## Homework 2

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**Problem 2.1.** Find MLE for the following parameters:

(a) Probability of success p in Bernoulli(p) model

**Solution** (a). Let X be a bernoulli random variable with parameter p. Let  $X_1,...,X_n$  be the random sample of X.

PDF for the bernoulli distribution with parameter p is

$$f(x) = p^{x}(1-p)^{1-x} x = 0, 1$$

Likelihood function of the sample:

$$= \prod_{i=1}^{n} f(x_i, p)$$

$$= \prod_{i=1}^{n} p^{x_i} (1-p)^{1-x_i}$$

$$= p^{\sum_{i=1}^{n} x_i} (1-p)^{n-\sum_{i=1}^{n} x_i}$$

Log-likelihood function of the sample:

$$= \ln \left( p^{\sum_{i=1}^{n} x_i} (1-p)^{n-\sum_{i=1}^{n} x_i} \right)$$
$$= \left( \sum_{i=1}^{n} x_i \right) \ln(p) + \left( n - \sum_{i=1}^{n} x_i \right) \ln(1-p)$$

Maximum value of this function can be found by taking the derivative of the above function with respect to p, and setting it to 0.

$$= \frac{\sum_{i=1}^{n} x_i}{p} - \left(\frac{n - \sum_{i=1}^{n} x_i}{1 - p}\right) = 0$$

gives that 
$$p = \frac{\sum\limits_{i=1}^{n} x_i}{n}$$

(b) Probability of success p in Binomial(n,p) model

**Solution** (b). Let X be a binomail random variable with parameters n and p. Let  $X_1,...,X_m$  be the random sample of X.

PDF for the binomial distribution with parameters n and p is

$$f(x) = \binom{n}{x} p^x (1-p)^{n-x} \qquad x = 0, 1, ..., n.$$

Likelihood function of the sample:

$$= \prod_{i=1}^{m} f(x_i, p)$$

$$= \prod_{i=1}^{m} \binom{n}{x} p^{x_i} (1-p)^{n-x_i}$$

Log-likelihood function of the sample:

$$= \ln\left(\prod_{i=1}^{m} \binom{n}{x} p^{x_i} (1-p)^{n-x_i}\right)$$
$$= \sum_{i=1}^{m} \ln\binom{n}{x_i} + \left(\sum_{i=1}^{n} x_i\right) \ln(p) + \left(mn - \sum_{i=1}^{m} x_i\right) \ln(1-p)$$

Maximum value of this function can be found by taking the derivative of the above function with respect to p, and setting it to 0.

$$= 0 + \frac{\sum_{i=1}^{m} x_i}{p} - \left(\frac{mn - \sum_{i=1}^{m} x_i}{1 - p}\right) = 0$$

gives that  $p = \frac{\sum\limits_{i=1}^{m} x_i}{mn}$ 

(c) Probability of success p in Geometric(p) model

**Solution** (c). Let X be a geometric random variable with parameter p. Let  $X_1,...,X_n$  be the random sample of X.

PDF for the geometric distribution with parameter p is

$$f(x) = p(1-p)^{x-1}$$
  $x = 1, 2, ...$ 

Likelihood function of the sample:

$$= \prod_{i=1}^{n} f(x_i, p)$$

$$= \prod_{i=1}^{n} p(1-p)^{x_i-1}$$

$$= p^n (1-p)^{\sum_{i=1}^{n} x_i - n}$$

Log-likelihood function of the sample:

$$= \ln \left( p^{n} (1-p)^{\sum_{i=1}^{n} x_{i} - n} \right)$$
$$= n \ln(p) + \left( \sum_{i=1}^{m} x_{i} - n \right) \ln(1-p)$$

Maximum value of this function can be found by taking the derivative of the above function with respect to p, and setting it to 0.

$$= \frac{n}{p} + \left(\frac{n - \sum_{i=1}^{m} x_i}{1 - p}\right) = 0$$

gives that 
$$p = \frac{n}{\sum\limits_{i=1}^{m} x_i}$$

(d) Intensity  $\lambda$  in Poisson( $\lambda$ ) model

**Solution** (d). Let X be a Poisson random variable with parameter  $\lambda$ . Let  $X_1,...,X_n$  be the random sample of X.

PDF for the poisson distribution with parameter  $\lambda$  is

$$f(x) = \frac{e^{-\lambda}\lambda^x}{x!} \qquad x = 0, 1, \dots$$

Likelihood function of the sample:

$$= \prod_{i=1}^{n} f(x_i, \lambda)$$

$$= \prod_{i=1}^{n} \frac{e^{-\lambda} \lambda^{x_i}}{x_i!}$$

$$= e^{-\lambda n} \cdot \lambda^{\sum_{i=1}^{n} x_i} C \qquad where \qquad C = \prod_{i=1}^{n} \frac{1}{x_i!}$$

Log-likelihood function of the sample:

$$= \ln \left( e^{-\lambda n} \cdot \lambda_{i=1}^{\sum_{i=1}^{n} x_i} C \right)$$
$$= -n\lambda + \left( \sum_{i=1}^{m} x_i \right) \ln(\lambda) + \ln(C)$$

Maximum value of this function can be found by taking the derivative of the above function with respect to  $\lambda$ , and setting it to 0.

$$= -n + \left(\frac{\sum_{i=1}^{m} x_i}{\lambda}\right) + 0 = 0$$

gives that  $\lambda = \frac{\sum\limits_{i=1}^{m} x_i}{n}$ 

**Problem 2.2.** Find the Fisher information for the models of problem 2.1 (a) Probability of success p in Bernoulli(p) model

Solution (a). The Fisher Information for parameter  $\hat{p}$  can be written two ways:  $I(\hat{p}) = E[(\frac{\partial}{\partial p} \ln f(x, p))^2 | p]$  or  $I(\hat{p}) = -E[\frac{\partial^2}{\partial p^2} \ln f(x, p) | p]$ .

From problem 2.1 (a) we know

$$\frac{\partial}{\partial p} \ln f(x, p) = \frac{\partial}{\partial p} [xp + (1 - x) \ln (1 - p)]$$
$$= \frac{x}{p} - \frac{1 - x}{1 - p}$$

Recall that E(X) = p where X is the bernoulli random variable.

Therefore,

$$= I(\hat{p}) = -E\left[\frac{\partial^2}{\partial p^2} \ln f(x, p)\right]$$

$$= -E\left[-\frac{X}{p^2} - \left(\frac{1 - X}{q^2}\right)\right]$$

$$= \frac{E(X)}{p^2} + \frac{E(1 - X)}{(1 - p)^2}$$

$$= \frac{1}{p} + \frac{1}{1 - p}$$

gives  $\frac{1}{p(1-p)}$ 

(b) Probability of success p in Binomial(n,p) model

**Solution (b).** The Fisher Information for parameter  $\hat{p}$  can be written two ways:  $I(\hat{p}) = E\left[\left(\frac{\partial}{\partial p}\ln f(x,p)\right)^2|p\right]$  or  $I(\hat{p}) = -E\left[\frac{\partial^2}{\partial p^2}\ln f(x,p)|p\right]$ .

From problem 2.1 (b) we know

$$\frac{\partial}{\partial p} \ln f(x, p) = \frac{\partial}{\partial p} \left[ \ln \binom{n}{x} + x \ln p + (n - x) \ln (1 - p) \right]$$

$$= \frac{x}{p} - \frac{n - x}{1 - q}$$

$$\left( \frac{x}{p} - \frac{n - x}{1 - q} \right)^2 = x^2 \frac{1}{p^2 q^2} - 2x \frac{n}{q^2 p} + \frac{n^2}{q^2} \qquad where \qquad q = (1 - p)$$

Thus,

$$E\left[X^{2}\frac{1}{p^{2}q^{2}}-2x\frac{n}{q^{2}p}+\frac{n^{2}}{q^{2}}\right]$$

Recall that E(X) = np and  $E(X^2) = npq + n^2p^2$  where X is the binomial random variable.

Therefore,

$$= \frac{npq}{p^2q^2} + \frac{n^2p^2}{p^2q^2} - 2\frac{n^2p}{q^2p} + \frac{n^2}{q^2}$$
$$= \frac{n}{pq} + \frac{n^2}{q^2} - 2\frac{n^2}{q^2} + \frac{n^2}{q^2}$$

 $\frac{n}{pq}$ 

(c) Probability of success p in Geometric(p) model

**Solution (c).** The Fisher Information for parameter  $\hat{p}$  can be written two ways:  $I(\hat{p}) = E\left[\left(\frac{\partial}{\partial p}\ln f(x,p)\right)^2|p\right]$  or  $I(\hat{p}) = -E\left[\frac{\partial^2}{\partial p^2}\ln f(x,p)|p\right]$ .

From problem 2.1 (c) we know

$$\frac{\partial}{\partial p} \ln f(x, p) = \frac{\partial}{\partial p} [\ln p + (x - 1) \ln (1 - p)]$$
$$= \frac{1}{p} - \left(\frac{x}{1 - p}\right)$$

Recall that  $E(X) = \frac{1}{p}$  where X is the geometric random variable.

Therefore,

$$I(\hat{p}) = -E\left[\frac{\partial^2}{\partial p^2} \ln f(x, p)\right]$$

$$= -E\left[\frac{-1}{p^2} - \left(\frac{X}{1 - p^2}\right)\right]$$

$$= \frac{1}{p^2} + \frac{E(X)}{(1 - p)^2}$$

$$= \frac{1}{p^2} + \frac{1}{p(1 - p)^2}$$

$$= \frac{1 - p}{p^2(1 - p)^2}$$

gives 
$$\frac{1}{p^2(1-p)}$$

(d) Intensity  $\lambda$  in Poisson( $\lambda$ ) model

**Solution (d).** The Fisher Information for parameter  $\hat{\lambda}$  can be written two ways:  $I(\hat{\lambda}) = E\left[\left(\frac{\partial}{\partial\lambda}\ln f(x,\lambda)\right)^2|\lambda\right]$  or  $I(\hat{\lambda}) = -E\left[\frac{\partial^2}{\partial\lambda^2}\ln f(x,\lambda)\right]$ .

From problem 2.1 (d) we know

$$\frac{\partial}{\partial p} \ln f(x, \lambda) = \frac{\partial}{\partial p} [-\lambda + x \ln \lambda + \ln \frac{x}{x!}]$$

$$= \frac{x}{\lambda} - 1$$

$$\frac{\partial^2}{\partial \lambda^2} \ln f(x, \lambda) = \frac{-x}{\lambda^2}$$

Recall that  $E(X) = \lambda$  where X is the poisson random variable with parameter  $\lambda$ .

Therefore,

$$I(\lambda) = -E\left[\frac{\partial^2}{\partial p^2} \ln f(x, p)\right]$$
$$= -E(\frac{-X}{\lambda^2})$$
$$= \frac{1}{\lambda^2} E(X)$$
$$= \frac{\lambda}{\lambda^2}$$

gives  $\frac{1}{\lambda}$ 

**Problem 2.3.** Assess the efficiency of the MLE estimators for the models of problem 2.1. (a) Probability of success p in Bernoulli(p) model

**Solution (a).** According to Cramer-Rao inequality  $Var(\hat{p}) \geq \frac{1}{nI(\hat{p})}$  It states that if there is an estimator that achieves this lower bound, then it is the efficient estimator for that parameter.

From problem 2.1 (a) we have that  $\hat{p} = \frac{\sum_{i=1}^{n} x_i}{n}$ . So, the variance of this parameter is

$$var(\hat{p}) = var\left(\frac{\sum_{i=1}^{n} x_i}{n}\right)$$

$$= \frac{var\sum_{i=1}^{n} x_i}{n^2}$$

$$= \frac{np(1-p)}{n^2}$$

$$= \frac{p(1-p)}{n}$$

Also, from problem 2.2 (a),  $I(\hat{p}) = \frac{1}{p(1-p)}$ . Hence  $\hat{p}$  is an efficient estimator.

(b) Probability of success p in Binomial (n, p) model

**Solution (b).** From problem 2.1 (b) we have that  $\hat{p} = \frac{\sum_{i=1}^{m} x_i}{mn}$ . So, the variance of this parameter is

$$var(\hat{p}) = var \left( \frac{\sum_{i=1}^{m} x_i}{mn} \right)$$

$$= \frac{var \sum_{i=1}^{m} x_i}{m^2 n^2}$$

$$= \frac{mnp(1-p)}{m^2 n^2}$$

$$= \frac{p(1-p)}{mn}$$

Also, from problem 2.2 (b),  $I(\hat{p}) = \frac{n}{p(1-p)}$ . Hence  $\hat{p}$  is an efficient estimator.

(c) Probability of success p in Geometric(p) model

**Solution** (c). From problem 2.1 (c) we have that  $\hat{p} = \frac{n}{\sum\limits_{i=1}^{m} x_i} = \frac{1}{\bar{X}}$ . Variance of this estimator can be calculated using delta method.

Delta method says that:

if  $\theta$  follows normal distribution asymptotically with mean  $= \theta$  and variance  $= \sigma^2$  then  $f(\theta)$  follows normal distribution asymptotically with mean  $f(\theta)$  and variance  $f'(\theta)^2 \cdot \sigma^2$ .

From CLT we know that  $\bar{X}$  follows normal distribution if n is large. Hence  $\bar{X}$  follows normal distribution asymptotically with mean  $=\frac{1}{p}$  and variance  $=\frac{1-p}{np^2}$ . Here the function is  $f(\theta)=\frac{1}{\bar{X}}$  and  $f'(\theta)=\frac{-1}{\bar{X}^2}$ . From the delta method, mean of  $f(\theta)$  is p and variance of  $f(\theta)$  is  $p^4.\frac{1-p}{np^2}=\frac{(1-p)p^2}{n}$ . So, the variance of this parameter is

$$var(\hat{p}) = \frac{(1-p)p^2}{n}$$

Also, from problem 2.2 (c),  $I(\hat{p}) = \frac{1}{p^2(1-p)}$ . Hence  $\hat{p}$  is an efficient estimator.

**Solution (d).** From problem 2.1 (d) we have that  $\hat{\lambda} = \frac{\sum_{i=1}^{n} x_i}{n}$ . So, the variance of this parameter is

$$var(\hat{\lambda}) = var\left(\frac{\sum_{i=1}^{n} x_i}{n}\right)$$

$$= \frac{var\sum_{i=1}^{n} x_i}{n^2}$$

$$= \frac{n\lambda}{n^2}$$

$$= \frac{\lambda}{n}$$

Also, from problem 2.2 (d),  $I(\hat{\lambda}) = \frac{1}{\lambda}$ . Hence  $\hat{\lambda}$  is an efficient estimator.