Homework 4

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3.8.12

a) Let X be the number of visits in a day. $X \sim Poisson(\lambda T) = Poisson(1.8)$

$$P(X > 5) = 1 - P(X \le 5)$$

$$= 1 - [P(X = 5) + P(X = 4) + P(X = 3) + P(X = 2) + P(X = 1) + P(X = 0)]$$

$$= 0.01037804$$

b) Let Y be the number of visits in a week. $Y \sim Poisson(\lambda T) = Possion(1.8 \times 7)$

$$\begin{split} P(Y<5) &= P(4) + P(3) + P(2) + P(1) + P(0) \\ &= \frac{e^{-12.6}12.6^4}{4!} + \frac{e^{-12.6}12.6^3}{3!} + \frac{e^{-12.6}12.6^2}{2!} + \frac{e^{-12.6}12.6^1}{1!} + \frac{e^{-12.6}12.6^0}{0!} \\ &\approx 0.004979028 \end{split}$$

c) We want to find how many days it would take for $P(Z \ge 1) \ge 0.99$ where Z is the number of visits in $x \in \mathbf{Z}$ days. We'll first solve for $P(Z \ge 1) = 0.99$:

$$P(Z \ge 1) = 0.99$$

$$1 - P(0) = 0.99$$

$$1 - \frac{e^{-\lambda \cdot x}(\lambda \cdot x)^{0}}{0!} = 0.99$$

$$e^{-\lambda \cdot x} = -0.01$$

$$e^{-\lambda \cdot x} = 0.01$$

$$\lambda \cdot x = \ln(0.01)$$

$$x = \frac{\ln(0.01)}{\lambda}, \text{ where } \lambda = 1.8$$

$$x \approx 2.558427$$

Since Poisson takes discrete numbers / whole numbers, we'll round x up to 3 days. Hence, it'll take at least 3 days such that the probability of at least one visit is 99%.

d) Let X be the number of visits per day, then $X \sim Pois(\lambda T)$. We want to find P(X > 5) = 0.1 which is equivalent to:

$$1 - [P(X = 0) + P(X = 1) + P(X = 2) + P(X = 3) + P(X = 4) + P(X = 5)] = 0.01$$

And we solve for the following, where
$$\lambda t = \mu$$
:
$$1 - \left[\frac{e^{-\mu}\mu^0}{0!} + \frac{e^{-\mu}\mu^1}{1!} + \frac{e^{-\mu}\mu^2}{2!} + \frac{e^{-\mu}\mu^3}{3!} + \frac{e^{-\mu}\mu^4}{4!} + \frac{e^{-\mu}\mu^5}{5!} \right] = 0.1$$

1 - ppois(5, 1.8)

[1] 0.01037804

ppois(4, 12.6)

[1] 0.004979028

1 - ppois(0, 1.8 * 3)

[1] 0.9954834

4.1.2

a)
$$P(X<2) = \int_1^2 \frac{2}{x^3} dx = \frac{2x^{-2}}{-2} |_1^2 = -[\frac{1}{x^2} |_1^2] = \frac{3}{4} = 0.75$$

b)
$$P(X>5) = 1 - \int_1^5 \frac{2}{x^3} dx = 1 - \left[\frac{-1}{x^2} \mid_2^5\right] = 1 + \left[\frac{1}{25} - 1\right] = \frac{1}{25}$$

c)
$$P(4 < X < 8)$$
 $\int_4^8 \frac{2}{x^3} dx = \frac{-1}{x^2} |_4^8 = -[\frac{1}{64} - \frac{1}{16}] = \frac{3}{64}$

d)
$$P(X < 4 \text{ or } X > 8) = 1 - (4 < X < 8) = 1 - \int_{4}^{8} \frac{2}{x^3} dx = 1 - (-\frac{1}{x^2} |_{4}^{8}) = 1 + (\frac{1}{64} - \frac{1}{16}) = 1 - \frac{3}{64} = \frac{61}{64}$$

e) Let x be the value such that F(X < x) = 0.95

$$\int_{1}^{x} \frac{2}{t^{3}} dt = 0.95$$

$$\frac{-1}{t^{2}} \Big|_{1}^{x} = 0.95$$

$$\frac{-1}{x^{2}} + 1 = 0.95$$

$$\frac{1}{x^{2}} = 0.05$$

$$\frac{1}{x^{2}} = \frac{5}{100}$$

$$\frac{1}{x^{2}} = \frac{1}{20}$$

$$x = \sqrt{20}$$

4.1.2 Extras

a) CDF for
$$x > 1 = P(X \le x) = \int_1^x \frac{2}{t^3} dt = \frac{-1}{t^2} \Big|_1^x = \frac{-1}{x^2} + 1$$

b)
$$\mu = \int_{1}^{\infty} x f(x) dx = \int_{1}^{\infty} x \cdot \frac{2}{x^{3}} dx = \int_{1}^{\infty} \frac{2}{x^{2}} dx = \frac{-2}{x} \Big|_{1}^{\infty} = \frac{-2}{\infty} - \frac{-2}{1} = 0 + 2 = 2$$

c)
$$\sigma^2 = E[X^2] - E[X]^2 = \int_1^\infty x^2 f(x) dx - 2^2 = \int_1^\infty x^2 \cdot \frac{2}{x^3} dx - 4 = \int_1^\infty \frac{2}{x} dx - 4 = 2ln(x) \mid_1^\infty - 4 = \infty$$

4.4.2

a) CDF for $0.95 < x < 1.05 = \int_{0.95}^{x} \frac{1}{(1.05 - 0.95)} dt = \int_{0.95}^{x} 10 dt = 10t \mid_{0.95}^{x} = 10x - 10(0.95)$

We then get:

$$F(x) = \begin{cases} 0, & \text{if } x < 0.95\\ 10x - 10(0.95), & \text{if } 0.95 \le x \le 1.05\\ 1, & \text{if } x > 1.05 \end{cases}$$

b)
$$P(X > 1.02) = \int_{1.02}^{1.05} \frac{1}{1.05 - 0.95} dx = 10x \mid_{1.02}^{1.05} = 10.5 - 10.2 = 0.3$$

c) let a be the lower bound of 0.95, and b be the upper bound of 1.05. To find the thickness exceeded by 90% of flanks, we just need to find where P(X < x) = 1 - 0.9. Since our probability is uniform, we can simply get the values of the lower 10% range and add it to the lower bound. Thus we get:

$$0.95 + 0.1(1.05 - 0.95) = 1.02(0.1 \cdot 0.1) = 0.95 + 0.01 = 0.96$$

4.5.6

$$a)X \sim N(260, 50^2)$$

We want to find P(X>4*60) = P(X>240). In order to do this, we'll first find Z, where $Z \sim N(0,1)$:

$$Z = \frac{X - \mu}{\sigma} = \frac{240 - 260}{50} = \frac{-20}{50} = \frac{-2}{5} = -0.4$$

$$P(X > 240) = P(Z > -0.4) = P(Z < 0.4) = 0.6554$$

b) We can use the z-score of the first and third quartile:

The z-score for the third quartile (75%) is approximately 0.67, which we will then convert to the cumulative distribution of the problem:

$$Z = \frac{X - \mu}{\sigma}$$

$$0.67 = \frac{X - 260}{50}$$

$$33.5 = X - 260$$

$$293.5 = X$$

We can also do the same to find our first quartile, who's z-score is just the negative of the third quartile. Hence, Z = -0.67:

$$Z = \frac{X - \mu}{\sigma}$$

$$-0.67 = \frac{X - 260}{50}$$

$$-33.5 = X - 260$$

$$226.5 = X$$

Thus, the first quartile (25%) has a battery life of 226.5 minutes while the third quartile (75%) has a battery life of 293.5 minutes.

c) Looking at our z-table, we see that the z-score with 95% probability is approximately 1.64. We will take Z = -1.64 since we're looking for the battery life that is surpassed with 95% probability, while 1.64 gives us the value surpassed by 5%. We will convert -1.64 to our current distribution where $X \sim N(260, 50^2)$:

$$Z = \frac{X - \mu}{\sigma}$$

$$-1.64 = \frac{X - 260}{50}$$

$$-82 = X - 260$$

$$178 = X$$

Thus, the value that is surpassed by 95% of battery lives is approximately 178 minutes.

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1 - pnorm(240, 260, 50)

## [1] 0.6554217

pnorm(293.5, 260, 50)

## [1] 0.7485711

pnorm(226.5, 260, 50)

## [1] 0.2514289

pnorm(178, 260, 50)
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[1] 0.05050258

4.5.12

a) $X \sim N(310, 45^2)$ (units are in millions of gallons). We want to find when P(X > 350). Thus, we'll just find 1 - P(X < 350). But first, we'll convert our distribution to a standardized distribution $Z \sim N(0, 1)$:

$$Z = \frac{X - \mu}{\sigma}$$

$$Z = \frac{350 - 310}{45}$$

$$Z = \frac{40}{45}$$

$$Z \approx 0.888889$$

Then P(Z < 0.89) = 0.81327, and 1 - P(Z < 0.89) = 0.18673. Thus, the probability that a day requires more water than what's stored in the city reservoir is 18.673%.

b) We will need to find where P(X < x) = 0.99. The z-score for 99% probability in a standard normal distribution is ≈ 2.33 . Then we'll convert from standard $Z \sim N(0,1)$ to $X \sim N(310,45^2)$.

$$Z = \frac{X - \mu}{\sigma}$$
$$2.33 = \frac{X - 310}{45}$$
$$104.85 = X - 310$$
$$414.85 = X$$

Thus, the city will need a reservoir capacity of 414.85 million gallons in order for the probability of it being exceeded to be 1%.

c) We need to find where P(Z < z) = 0.05. The z-score for 0.95 is ≈ 1.64 , thus we'll our z-score will need to convert x = -1.65 to the distribution where $X \sim N(310, 45^2)$:

$$Z = \frac{X - \mu}{\sigma}$$

$$-1.65 = \frac{X - 310}{45}$$

$$-74.25 = X - 310$$

$$235.75 = X$$

Thus, the amount of water needed for water use to be exceeded with 95% probability is 235.75 million gallons.

d) We want to find the mean when P(X > 350) = 0.01 where $X \sim N(310, 45^2)$. This is equivalent to finding when P(X < 350) = 0.99. The z-score in the standard normal distribution for 0.99 is approximately 2.33:

$$Z = \frac{X - \mu}{\sigma}$$
$$2.33 = \frac{350 - \mu}{45}$$
$$104.85 = 350 - \mu$$
$$245.15 = \mu$$

Since we're looking for consumption per person, we'll get $\frac{245.15}{1.4} \approx 175.10714$. Thus, the mean daily consumption per person will need to be approximately 175.10714 gallons.

4.7.5

a) Let Y be the number of calls in 30 minutes, then $Y \sim Pois(2)$: $P(Y=0) = e^{-2} \approx 0.135335$

b) Let X be the number of calls in 10 minutes, then $X \sim Pois(\frac{2}{3})$. We want to find $P(X \le 1)$: $P(Z \le 1) = 1 - P(Z > 1) = 1 - e^{\lambda t} = 1 - e^{\frac{2}{3}} \approx 0.48658$

c)
$$P(\frac{1}{3} \le X \le \frac{2}{3}) = F(\frac{2}{3}) - F(\frac{1}{3}) = (1 - e^{-2/3}) - (1 - e^{-1/3}) = 0.203114$$

d) We're trying to find when F(t) = 0.9 since F(t) gives the probability of a call within a certain time interval:

$$\begin{split} F(t) &= 0.9 \\ P(X \leq t) &= 0.9 \\ 1 - e^{-\lambda t} &= 0.9 \\ -e^{-\lambda t} &= -0.1 \\ e^{-\lambda t} &= 0.1 \\ -\lambda t &= \ln(0.1) \\ \lambda t &= -\ln(0.1) \\ t &= \frac{-\ln(0.1)}{\lambda} \\ t &= \frac{-\ln(0.1)}{\frac{1}{15}}, \text{let } \lambda \text{ be } 1/15 \text{ since its the rate per minute} \\ t &\approx 34.538776 \end{split}$$

Thus, the time interval would have to be approx. 34.538776 minutes in order for the probability of at least one call in interval to be 0.90.

dpois(0, 2)

[1] 0.1353353

1 - dpois(0, 2/3)

[1] 0.4865829

(1 - dpois(0, 2/3)) - (1 - dpois(0, 1/3))

[1] 0.2031142

qexp(0.9, 1/15)

[1] 34.53878

4.7.11

a) Let Y be the number of calls in 30 minutes, then $Y \sim Pois(3)$. We want to find $P(X \leq 3)$. This will be equivalent to 1 - [P(0) + P(1) + P(2) + P(3)] since we're subtracting the probability of getting 3 or less calls in half an hour:

$$P(0) = \frac{e^{-\lambda t}(\lambda t)^0}{0!} = e^{-3}$$

$$P(1) = \frac{e^{-\lambda t}(\lambda t)^1}{1!} = 3e^{-3}$$

$$P(2) = \frac{e^{-\lambda t}(\lambda t)^2}{2!} = \frac{9e^{-3}}{2}$$

$$P(3) = \frac{e^{-\lambda t}(\lambda t)^3}{3!} = \frac{27e^{-3}}{6}$$

$$1 - [P(0) + P(1) + P(2) + P(3)] \approx 0.3527681$$

Thus the probability of getting more than 3 calls per half hour is approximately 0.3528.

b) We're trying to find P(X > 1) = P(N=0): $P(X > 1) = e^{-\lambda t} = e^{-3} \approx 0.0498$

c) We're solving for P(X=0)=0.01 where X is the number of calls in 1 hr, thus $X \sim Pois(6)$:

$$P(X = 0) = 0.01$$

$$e^{-\lambda x} = 0.01$$

$$-\lambda x = \ln(0.01)$$

$$x = \frac{\ln(0.01)}{-\lambda}$$

$$x = \frac{\ln(0.01)}{-6}$$

$$x \approx 0.76753$$

Thus, the x hours required for the probability of no calls to be 0.01 is x = 0.76753 hours.

d) We're looking for P(X > 2) where t=2 and X is time between calls. This is equivalent to P(N = 0), where N is the number of calls in 2 hours, and $N \sim Pois(12)$:

$$P(X > 2) = P(N = 0) = e^{-12} \approx 0.0000061442123$$

Thus, the probability that we get no call in 2 hours is 0.0000061442123.

$$1 - (dpois(3, 3) + dpois(2, 3) + dpois(1, 3) + dpois(0, 3))$$

[1] 0.3527681

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dpois(0, 3)
## [1] 0.04978707
qexp(0.99, 2*3)
## [1] 0.7675284
dpois(0, 6*2)
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[1] 6.144212e-06