

# Seminar 8: Bootstrap and Asymptotic Confidence Intervals

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### Common mistakes

- $\text{mean}(\text{FireLosses})$  instead of  $\text{mean}(\text{FireLosses}[\text{FireLosses} > u])$ . First compute the mean on the full dataset while the second compute the mean on the tail. Same apply for  $\text{var}$

Use of simulation to compute Statistical quantities (i.e. Confidence intervals)

## Why using simulation

- Asymptotic might be a poor approximation when the sample size is small.
- Confidence intervals based on asymptotic normality may be too wide or too narrow and exhibit a wrong coverage probability.

if  $\sqrt{T}(\hat{\theta} - \theta) \sim \mathcal{N}(0, \Sigma)$ , then  $\sqrt{T}(g(\hat{\theta}) - g(\theta)) \sim \mathcal{N}\left(0, \frac{\partial g}{\partial \theta} \Sigma \frac{\partial g'}{\partial \theta}\right)$ ,

where  $g$  is a continuous function of the parameter vector  $\theta$ . In practice,  $\Sigma$  is estimated by the estimated covariance matrix of  $\hat{\theta}$  and  $\frac{\partial g}{\partial \theta}$  is evaluated at  $\hat{\theta}$ .

You will also need the following property:

if  $\varphi \sim \mathcal{N}(\mu, \sigma^2)$ , then:

$$\sqrt{T}(\hat{\mu} - \mu) \sim \mathcal{N}(0, \sigma^2), \quad \sqrt{T}(\hat{\sigma}^2 - \sigma^2) \sim \mathcal{N}(0, \sigma^4), \quad \text{and} \\ \hat{\mu} \text{ and } \hat{\sigma} \text{ are orthogonal.}$$

## Breakdown of the exercise 1

- ① Use `np.random.seed` to fix a specific seed (so the results are reproducible) and `np.random.normal` to simulate data from a normal distribution.

- ② Compute the empirical quantity  $(\hat{\mu}, \hat{\sigma} \text{ and } \widehat{SR})$

- ③ Delta methods:

- ①  $g = \widehat{SR} = \frac{\hat{\mu}}{\sqrt{\hat{\sigma}^2}}.$

- ②

$$\Sigma = \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^4 \end{pmatrix}.$$

- ③ Compute the sandwich formula  $\frac{\partial g}{\partial \theta} \Sigma \frac{\partial g'}{\partial \theta}$ . where  $\hat{\theta} = (\hat{\mu}, \hat{\sigma})$ , to get the asymptotic variance.

- 1 Start from initial data  $Y_1, \dots, Y_T$  and compute the empirical cdf.
- 2 Generate  $ST$  independent drawings by sampling randomly in the initial data with replacement  $Y_1^s, \dots, Y_T^s$ ,  $s = 1, \dots, S$  (draw  $ST$  realizations of a uniform  $[0, 1]$  variable and invert the empirical cdf).
- 3 For each simulated sample  $s = 1, \dots, S$  of length  $T$ , compute the estimate  $\hat{\theta}^s$ , for example the empirical mean  $\hat{\theta}^s = \hat{m}_s = \frac{1}{T} \sum_{t=1}^T Y_t^s$ .
- 4 The empirical distribution of the estimates  $\hat{\theta}^s$ ,  $s = 1, \dots, S$ , is a good approximation to the true distribution of the estimator in small sample (consistent as  $s \rightarrow \infty$ ).

## Breakdown of the exercise 2

- 1 Use `np.random.randint` to sample  $B = 2'000$  vectors of random integer with replacement from 1 to  $T$ , where  $T$  is the sample size. Those integer will be the indices representing the observation we "select" to compute our Bootstrapped estimator.
- 2 Define a new variable  $X_{star}$ , which represent a  $T \times B$  matrix of re-sampled observation.
- 3 Compute the statistics of interests on  $X_{star}$  (i.e.  $\widehat{SR}$ )
- 4 Construct a 95% confidence interval on  $\widehat{SR}$  based on the following formula:

$$\left[ 2\widehat{SR} - q^*(1 - \alpha/2), 2\widehat{SR} - q^*(\alpha/2) \right].$$