# Reportes resultados sobre modelos panel

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### Abstract

In these note there is a summary of the relations between several models that can use panel data.

The objective of these note is to give a broad overview of the possible models that can use panel data. There are several usual features to consider in a model with panel data, for example, changes on parameters for time or individual. Also, specification on error term is relevant for interpretation.

These notes are based on Hsiao (2014) It goes from the theory in the text, to the application.

## 1 Dummies for trend-spline

Each dummy is one for the period that ends in the first month of the year.

$$y_{itm} = \alpha_i^* + \lambda * t + \beta'_{2012} Dummy 2012_m + \ldots + \beta'_{2020} Dummy 2020_m + \beta'_0 jan Dummy_m + \beta'_1 tax Dummy_m + u_{it};$$
  
 $i = 1, \ldots, N; t = 1, \ldots, T.$ 

### 1.1 Comparisons by segment

Results for brand type 2 is medium.

The interpretation of the spline with negative signs?

	(1)	(2)
VARIABLES	ppu4	ppu7
m1	0.013	0.039***
	(0.012)	(0.007)
$m1_{-}20$	0.256	0.192
	(0.000)	(0.000)
Observations	1,356	1,356
0 .0 .0		*
Number of cve_ciudad	43	43

Standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

# 2 Quadratic trend

Each dummy is one for the period that ends in the first month of the year.

$$y_{itm} = \alpha_i^* + \lambda_1 * t + \lambda_2 * t^2 + \beta_0' janDummy_m + \beta_1' taxDummy_m + u_{itm};$$
  
$$i = 1, \dots, N; t = 1, \dots, T.$$

### 2.1 Comparisons by segment

Results for medium brands.

	(1)	(2)	
VARIABLES	ppu4	ppu7	
m1	-0.035***	-0.039***	
	(0.012)	(0.009)	
$m1_{-}20$	0.163	0.097	
	(0.000)	(0.000)	
ym	-0.035***	0.027***	
	(0.007)	(0.010)	
Observations	1,356	1,356	
Number of cve_ciudad	43	43	
Standard errors in parentheses			
*** p<0.01, ** p<0.05, * p<0.1			

# 3 Sample by brand

There is no common sample.

$$y_{itm} = \alpha_i^* + \lambda * t + \beta_0' janDummy_m + \beta_1' taxDummy_m + u_{itm};$$
 
$$i = 1, \dots, N; t = 1, \dots, T.$$

### 3.1 Comparisons by segment

Results for medium brands: Lucky , PallMall.

	(1)	(2)	(3)	(4)
VARIABLES	ppu4	ppu4	ppu7	ppu7
m1	-0.042***	-0.048***	-0.043***	-0.039***
	(0.012)	(0.015)	(0.009)	(0.012)
$m1_{-}20$	0.153***	0.154***	0.193***	0.162***
	(0.031)	(0.037)	(0.023)	(0.031)
$_{ m ym}$	0.010***	0.010***	0.010***	0.011***
	(0.000)	(0.000)	(0.000)	(0.000)
Observations	1,837	1,356	3,157	1,356
Number of cve_ciudad	36	29	41	29

Standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

### Specification with dummies for each level 4

One simple regression with indicators for city, brand and time. Same effect on all the brands.

$$y_{itm} = \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0' janDummy_m + \beta_1' taxDummy_m + u_{itm};$$

 $i = \underbrace{1, \dots, N; t = 1, \dots, T; m = 1, \dots, M}_{\text{(1)}}$ . Results for principal brands:

	(1)	
VARIABLES	ppu	
m1	-0.023***	
	(0.004)	
$m1_{-}20$	0.214***	
	(0.010)	
ym	0.009***	
	(0.000)	
Observations	22,771	
R-squared	0.898	
Standard errors	s in parentheses	
*** n<0.01 ** 1	n < 0.05 * n < 0.1	

p<0.01, \*\* p<0.05, \* p<0.1

#### 4.1 Comparisons by segment

Results for brand type: 1 is premium, 2 is medium, 3 is low.

	(1)	(2)	(3)
VARIABLES	ppu	ppu	ppu
m1	-0.018***	-0.046***	-0.012
	(0.004)	(0.008)	(0.008)
$m1_{-}20$	0.233***	0.175***	0.193***
	(0.011)	(0.019)	(0.029)
ym	0.010***	0.010***	0.007***
	(0.000)	(0.000)	(0.000)
Observations	$13,\!598$	$5,\!256$	3,917
R-squared	0.916	0.868	0.774

Standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

## 5 Parameters are different for each brand

Separate regression for each brand. The main estimation routine is a random coefficients result.

$$\begin{aligned} y_{itm} &= \alpha_{im}^* + \delta_m^{'} janDummy_m + \beta_m^{'} taxDummy_m + \lambda_m^{'} t_m + u_{itm}; \\ i &= 1, \dots, N; t = 1, \dots, T; m = 1, \dots, M. \end{aligned}$$

### 5.1 Comparisons by segment

Results for premium brands

	(1)	(2)	(3)
VARIABLES	ppu1	ppu2	ppu5
m1	-0.001	-0.000	-0.031***
	(0.003)	(0.003)	(0.005)
$m1_{-}20$	0.282***	0.106***	0.047*
	(0.017)	(0.018)	(0.025)
ym	0.009***	0.009***	0.009***
	(0.000)	(0.000)	(0.000)
Observations	2,488	2,488	2,488
Number of cve_ciudad	45	45	45

Standard errors in parentheses \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Results for lower segment brands

	(1)	(2)				
VARIABLES	ppu3	ppu6				
m1	0.019**	-0.051***				
	(0.009)	(0.012)				
$m1_{-}20$	0.194***	0.151***				
	(0.028)	(0.038)				
ym	0.008***	0.005***				
	(0.000)	(0.000)				
Observations	587	587				
Number of cve_ciudad	43	43				
Standard errors	in parenthe	Standard errors in parentheses				
*** p<0.01, ** p<0.05, * p<0.1						
*** p<0.01, ** p	0<0.05, *p	< 0.1				
*** p<0.01, ** p Results for mid-range seg	, -					
1 , 1	, -					
1 , 1	gment bran	ds				
Results for mid-range seg	gment brane (1)	(2)				
Results for mid-range seg	gment brane (1)	(2)				
Results for mid-range seg  VARIABLES	gment branch (1) ppu4	(2) ppu7				
Results for mid-range seg  VARIABLES	gment branc (1) ppu4 0.098***	(2) ppu7 -0.117***				
Results for mid-range seg  VARIABLES  m1	(1) ppu4 0.098*** (0.012)	(2) ppu7 -0.117*** (0.010)				
Results for mid-range seg  VARIABLES  m1	(1) ppu4 0.098*** (0.012) -0.420	(2) ppu7 -0.117*** (0.010) 0.735				
Results for mid-range seg  VARIABLES  m1  m1_20	(1) ppu4 0.098*** (0.012) -0.420 (0.000)	(2) ppu7 -0.117*** (0.010) 0.735 (0.000)				
Results for mid-range seg  VARIABLES  m1  m1_20	0.098*** (0.012) -0.420 (0.000) 0.004***	(2) ppu7 -0.117*** (0.010) 0.735 (0.000) 0.004***				

 $\begin{tabular}{lll} Number of cve\_ciudad & 43 \\ \hline Standard errors in parentheses \\ *** p<0.01, ** p<0.05, * p<0.1 \\ \hline \end{tabular}$ 

## 6 Parameters are constant over time

Section based on Hsiao (2014) Separate regression for each individual

$$y_{it} = \alpha_i^* + \beta_i' x_{it} + u_{it}; i = 1, \dots, N; t = 1, \dots, T.$$

### 6.1 Restrictions to the constant parameters over time model

Separate regression for each individual

$$y_{it} = \alpha_i^* + \beta_i' x_{it} + u_{it}; i = 1, \dots, N; t = 1, \dots, T.$$

It can be restricted in several ways. Only slope coefficients are identical, intercepts are individual.

$$y_{it} = \alpha_i^* + \beta' x_{it} + u_{it}.$$

Both slope coefficients and intercepts are identical.

$$y_{it} = \alpha^* + \beta' x_{it} + u_{it}.$$

Following Hsiao(2014) the first model is called unrestrincted, the second as individual-mean regression model and the last model is known as pooled model.

## 7 Simple Regression with Variable Intercept

This models are one way to consider unobserved heterogeneity across individuals and/or through time. The assumption is that the effects of that heterogeneity come from three types of variables: time-invariant, individual-invariant and individual time-variant. The model can be written:

$$y_{it} = \alpha_i^* + x_{it}' \beta_i + u_{it}; i = 1, \dots, N; t = 1, \dots, T.$$

With the assumption that  $u_{it}$  is uncorrelated with  $(x_{i1}, \ldots, x_{iT})$  and have an independent identically distributed random variable with mean 0 and constant variance. Following Hsiao(2014), the OLS estimator is called least-squares dummy variable (LSDV), covariance (CV) estimator or within-group estimator. If the variance is constant for every individual an efficient estimator can be obtained using weighted least-squares with the initial estimator for individual variance from the individual errors.

### 7.1 Estimations of Variance-Components models

The individual-specific effects as random variables. The residual can be assumed to consist three components:

$$v_{it} = \alpha_i + \lambda_t + u_{it}.$$

The estimation for random-effects model assume a distribution on the conditional distribution of  $f(\alpha_i, \lambda_t | x_i)$ . With the assumption of constant  $\lambda_t$  for all t, the presence of  $\alpha_i$  produces correlations in  $v_{it}$  over time for a given individual. Consistent estimates in finite samples can be obtained by Generalized Least-Squares (GLS). The GLS estimator is a weighted average of the betweengroup and the within-group estimator. In a practical situation, without knowing the constants from variance components, the estimation uses feasible GLS or two-step GLS.

### 7.2 Fixed or Random effects

When N is fixed and T is large LSDV and GLS are the same estimator. [The time-specific effects could be a problem?] The residual can be assumed to consist three components:

$$v_{it} = \alpha_i + \lambda_t + u_{it}.$$

The estimation for random-effects model assume a distribution on the conditional

# 8 Dynamic models

Given that some assumptions of the static model are not valid. It can be proposed a dynamic alternative.

$$y_{it} = \gamma y_{i,t-1} + x'_{it}\beta_i + \alpha_i^* + \lambda_t + u_{it}; i = 1, \dots, N; t = 1, \dots, T.$$

The assumption of strict exogeneity is no longer valid. The initial values become relevant. The way in which the T and N tend to infinity become relevant for asymptotic properties, like consistency.

With the assumption that  $u_{it}$  is uncorrelated with  $(x_{i1}, \ldots, x_{iT})$  and have an independent identically distributed random variable with mean 0 and constant variance.

lance.	(1)	(2)	(3)
VARIABLES	price	price	price
weight	-7.273***	-3.715	-0.513
wtsq	(2.326) $0.002***$	(2.481) $0.001***$	(3.147) 0.001**
foreign	(0.000)	(0.000) $3,247.034***$	(0.000) $3,263.819****$
length		(694.805)	(694.275) $-73.435*$
			(43.747)
Observations	74	74	74
R-squared	0.394	0.554	0.584

Robust standard errors in parentheses
\*\*\* p<0.01, \*\* p<0.05, \* p<0.1