

Markov Processes.

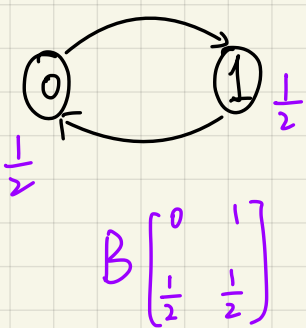
i. Recurrent, Long-Run Behavior \Rightarrow j Non-Recurrent: $\forall i, P_{ij}(n) \xrightarrow{n \rightarrow \infty} 0$

$$P_{ij}(n) \xrightarrow{n \rightarrow \infty} \pi_j \quad \text{Exist?}$$

$$\begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

Long-term behavior (limit): Process Property \downarrow . Randomness stays

(on diff. states, prob. \leftarrow freq. = $\frac{\# \text{ of this states}}{\text{length of the chain}}$)



$$P_{10}(n) = P_{01}(n) = \{1, 0, 1, 0, \dots\}$$

$$P_{00}(n) = P_{11}(n) = \{0, 1, 0, 1, \dots\}$$

does not converge (oscillating periodically)

Define a concept about the existence of $\vec{\pi} = [\pi_1, \dots, \pi_n]^T \Rightarrow$ Periodic Property

Define a concept encompassing the above example into long-term behavior.

Periodic Property: $i, d_i = \text{GCD} \{k : P_{ii}(k) > 0\}$ Greatest Common Divisor

$d_i = 1 \Rightarrow i = \text{Non-Periodic}$

Theorem.

$$i \leftrightarrow j \iff d_i = d_j \quad (S \text{ is irreducible } d \rightarrow S \mid d \text{ is a property of } S)$$

$i \leftrightarrow j, \exists m, P_{ij}(m) > 0, \exists n, P_{ji}(n) > 0 \Rightarrow P_{ii}(m+n) > 0$, Denote $\{k : P_{ii}(k) > 0\}$ as R_i

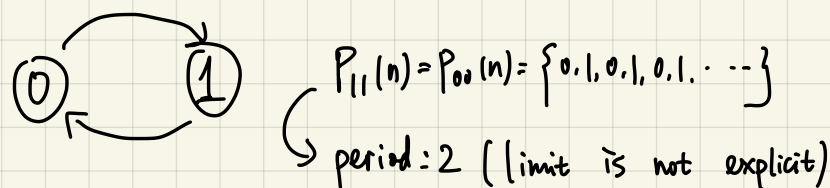
i) $P_{ii}(m+n) > 0 \Rightarrow m+n \in R_i \Rightarrow d_i \mid (m+n)$

ii) $\forall k \in R_j, P_{ii}(m+k+n) > 0 \Rightarrow m+k+n \in R_i \Rightarrow d_i \mid (m+n+k)$

iii) $d_i \mid (m+n+k) - (m+n)$, i.e. $d_i \mid k \Rightarrow \forall k \in R_j, d_i \mid k$, d_i is a common divisor of $R_j \mid d_i = \text{CD}\{R_j\}$

iv) $d_i = \text{CD}\{R_j\}, d_j = \text{GCD}\{R_j\} \Rightarrow d_i \mid d_j$

vice versa ($d_j \mid d_i$) $\Rightarrow \begin{cases} d_i \mid d_j \\ d_j \mid d_i \end{cases} \Rightarrow d_i = d_j$



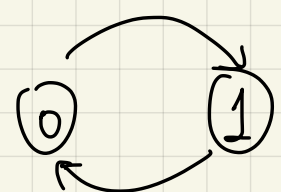
Theorem. Markov Chain S irreducible, non-periodic ($\forall i \in [1, n], d_i = 1$)

$$\Rightarrow P_{ij}(n) \xrightarrow{n \rightarrow \infty} \pi_j \quad (\text{limit exists})$$

periodic is not a general property unless S is irreducible. (if S is reducible, period is only a local property)

$\begin{cases} \text{Positive Recurrent} \\ \text{Null Recurrent} \end{cases}$

$i \xrightarrow{\text{Criterion}} \text{研究状态 } j$
 $\forall i \in \mathbb{N}, \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N P_{ij}(n) = ?$



definitely Recurrent.

稍弱

$(a_n \rightarrow a \Rightarrow \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n a_i = a)$ (做一个较弱的定义)
(weaker definition)

$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N P_{ij}(n) = \begin{cases} > 0, & j \text{ Positive Recurrent} \\ = 0, & j \text{ Null Recurrent} \end{cases}$

$P_{i0}(n) \rightarrow ? (i=0,1) \text{ diverge}$
 $\frac{1}{N} \sum_{n=1}^N P_{i0}(n) \xrightarrow{N \rightarrow \infty} \frac{1}{2}, \text{ Positive Recurrent}$

$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N P_{ij}(n)$ Cesaro Sum

Weak Ergodic

$\frac{1}{\sum_{k=1}^{\infty} k f_{ii}(k)}$

$f_{00}(k) = \{0, 1, 0, 0, \dots\} \quad \sum_{k=1}^{\infty} k f_{00}(k) = 2$

Even if Transitional Prob. is not convergent. we can go to Cesaro Sum / Weak Ergodic (Pos./Null Recurrent)

How to calculate $\lim_{n \rightarrow \infty} P_{ij}(n)$? (if limit does exist)

$\begin{cases} P(n) = P(1)P(n-1) & \text{后退方程} \\ P(n) = P(n-1)P(1) & \text{前进方程} \end{cases}$

Suppose $P_{ij}(n) \xrightarrow{n} \pi_j$

$P(n) \rightarrow \begin{bmatrix} \pi_1 & \pi_2 & \dots & \pi_n \\ \vdots & \vdots & & \vdots \\ \pi_1 & \pi_2 & \dots & \pi_n \end{bmatrix} = \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} [\pi_1 \dots \pi_n]$

前进方程更本质!

$\begin{cases} \pi = P(1) \cdot \pi = P(1) \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} [\pi_1, \dots, \pi_n] \\ \pi = \pi \cdot P(1) = \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} [\pi_1, \dots, \pi_n] P(1) \end{cases}$
not always exist

$(\begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} - P(1) \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}) (\pi_1, \dots, \pi_n) = 0$
 $\begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} ([\pi_1, \dots, \pi_n] - [\pi_1, \dots, \pi_n] P(1)) = 0$

行和为1, trivial

Let $\pi = (\pi_1, \dots, \pi_n)$ row vector

$\pi = \pi P$

$\pi \rightarrow$ Limit Distribution ($\sum_k \pi_k = 1$, finite states)

i) Stationary Distribution.

$P(X_0=k) = \pi_k \Rightarrow P(X_n=k) = \pi_k, \forall n$

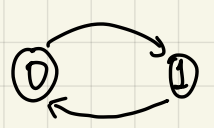
$P(X_n=k) = \sum_l P(X_n=k, X_{n-1}=l) = \sum_l P(X_{n-1}=l) P(X_n=k | X_{n-1}=l) = \sum_l P_{lk} P(X_{n-1}=l)$

Suppose $P(X_0=k) = \pi_k \Rightarrow P(X_1=k) = \sum_l P_{lk} P(X_0=l) = \sum_l P_{lk} \pi_l$

$\pi = \pi P \Rightarrow [\pi_1, \dots, \pi_k, \dots, \pi_n] = [\pi_1, \dots, \pi_n] \begin{bmatrix} P_{11} & \dots & P_{1k} & \dots & P_{1n} \\ \vdots & & \vdots & & \vdots \\ P_{n1} & \dots & P_{nk} & \dots & P_{nn} \end{bmatrix}$
 $\Rightarrow \forall k \in [1, n], \pi_k = \sum_i \pi_i \cdot P_{ik} \Rightarrow P(X_1=k) = \pi_k$

if Limit Distribution does not exist, what's the implication of $\pi (\pi = \pi P)$?

$\Rightarrow \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N P_{ij}(n)$



$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$

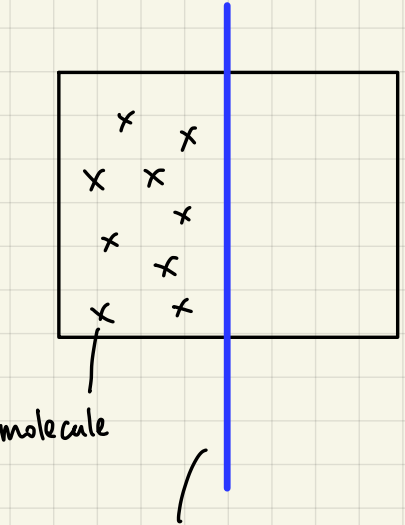
$\begin{cases} \pi = \pi P \\ \sum \pi_i = 1 \end{cases} \Rightarrow \pi = [\frac{1}{2}, \frac{1}{2}] = \lim_{N \rightarrow \infty} \frac{1}{N} \sum P_{ij}(n)$

π_j does not exist

ii) $\{\exists \pi \neq 0, \pi = \pi P\} \Leftrightarrow \{\exists \pi \neq 0, P^T \pi^T = \pi^T\} \Leftrightarrow \{\exists \pi \neq 0, (P-I)^T \pi^T = 0\} \Leftrightarrow \det(P-I) = 0 \Leftrightarrow 1 = \text{eig}(P)$
 \downarrow
 Singular
 $\{\exists \lambda = 1, \det(\lambda I - P) = 0\}$
 \uparrow
 $\{P \cdot \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} = I \cdot \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}\}$

行和为1 \Rightarrow 每行减1, 一定奇异
 $\sum_j P_{ij}(1) = 1$

Example: **Shrenfest Model** (Markov Chain: characterize molecular motion)



i) $S = \{1, 2, \dots, N\}$ # of particles. State Space: # of particles in left half space.

ii) Discrete Time: only 1 particle change its state in a time unit.

X_n : # of particles in left half on time n
 $\{X_n\}$: Markov Chain

撤去板子, 如何演化?

Assumption: equal chance of moving to the other side.

	0	1	2	...	N
0	0	1	0	...	0
1	$\frac{1}{N}$	0	$\frac{N-1}{N}$...	$\frac{1}{N}$
2	0	$\frac{2}{N}$	0	$\frac{N-2}{N}$	0
...
N	0	...	$\frac{N-1}{N}$	0	1

Transition Prob.
 P

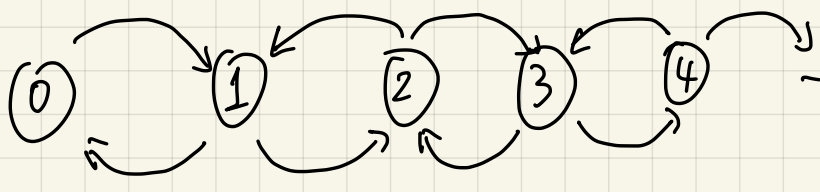
$\pi = \pi P$
 $P^T \pi^T = \pi^T$

$\pi^T = \pi^T$

$P - I = \begin{bmatrix} -1 & \frac{1}{N} & & & \\ & -1 & \frac{2}{N} & & \\ & \frac{N-1}{N} & -1 & \frac{3}{N} & \\ & & \frac{2}{N} & -1 & 1 \\ & & & \frac{1}{N} & -1 \end{bmatrix}$

$\pi_0 = \frac{1}{N} \pi_1 \quad (\pi_1 = \frac{N}{1} \pi_0)$
 $\pi_0 + \frac{2}{N} \pi_2 = \pi_1 \Rightarrow \pi_2 = \frac{N-1}{2} \pi_1$
 $\frac{N-1}{N} \pi_1 + \frac{3}{N} \pi_3 = \pi_2 \Rightarrow \pi_3 = \frac{N-2}{3} \pi_2$
 \vdots
 $\Rightarrow \pi_k = \frac{N-k+1}{k} \pi_{k-1}$
 $\pi_N = \frac{1}{N} \pi_{N-1}$

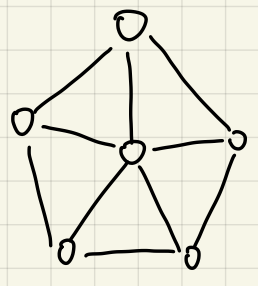
$\pi_1 = \binom{N}{1} \pi_0$
 $\pi_2 = \frac{N(N-1)}{2!} \pi_0 = \binom{N}{2} \pi_0 \quad \therefore \pi_0 = \frac{1}{\sum_{k=0}^N \binom{N}{k}} = \frac{1}{2^N}$
 $\pi_3 = \binom{N}{3} \pi_0$
 \vdots
 $\pi_N = \binom{N}{N} \pi_0$
 $\therefore i \in [0, N], \pi_i = \frac{\binom{N}{i}}{\sum_{k=0}^N \binom{N}{k}} = \frac{\binom{N}{i}}{2^N}$
 $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N P_{ij}(n)$



$d=2$, periodic \therefore Limit Distribution does not exist.

$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N P_{i0}(n) = \frac{1}{\sum_{k=1}^{\infty} k f_{00}(k)} = \pi_0 = \frac{1}{2^N}$
 Expectation of 1st Passage = 2^N , 很长
 N : 阿伏加德罗常数.

Example: Random Walk on Graph. \Rightarrow **PageRank Algorithm (Google)**



degree: 边数 d_i , prob. of each vertex = $\frac{1}{d_i}$

N nodes. $\pi_k = \frac{d_k}{\sum_i d_i}$. Examine whether $\pi = \pi P$
 $\sum_{k=1}^N \pi_k P_{kl} = \sum_{k=1}^N \frac{d_k}{\sum_i d_i} \cdot P_{kl}$
 Uniform Dist. on node k

$$= \sum_{k=1, P_{k\ell} \neq 0}^N \frac{dk}{\sum_i d_i} \cdot \frac{1}{dk} (P_{k\ell} = 0 \cdot \frac{1}{dk}, \text{ non-zero term } \# = d_\ell)$$

$$= d_\ell \cdot \frac{dk}{\sum_i d_i} \cdot \frac{1}{dk} = \frac{d_\ell}{\sum_i d_i} = \pi_\ell \quad \text{End of Proof.}$$

Detailed Balance Relation: $\pi = (\pi_0, \dots; \pi_k, \dots)$, $\pi_i P_{ij} = \pi_j P_{ji}$

$$\sum_k \pi_k P_{k\ell} = \sum_k \pi_\ell P_{\ell k} = \pi_\ell, \quad \text{Markov Chain Monte Carlo}$$

Construct P from π :

$$i) \forall P_{ij}, \quad \pi_i P_{ij} \stackrel{?}{=} \pi_j P_{ji} \quad \tilde{P}_{ij} = P_{ij} \min(1, \frac{\pi_j P_{ji}}{\pi_i P_{ij}}) \rightarrow \text{Satisfy Detailed Balance Relation.}$$

$$\pi_i \tilde{P}_{ij} = \pi_i P_{ij} \min(1, \frac{\pi_j P_{ji}}{\pi_i P_{ij}}) = \min(\pi_i P_{ij}, \pi_j P_{ji}) = \pi_j P_{ji} \min(1, \frac{\pi_i P_{ij}}{\pi_j P_{ji}}) = \pi_j \tilde{P}_{ji} \quad \text{Metropolis-Hastings Algorithms.}$$

MCMC

Next Lecture: Continuous-Time Markov Chain

Birth & Death Process