Lecture 19: Multi-Armed Bandit

Course: Reinforcement Learning Theory Instructor: Lei Ying Department of EECS University of Michigan, Ann Arbor

Suppose that none of the events occurs (i.e. all three equations are false). Then we have

$$\mu_1(t-1) + \sqrt{\frac{\alpha \log(t)}{N_1(t-1)}} > \mu_1$$
 (condition 1)

$$\mu_{1} = \mu_{i} + \Delta_{i}$$

$$\geq \mu_{i} + 2\sqrt{\frac{\alpha \log(t)}{N_{i}(t-1)}}$$
(condition 3)

$$\geq \mu_1(t-1) + \sqrt{\frac{\alpha \log(t)}{N_i(t-1)}} \qquad \qquad \text{(condition 2)}$$

But then the algorithm should have picked arm 1 over arm i (contradiction).

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 \implies {arm i played at t AND (3) false} \subseteq {(1) true OR (2) true}

If arm i played at time t and (3) false, then (1) true or (2) true.

$$E[N_i(T)] = E\left[\sum_{t=1}^T \mathbb{I}\{\text{arm } i \text{ is played at time } t\}\right]$$

Define $u = \left\lceil \frac{4\alpha \log(T)}{\Delta_i^2} \right\rceil$, then

$$E[N_i(T)] \leq u + E\left[\sum_{t=u+1}^T \mathbb{I}\{\text{arm } i \text{ is played at time } t \text{ AND } N_i(t-1) \geq u\}\right]$$

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Example

Consider each realization, 1 indicates arm $\it i$ is played and 0 otherwise.

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Example

Consider each realization, 1 indicates arm i is played and 0 otherwise.

Mark first u 1's to be blue (u = 2).

$$E[N_i(T)] \leq \underbrace{u} + E[\sum_{t=u+1}^T \mathbb{I}\{\text{arm } i \text{ is played \& } \underbrace{N_i(t-1) \geq u}\}]$$
 Marked 1's Unmarked 1 requires (3) false $\iff N_i(t-1) \geq u$

Remark: " \leq " because it is possible less than u times.

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$$\begin{split} E[N_i(T)] &\leq u + E\left[\sum_{t=u+1}^T \mathbb{I}\{\text{arm } i \text{ is played \& } N_i(t-1) \geq u\}\right] \\ &= u + \sum_{t=u+1}^T \Pr(\text{arm } i \text{ played \& (3) false}) \\ &\leq u + \sum_{t=u+1}^T \Pr(\text{(1) true or (2) true}), \end{split}$$

where the last inequality holds because

$$N_i(t-1) \ge u = \left\lceil \frac{4\alpha \log(t)}{\Delta_i^2} \right\rceil \implies$$
 (3) false

Thus,

$$E[N_i(T)] \le u + \sum_{t=u+1}^T \Pr((1) \text{ true}) + \Pr((2) \text{ true})$$

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We now focus on the probability of (1) being true, i.e.

$$\Pr\left(\mu_1(t-1) + \sqrt{\frac{\alpha \log t}{N_1(t-1)}} \le \mu_1\right).$$

The difficulty of analyzing the probability above, e.g. applying Azuma-Hoeffding's inequality is $N_1(t-1)$ is a random variable and is correlated with $\mu_1(t-1)$.

Difficulty

We need to somehow decouple the randomness of generating the rewards from the algorithm itself.

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Azuma-Hoeffding Inequality

Consider Bernoulli random variables $X_i \in \{0,1\}$. Define

$$\mu_i(t) = \frac{\sum_{s=1}^t X_i(s)}{t}$$

Imagine that we are estimating μ_i by pulling arm i t times. Define $E[X_i] = \mu_i$.

Azuma-Hoeffding inequality for Bernoulli random variables:

$$\Pr(\mu_i - \mu_i(t) > \epsilon) \le e^{-\frac{t\epsilon^2}{2\mu_i}} \le e^{-\frac{t\epsilon^2}{2}}.$$

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Optimality of UCB: Example

Consider two arms and t = 10,

Generate 10 samples from Bernoulli (μ_1):

$$X_1(1), X_1(2), \ldots, X_1(10)$$

10 samples from Bernoulli (μ_2):

$$X_2(1), X_2(2), \ldots, X_2(10)$$

When arm i is played for the kth time, the reward is $X_i(k)$.

Note that $\{X_i(k)\}_{i=1,2,\,k=1,\dots,10}$ are generated once at the beginning, so are not decoupled from the MAB algorithm.

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Optimality of UCB: Example

When $\left\{\mu_1(t-1) + \sqrt{\frac{\alpha \log(t)}{N_1(t-1)}} \le \mu_1\right\}$ occurs, it means that when we take the first $N_1(t-1)(\le t-1)$ values of $X_1(k)$,

$$\frac{1}{N_1(t-1)} \sum_{k=1}^{N_1(t-1)} X_1(k) + \sqrt{\frac{\alpha \log(t)}{N_1(t-1)}} \le \mu_1.$$

A Necessary Condition

There exists $s \le t - 1$ such that

$$\frac{1}{s} \sum_{k=1}^{s} X_1(k) + \sqrt{\frac{\alpha \log(t)}{s}} \le \mu_1.$$
 (1)

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Note that for a fixed s, we can apply concentration inequality such as the Azuma-Hoeffding inequality to understand the probability of (1).

$$\Pr\left(\mu_{1}(t-1) + \sqrt{\frac{\alpha \log(t)}{N_{1}(t-1)}} \leq \mu_{1}\right)$$

$$\leq \Pr\left(\exists s, \frac{1}{s} \sum_{k=1}^{s} X_{1}(k) + \sqrt{\frac{\alpha \log(t)}{s}} \leq \mu_{1}\right)$$

$$\leq \sum_{s=1}^{t-1} \Pr\left(\frac{1}{s} \sum_{k=1}^{s} X_{1}(k) + \sqrt{\frac{\alpha \log(t)}{s}} \leq \mu_{1}\right)$$

$$< \sum_{s=1}^{t-1} e^{-s(\frac{\alpha \log(t)}{s}) \times \frac{1}{2}} = \sum_{s=1}^{t-1} e^{-\frac{1}{2}\alpha \log(t)}$$

$$= \sum_{s=1}^{t-1} \frac{1}{t^{\alpha/2}} = (t-1) \times \frac{1}{t^{\alpha/2}} \leq \frac{1}{t^{\frac{\alpha}{2}-1}}.$$

We analyze P((1) true) for given t, and a similar analysis can be done for P((2) true) for given t. We then obtain for some b > 0 and $\alpha > 4$,

$$\begin{split} E[N_i(T)] &\leq u + \sum_{t=u+1}^T \Pr((1) \text{ true}) + \Pr((2) \text{ true}) \\ &\leq \left\lceil \frac{4\alpha \log(T)}{\Delta_i^2} \right\rceil + \sum_{t=u+1}^T \frac{b}{t^{\frac{\alpha}{2}-1}} \\ &\leq \left\lceil \frac{4\alpha \log(T)}{\Delta_i^2} \right\rceil + \int_{u+1}^\infty \frac{b}{\tau^{\frac{\alpha}{2}-1}} \, d\tau \\ &= \left\lceil \frac{4\alpha \log(T)}{\Delta_i^2} \right\rceil + \frac{b}{\frac{\alpha}{2}-2} \frac{1}{(u+1)^{\frac{\alpha}{2}-2}} \\ &\leq \left\lceil \frac{4\alpha \log(T)}{\Delta_i^2} \right\rceil + o(1). \end{split}$$

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Thus we have derived a lower bound for the regret,

$$R_T \le \sum_{i=2}^K \Delta_i \left\lceil \frac{4\alpha \log(T)}{\Delta_i^2} \right\rceil + o(1)$$
$$\le 4\alpha \left(\sum_{i=2}^K \frac{1}{\Delta_i}\right) \log(T) + O(1).$$

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Thompson Sampling (Thompson '33)

Start with prior over parameters

Example: Bernoulli(μ_i) with prior $\mu_i \sim \text{Beta}(\alpha_i(0), \beta_i(0))$

Note:

$$X \sim \mathsf{Beta}(\alpha, \beta)$$

$$E[X] = \frac{\alpha}{\alpha + \beta}$$

X concentrates as $\alpha + \beta$ increases.

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Thompson Sampling (Thompson '33)

- Sample $\mu_i(t)$ from distribution Beta $(\alpha_i(t), \beta_i(t))$ for each arm.
- Select arm i(t) such that $i(t) \in \arg \max_i \mu_i(t)$
- ullet Observe the reward, and update distribution of $\mu_i(t)$

If the reward is 1,
$$\alpha_i(t+1) = \alpha_i(t) + 1$$
 $\beta_i(t+1) = \beta_i(t)$

If the reward is 0,
$$\alpha_i(t+1) = \alpha_i(t)$$
 $\beta_i(t+1) = \beta_i(t) + 1$

Regret under Thompson sampling: $O(\log T)$.

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Reference

• Chapter 2.2 of Bubeck, Sébastien, and Nicolo Cesa-Bianchi. "Regret analysis of stochastic and nonstochastic multi-armed bandit problems." Foundations and Trends® in Machine Learning 5, no. 1 (2012): 1-122.

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