

Time Series Analysis

Lecture **X**: Summary Questions and Answers

Tohid Ardeshiri

Linköping University
Division of Statistics and Machine Learning

October 16, 2019



Course topics

- Time series, time series regression and exploratory analysis
 - ▶ Autocovariance, ACF
 - ▶ Sample ACF
 - ▶ Stationarity, detrending, differencing, transformation and smoothing
- ARIMA models
 - ▶ AR, MA, ARMA, ARIMA, seasonal ARIMA
 - ▶ PACF
 - ▶ Model selection
 - ▶ Estimation
 - ▶ Forecasting
- State space models
 - ▶ Linear and Gaussian state space models
 - ▶ Kalman filtering, Kalman smoothing and Forecasting
 - ▶ Maximum likelihood estimate of the state space models
 - ▶ Stochastic volatility
- Recurrent Neural Networks (RNNs)

Stationarity

- Time series x_t is **weakly stationary (stationary)** if
 - ▶ $Ex_t = \text{const}$
 - ▶ $\gamma(s, t) = \gamma(|s - t|)$
 - ▶ $\text{var}(x_t) < \infty$
- $\gamma(t, t + h) = \gamma(|t + h - t|) = \gamma(h)$
 - ▶ **Autocovariance depends on lag only!**
- Autocovariance for stationary process $\gamma(h) = \text{cov}(x_t, x_{t+h})$
- ACF for stationary process $\rho(h) = \frac{\gamma(h)}{\gamma(0)}$

Time domain: The Big Picture

Time Series data

Exploratory data analysis

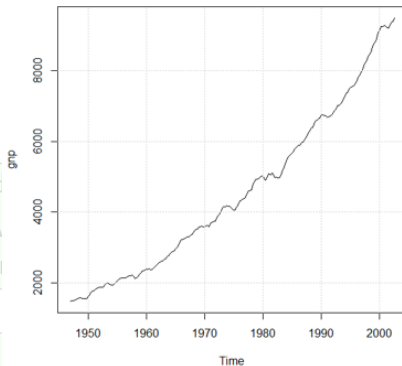
Make data stationary

Suggest a model $M(\phi, \theta)$

$$\phi = (\phi_1, \dots, \phi_p)$$

$$\theta = (\theta_1, \dots, \theta_q)$$

Prediction



Model

• A
a

• Estimate ϕ' s, θ' s, ...

Time domain: The Big Picture

Time Series data

Exploratory data analysis

Make data stationary

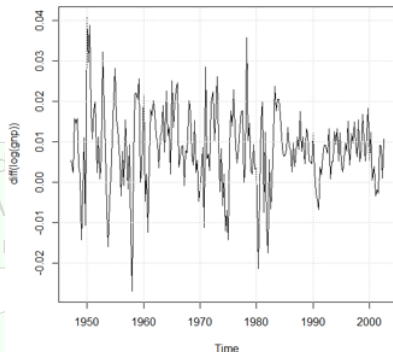
Suggest a model $M(\phi, \theta)$

$\phi = (\phi_1, \dots, \phi_p)$

$\theta = (\theta_1, \dots, \theta_q)$

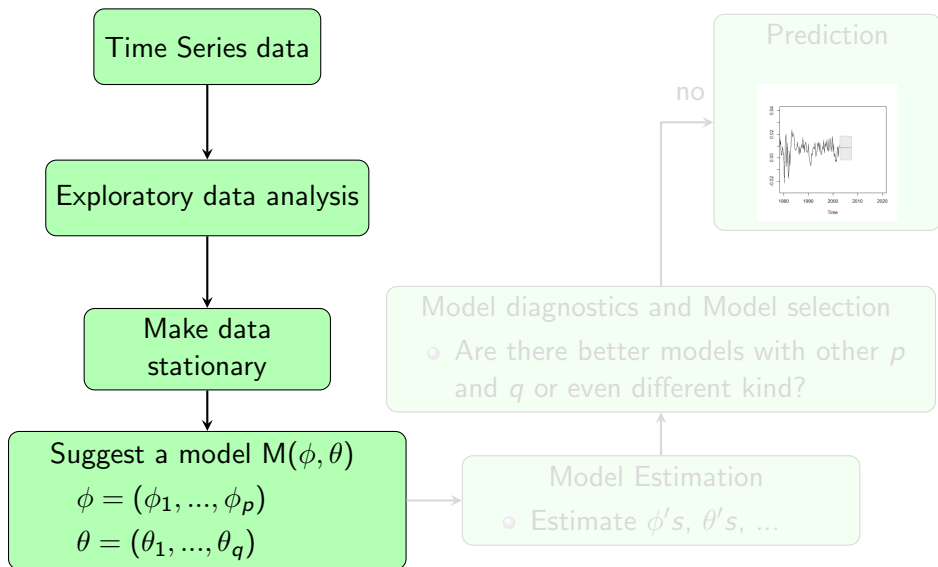
$$Y_t = \nabla(\log(X_t))$$

Prediction

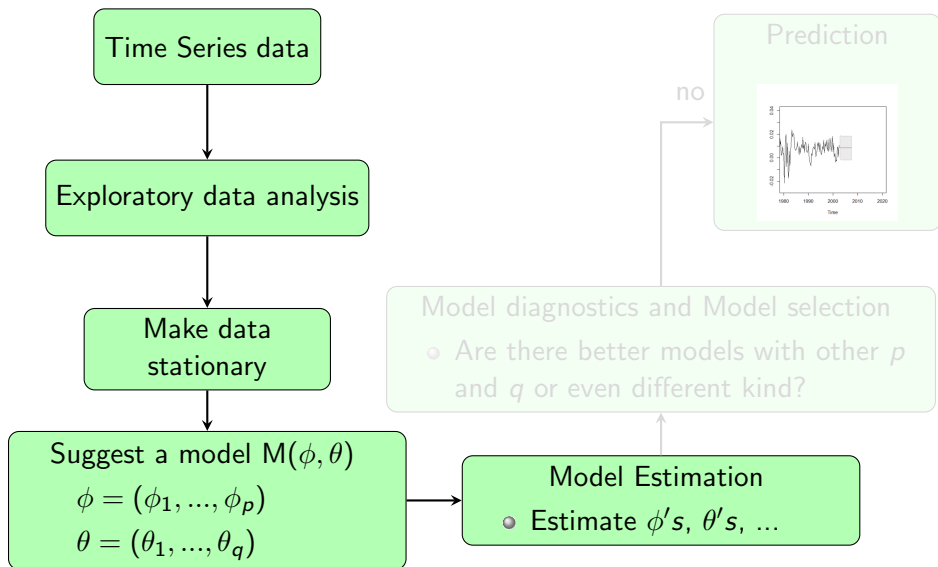


• Estimate ϕ 's, θ 's, ...

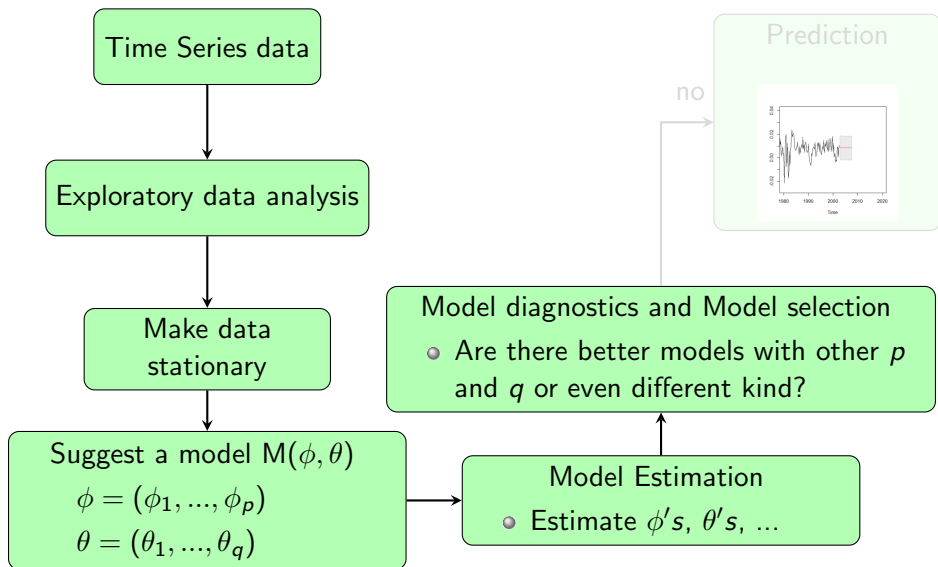
Time domain: The Big Picture



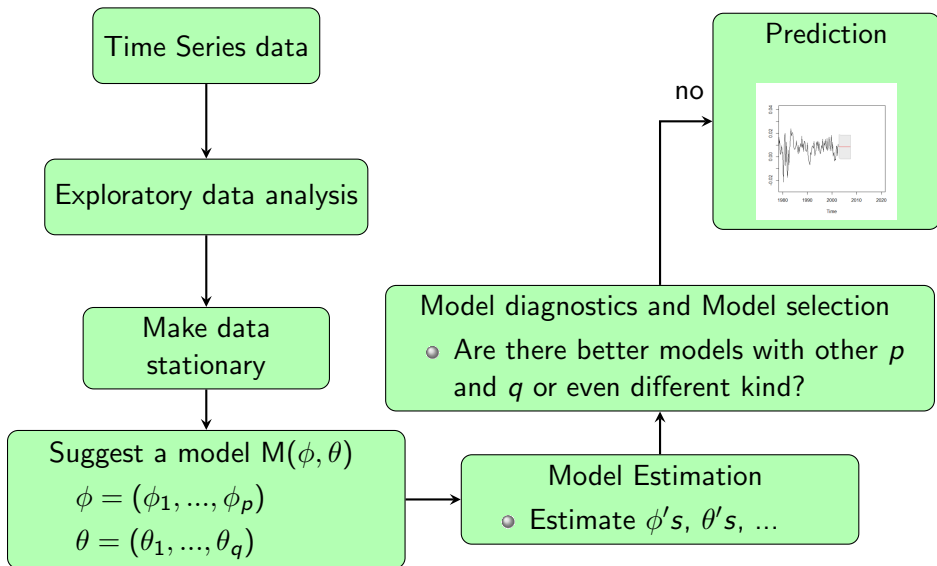
Time domain: The Big Picture



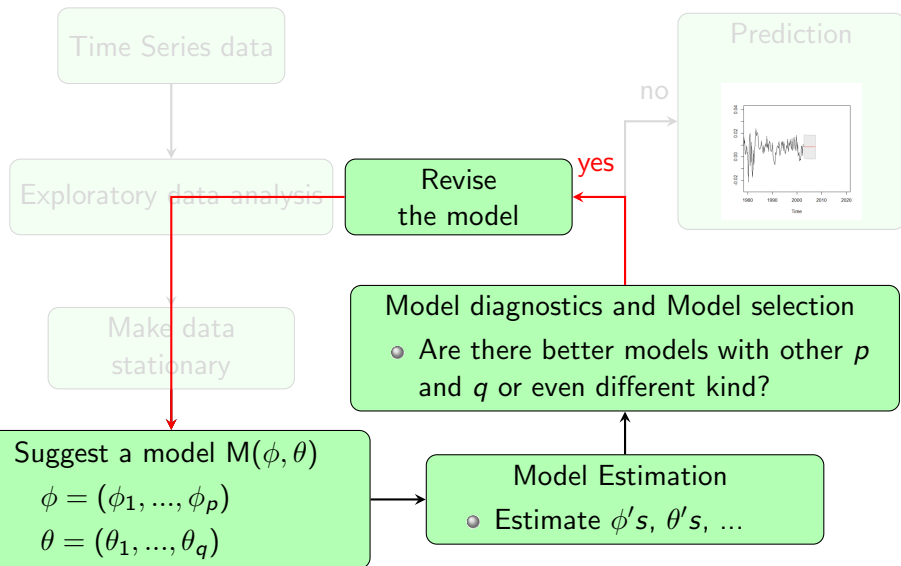
Time domain: The Big Picture



Time domain: The Big Picture



Time domain: The Big Picture



ARIMA modelling

- ARIMA models
 - ▶ AR, MA, ARMA, ARIMA, seasonal ARIMA
 - ▶ PACF
 - ▶ Model selection
 - ▶ Estimation
 - ▶ Forecasting

ARIMA models

Time series models so far

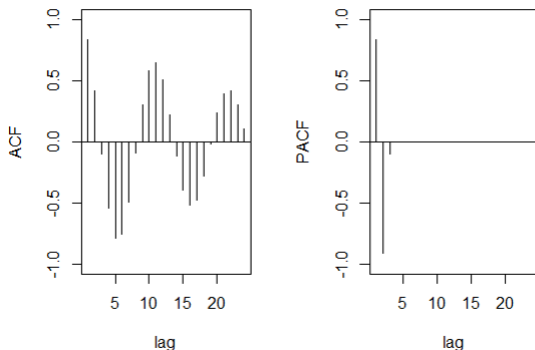
$$\phi^p(B)x_t = \theta^q(B)w_t$$

Model	Concise form
AR(p)	$\phi^p(B)x_t = w_t$
MA(q)	$x_t = \theta^q(B)w_t$
ARMA(p, q)	$\phi^p(B)x_t = \theta^q(B)w_t$
ARIMA(p, d, q)	$\phi^p(B)(1 - B)^d x_t = \theta^q(B)w_t$
ARMA(P, Q) _s	$\Phi^P(B^s)x_t = \Theta^Q(s)w_t$
ARIMA(P, D, Q) _s	$\Phi^P(B^s)(1 - B^s)^D x_t = \Theta^Q(B^s)w_t$
ARMA(p, q) \times (P, Q) _s	$\Phi^P(B^s)\phi^p(B)x_t = \Theta^Q(B^s)\theta^q(B)w_t$
ARIMA(p, d, q) \times (P, D, Q) _s	$\Phi^P(B^s)\phi^p(B)(1 - B^s)^D(1 - B)^d x_t = \Theta^Q(B^s)\theta^q(B)w_t$

* The notation used in this slide deviates from the notation used in the course literature so far.

PACF for AR(p)

- **Example:** AR(3) $\phi_1 = 1.5$, $\phi_2 = -0.75$, $\phi_3 = -0.1$



Seasonal?

ACF and PACF

	$AR(p)$	$MA(q)$	$ARMA(p, q)$
ACF	Tails off	Cuts off after lag q	Tails off
PACF	Cuts off after lag p	Tails off	Tails off

How to differentiate between $ARMA(p,q)$?

Empirical ACF (EACF)

Idea:

- ARMA(p,q): $x_t = \sum_{j=1}^p \phi_j x_{t-j} + \sum_{j=1}^q \theta_j w_{t-j} + w_t$
- If we can estimate $\phi_j \rightarrow x'_t = x_t - \sum_{j=1}^p \phi_j x_{t-j}$ is linear function in w_t, \dots, w_{t-q}
- If we run regression x'_t against $w_t \dots w_{t-q}$:
 - ▶ Residuals are white noise, $j \geq q \rightarrow$ ACFs not significant
 - ★ Some of the coefficients will be 0
 - ▶ Residuals are not white noise, $j < q \rightarrow$ ACFs significant
 - ▶ Note: w_t s substituted by lagged residuals from a series of regressions
- If $x'_t = x_t - \sum_{j=1}^k \phi_j x_{t-j}$, $k < p \rightarrow$ white noise will never be achieved \rightarrow ACFs are not zero

Empirical ACF (EACF)

- $k > p$ General result: ACFs are 0 for $j > q + (k - p)$
 - ▶ Example: ARMA(0,1)
- General conclusion for AR,MA =(k,j):
 - ▶ This is theoretical one! → not exactly the same for the samples

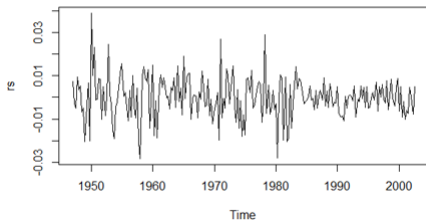
AR/MA	0	1	2
0	X	X	X	X	X	X	X
1	X	X	X	X	X	X	X
2	X	X	X	X	X	X	X
...	X	X	X	X	X	X	X
...	X	X	X	X	X	X	X
...	X	X	0	0	0	0	0
...	X	X	X	0	0	0	0
...	X	X	X	X	0	0	0
...	X	X	X	X	X	0	0

Residual analysis

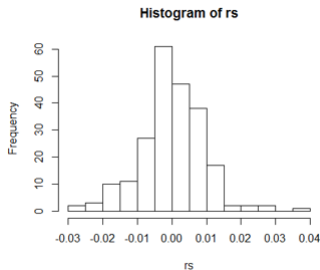
- Residuals $r_t = x_t - \hat{x}_t^{t-1}$? they are innovations
 - ▶ Note: computed from one-step-ahead predictions!
 - ▶ Measures predictive quality of the model (compare OLS)
- Residual analysis
 - ▶ Visual inspection: stationary? Patterns?
 - ▶ Histograms, Q-Q plots
 - ▶ ACF, PACF
 - ▶ Runs test
 - ▶ Box-Ljung test

Residual analysis - Visual inspection

Histogram and visual inspection



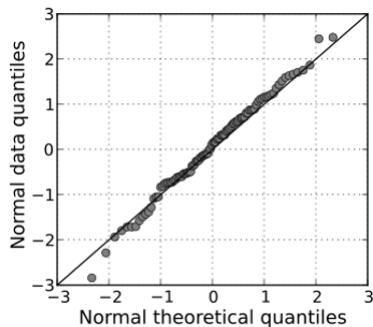
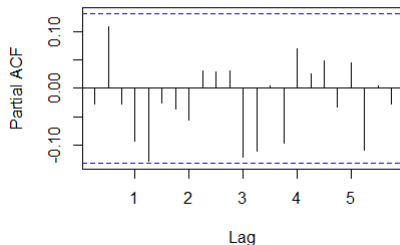
If looks white is good



If looks Normal is good

Residual analysis - ACF /PACF Q-Q plots

Series rs



Wikipedia

If between the blue lines good

If along the diagonal line GOOD

Statistical tests

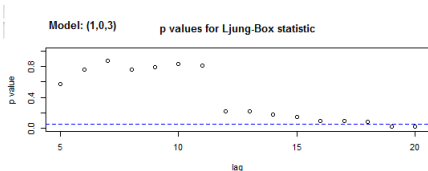
Tests are used to test independence

Runs test

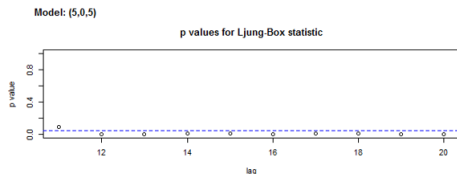
- H_0 : x_t values are i.i.d. **p-value NOT small**
- H_a : x_t values are not i.i.d. **p-value small**

Box-Ljung test

- H_0 : data are independent **p-value NOT small**
- H_a : data are not independent **p-value small**



GOOD



BAD

SARIMA

- Multiplicative seasonal autoregressive integrated moving average model $ARIMA(p, d, q) \times (P, D, Q)_s$

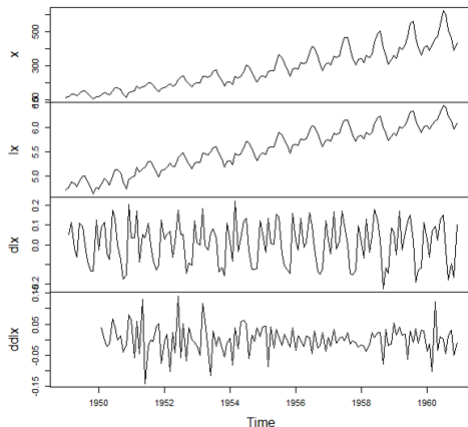
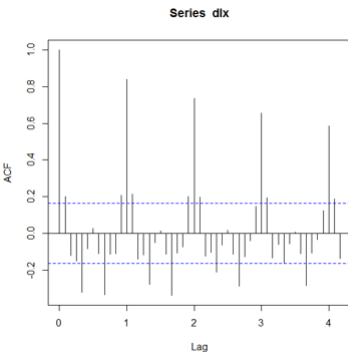
$$\Phi_p(B^s)\phi(B)\nabla_s^D\nabla^d x_t = \delta + \Theta_Q(B^s)\theta(B)w_t$$

$$\nabla_s^D = (1 - B^s)^D$$

- How to identify SARIMA?
 - ① Perform differencing first (trend)
 - ② Investigate ACF \rightarrow slowly decays at peaks
 - ① Yes \rightarrow Additional differencing by ∇_s^D
 - ③ Model non-seasonal part
 - ④ Model seasonal part (check peaks), check ACF and PACF of residuals

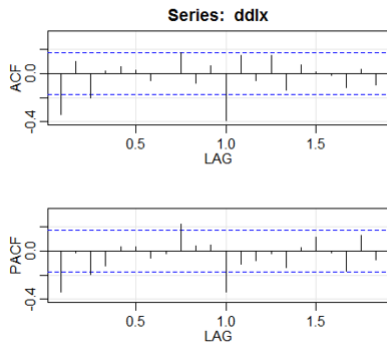
SARIMA

- **Example:** Air passengers



SARIMA

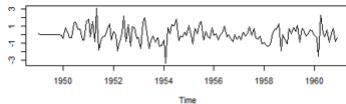
- **Example:** Air passengers



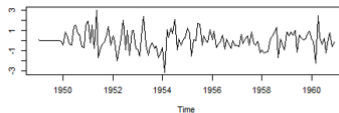
$(0, 1, 1)_{12}$ or $(1, 1, 0)_{12}$

SARIMA

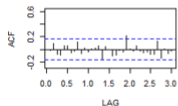
Model: (1,1,1) (0,1,1) Standardized Residuals



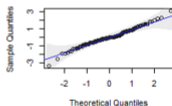
Model: (1,1,1) (1,1,0) Standardized Residuals



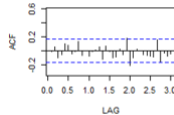
ACF of Residuals



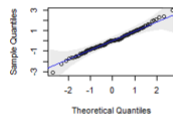
Normal Q-Q Plot of Std Residuals



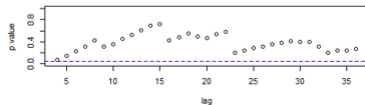
ACF of Residuals



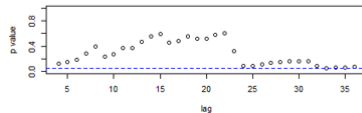
Normal Q-Q Plot of Std Residuals



p values for Ljung-Box statistic



p values for Ljung-Box statistic



SARIMA

- Remove AR term!

Is one model much better the other one?

$$(1, 1, 1) \times (1, 1, 0)_{12}$$

```
> m1$fit
```

```
Call:
```

```
stats::arima(x = xdata, order = c(p, d, q), seasonal = list(order = c(P, D, Q), period = S), include.mean = !no.constant, optim.control = list(trace = trc, REPORT = 1, reltol = tol))
```

```
Coefficients:
```

	ar1	ma1	sar1
	0.0547	-0.4886	-0.4731
s.e.	0.2161	0.1933	0.0800

```
sigma^2 estimated as 0.001425: log likelihood = 241.73, aic = -475.47
```

```
> m2$fit
```

```
Call:
```

```
stats::arima(x = xdata, order = c(p, d, q), seasonal = list(order = c(P, D, Q), period = S), include.mean = !no.constant, optim.control = list(trace = trc, REPORT = 1, reltol = tol))
```

```
Coefficients:
```

	ar1	ma1	sma1
	0.1960	-0.5784	-0.5643
s.e.	0.2475	0.2132	0.0747

```
sigma^2 estimated as 0.001341: log likelihood = 244.95, aic = -481.9
```

$$(1, 1, 1) \times (0, 1, 1)_{12}$$

State space modelling

- State space models
 - ▶ Linear and Gaussian state space models
 - ▶ Kalman filtering, Kalman smoothing and Forecasting
 - ▶ Maximum likelihood estimate of the state space models
 - ▶ Stochastic volatility

Consider an AR(2) model

$$x_t = \phi_1 x_{t-1} + \phi_2 x_{t-2} + w_t$$

Let $\mathbf{z}_t = \begin{bmatrix} x_t \\ x_{t-1} \end{bmatrix}$ and $\mathbf{e}_t = \begin{bmatrix} w_t \\ 0 \end{bmatrix}$.

Show that we rewrite the AR(2) model in the state space form:

$$\mathbf{z}_t = \begin{bmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{bmatrix} \mathbf{z}_{t-1} + \mathbf{e}_t$$
$$x_t = \begin{bmatrix} 1 & 0 \end{bmatrix} \mathbf{z}_t,$$

$$\phi^p(B)x_t = \theta^q(B)w_t$$

Can we rewrite any model of this form as a state space model?

$$\mathbf{z}_t = A\mathbf{z}_{t-1} + \mathbf{e}_t,$$

$$\mathbf{x}_t = C\mathbf{z}_t + \nu_t,$$

$$\phi^p(B)x_t = \theta^q(B)w_t$$

Outline of the solution:

Let $r = \max(p, q + 1)$,

$$\phi^r(B) = 1 - \phi_1 B - \dots - \phi_r B^r,$$

$$\theta^r(B) = 1 + \theta_1 B - \dots - \theta_{r-1} B^{r-1},$$

$\phi^r(B)(\theta^r(B))^{-1}x_t = w_t$. Hence, for $z_t = (\theta^r(B))^{-1}x_t$ we can have

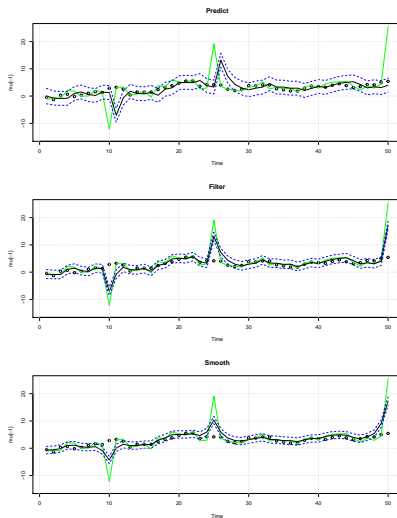
$$\phi^r(B)z_t = w_t$$

$$\mathbf{z}_t = \begin{bmatrix} z_t \\ z_{t-1} \\ z_{t-2} \\ \vdots \\ z_{t-r+1} \end{bmatrix} \quad \text{and} \quad \mathbf{z}_t = \begin{bmatrix} \phi_1 & \phi_2 & \cdots & \phi_r \\ 1 & 0 & \cdots & 0 \\ 0 & 1 & & 0 \\ \vdots & \vdots & & \vdots \\ 0 & 0 & \cdots & 1 \end{bmatrix} \mathbf{z}_{t-1} + \begin{bmatrix} w_t \\ 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

$$x_t = \begin{bmatrix} 1 & \theta_1 & \theta_2 & \cdots & \theta_r \end{bmatrix} \mathbf{z}_t$$

Robustness to outliers: filter versus smoother

Live example in Rstudio



Stochastic Volatility : Gaussian sum filter

The problem is finding the filtering distribution of $\mathbf{z}_t | \mathbf{x}_{1:t}$ when

$$\mathbf{z}_t = A\mathbf{z}_{t-1} + w_t$$

$$\mathbf{x}_t = C\mathbf{z}_t + \eta_t$$

and

$$w_t \sim iidN(0, Q)$$

$$\eta_t \sim \pi_0 N(\mu_0, R_1) + \pi_1 N(\mu_1, R_2)$$

where $\pi_0 + \pi_1 = 1$

Examination

- Most of the examination will be your Computer labs and assignments from the teaching sessions with a twist.
- You need to have a deep knowledge of the subjects covered in the lectures to get a B+ score.
- Study them over and over and make sure you have the correct solutions with you on the examination day.