Saussian Processes X(t) Gaussian Process:  $\forall n$ ,  $X=(X(t), \dots, X(t))^T$ ,  $X \sim N(M, \Sigma)$ .  $\{M=E(X)\}$ n = 1,  $\times \sim N(M.6^2)$ .  $f_{\times}(x) = \frac{1}{\sqrt{226^2}} exp[-\frac{(x-\mu)^2}{26^2}]$  $N=2, \quad X \sim \mathcal{N}(\mu_1, \mu_2, 6.^2, 6.^2, \rho), \quad f_{X, X_2}(x, x_2) = \frac{1}{2\lambda 6.62\sqrt{1-\rho^2}} exp\left[-\frac{1}{2(1-\rho^2)}\left(\frac{x_1-\mu_1}{61}\right)^2 + \left(\frac{x_2-\mu_2}{62}\right)^2 - 2\rho\left(\frac{x_1-\mu_1}{61}\right)\left(\frac{x_2-\mu_2}{62}\right)\right]$  $Matriz-Vector Notation. X \in \mathbb{R}^n$ ,  $X \sim N(\mu, \Sigma)$ .  $\mu \in \mathbb{R}^n$ ,  $\Sigma \in \mathbb{R}^{n \times n}$ .  $f_{X}(x) = \frac{1}{(2\pi)^{\frac{n}{2}} \left[ \det \Sigma \right]^{\frac{1}{2}}} \exp \left[ -\frac{1}{2} \left[ X - \mu \right]^{\frac{1}{2}} \sum_{i=1}^{n} \left[ X - \mu \right] \right]$ Examine  $f_{X}(x)$  is indeed a distribution:  $f_{X}(x) \ge 0$ ,  $\int_{\mathbb{R}^n} f_{X}(x) dx = 1$ i)  $f_X(x) \ge 0$ .  $\sum pos. def. = ) det <math>\sum > 0$ ,  $f_X(x) \ge 0$   $\frac{1}{2} \sum_{x \in X_k^2} + \sum_{x \in X_j} k_{ij}$ ii)  $\int_{\mathbb{R}^n} f_{\chi}(x) dx = 1$ ,  $\int_{\mathbb{R}^n} f_{\chi}(x) dx = \frac{1}{(2x)^{\frac{n}{2}} [\det \Sigma]^{-1}} \int_{\mathbb{R}^n} \exp[-\frac{1}{2} [\chi - \mu]^T \sum^{-1} [\chi - \mu]] dx$ Diagonalize = 5 Covariance Matrix, = 5', 5 pos. def.  $\begin{cases} \sum_{i=1}^{\infty} \int_{0}^{\infty} \int_{0}^{\infty} \left( \sum_{i=1}^{\infty} \int_{0}^{\infty} \int_{0}$  $\frac{1}{(2\lambda)^{\frac{1}{2}}} \int_{\mathbb{R}^{n}} \exp\left[-\frac{1}{2}(x-\mu)^{T}B^{T}B(x-\mu)\right] dx = 0$   $\frac{1}{(2\lambda)^{\frac{1}{2}}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} \int_{\mathbb{R}^{n}} \exp\left[-\frac{1}{2}(x-\mu)^{T}B^{T}B(x-\mu)\right] dx = 0$   $\frac{1}{(2\lambda)^{\frac{1}{2}}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} dx$   $\frac{1}{(2\lambda)^{\frac{1}{2}}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} \left(\operatorname{det} \Sigma\right)^{\frac{1}{2}} dx$ 

 $\int dx = \frac{\partial X}{\partial Y} dY \qquad dX = \left[ \text{olet B} \right]^{-1} dY = \left[ \left( \prod_{k} \lambda_{k} \right)^{-1/2} \right]^{-1} dY = \left( \prod_{k} \lambda_{k} \right)^{1/2} dY = \left( \text{det } \sum_{k} \sum_{k} \lambda_{k} \right)^{1/2} dY = \left( \text{det } \sum_{k} \sum_{k} \lambda_{k} \right)^{1/2} dY = \left( \prod_{k} \lambda_{k} \right)^{-1/2} dY = \left($ 

3 Integration Range: Still IRn

 $\frac{1}{(2\pi)^{\frac{1}{2}}} \int_{\mathbb{R}^n} \exp(-\frac{1}{2} \sum_{k} y_k^2) \, dy, \dots \, dy_n = \frac{1}{[1]} \left[ \frac{1}{(2\pi)^{\frac{1}{2}}} \int_{\mathbb{R}^n} \exp(-\frac{1}{2} y_k^2) \, dy_k \right]$ 

= 1 - is a distribution

 $I = \int_{\mathbb{R}} \exp(-\frac{y^2}{2}) dy$   $I^2 = \int_{\mathbb{R}^2} \exp(-\frac{x^2 + y^2}{2}) dx dy \qquad [n \text{ polar coord.}]$   $I^2 = \int_0^{2\pi} \exp(-\frac{p^2}{2}) p dp dp = 2\pi. \int_0^{\infty} \exp(-\frac{p^2}{2}) p dp$   $= 2\pi$   $\Rightarrow I = (2\pi)^{-\frac{1}{2}}$ 

```
N-dimensional Characteristic Function:
               X \in \mathbb{R}^n r.v. \phi_X(w) = E[\exp(jw^T X)]
                     W = (w_1, \dots, w_n)^T
\times \sim N(M, \Sigma), \phi_{X}(w) = \int_{\mathbb{R}^n} \exp(jw^T x) \cdot f_{X}(x) dx = \frac{1}{(2\pi)^{N/2} (det \Sigma)^{N/2}} \int_{\mathbb{R}^n} \exp(jw^T x) \exp(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)) dx
 Informal: 1-Dim
                   -\frac{1}{2}(x-\mu)^{T} \sum^{-1} (x-\mu) + j w^{T} x \xrightarrow{1-\Omega m} -\frac{1}{26^{2}} (x-\mu)^{2} + j w x = -\frac{1}{26^{2}} (x-\mu)^{2} - \frac{1}{26^{2}} w^{2} + j w \mu
-\frac{1}{2}(x-\mu-j\Sigma\omega)^{T}\Sigma^{-1}(x-\mu-j\Sigma\omega)-\frac{1}{2}\omega^{T}\Sigma\omega+j\omega^{T}\mu \qquad \qquad \qquad \qquad \qquad -\frac{1}{26^{2}}(x-\mu-j6^{2}\omega)^{2}-\frac{1}{2}6^{2}\omega^{2}+j\omega\mu
 indep. to \chi.

Temahs

\frac{1}{(2\pi)^{N_1}} \frac{1}{(\text{olet } \Sigma)^{N_2}} \int_{\mathbb{R}^N} \exp\left[-\frac{1}{2}(x-\mu-j\Sigma\omega)^T \Sigma^{-1}(x-\mu-j\Sigma\omega)\right] dx = 1 \quad \therefore \quad (\chi) = \exp\left(j\omega^T \mu - \frac{1}{2}\omega^T \Sigma \omega\right)

   Linearity Property. X \in \mathbb{R}^n, X \sim \mathcal{N}(M, \Sigma), A \in \mathbb{R}^{m \times n}. Universal.
        Y= AXEIR" => Y~ NIAM, AIAT)
      \phi_{Y(W)} = E\left[\exp\left(j_{W}^{T}Y\right)\right] = E\left[\exp\left(j_{W}^{T}AX\right)\right] = E\left[\exp\left(j_{W}^{T}AX\right)\right] = \phi_{X}(A^{T}W) = \exp\left(j_{W}^{T}A^{T}W\right) = \exp\left(j_{W}^{T}A^{T}W\right)
                     = exp(jwT(AM) - \frac{1}{2}wT(A\SAT)W) => Gaussian Dist.
   · MY = AM. IY = ASAT, Y~NIAM, ASAT)
  X = (X_1, \dots, X_n)^T \sim N, X = (X_{n_1}, \dots, X_{n_k}), (n_1, \dots, n_k) \leq (1, \dots, n) X \sim N (Y margin \rightarrow normal dist.)
 \widetilde{X} = \begin{bmatrix} X_{n_1} \\ \vdots \\ X_{n_k} \end{bmatrix} = \begin{bmatrix} \vdots \\ \vdots \\ \vdots \\ X_n \end{bmatrix} = \begin{bmatrix} X_{n_1} \\ \vdots \\ X_n \end{bmatrix} = \begin{bmatrix} X_{n_1} \\ \vdots \\ X_n \end{bmatrix}
     Joint Gaussian Bourdary Gaussian

\begin{cases}
\int_{\mathbb{R}} g(x_1, x_2) dx_1 = 0 \\
\int_{\mathbb{R}} g(x_1, x_2) dx_2 = 0
\end{cases}

\sum xample: f(X_1, X_1) = \frac{1}{2\pi} exp(-\frac{1}{2}(X_1^2 + X_2^2)) + g(X_1, X_1)
                                                                                                                                                         g(x,x)= Sinx, Sin x2.
   f'(x_1, \chi_2) = \frac{1}{2\lambda} \exp\left[-\frac{1}{2}(\chi_1^2 + \chi_2^2)\right] \left[+ \sinh \chi_1 \sinh \chi_2\right] > 0
           Boundary V. Joint X
```

```
Criterion for Joint Gaussian-
     XER", X~N => { V « ER", 2 X~N}
                                                                                                                                                                                                                                                                                                                                              \int \mu_{w^{T}X} = E[w^{T}X] = w^{T}\mu_{X}
                                                                                                                                                                                                                                                                                                                                              \left( 6^{2} w^{T} \times - E \left[ w^{T} \times - w^{T} \mu_{X} \right]^{2} \right)
     \phi_{\mathbf{x}}(\mathbf{w}) = \mathbb{E}\left[\exp(j\mathbf{w}^{\mathsf{T}}\mathbf{x})\right] = \phi_{\mathbf{w}^{\mathsf{T}}\mathbf{x}}(1) = \exp(j\mathbf{w}^{\mathsf{T}}\mathbf{x}) = \exp(j\mathbf{w}^{\mathsf{T}}\mathbf{x}) = \exp(j\mathbf{w}^{\mathsf{T}}\mathbf{x})
                                                                                                                                                                                                                                                                                                                                                                        = E[M_L(X-W^{X})]_{s}
                                                                                                                                                                                                                                                                                                                                                                         = \omega^T \sum_{\mathbf{X}} \omega
     X = (X_1, \dots, X_n)^T \sim N, E[X_i X_j] = EX_i EX_j \Rightarrow X_i, X_j independence.
     \sum_{ij} = E\left[X_i - EX_i\right]\left[X_j - EX_j\right] = E\left[X_i X_j\right] - EX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i EX_j\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} is \text{ diagonal. } \left(\sum_{i=1}^{n} eX_i\right) - eX_i EX_j = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=1}^{n} eX_i = 0 \quad \text{$(i \neq j)$} \Rightarrow \sum_{i=
     f_{\chi} - \chi_n = \frac{1}{(2\pi)^{\frac{n}{2}} \left( \prod_{k} 6_k^2 \right)^{\frac{n}{2}}} \exp \left( -\frac{1}{2} \sum_{k} \frac{(\chi_k - \mu_k)^2}{6_k^2} \right)
                                                                                                                                                                                                                                                            In Gaussian.
                                                                                                                                                                                                                                                            Uncorrelated (=> Independent.
                                  = \prod_{k} \left[ \frac{1}{\sqrt{2\pi 6 k^{2}}} \exp \left[ -\frac{1}{2} \frac{(\chi_{k} - \mu_{k})^{2}}{6 k^{2}} \right] \right] = \prod_{k} f_{\chi_{k}}(\chi_{k})
PCA => ICA (Independent)

Blind Source Separation

Uncorrelated
                                                                                                                                                                                                                                                                                           AlGC
                                                                                                                                                                                                                                                                                            Diffusion Mid Journey.
                                                                                                                                                                                                                                                                                            highly relevant to Stochastic Processes
       X \sim N[0,1] \Rightarrow \widetilde{X} = \Sigma^{1/2}(X + \overline{\Sigma}^{1/2}\mu) \sim N(\mu, \Sigma)
       Xo, X, X2, ---, XNER" (Xx= 11-0x Xx-1+ 10x Ex. Ex~ N(0, I). i.i.d.)
        X_k \sim N! \rightarrow \text{Determine } X_m.!
                                                                                                                 Br= 1-xx, Xk= TBk Xx++ JI-BK Ex= JBK [Bk+ Xk-2+JI-Bk-1 Ex-1)+JI-BK Ex
                                                                                                                                                                              = JBKBK-1 XK-2 + JBKC1-BK-1) EK-1+J1-BK EK
       "Reparametric Trick"
                                                                                                                                                                                                                                                                             indep. Gaussian
                                                                                                                                                                                                            NO. BELL-BEN 1) + NO. I-BEI) => NO. I-BEBE-11)
      \frac{1}{3} = \pi \beta k
                                                                                                                                                                              = JBKBK-1BK-2 XK-2 + JI-BKBK-1BK-2 E0 =
```