha^{\top} K\alpha$ 

# **KNN Classification**

- Pick k and distance metric d
- For given x, find among  $x_1,...,x_n \in D$ the k closest to  $x \to x_{i_1}, ..., x_{i_k}$
- Output the majority vote of labels

#### **Neural Networks**

*w* are the weights and  $\varphi : \mathbb{R} \mapsto \mathbb{R}$  is a nonlinear activation function:  $\phi(x, w) = \phi(w^{\top}x)$ 

**ReLU:** max(0,z), **Tanh:**  $\frac{\exp(z) - \exp(-z)}{\exp(z) + \exp(-z)}$ **Sigmoid:**  $\frac{1}{1+\exp(-z)}$ 

Universal Approximation Theorem: We can approximate any arbitrary smooth target function, with 1+ layer with sufficient width.

# Forward Propagation

Input:  $v^{(0)} = [x; 1]$  Output:  $f = W^{(L)}v^{(L-1)}$ Hidden:  $z^{(l)} = W^{(l)}v^{(l-1)}, v^{(l)} = [\varphi(z^{(l)}); 1]$ 

# Backpropagation

Non-convex optimization problem:

$$\sum_{i=1}^{n} \max(0, 1 - y_i w^{\top} x_i) = \frac{\partial \ell}{\partial W^{(L)}} = \frac{\partial \ell}{\partial W^{(L)}} = \frac{\partial \ell}{\partial f} \frac{\partial f}{\partial W^{(L)}}$$

$$\sum_{i=1}^{n} \max(0, 1 - y_i w^{\top} x_i) = \frac{\partial \ell}{\partial W^{(L)}} = \frac{\partial \ell}{\partial W^{(L)}} = \frac{\partial \ell}{\partial f} \frac{\partial f}{\partial W^{(L)}} = \frac{\partial \ell}{\partial G} \frac{\partial f}{\partial G} \frac{\partial f}{\partial W^{(L)}} = \frac{\partial \ell}{\partial G} \frac{\partial f}{\partial G} \frac{\partial f}{\partial G} \frac{\partial f}{\partial G} = \frac{\partial \ell}{\partial G} \frac{\partial f}{\partial G} \frac{\partial f}{\partial$$

**Regularization**; Early Stopping; Dropout: Assume that data is generated iid. by some ignore hidden units with prob. p, after train- p(x,y). We want to find  $f: X \mapsto Y$  that miniing use all units and scale weights by p; mizes the population risk. Batch Normalization: normalize the input Opt. Predictor for the Squared Loss data (mean 0, variance 1) in each layer

**CNN** 
$$\varphi(W * v^{(l)})$$

Overfitting

The output dimension when applying m dif- Estimate  $\hat{p}(y \mid x)$  with MLE: ferent  $\hat{f} \times f$  filters to an  $n \times n$  image with padding p and stride s is:  $l = \frac{n+2p-f}{s} + 1$ 

# **Unsupervised Learning**

# k-Means Clustering

Optimization Goal (non-convex):

$$\hat{R}(\mu) = \sum_{i=1}^{n} \min_{j \in \{1, \dots, k\}} ||x_i - \mu_j||_2^2$$

 $k = k_1 + k_2$ ,  $k = k_1 \cdot k_2$   $\forall c > 0$ .  $k = c \cdot k_1$ , Lloyd's heuristics: Init. cluster centers  $\mu^{(0)}$ :

- Assign points to closest center
- Update  $\mu_i$  as mean of assigned points
  - Random data point  $\mu_1 = x_i$
  - Add  $\mu_2, ..., \mu_k$  rand., with prob: given  $\mu_{1:i}$  pick  $\mu_{i+1} = x_i$

where 
$$p(i) = \frac{1}{z} \min_{l \in \{1,...,j\}} ||x_i - \mu_l||_2^2$$

Find k by negligible loss decrease or reg.

# **Principal Component Analysis**

The optimal solution is given by  $z_i = w^{\top} x_i$ . Optimization goal:

$$\hat{w} = \underset{||w||_2=1}{\operatorname{argmin}} \sum_{i=1}^n (w^\top x_i)^2 = \underset{||w||_2=1}{\operatorname{argmin}} w^\top \Sigma w$$

Where  $\Sigma = \frac{1}{n} \sum_{i=1}^{n} x_i x_i^{\top}$  is the empirical covariance. Closed form solution given by the principal eigenvector of  $\Sigma$ , i.e.  $w = v_1$  for  $\lambda_1 \geq ... \geq \lambda_d \geq 0$ :  $\Sigma = \sum_{i=1}^d \lambda_i v_i v_i^{\top}$ 

For k > 1 we have to change the normalization to  $W^{\top}W = I$  then we just take the first k principal eigenvectors so that  $W = [v_1, ..., v_k]$ . PCA through SVD

The first *k* columns of *V* where  $X = USV^{\top}$ . Kernel PCA

$$\Sigma = \frac{1}{n} \sum_{i=1}^{n} x_i x_i^{\top} = X^{\top} X \Rightarrow \text{kernel trick:}$$

$$\hat{\alpha} = \operatorname{argmax}_{\alpha} \frac{\alpha^{\top} K^{\top} K \alpha}{\alpha^{\top} K \alpha}$$

Closed form solution:

$$\alpha^{(i)} = \frac{1}{\sqrt{\lambda_i}} v_i \quad K = \sum_{i=1}^n \lambda_i v_i v_i^{\top}$$

A point *x* is projected as:  $z_i = \sum_{j=1}^n \alpha_j^{(i)} k(x_j, x)$ Autoencoders

 $\hat{x} = f_{dec}(f_{enc}(x, \theta_{enc}); \theta_{dec})$ for ReLu and  $1/n_{in}$  or  $1/(n_{in}+n_{out})$  for Tanh) Lin. activation func. & square loss => PCA

# Statistical Perspective

f minimizing the population risk:

$$f^*(x) = \mathbb{E}[y \mid X = x] = \int y \cdot p(y \mid x) dy$$
  
Estimate  $\hat{p}(y \mid x)$  with MLE:

$$\theta^* = \underset{\theta}{\operatorname{argmax}} \hat{p}(y_1, ..., y_n \mid x_1, ..., x_n, \theta)$$
$$= \underset{\theta}{\operatorname{argmin}} - \sum_{i=1}^n \log p(y_i \mid x, \theta)$$

The MLE for linear regression is unbiased and has minimum variance among all unbiased estimators. However, it can overfit.

# Ex. Conditional Linear Gaussian

Assume Gaussian noise  $y = f(x) + \varepsilon$  with  $\varepsilon \sim$  $\mathcal{N}(0, \sigma^2)$  and  $f(x) = w^{\top}x$ :

$$\hat{p}(y \mid x, \theta) = \mathcal{N}(y; w^{\top} x, \sigma^2)$$

The optimal  $\hat{w}$  can be found using MLE:

 $\hat{w} = \underset{w}{\operatorname{argmax}} \ p(y|x, \theta) = \underset{w}{\operatorname{argmin}} \sum_{i=1} (y_i - w^{\top} x_i)^2$ 

# Maximum a Posteriori Estimate

Introduce bias to reduce variance. The small weight assumption is a Gaussian prior  $w_i \sim$  $\mathcal{N}(0,\beta^2)$ . The posterior distribution of w is

given by: 
$$p(w \mid x, y) = \frac{p(w) \cdot p(y \mid w, x)}{p(y \mid x)}$$
  
Now we want to find the MAP for  $w$ :  
 $\hat{w} = \underset{w}{\operatorname{argmax}} p(w \mid \bar{x}, \bar{y})$ 

$$= \underset{w}{\operatorname{argmin}} \frac{\sigma^{2}}{\beta^{2}} ||w||_{2}^{2} + \sum_{i=1}^{n} (y_{i} - w^{\top} x_{i})^{2}$$

Regularization can be understood as MAP inference, with different priors (= regularizers) and likelihoods (= loss functions).

# Statistical Models for Classification

f minimizing the population risk:

$$f^*(x) = \operatorname{argmax}_{\hat{y}} p(\hat{y} \mid x)$$

This is called the Bayes' optimal predictor for the 0-1 loss. Assuming iid. Bernoulli noise, the conditional probability is:

$$p(y \mid x, w) \sim \text{Ber}(y; \sigma(w^{\top}x))$$

Where  $\sigma(z) = \frac{1}{1 + \exp(-z)}$  is the sigmoid function. Using MLE we get:

$$\hat{w} = \underset{w}{\operatorname{argmin}} \sum_{i=1}^{n} \log(1 + \exp(-y_i w^{\top} x_i))$$

Which is the logistic loss. Instead of MLE we can estimate MAP, e.g. with a Gaussian prior:

$$\hat{w} = \underset{w}{\operatorname{argmin}} \lambda ||w||_{2}^{2} + \sum_{i=1}^{n} \log(1 + e^{-y_{i}w^{T}x_{i}})$$

# **Bayesian Decision Theory**

mum expected utility.

$$a^* = \operatorname{argmin}_{a \in A} \mathbb{E}_y[C(y, a) \mid x]$$

Can be used for asymetric costs or abstention. Assume that data is generated from a convex-

# Generative Modeling

using Bayes' rule.

# **Naive Bayes Model**

GM for classification tasks. Assuming for a for the Gaussian distributions. class label, each feature is independent. This helps estimating  $p(x \mid y) = \prod_{i=1}^{d} p(x_i \mid y_i)$ .

# **Gaussian Naive Bayes Classifier**

Naive Bayes Model with Gaussians features. Estimate the parameters via MLE:

MLE for class prior: 
$$p(y) = \hat{p}_y = \frac{\text{Count}(Y=y)}{n}$$
  
MLE for feature distribution:

Where: 
$$p(x_i \mid y) = \mathcal{N}(x_i; \hat{\mu}_{y,i}), \sigma_{y,i}^2$$

$$\mu_{y,i} = \frac{1}{\text{Count}(Y=y)} \sum_{j \mid y_j = y} x_{j,i}$$

$$\mu_{y,i} = \frac{1}{\text{Count}(Y=y)} \sum_{j \mid y_j = y} x_{j,i} 
\sigma_{y,i}^2 = \frac{1}{\text{Count}(Y=y)} \sum_{j \mid y_j = y} (x_{j,i} - \hat{\mu}_{y,i})^2$$

Predictions are made by:

$$y = \underset{\hat{y}}{\operatorname{argmax}} p(\hat{y} \mid x) = \underset{\hat{y}}{\operatorname{argmax}} p(\hat{y}) \cdot \prod_{i=1}^{d} p(x_i \mid \hat{y})$$

Equivalent to decision rule for bin. class.:

$$y = \operatorname{sgn}\left(\log \frac{p(Y=+1\mid x)}{p(Y=-1\mid x)}\right)$$

If the conditional independence assumption is lent to k-Means with Lloyd's heuristics. violated, the classifier can be overconfident.

## **Gaussian Bayes Classifier**

No independence assumption, model the weights for each point: features with a multivariant Gaussian  $\mathcal{N}(x; \mu_{\nu}, \Sigma_{\nu})$ :

$$\mu_{y} = \frac{1}{\text{Count}(Y=y)} \sum_{j \mid y_{j}=y} x_{j}$$

$$\sum_{y} = \frac{1}{\text{Count}(Y=y)} \sum_{j \mid y_{j}=y} (x_{j} - \hat{\mu}_{y}) (x_{j} - \hat{\mu}_{y})^{\top}$$

This is also called the quadratic discriminant analysis (QDA).

Linear Discriminant Analysis:  $\Sigma_{+} = \Sigma_{-}$ Fisher LDA:  $p(y) = \frac{1}{2}$ 

Outlier detection with:  $p(x) \le \tau$ .

## **Avoiding Overfitting**

MLE is prone to overfitting. Avoid this by restricting model class (fewer parameters, e.g. GNB) or using priors (restrict param. values).

# Generative vs. Discriminative

**Discriminative models:** 

p(y|x), can't detect outliers, more robust Generative models:

Given  $p(y \mid x)$ , a set of actions A and a cost p(x,y), can be more powerful (dectect out- $C: Y \times A \mapsto \mathbb{R}$ , pick the action with the maxiliers, missing values) if assumptions are met, modelled by a GMM. are typically less robust against outliers

## **Gaussian Mixture Model**

combination of Gaussian distributions:

Aim to estimate 
$$p(x,y)$$
 for complex situations  $p(x|\theta) = p(x|\mu, \Sigma, w) = \sum_{j=1}^k w_j \mathcal{N}(x; \mu_j, \Sigma_j)$  using Bayes' rule. We don't have labels and want to cluster this data. The problem is to estimate the param. GM for classification tasks. Assuming for a for the Gaussian distributions.

$$\underset{w \text{ if } \Sigma}{\operatorname{argmin}} - \sum_{i=1}^{n} \log \sum_{j=1}^{k} w_{j} \cdot \mathcal{N}(x_{i} \mid \mu_{j}, \Sigma_{j})$$

This is a non-convex objective. Similar to training a GBC without labels. Start with guess for our parameters, predict the unknown variables to generate a likelihood function Q: labels and then impute the missing data. Now we can get a closed form update.

# Hard-EM Algorithm

**E-Step**: predict the most likely class for each data point:

$$z_i^{(t)} = \underset{z}{\operatorname{argmax}} \ p(z \mid x_i, \boldsymbol{\theta}^{(t-1)})$$
$$= \underset{z}{\operatorname{argmax}} \ p(z \mid \boldsymbol{\theta}^{(t-1)}) \cdot p(x_i \mid z, \boldsymbol{\theta}^{(t-1)})$$

**M-Step**: compute MLE of  $\theta^{(t)}$  as for GBC.

Problems: labels if the model is uncertain, GANs tries to extract too much inf. Works poorly Learn f: "simple" distr.  $\mapsto$  non linear distr. if clusters are overlapping. With uniform Computing likelihood of the data becomes Where f(x) is called the discriminant function. weights and spherical covariances is equiva- hard, therefore we need a different loss.

# **Soft-EM Algorithm**

E-Step: calculate the cluster membership

$$\gamma_j^{(t)}(x_i) = \frac{w_j \cdot p(x_i; \theta_j^{(t-1)})}{\sum_k w_k \cdot p(x_i; \theta_k^{(t-1)})}$$
**M-Step**: compute MLE with closed form:

$$\Sigma_{y} = \frac{1}{\text{Count}(Y=y)} \sum_{j \mid y_{j}=y} (x_{j} - \hat{\mu}_{y})(x_{j} - \hat{\mu}_{y})^{\top} \quad w_{j}^{(t)} = \frac{1}{n} \sum_{i=1}^{n} \gamma_{j}^{(t)}(x_{i}) \quad \mu_{j}^{(t)} = \frac{\sum_{i=1}^{n} x_{i} \gamma_{j}^{(t)}(x_{i})}{\sum_{i=1}^{n} \gamma_{j}^{(t)}(x_{i})}$$
his is also called the **quadratic discrimi-**
ant analysis (QDA).

$$\Sigma_{j}^{(t)} = \frac{\sum_{i=1}^{n} \gamma_{j}^{(t)}(x_{i})(x_{i} - \mu_{j}^{(t)})(x_{i} - \mu_{j}^{(t)})^{\top}}{\sum_{i=1}^{n} \gamma_{i}^{(t)}(x_{i})}$$
inear Discriminant Analysis:  $\Sigma_{\perp} = \Sigma_{\perp}$ 

Init. the weights as uniformly distributed,  $DG = \max_{w'_D} M(w_G, w'_D) - \min_{w'_G} M(w'_G, w_D)$  are the weights as uniformly distributed,  $w'_D = \max_{w'_D} M(w_G, w_D)$  and  $w'_G = \max_{w'_D} M(w'_G, w_D)$  is the training objective. use spherical init. or empirical covariance of Various the data. Select k using cross-validation.

#### **Degeneracy of GMMs**

GMMs can overfit with limited data. Avoid this by add  $v^2I$  to variance, so it does not collapse (equiv. to a Wishart prior on the covariance matrix). Choose v by cross-validation.

# Gaussian-Mixture Bayes Classifiers

Assume that  $p(x \mid y)$  for each class can be **Normal Distribution**:

$$p(x \mid y) = \sum_{j=1}^{k_y} w_j^{(y)} \mathcal{N}(x; \mu_j^{(y)}, \Sigma_j^{(y)})$$
wing highly complex decision boundaries

Giving highly complex decision boundaries:

$$p(y|x) = \frac{1}{z}p(y)\sum_{j=1}^{k_{y}} w_{j}^{(y)} \mathcal{N}(x; \mu_{j}^{(y)}, \Sigma_{j}^{(y)})$$

# **GMMs for Density Estimation**

Can be used for anomaly detection or data imputation. Detect outliers, by comparing the estimated density against τ. Allows to control the FP rate. Use ROC curve as evaluation criterion and optimize using CV to find  $\tau$ .

# **General EM Algorithm**

E-Step: Take the expected value over latent

$$Q(\theta; \theta^{(t-1)}) = \mathbb{E}_{Z}[\log p(X, Z \mid \theta) \mid X, \theta^{(t-1)}]$$
$$= \sum_{i=1}^{n} \sum_{z_i=1}^{k} \gamma_{z_i}(x_i) \log p(x_i, z_i \mid \theta)$$

with 
$$\gamma_z(x) = p(z \mid x, \theta^{(t-1)})$$

M-Step: Compute MLE / Maximize:

$$\theta^{(t)} = \underset{\theta}{\operatorname{argmax}} Q(\theta; \theta^{(t-1)})$$

We have monotonic convergence, each EMiteration increases the data likelihood.

$$\min_{w_G} \max_{w_D} \mathbb{E}_{x \sim p_{\text{data}}}[\log D(x, w_D)] \\ + \mathbb{E}_{z \sim p_z}[\log(1 - D(G(z, w_G), w_D))]$$

Training requires finding a saddle point. For a fixed G, the optimal discriminator is:

$$D_G(x) = \frac{p_{\rm data}(x)}{p_{\rm data}(x) + p_G(x)}$$
 Powerful discriminator could lead to memo-

rization of finite data. Other issues are oscillations/divergence or mode collapse.

One possible performance metric:

$$DG = \max_{w'_D} M(w_G, w'_D) - \min_{w'_G} M(w'_G, w_D)$$

**Derivatives:** 

$$\nabla_{x}x^{\top}A = A \quad \nabla_{x}a^{\top}x = \nabla_{x}x^{\top}a = a$$

$$\nabla_{x}b^{\top}Ax = A^{\top}b \quad \nabla_{x}x^{\top}x = 2x \quad \nabla_{x}x^{\top}Ax = 2Ax$$

$$- \nabla_{w}||y - Xw||_{2}^{2} = 2X^{\top}(Xw - y)$$

**Bayes Theorem:** 

neorem:  

$$p(y \mid x) = \frac{1}{p(x)} \underbrace{p(y) \cdot p(x \mid y)}_{\text{position}}$$

$$f(x) = \frac{1}{\sigma\sqrt{2}} \exp(-\frac{(x-\mu)^2}{\sigma^2})$$