

Gaussian Processes 2

Recap.

$$X \in \mathbb{R}^n, X \sim N(\mu, \Sigma), Y = AX, A \in \mathbb{R}^{m \times n}, Y \sim N(A\mu, A\Sigma A^T)$$

Example.

$$X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2), \bar{X} = \frac{1}{n} \sum_k X_k \Rightarrow \hat{\mu} \quad \bar{S} = \frac{1}{n-1} \sum_k (X_k - \bar{X})^2 \Rightarrow \hat{\sigma}^2$$

Contains \bar{X}

\bar{X} and \bar{S} are independent. given $X_k \sim N(\mu, \sigma^2)$

$$(n-1)\bar{S} = \sum_k (X_k - \bar{X})^2 = \sum_k (X_k^2 - 2\bar{X}X_k + \bar{X}^2) = \sum_k X_k^2 - 2\bar{X} \cdot n\bar{X} + n\bar{X}^2 = \sum_k X_k^2 - n\bar{X}^2$$

Construct an unitary matrix $A = \begin{bmatrix} \frac{1}{\sqrt{n}} & \frac{1}{\sqrt{n}} & \dots & \frac{1}{\sqrt{n}} \\ & * & & \\ & & \ddots & \\ & & & \text{don't care} \end{bmatrix}$ $Y_1 = \frac{1}{\sqrt{n}} \sum_k X_k$ $\Rightarrow Y_1 = \sqrt{n} \bar{X}$ $\sum_i Y_i^2 = \sum_j X_j^2$ $Y_1 \sim Y_n \text{ indep.}$ $\bar{S}_Y = A \Sigma_X A^T = A(\sigma^2 I)A^T = \sigma^2 (A I A^T) = \sigma^2 I$ $A \text{ (unitary mat.) doesn't change energy}$

$$\therefore (n-1)\bar{S} = \sum_k X_k^2 - n\bar{X}^2 = \sum_k Y_k^2 - Y_1^2 = \sum_{k=2}^n Y_k^2 \text{ indep. to } Y_1 = \sqrt{n} \bar{X}$$

Degree of Freedom: $(n-1) \Rightarrow \bar{S}$ should be divided by $(n-1)$, not n

i.e. \bar{S} indep. to \bar{X} (works only in Gaussian Distribution) (Cochran)

Why \bar{S} has only $(n-1)$ DoF? Because $\bar{S} = \sum (X_k - \bar{X})^2$ Sample mean, not actual mean
1 DoF is subtracted here

Another Important Property:

$$(X_1, X_2) \sim N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}\right) \quad X_1, \mu_1 \in \mathbb{R}^m \quad X_2, \mu_2 \in \mathbb{R}^n \quad \text{Conditional Distribution? } (X_1|X_2)$$

$$f_{X_2|X_1} = \frac{f_{X_1, X_2}}{f_{X_1}} \Rightarrow \text{Suppose to be a Gaussian Distribution, } \Sigma_{X_2|X_1} = ?, \mu_{X_2|X_1} = ?$$

$$= C \exp\left(-\frac{1}{2}(X_1^T - \mu_1^T, X_2^T - \mu_2^T) \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}^{-1} \begin{pmatrix} X_1 - \mu_1 \\ X_2 - \mu_2 \end{pmatrix} + \frac{1}{2}(X_1^T - \mu_1^T) \Sigma_{11}^{-1} (X_1 - \mu_1)\right)$$

Diagonalize

$$\begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \xrightarrow[\text{left}]{\begin{pmatrix} I & 0 \\ -\Sigma_{21}\Sigma_{11}^{-1} & I \end{pmatrix} \cdot \text{matrix}} \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ 0 & -\Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12} + \Sigma_{22} \end{pmatrix} \xrightarrow[\text{right}]{\text{matrix} \cdot \begin{pmatrix} I & -\Sigma_{11}^{-1}\Sigma_{12} \\ 0 & I \end{pmatrix}} \begin{pmatrix} \Sigma_{11} & 0 \\ 0 & \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12} \end{pmatrix}$$

$$\Rightarrow \begin{pmatrix} I & 0 \\ -\Sigma_{21}\Sigma_{11}^{-1} & I \end{pmatrix} \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \begin{pmatrix} I & -\Sigma_{11}^{-1}\Sigma_{12} \\ 0 & I \end{pmatrix} = \begin{pmatrix} \Sigma_{11} & 0 \\ 0 & \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12} \end{pmatrix}$$

$$\Rightarrow \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} = \begin{pmatrix} I & 0 \\ \Sigma_{21}\Sigma_{11}^{-1} & I \end{pmatrix} \begin{pmatrix} \Sigma_{11} & 0 \\ 0 & \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12} \end{pmatrix} \begin{pmatrix} I & \Sigma_{11}^{-1}\Sigma_{12} \\ 0 & I \end{pmatrix}$$

$$\Rightarrow \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}^{-1} = \begin{pmatrix} I & -\Sigma_{11}^{-1}\Sigma_{12} \\ 0 & I \end{pmatrix} \begin{pmatrix} \Sigma_{11}^{-1} & 0 \\ 0 & (\Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12})^{-1} \end{pmatrix} \begin{pmatrix} I & 0 \\ -\Sigma_{21}\Sigma_{11}^{-1} & I \end{pmatrix}$$

$$\mathcal{J}_{\mathcal{X}} \mathcal{K}, f_{X_2|X_1} = \left(\exp\left(-\frac{1}{2}(X_1^T - \mu_1^T, X_2^T - \mu_2^T) \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}^{-1} \begin{pmatrix} X_1 - \mu_1 \\ X_2 - \mu_2 \end{pmatrix} + \frac{1}{2}(X_1^T - \mu_1^T) \Sigma_{11}^{-1} (X_1 - \mu_1) \right)$$

$$= (X_1^T - \mu_1^T, X_2^T - \mu_2^T) \begin{pmatrix} I & -\Sigma_{11}^{-1}\Sigma_{12} \\ 0 & I \end{pmatrix} \begin{pmatrix} \Sigma_{11}^{-1} & 0 \\ 0 & (\Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12})^{-1} \end{pmatrix} \begin{pmatrix} I & 0 \\ -\Sigma_{21}\Sigma_{11}^{-1} & I \end{pmatrix} \begin{pmatrix} X_1 - \mu_1 \\ X_2 - \mu_2 \end{pmatrix}$$

$$= (X_1^T - \mu_1^T, -(X_1^T - \mu_1^T)\Sigma_{11}^{-1}\Sigma_{12} + (X_2^T - \mu_2^T)) \begin{pmatrix} \Sigma_{11}^{-1} & 0 \\ 0 & (\Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12})^{-1} \end{pmatrix} \begin{pmatrix} X_1 - \mu_1 \\ X_2 - \mu_2 - \Sigma_{21}\Sigma_{11}^{-1}(X_1 - \mu_1) \end{pmatrix}$$

$$= (X_1^T - \mu_1^T) \Sigma_{11}^{-1} (X_1 - \mu_1) + [(X_2^T - \mu_2^T) - (X_1^T - \mu_1^T)\Sigma_{11}^{-1}\Sigma_{12}] (\Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12})^{-1} [X_2 - \mu_2 - \Sigma_{21}\Sigma_{11}^{-1}(X_1 - \mu_1)]$$

$$\therefore f_{X_2|X_1} = \left(\exp\left(-\frac{1}{2} \left[\underbrace{X_2^T - \mu_2^T - (X_1^T - \mu_1^T)\Sigma_{11}^{-1}\Sigma_{12}}_{\text{mean}} \underbrace{(\Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12})^{-1}}_{\text{covariance matrix}} \left[X_2 - \mu_2 - \Sigma_{21}\Sigma_{11}^{-1}(X_1 - \mu_1) \right] \right) \right)$$

mean

covariance matrix

$E[X_2|X_1]$ is a r.v. on X_1

$$\therefore X_2|X_1 \sim \mathcal{N}(\mu_2 + \Sigma_{21}\Sigma_{11}^{-1}(X_1 - \mu_1), \underbrace{\Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12}}_{\text{pos. def.}})$$

$$X_1, X_2 \in \mathbb{R}^1, E[X_2|X_1] = \mu_2 + \frac{\sigma_{21}}{\sigma_{11}}(X_1 - \mu_1) \text{ — linear adjustment. (Gaussian = Optimal!)} \\ \underbrace{\mu_2 + \frac{\sigma_{21}}{\sigma_{11}}(X_1 - \mu_1)}_{\text{projection } X_2} \quad \underbrace{X_1 - \mu_1}_{X_1: \text{prior knowledge}}$$

$$\text{Var}(X_2|X_1) = \sigma_{22} - \frac{\sigma_{21}\sigma_{12}}{\sigma_{11}} = \sigma_{22} - \frac{\sigma_{12}^2}{\sigma_{11}} \quad (\text{Randomness } \downarrow : \text{Gain Information, Lose Uncertainty})$$

$$X_1, X_2 \text{ uncorrelated} : \sigma_{12} = \sigma_{21} = 0 \Rightarrow \begin{cases} E[X_2|X_1] = \mu_2 = E[X_2] \\ \text{Var}(X_2|X_1) = \sigma_{22} = \text{Var}(X_2) \end{cases}$$

$$\text{Example: } X_1, X_2 \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0,1) \quad E[X_1 + X_2 | X_1 - X_2] = ? \quad , \quad E[(X_1 + X_2)^2 | X_1 - X_2] = ?$$

$$\textcircled{1} (X_1 + X_2, X_1 - X_2) \sim \mathcal{N}(?, ?)$$

$$\begin{pmatrix} X_1 + X_2 \\ X_1 - X_2 \end{pmatrix} = \begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \sim \mathcal{N}\left(\begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix} \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix} I \begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix}\right) = \mathcal{N}\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \underbrace{\begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}}_{\text{diagonal}}\right)$$

diagonal $\Rightarrow X_1', X_2'$ indep.

$$E[X_1' | X_2'] = E[X_1'] = 0.$$

$$\textcircled{2} E[X_1' | X_2'] = E[X_1'] + \Sigma_{12} \Sigma_{22}^{-1} (X_2' - EX_2') = 0 + \frac{0}{2} (X_1 - X_2 - 0) = 0$$

$$X_1', X_2' \text{ indep. } E[X_1'^2 | X_2'] = E[X_1'^2] = 2$$

$E[X_2 | X_1]$ has randomness, a r.v.

$$(X_1, X_2)^T \sim \mathcal{N}\left(\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}\right), \quad \underbrace{X_2 | X_1}_{Y_{X_2|X_1}} \sim \mathcal{N}(\mu_2 + \Sigma_{21} \Sigma_{11}^{-1} (X_1 - \mu_1), \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12})$$

$$E[X_1'^2 | X_2'] = E[\underbrace{Y_{X_1'|X_2'}^2}_{\text{Conditional Variable}}] = \text{Var}(Y_{X_1'|X_2'}) + E^2(Y_{X_1'|X_2'})$$

Conditional Variable $\sim \mathcal{N}(\dots)$ conditional distribution.

$$E[(X_1 + X_2)^4 | X_1 - X_2] = ? \quad Y_{X_1'|X_2'} = X_2 + X_1 | X_2 - X_1 \sim \mathcal{N}(0, 2)$$

$$E[Y_{X_1'|X_2'}^4] \longrightarrow \text{more general, } Y \sim \mathcal{N}(0, \sigma^2), \quad E[Y^n] = \begin{cases} 0, & n = 2k-1 \\ ? & n = 2k \end{cases} \quad \int y^n \exp(-\frac{y^2}{2\sigma^2}) dy$$

$$E[Y^{2k}] = \frac{1}{\sqrt{2\pi}\sigma^2} \int_{\mathbb{R}} y^{2k} \exp(-\frac{y^2}{2\sigma^2}) dy = \frac{1}{\sqrt{2\pi}\sigma^2} \int_{\mathbb{R}} y^{2k-1} d(-\sigma^2 \exp(-\frac{y^2}{2\sigma^2}))$$

$$= \frac{1}{\sqrt{2\pi}\sigma^2} \left[\underbrace{\int_{\mathbb{R}} d(-\sigma^2 y^{2k-1} \exp(-\frac{y^2}{2\sigma^2}))}_0 - (-\sigma^2)(2k-1) \cdot \int_{\mathbb{R}} y^{2k-2} \exp(-\frac{y^2}{2\sigma^2}) dy \right]$$

$$E[Y^n] = \begin{cases} 0, & n = 2k-1 \\ (2k-1)!! \cdot \sigma^{2k}, & n = 2k \end{cases}$$

$$= \sigma^2(2k-1) \cdot \frac{1}{\sqrt{2\pi}\sigma^2} \int_{\mathbb{R}} y^{2k-2} \exp(-\frac{y^2}{2\sigma^2}) dy = \sigma^2(2k-1) \cdot E[Y^{2k-2}] = (\sigma^2)^{k-1} (2k-1)!! \cdot E[Y^2] = \sigma^{2k} \cdot (2k-1)!!$$

$$E[3X_1 + 2X_2 | 2X_1 + 3X_2] = ?$$

$$\begin{bmatrix} 3X_1 + 2X_2 \\ 2X_1 + 3X_2 \end{bmatrix} = \begin{bmatrix} 3 & 2 \\ 2 & 3 \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} \sim \mathcal{N}\left(\begin{bmatrix} 3 & 2 \\ 2 & 3 \end{bmatrix} \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 3 & 2 \\ 2 & 3 \end{bmatrix} I \begin{bmatrix} 3 & 2 \\ 2 & 3 \end{bmatrix}\right) = \mathcal{N}\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 13 & 12 \\ 12 & 13 \end{bmatrix}\right)$$

$$3X_1 + 2X_2 | 2X_1 + 3X_2 \sim \mathcal{N}(\mu_1 + \Sigma_{12} \Sigma_{22}^{-1} (X_2' - \mu_2), \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21})$$

$$X_1' | X_2' = \mathcal{N}(0 + 12 \cdot \frac{1}{13} X_2', 13 - \frac{144}{13})$$

$$= \mathcal{N}(\frac{12}{13}(2X_1 + 3X_2), \frac{25}{13})$$

$$\Rightarrow E[3X_1 + 2X_2 | 2X_1 + 3X_2] = \frac{12}{13}(2X_1 + 3X_2)$$

$$E[\cos(3X_1 + 2X_2) | 2X_1 + 3X_2] = ?$$

$$Y_{1/2} = 3X_1 + 2X_2 | 2X_1 + 3X_2 \sim \mathcal{N}(\frac{12}{13}(2X_1 + 3X_2), \frac{25}{13})$$

$$E[\cos Y_{1/2}] = E[\frac{1}{2}(\exp(jY) + \exp(-jY))]$$

$$= \frac{1}{2} \exp(-\frac{\frac{25}{13}}{2}) \cos(\frac{12}{13}(2X_1 + 3X_2))$$

$$= \frac{1}{2} \exp(-\frac{25}{26}) \cos(\frac{12}{13}(2X_1 + 3X_2))$$

$$Y \sim \mathcal{N}(\mu, \sigma^2), \quad E[\cos Y] = ?$$

Use characteristic function:

$$E[\cos Y] = E[\frac{1}{2}(\exp(jY) + \exp(-jY))] = \frac{1}{2}[\phi_Y(1) + \phi_Y(-1)]$$

$$\phi_Y = \exp(j\omega\mu - \frac{\omega^2\sigma^2}{2})$$

$$\Rightarrow E(\cos Y) = \exp(-\frac{\sigma^2}{2}) \cdot \cos \mu$$

$$X \sim \mathcal{N}(\mu_x, \Sigma_x), \quad Y = BX + z, \quad z, X \text{ independent}, \quad z \sim \mathcal{N}(0, \Sigma_z)$$

$$1) (X, z) \sim \mathcal{N}, \quad f_{X,z}(x, z) = f_X(x) \cdot f_z(z) = C \cdot \exp\left(-\frac{1}{2}(x-\mu_x)^T \Sigma_x^{-1}(x-\mu_x) - \frac{1}{2}z^T (6^{-2}I)z\right) \\ = C \cdot \exp\left[-\frac{1}{2} \begin{pmatrix} (x-\mu_x)^T & z^T \end{pmatrix} \begin{pmatrix} \Sigma_x^{-1} & 0 \\ 0 & \Sigma_z^{-1} \end{pmatrix} \begin{pmatrix} x-\mu_x \\ z \end{pmatrix}\right] \Rightarrow \text{Gaussian}$$

$$\begin{pmatrix} X \\ Y \end{pmatrix} = \begin{pmatrix} I & 0 \\ B & I \end{pmatrix} \begin{pmatrix} X \\ z \end{pmatrix} \sim \mathcal{N}$$

$$2) X|Y \sim ? \quad (X: \text{state}, \quad z: \text{noise}, \quad Y: \text{observation})$$

$$X_2|X_1 \sim \mathcal{N}(\mu_2 + \Sigma_{21}\Sigma_{11}^{-1}(x_1 - \mu_1), \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12})$$

$$E[X|Y] = \mu + \Sigma_{XY}\Sigma_{YY}^{-1}(Y - \mu_Y)$$

X, z indep.

$$\Sigma_{XY} = E[(X - \mu_X)(Y - \mu_Y)^T] = E[(X - \mu_X)(BX + z - B\mu_X)^T] = \Sigma_X B^T + E(X - \mu_X)E(z)^T = \Sigma_X \cdot B^T$$

$$\Sigma_{YY} = E[(BX + z - B\mu_X)(BX + z - B\mu_X)^T] = B\Sigma_X B^T + \Sigma_z$$

$$\therefore E[X|Y] = \mu + \underbrace{\Sigma_{XY}}_{\Sigma_{XY}} \cdot \underbrace{\Sigma_{YY}^{-1}}_{\Sigma_{YY}} (Y - B\mu_X) \quad \text{Bayesian.}$$

$$\Sigma_{X|Y} = \Sigma_X - \Sigma_{XY}\Sigma_{YY}^{-1}\Sigma_{YX} = \Sigma_X - \Sigma_X B^T (B\Sigma_X B^T + \Sigma_z)^{-1} \underbrace{B\Sigma_X^T}_{\Sigma_X}$$