STAT406- Methods of Statistical Learning Lecture 5

Matias Salibian-Barrera

UBC - Sep / Dec 2018









http://xkcd.com/292/

Comparing models

- AIC suggests a submodel
- Prediction-wise the full model is better
- AIC can be highly variable

"Smoother" model selection

- Ridge regression
- Can be thought as a type of feature selection
- It is a member of a larger class called "shrinkage methods"
- However, its origins are rather different

Without loss of generality...

• If covariates are centered, $\sum_{i=1}^{n} \mathbf{x}_{i} = \mathbf{0}$

$$\arg\min_{\alpha,\beta} \sum_{i=1}^{n} (y_i - \alpha - \beta' \mathbf{x}_i)^2$$

satisfies

$$\hat{\alpha}_n = \frac{1}{n} \sum_{i=1}^n y_i = \bar{y}_n,$$

and

$$\hat{\boldsymbol{\beta}}_{IS} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y},$$

Without loss of generality...

We can always assume that

$$\sum_{i=1}^{n} \mathbf{x}_i = \mathbf{0}$$

and hence

$$\hat{\alpha}_n = \frac{1}{n} \sum_{i=1}^n y_i = \bar{y}_n,$$

• In what follows, there is no intercept

Shrinkage methods

When covariates are correlated, LS estimators can be highly variable

$$\hat{\boldsymbol{\beta}}_{LS} = \left(\mathbf{X}' \mathbf{X} \right)^{-1} \mathbf{X}' \mathbf{Y}$$

$$\operatorname{var}\left(\hat{\boldsymbol{\beta}}_{n}\right) = \sigma^{2} \left(\mathbf{X}'\mathbf{X}\right)^{-1}$$

When X'X is close to singular...

Ridge Regression

 One way to "avoid" this problem is to add a "ridge" to X'X...

$$\hat{\boldsymbol{\beta}}_{RR} = \left(\mathbf{X}' \mathbf{X} + \lambda \, \mathbf{I}_p \right)^{-1} \, \mathbf{X}' \, \mathbf{Y}$$

where $\lambda > 0$ and

$$oldsymbol{\mathsf{I}}_{\mathcal{P}} = \left(egin{array}{cccc} 1 & 0 & \cdots & 0 \ 0 & 1 & \cdots & 0 \ 0 & \cdots & \ddots & 0 \ 0 & \cdots & \cdots & 1 \end{array}
ight)$$

Ridge Regression

• This is equivalent to solving

$$\min_{\boldsymbol{\beta}} \sum_{i=1}^{n} (y_i - \boldsymbol{\beta}' \mathbf{x}_i)^2 + \lambda \|\boldsymbol{\beta}\|_2^2$$

Ridge Regression

And also equivalent to solving

$$\min_{\beta} \sum_{i=1}^{n} (y_i - \beta' \mathbf{x}_i)^2$$

subject to

$$\sum_{j=1}^{p} \beta_j^2 \leq C$$

for some C > 0

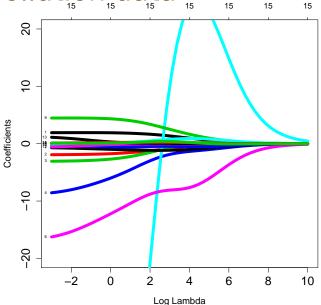
Bias / variance trade-off

- Ridge regression was originally proposed as a "hack" to "push" X'X away from singularity
- It can also be thought as a way of reducing the variance of $\hat{\beta}_n$
- This may increase the bias of the estimator, but if the variance is reduced even more, we might gain overall in expected squared error performance...

Ridge regression

- We now have a sequence ("path") of estimators (one for each $\lambda > 0$)
- $\mathbf{X}'\mathbf{X} + \lambda \mathbf{I}_p$ is always non-singular for $\lambda > 0$ (why?)
- Why are they called "shrinkage methods"?

Air pollution data



Questions

- What does λ measure?
- How do I choose one among these infinitely many "solutions"?

- How many "effective" parameters are we using?
- In linear regression, we have p parameters
- A more general definition is as follows. For a fitting method producing \hat{y}_1 , \hat{y}_2 , ..., \hat{y}_n ,

$$edf = \frac{1}{\sigma^2} \sum_{i=1}^{n} cov(\hat{y}_i, y_i)$$

Efron, B. (1986). How biased is the apparent error rate of a prediction rule? Journal of the

American Statistical Association, 81(394):461-470.

 It is easy to see that for least squares predictors, we have

$$\hat{\mathbf{y}} = \mathbf{H} \mathbf{y}$$

with

$$\mathbf{H} = \mathbf{X} (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'$$

and

$$edf = \frac{1}{\sigma^2} \sum_{i=1}^{n} cov(\hat{y}_i, y_i) = trace(\mathbf{H}) = p$$

 More in general, for any linear predictor

$$\hat{\mathbf{y}} = \mathbf{S} \, \mathbf{y}$$

we have

$$edf = trace(\mathbf{S}) = \sum_{i=1}^{n} \mathbf{S}_{i,i}$$

The ridge regression fit satisfies

$$\hat{\mathbf{y}} = \mathbf{S}_{\lambda} \mathbf{y}$$

where

$$\mathbf{S}_{\lambda} = \mathbf{X} \left(\mathbf{X}' \mathbf{X} + \lambda \mathbf{I}_{p} \right)^{-1} \mathbf{X}'$$

$$trace(S) = ?$$

 Using the singular value decomposition (SVD) of X

$$X = U \Lambda V'$$

where $\mathbf{U} \in \mathbb{R}^{n \times p}$, $\mathbf{V} \in \mathbb{R}^{p \times p}$ with

$$\mathbf{U}'\mathbf{U} = \mathbf{I}_{p} = \mathbf{V}'\mathbf{V}$$

and

$$\Lambda = \operatorname{diag}(d_1,\ldots,d_p)$$
,

we have

trace (**S**) =
$$\sum_{i=1}^{p} \left(\frac{d_i^2}{d_i^2 + \lambda} \right)$$

 For example, in the Air Pollution data example, if we use

$$\lambda = \exp(6)$$

we get

$$edf = 9.9$$

How do we select λ ?

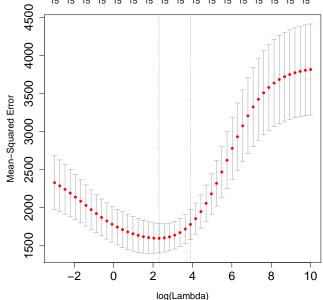
How can we select λ ?

How do we select λ ?

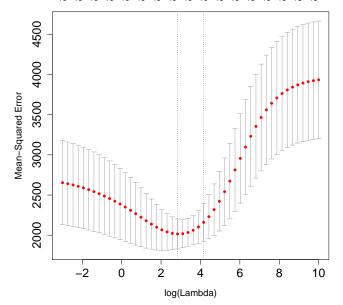
CV!

```
library (glmnet)
airp <- read.table('rutgers-lib-30861_CSV-1.csv'
      header=TRUE, sep=',')
v <- as.vector(airp$MORT)</pre>
xm < -as.matrix(airp[, -16])
lambdas \leftarrow exp(seq(-3, 10, length=50))
set.seed(123)
tmp <- cv.glmnet(x=xm, y=y, lambda=lambdas,</pre>
      nfolds=5, alpha=0,
      family='qaussian')
```

Air pollution data 15 15 15 15 15 15 15 15 15 15 15



Air pollution data - another run



Questions

- How are the standard errors estimated?
- Can we use AIC to compare these models?
- Why or why not?
 - If the answer is yes, how?
 - If the answer is no, why not?

CV

Cross validation selects

$$\lambda_{\sf op} \approx \exp(3)$$
 edf ≈ 13

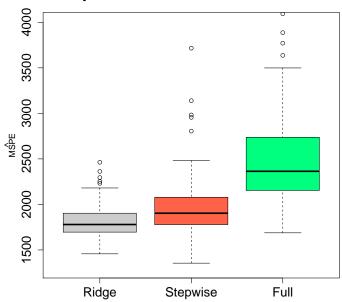
Stepwise selects

```
Call:
lm(formula = MORT ~ NONW + EDUC + JANT + SO.
+ PREC + JULT + POPN, data = airp)

Coefficients:
(Intercept) NONW EDUC JANT
1429.1866 5.2161 -16.9656 -1.8934

SO. PREC JULT POPN
0.2253 1.6485 -2.3006 -62.0118
```

Air pollution - 100 5-fold CV runs



Sometimes...

- Selecting variables is not always necessary in terms of prediction accuracy.
- One such an example is discussed on Github (Lecture 4)...
- ...and revised in Lecture 5. Read it carefully.