

# Seminar 9: ACF, PACF and ARMA models

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## Correction of HW8

From the property, of the distribution of  $\mu$  and  $\sigma$  we have:

$$\Sigma = \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^4 \end{pmatrix}.$$

Then, using the hint we obtain:

$$\frac{\partial g}{\partial \theta} = \begin{pmatrix} \frac{1}{\sqrt{\sigma^2}} & -\frac{1}{2}\mu\sigma^{-3} \end{pmatrix}.$$

where  $g = \frac{\mu}{\sqrt{\sigma^2}}$

## Correction of HW8

Hence:

$$\begin{aligned}\frac{\partial g}{\partial \theta} \Sigma \frac{\partial g'}{\partial \theta} &= \begin{pmatrix} \frac{1}{\sqrt{\sigma^2}} \\ -\frac{1}{2}\mu\sigma^{-3} \end{pmatrix} \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^4 \end{pmatrix} \begin{pmatrix} \frac{1}{\sqrt{\sigma^2}} & -\frac{1}{2}\mu\sigma^{-3} \end{pmatrix} \\ &= 1 + \frac{1}{4} \left( \frac{\mu}{\sigma} \right)^2.\end{aligned}$$

Denoting by  $z_{\alpha/2}$  and  $z_{1-\alpha/2}$  the  $\alpha/2$  and  $1 - \alpha/2$  quantiles of  $\mathcal{N}(0, 1 + \frac{1}{4}(\frac{\hat{\mu}}{\hat{\sigma}})^2)$ , and using the delta method, we can write:

$$P\left(z_{\alpha/2} \leq \sqrt{T}(g(\hat{\theta}) - g(\theta)) \leq z_{1-\alpha/2}\right) = 1 - \alpha.$$

Rearranging:

$$P\left(\widehat{SR} - \frac{z_{1-\alpha/2}}{\sqrt{T}} \leq SR \leq \widehat{SR} - \frac{z_{\alpha/2}}{\sqrt{T}}\right) = 1 - \alpha,$$

which gives the asymptotic confidence interval for  $SR$ .

## Delta methods 95% confidence intervals (Solutions)

### Solution

- Define the variance of the asymptotic distribution of  $SR$  as  $VAR = 1 + 0.25 * (\mu_{\hat{}} / \sigma_{\hat{}})^2$
- Compute  $z = \text{norm.ppf}(0.975, 0, VAR^{**0.5})$  from a normal PDF.
- $\text{delta\_method} = [SR_{\hat{}} - z / \text{np.sqrt}(T), SR_{\hat{}} + z / \text{np.sqrt}(T)]$

## Common mistakes

- Square of  $T$  `deltamethod = [SRobs - z/(math.pow (T, 2)), SRobs + z/(math.pow (T, 2))]`.
- Not taking the square-root of 16 for `np.random.normal(loc=1, scale=16, size=T)`.
- Doing the derivatives of  $SR$  with symbolic toolbox in Python. Better to do it by hand.
- Use MLE estimator instead of MM. Easier the second one.
- Using .95 as quantile instead of .975 like this: `norm.ppf(0.95, mu, math.sqrt(sigma_2))`. CI are two tailed.
- For delta method CI, not using the sandwich formula  $\frac{\partial g}{\partial \theta} \Sigma \frac{\partial g'}{\partial \theta}$  as the asymptotic variance.

# Results must make sense

```
In [59]: #use conf_interval command  
conf_interval = np.percentile(array2, [2.5, 97.5])  
print(conf_interval)  
[-5.90513041  8.80310158]
```

# Notion of Dependence

Covariance (or correlation) = measure of *linear dependence* between two random variables. Will be used to measure dependence between current observations and past observations. Current return:  $Y_t$  Return with lag  $\tau$ :  $Y_{t-\tau}$  The covariance between current and lagged return

$$\gamma(\tau) = \text{Cov}(Y_t, Y_{t-\tau})$$

is called the *autocovariance of order  $\tau$* .

The *autocorrelation function* (ACF) =  $\rho(\tau)$ ,  $\tau = 1, 2, \dots$  of successive orders will allow to detect linear temporal dependence.

Autocorrelation of order  $\tau$

$$\rho(\tau) = \frac{\text{Cov}(Y_t, Y_{t-\tau})}{V(Y_t)} = \frac{\gamma(\tau)}{\gamma(0)}$$



# ARMA process

## Pure AR

A pure Auto-Regressive (AR) process of order 1 is defined as followed:

$$Y_t = \mu + \omega_1 Y_{t-1} + \varepsilon_t$$

where  $\varepsilon_t$  is a noise term. Hence we can define an AR of order  $p$  as:

$$Y_t = \mu + \omega_1 Y_{t-1} + \dots + \omega_p Y_{t-p} + \varepsilon_t.$$

## Pure MA

A pure Moving-Average (MA) process of order 1 is defined as followed:

$$Y_t = \mu + \varepsilon_t - \theta_1 \varepsilon_{t-1} - \varepsilon_{t-q}$$

where  $\varepsilon_t$  is a noise term. Hence we can define an MA of order  $q$  as:

$$Y_t = \mu + \varepsilon_t - \theta_1 \varepsilon_{t-1} - \dots - \theta_q \varepsilon_{t-q}$$

## ARMA $(p, q)$

Hence is it now easy to see that  $Y_t$  can be represented a combination of pure AR( $p$ ) and and MA( $q$ ) process in the following manner:

$$Y_t = \mu + \omega_1 Y_{t-1} + \dots + \omega_p Y_{t-p} + \varepsilon_t - \theta_1 \varepsilon_{t-1} - \dots - \theta_q \varepsilon_{t-q}.$$

## Order $p$ and $q$ detection

### Reminder: Autocorrelation

$$\rho(\tau) = \frac{\text{Cov}(Y_t, Y_{t-\tau})}{V(Y_t)} = \frac{\gamma(\tau)}{\gamma(0)}$$

### Partial autocorrelation

The partial autocorrelation  $r(\tau)$  is defined as the last coefficient in the autoregression of order  $\tau$ ,

$$Y_t = \mu + \omega_1 Y_{t-1} + \dots + \omega_\tau Y_{t-\tau} + \varepsilon_t,$$

i.e.,  $r(\tau) = \omega_\tau$ .

# Order $p$ and $q$ detection

## Order detection

if pure MA( $q$ ), autocorrelations  $\rho(\tau)$  are zero after order  $q$ ,

$$\rho(\tau) = 0, \quad \tau > q$$

if pure AR( $p$ ), partial autocorrelations  $\rho(\tau)$  are zero after order  $p$ ,

$$r(\tau) = 0, \quad \tau > p$$

## Detection?

- Plotting  $r(\tau)$  and  $\rho(\tau)$  function for orders  $\tau = 1, 2, \dots$
- Assess if  $r(\tau)$  and  $\rho(\tau)$  are significantly  $\neq 0$  through confidence intervals.

## Usefull functions

- Import `import statsmodels.api as sm` and use `sm.graphics.tsa.plot_acf` to plot the *acf*.
- Repeat same procedure for *pacf*.
- Import `import from statsmodels.stats.diagnostic import acorr_ljungbox` and use `acorr_ljungbox` to compute the Ljung-Box test
- `ols` function to fit the models and `resid` to extract the residuals and plot their *acf* and *pacf*.