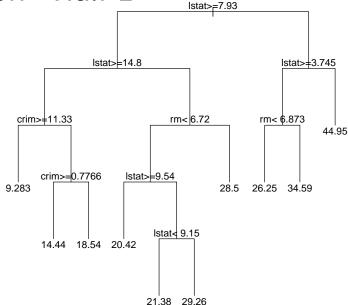
STAT406- Methods of Statistical Learning Lecture 12

Matias Salibian-Barrera

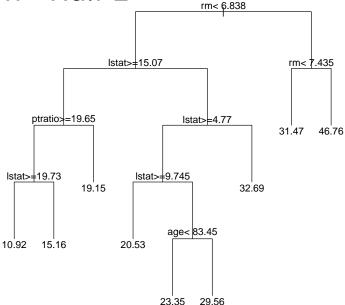
UBC - Sep / Dec 2017

- Trees can be highly variable
- Trees computed on samples from the sample population can be quite different from each other
- For example, we split the Boston data in two...

Boston - Half 1



Boston - Half 2



- Linear regression, for example, is not so variable
- Estimated coefficients computed on the same two halfs

```
(Intercept) crim zn indus chas
[1,] 39.21 -0.13 0.04 0.04 2.72
[2,] 33.12 -0.10 0.05 -0.01 2.80

nox rm age dis rad tax
[1,] -20.07 3.45 0 -1.44 0.28 -0.01
[2,] -14.18 4.15 0 -1.46 0.34 -0.02

ptratio black lstat
[1,] -1.01 0.01 -0.56
[2,] -0.90 0.01 -0.50
```

- If we could average many trees trained on independent samples from the same population, we would obtain a predictor with lower variance
- If \hat{f}_1 , \hat{f}_2 , ..., \hat{f}_B are B regression trees, then their average is

$$\hat{f}_{av}(\mathbf{x}) = \frac{1}{B} \sum_{i=1}^{B} \hat{f}_{i}(\mathbf{x})$$

- However, we generally do not have B training sets...
- We can **bootstrap** the training set to obtain B pseudo-new-training sets
- Let (Y_1, \mathbf{X}_1) , (Y_2, \mathbf{X}_2) , ..., (Y_n, \mathbf{X}_n) be the training sample, where

$$(Y_j, \mathbf{X}_j) \sim F_0$$

- If we knew F₀, then we could generate / simulate new training sets, and average the resulting trees...
- We do not know F₀, but we have an estimate for it
- Let F_n be the empirical distribution of our only training set (Y_1, \mathbf{X}_1) , (Y_2, \mathbf{X}_2) , ..., (Y_n, \mathbf{X}_n)

We know that

$$F_n \xrightarrow[n\to\infty]{} F_0$$

(in what sense?)

- Bootstrap generates / simulates samples from F_n
- Taking a sample of size n from F_n is the same as sampling with replacement from the training set $(Y_1, \mathbf{X}_1), (Y_2, \mathbf{X}_2), \ldots, (Y_n, \mathbf{X}_n)$

- To apply bagging to a regression tree, take B independent samples (with replacement) from the training set
- Obtain the B trees: \hat{f}_1^* , \hat{f}_2^* , ..., \hat{f}_B^*
- and average their predictions

$$\hat{f}_{\text{bag}}(\mathbf{x}) = \frac{1}{B} \sum_{i=1}^{B} \hat{f}_{i}^{*}(\mathbf{x})$$

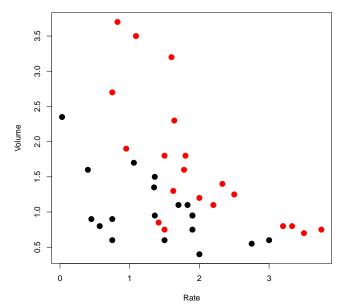
- Generally, we apply bagging on "large" trees, without pruning them (try to retain their low-bias and reduce their variance by averaging)
- With the Boston data set, if we apply bagging to the regression tree computed on the training set, and then use it to predict on the test set, we obtain:

• B = 1> mean((dat.te\$medv - pr.ba)^2) [1] 16.44972 • B = 5> mean((dat.te\$medv - pr.ba)^2) [1] 15.12332 • B = 100> mean((dat.te\$medv - pr.ba)^2) [1] 12.30543 • B = 500> mean((dat.te\$medv - pr.ba)^2) [1] 12.32504

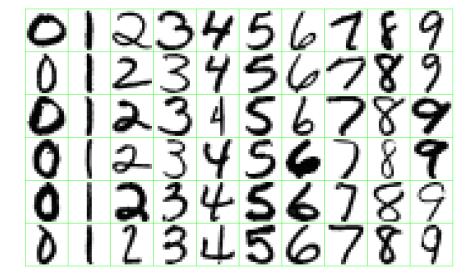
B = 2000
mean((dat.te\$medv - pr.ba)^2)
[1] 11.8116
B = 5000
mean((dat.te\$medv - pr.ba)^2)
[1] 11.85943

- This approach applies to any predictor (not only trees)
- It will be particularly useful for low-bias / high-variance predictors

Classification



Predict hand-written digits



- In general, we have n observations (training)
- $(g_1, \mathbf{x}_1), (g_2, \mathbf{x}_2), \ldots, (g_n, \mathbf{x}_n)$
- we would like to build a classifier, a function $\hat{g}(\mathbf{x})$ to predict the true class g of a future observation (g, \mathbf{x}) (for which g is unknown)

- In general, there are K possible classes, c_1, c_2, \ldots, c_K . In other words $g \in \{c_1, c_2, \ldots, c_K\}$
- Consider the following loss function

$$L(a,b) = \begin{cases} 0 & \text{if } a = b \\ 1 & \text{if } a \neq b \end{cases}$$

• Find a classifier $\hat{g}(\mathbf{x})$ such that

$$E_{(G,\mathbf{X})}[L(G,\hat{g}(\mathbf{X}))] \leq E_{(G,\mathbf{X})}[L(G,h(\mathbf{X}))]$$

for any other function h

$$E_{(G,\mathbf{X})}\left[L\left(G,\hat{\mathbf{g}}(\mathbf{X})\right)\right] = E_{\mathbf{X}}\left\{E_{G|\mathbf{X}}\left[L\left(G,\hat{\mathbf{g}}(\mathbf{X})\right)\right]\right\}$$
$$= E_{\mathbf{X}}\left\{\sum_{j=1}^{K}L\left(c_{j},\hat{\mathbf{g}}(\mathbf{X})\right)P\left(G=c_{j}|\mathbf{X}\right)\right\}$$

• It is sufficient to find $\hat{g}(\mathbf{X})$ that minimizes

$$egin{aligned} \sum_{j=1}^{K} L\left(c_{j}, \hat{g}(\mathbf{X})
ight) P\left(G = c_{j} \middle| \mathbf{X}
ight) \ &= \sum_{c_{j}
eq \hat{g}(\mathbf{X})} P\left(G = c_{j} \middle| \mathbf{X}
ight) \ &= 1 - P\left(G = \hat{g}(\mathbf{X}) \middle| \mathbf{X}
ight) \end{aligned}$$

• Hence, the optimal classifier satisfies

$$P(G = \hat{g}(\mathbf{X})|\mathbf{X}) \geq P(G = c_i|\mathbf{X})$$
 for all c_i

More than 2 groups

• In other words, $\hat{g}(\mathbf{X})$ should be the class with the highest probability

$$\hat{g}(\mathbf{X}) = \arg \max_{\mathbf{g} \in \{c_1, \dots, c_K\}} P(G = \mathbf{g} | \mathbf{X})$$

 "Assign X to the class with largest posterior probability given X"

 Most classifiers can be thought of as different ways to estimate or model

$$\mathbf{f_j}(\mathbf{x}) = P(G = \mathbf{c}_j | \mathbf{X} = \mathbf{x})$$

 For example, logistic classifiers propose a model for f_i:

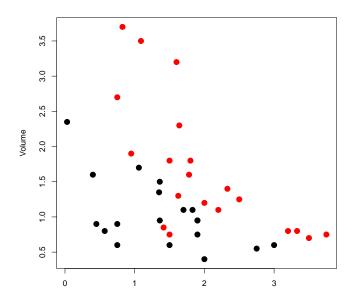
$$\mathbf{f_j}(\mathbf{x}) = \frac{\exp\left(eta_j \mathbf{x}\right)}{1 + \exp\left(eta_i \mathbf{x}\right)}$$

These posterior probabilities

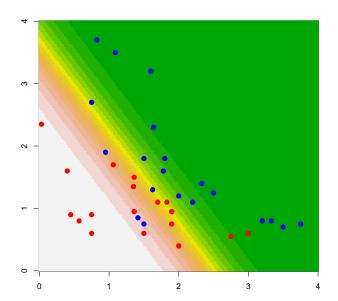
$$P(G = \mathbf{c}_j \mid \mathbf{X} = \mathbf{x})$$

can also be used

- to quantify uncertainty in the classification for a particular value of x
- to identify regions of the feature space where classification isn't so clear



Logistic based probabilities



A model for $\mathbf{X}|g$

If we **model** the feature **distribution** in each **group**:

$$f(\mathbf{X}|G=c_{\mathbf{k}})=f_{\mathbf{k}}(\mathbf{X})$$
 $\mathbf{k}=1,\ldots,\mathbf{K}$

then

$$P(G = c_{\mathbf{k}} | \mathbf{X}) = \frac{f(\mathbf{X} | G = c_{\mathbf{k}}) p_{\mathbf{k}}}{f(\mathbf{X})} = \frac{f_{\mathbf{k}}(\mathbf{X}) p_{\mathbf{k}}}{f(\mathbf{X})}$$

thus

$$\hat{\mathbf{g}}(\mathbf{X}) = \arg \max_{1 \le \mathbf{k} \le \mathbf{K}} f_{\mathbf{k}}(\mathbf{X}) p_{\mathbf{k}}$$

A model for $\mathbf{X}|g$

For example, we can assume that

$$|\mathbf{X}| G = c_{\mathbf{k}} \sim \mathcal{N}(\boldsymbol{\mu}_{\mathbf{k}}, \boldsymbol{\Sigma})$$

then, we can estimate

$$\hat{f}_{f k}({f X}) \sim \mathcal{N}\left(\hat{\mu}_{f k},\widehat{f \Sigma}
ight)$$

using the sample mean of each group and the pooled sample covariance matrix.

We can then find the class \mathbf{k} that has the largest $\hat{f}_{\mathbf{k}}(\mathbf{X}) p_{\mathbf{k}}$

Note that if $f_j \sim \mathcal{N}_{p}\left(\mu_j, \mathbf{\Sigma}\right)$, j=1,2

$$egin{aligned} f_1(\mathbf{x}) \, \rho_1 &> f_2(\mathbf{x}) \, \rho_2 &\Leftrightarrow \\ &\log \left(rac{f_1(\mathbf{x}) \, \rho_1}{f_2(\mathbf{x}) \, \rho_2}
ight) > 0 &\Leftrightarrow \\ &\mathbf{a}' \mathbf{x} + \mathbf{b} \, > \, 0 \end{aligned}$$

for some $\mathbf{a} \in \mathbb{R}^p$ and $\mathbf{b} \in \mathbb{R}$.

In other words, boundaries between classes are **linear**.

Furthermore, we can estimate this linear boundary because

$$\mathbf{a} = \mathbf{\Sigma}^{-1} \; (\mu_1 - \mu_2)$$

and

$$\mathbf{b} = -rac{1}{2} \left(\mathbf{\mu}_1 - \mathbf{\mu}_2
ight)' \mathbf{\Sigma}^{-1} \left(\mathbf{\mu}_1 + \mathbf{\mu}_2
ight) - \log \left(rac{\mathbf{p}_2}{\mathbf{p}_1}
ight)$$

We can also write this in term of class probabilities

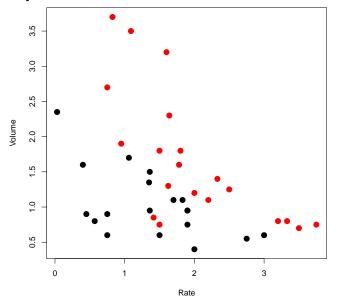
$$\frac{P(G = c_1 | \mathbf{X})}{P(G = c_2 | \mathbf{X})} > 1 \quad \Leftrightarrow \quad f_1(\mathbf{X}) p_1 > f_2(\mathbf{X}) p_2$$

$$\Leftrightarrow \log\left(\frac{f_1(\mathbf{x})\,\rho_1}{f_2(\mathbf{x})\,\rho_2}\right) > 0 \quad \Leftrightarrow \quad \mathbf{a}'\mathbf{x} + \mathbf{b} \, > \, 0$$

In fact, for normally distributed features we have

$$\begin{split} \log \left(\frac{P(G = c_1 | \mathbf{X})}{P(G = c_2 | \mathbf{X})} \right) &= \\ \log \left(\frac{P(G = c_1 | \mathbf{X})}{1 - P(G = c_1 | \mathbf{X})} \right) &= \mathbf{a}' \mathbf{x} + \mathbf{b} \end{split}$$

With two classes, we have also estimated a and b using logistic regression.



- First assume that Volume and Rate are normally distributed in each class
- Then, the optimal classifier classifies a point x = (Volume, Rate)' in class 1 (red) if

$$a'x + b > 0$$

where

$$\mathbf{a} = \mathbf{\Sigma}^{-1} \; (\mu_1 - \mu_2)$$

and

$$oldsymbol{b} = -rac{1}{2} \left(oldsymbol{\mu}_1 - oldsymbol{\mu}_2
ight)' \, oldsymbol{\Sigma}^{-1} \, \left(oldsymbol{\mu}_1 + oldsymbol{\mu}_2
ight) - \log \left(rac{oldsymbol{p}_2}{oldsymbol{p}_1}
ight)$$

- We can estimate μ_1 , μ_2 and Σ (and even p_1 and p_2). **How?**
- We get $\hat{\mathbf{a}} = (-2.77, -2.37)'$ and $\hat{\mathbf{b}} = 7.72$
- Then, the estimated optimal classifier classifies a point x = (Volume, Rate)' in class 1 (red) if
 - $-2.77 \, Volume 2.37 \, Rate + 7.72 > 0$

Furthermore

$$\widehat{P}(G=1| ext{(Volume, Rate)}) = \\ \frac{\exp\left(-2.77 ext{ Volume} - 2.37 ext{ Rate} + 7.72\right)}{1 + \exp\left(-2.77 ext{ Volume} - 2.37 ext{ Rate} + 7.72\right)}$$

and

$$\widehat{P}(G = 2 | (Volume, Rate)) = 1 - \widehat{P}(G = 1 | (Volume, Rate)) =$$

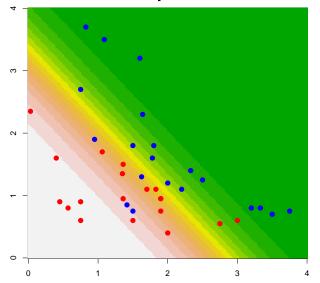
 Now, create a fine grid of Volume and Rate values, and use the previous formulas to predict

$$P(G = j | (Volume, Rate)), \quad j = 1, 2$$

- Plot these posterior probabilities
- We can do this by hand, or using the function lda in package MASS and its predict method

```
library (MASS)
data(vaso, package='robustbase')
plot(Volume ~ Rate, pch=19, col=c('red', 'blue')[Y+1],
      data=vaso, cex=1.3)
a.lda \leftarrow lda(Y \sim Volume + Rate, prior = c(.5, .5),
     data=vaso)
aa < - seg(0, 4, length=200)
bb < - seg(0, 4, length=200)
dd <- expand.grid(aa, bb)
names(dd) <- c('Volume', 'Rate')</pre>
pr.lda <- predict(a.lda, newdata=dd)$posterior[,1]</pre>
image(aa, bb, matrix(pr.1da, 200, 200),
     col=terrain.colors(15), xlab='', ylab='')
points(Volume ~ Rate, pch=19, col=c('red', 'blue')[Y+1],
     data=vaso, cex=1.3)
```

Gaussian-based probabilities



- Vaso example Logistic linear model
- Data (y_1, \mathbf{x}_1) , (y_2, \mathbf{x}_2) , ..., (y_n, \mathbf{x}_n)
- $y_j = 0, 1, \mathbf{x} = (rate, volume)'$
- A possible model is

$$P\left(y_{j}=1\big|\,oldsymbol{x}_{j}
ight) \,=\, rac{\exp\left(eta'\,oldsymbol{x}_{j}
ight)}{1+\exp\left(eta'\,oldsymbol{x}_{j}
ight)}$$

- We can estimate β using MLE
- Function glm in R
- Given values of rate and volume we predict a 1 if

$$\hat{P}(y_i = 1 | \text{rate}, \text{volume}) > 0.5$$

 Note that if we do not assume Gaussian features but insist that

$$egin{split} \log\left(rac{P(G=1|\mathbf{X})}{P(G=2|\mathbf{X})}
ight) = \ &\log\left(rac{P(G=1|\mathbf{X})}{1-P(G=1|\mathbf{X})}
ight) = \ &\mathbf{a}'\mathbf{x} + \mathbf{b} \end{split}$$

we can use glm to estimate \hat{a} and \hat{b} :

$$\hat{\mathbf{a}} = (-3.88, -2.65)'$$
 and $\hat{\mathbf{b}} = 9.53$

Logistic-based probabilities

```
data(vaso, package='robustbase')
a <- glm(Y ~ Volume + Rate, data=vaso, family=binomial
aa <- seg(0, 4, length=200)
bb < - seq(0, 4, length=200)
dd <- expand.grid(aa, bb)</pre>
names(dd) <- c('Volume', 'Rate')</pre>
yy <- predict(a, newdata=dd, type='response')</pre>
image(aa, bb, matrix(1-yy, 200, 200),
     col=terrain.colors(15), xlab='', ylab='')
points(Volume ~ Rate, pch=19, col=c('red', 'blue')[Y+1]
     data=vaso, cex=1.3)
```

Logistic-based probabilities

