

P1:

Since  $Y$  is a uniform r.v. over the interval  $[0, 1]$  and  $X$  is a Binomial r.v. with parameters  $N, Y$

$$Y \sim U(0, 1)$$

$$\therefore E[Y] = \frac{1}{2}(0+1) \\ = \frac{1}{2}$$

$$X|Y \sim B(N, Y)$$

$$\therefore E[X|Y] = NY$$

By the law of iterated Expectation:

$$\begin{aligned} E[X] &= E[E(X|Y)] \\ &= E[NY] \\ &= N \cdot E[Y] \\ &= N \cdot \frac{1}{2} \\ &= \frac{N}{2} \end{aligned}$$

P2.

We are given that  $V = U_1 + U_2 + \dots + U_N$  where  $V$  denote the time spent by the engine to find a copy  
 $U_i$  denote the time spent on website  $i$

$$U_i \sim \exp(\lambda)$$

$$P(U_i) = \lambda \cdot e^{-\lambda u_i}, u_i > 0$$

$$E[U_i] = \int_0^{\infty} u_i \cdot \lambda e^{-\lambda u_i} du_i$$

$$= \frac{1}{\lambda} \int_0^{\infty} t e^{-t} dt \quad \text{let } t = \lambda u_i$$

$$= \frac{1}{\lambda} [-e^{-t} - t e^{-t}]$$

$$= \frac{1}{\lambda}$$

$N$  is geometric r.v with  $p$

$$f(n) = p(1-p)^{n-1}, n=1, 2, 3, \dots$$

$$E(N) = \frac{1}{p}$$

$U_1, U_2, \dots$  are independent to each other

$$V = U_1 + U_2 + U_3 + \dots + U_N$$

$$\therefore E[V] = E[U_1 + U_2 + \dots + U_N]$$

$$= E_N(E[V|N]) \quad (\text{By Gem \#1: Smoothing} - E_x[E[Y|x]] = E[Y])$$

$$= E_N E[U_1 + U_2 + \dots + U_N | N]$$

$$= E_N \left( \frac{N}{\lambda} \right)$$

$$= \frac{1}{\lambda} \cdot E_N(N)$$

$$= \frac{1}{\lambda} \cdot \frac{1}{p}$$

$$= \frac{1}{\lambda p}$$

P3:

We are given that

$$f(x) = \begin{cases} 0 & x < 100 \\ \frac{100}{x^2} & x > 100 \end{cases}$$

and assume that the event of failure are independent.

By the formula of CDF,

$$F(x) = \int_{-\infty}^{\infty} f(x) dx$$

$$F(x) = \int_0^{100} 0 dx + \int_{100}^x \frac{100}{x^2} dx \quad \text{for } x > 100$$

$$= -\frac{100}{x} \Big|_{100}^x$$

$$= -\frac{100}{x} + 1$$

$$= 1 - \frac{100}{x}$$

$$\therefore P(X < 150) = 1 - \frac{100}{150}$$

$$= \frac{1}{3}$$

Let A denote that exactly 2 of the lamps fail within the first 150 hours of operation

$$P(A) = \binom{5}{2} \left(\frac{1}{3}\right)^2 \left(1 - \frac{1}{3}\right)^3$$

$$= 0.329$$

P4.

$$(a) \quad E(x) = \int_0^{\infty} P(x > t) \cdot dt \quad \text{where } x \text{ is non-negative r.v.}$$

proof.

Let  $F(x) = P(x \leq x)$  is the CDF of  $x$ , and  $f(x) = \frac{dF(x)}{dx}$  by definition of PDF and CDF

$$\text{LHS} = E(x) = \int_0^{\infty} x \cdot f(x) \cdot dx \quad \downarrow \quad dF(x) = f(x) \cdot dx$$

$$= \int_0^{\infty} x \cdot f(x) \cdot dx \quad (x \text{ is non-negative})$$

$$= \int_0^{\infty} x \, dF(x)$$

$$\text{Let } 1 - F(x) = R(x)$$

$$\therefore \text{LHS} = E(x) = \int_0^{\infty} x \, d(1 - R(x))$$

$$= - \int_0^{\infty} x \cdot dR(x)$$

$$= - \int_0^{\infty} x R(x) \cdot dx + \int_0^{\infty} R(x) \cdot dx$$

By the def of CDF, we have

$$F(0) = 0 \quad \text{and} \quad F(\infty) = 1$$

$$\Rightarrow R(0) = 1 \quad R(\infty) = 0$$

$$\Rightarrow - \int_0^{\infty} x R(x) \cdot dx = 0$$

$$\therefore \text{LHS} = E(x) = \int_0^{\infty} R(x) \cdot dx$$

$$= \int_0^{\infty} P(x > x) \cdot dx$$

$$= \int_0^{\infty} P(x > t) \cdot dt$$

4(b) we are given the CDF of  $x$   $P(X \leq x) = 1 - e^{-x^2}$  for  $x \geq 0$

Since  $X$  here is non-negative r.v., we can use part (a)

$$E(X) = \int_0^{\infty} P(X > t) \cdot dt$$

$$= \int_0^{\infty} (1 - P(X \leq t)) \cdot dt$$

$$= \int_0^{\infty} 1 - (1 - e^{-t^2}) \cdot dt$$

$$= \int_0^{\infty} e^{-t^2} \cdot dt$$

$$\text{let } x = t^2 \quad dx = 2t \cdot dt \Rightarrow dt = \frac{1}{2t} \cdot dx = \frac{1}{2\sqrt{x}} \cdot dx$$

$$\therefore E(X) = \int_0^{\infty} \frac{1}{2\sqrt{x}} \cdot e^{-x} \cdot dx$$

$$= \frac{1}{2} \int_0^{\infty} \underbrace{x^{\frac{1}{2}-1}} e^{-x} \cdot dx \quad (\text{By integral of Gamma function})$$

$$= \frac{1}{2} \Gamma\left(\frac{1}{2}\right)$$

$$= \frac{1}{2} \sqrt{\pi}$$

Ps.

(a) We are given that  $X$  and  $Y$  are independent r.v. with  $\text{Var}(X)=1$  and  $\text{Var}(Y)=2$

Let  $Z = 3X + 4Y$ .

By the property of Variance.

$$\begin{aligned}\text{Var}(Z) &= \text{Var}(3X + 4Y) \\ &= 3^2 \text{Var}(X) + 4^2 \text{Var}(Y) \\ &= 9 \cdot 1 + 16 \cdot 2 \\ &= 41\end{aligned}$$

(b)  $Y$  is r.v. which is uniformly distributed over the set  $\{1, 2, 3\}$

$$E[X|Y=i] = i \quad \text{and} \quad E[X^2|Y=i] = i^2 + 1$$

By formula of Variance.

$$\begin{aligned}\text{Var}(X|Y=i) &= E[X^2|Y=i] - E[X|Y=i]^2 \\ &= i^2 + 1 - i^2 \\ &= 1\end{aligned}$$

$$\begin{aligned}E[Y] &= \sum_i y \Pr(Y) \quad \text{where } \Pr(Y=i) = \frac{1}{3} \text{ for } i \in \{1, 2, 3\} \text{ since } Y \text{ is uniformly distributed} \\ &= \frac{1}{3} \cdot 1 + \frac{1}{3} \cdot 2 + \frac{1}{3} \cdot 3 \\ &= 2\end{aligned}$$

Similarly,  $E[Y^2] = \sum_i y^2 \Pr(Y)$

$$\begin{aligned}&= \frac{1}{3} \times 1^2 + \frac{1}{3} \times 2^2 + \frac{1}{3} \times 3^2 \\ &= \frac{14}{3}\end{aligned}$$

$$\begin{aligned}\Rightarrow \quad \therefore \text{Var}(Y) &= E[Y^2] - E[Y]^2 \\ &= \frac{14}{3} - 2^2 \\ &= \frac{2}{3}\end{aligned}$$

By Law of Total Variance.

$$\begin{aligned}\text{Var}(X) &= \text{Var}(E(X|Y=i)) + E(\text{Var}(X|Y=i)) \\ &= \text{Var}(Y) + E[1] \\ &= \text{Var}(Y) + 1 \\ &= \frac{2}{3} + 1 \\ &= \frac{5}{3}\end{aligned}$$

5(c) Since  $X_1, X_2, \dots, X_n$  are independent r.v. and  $\text{Var}(X_1) = \text{Var}(X_2) = \dots = \text{Var}(X_n) = \sigma^2$

$$\text{Thus, } \text{Var}(Y) = \text{Var}\left(\frac{1}{n}(X_1 + X_2 + \dots + X_n)\right)$$

$$= \text{Var}\left(\frac{1}{n}X_1 + \frac{1}{n}X_2 + \dots + \frac{1}{n}X_n\right)$$

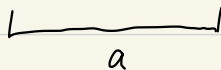
$$= \frac{1}{n^2} \text{Var}(X_1) + \frac{1}{n^2} \text{Var}(X_2) + \dots + \frac{1}{n^2} \text{Var}(X_n) \quad (\text{cov}(X_i, X_j) = 0 \quad \forall i \neq j)$$

$$= \frac{1}{n^2} (n \cdot \sigma^2)$$

$$= \frac{\sigma^2}{n}$$

Pb:

(a) Assume the length of yard stick is  $a$   
 A point is chosen at random on the yard stick  
 $X \sim U(0, a)$ ,  $f(x) = \frac{1}{a}$  for  $0 < x < a$



①

If The length of shorter piece is  $x$ ,

$$\begin{aligned} P\left(\frac{\text{shorter}}{\text{longer}} < \frac{1}{4}\right) &= P\left(\frac{x}{a-x} < \frac{1}{4}\right) \\ &= P(4x < a-x) \\ &= P(5x < a) \\ &= P\left(x < \frac{a}{5}\right) \\ &= \int_0^{\frac{a}{5}} \frac{1}{a} \cdot dx \\ &= \frac{1}{a}x \Big|_0^{\frac{a}{5}} \\ &= \frac{1}{5} \end{aligned}$$

②

If the length of shorter piece is  $a-x$

$$\begin{aligned} P\left(\frac{\text{shorter}}{\text{longer}} < \frac{1}{4}\right) &= P\left(\frac{a-x}{x} < \frac{1}{4}\right) \\ &= P(4a-4x < x) \\ &= P(4a < 5x) \\ &= P\left(x > \frac{4}{5}a\right) \\ &= 1 - P\left(x \leq \frac{4}{5}a\right) \\ &= 1 - \int_0^{\frac{4}{5}a} \frac{1}{a} dx \\ &= 1 - \frac{1}{a}x \Big|_0^{\frac{4}{5}a} \\ &= \frac{1}{5} \end{aligned}$$

Overall, by the law of probability addition,

$$P\left(\frac{\text{shorter}}{\text{longer}} < \frac{1}{4}\right) = \frac{1}{5} + \frac{1}{5} = \frac{2}{5}$$



(b)  $Y$  is uniformly distributed r.v. over  $(0, 5)$   
 $Y \sim U(0, 5)$   $f(y) = \frac{1}{5}$  for  $0 < y < 5$

$$4x^2 + 4xY + Y + 2 = 0$$

By formula of solution to equation,

$$x = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a}$$

$$= \frac{-8Y \pm \sqrt{16Y^2 - 4 \cdot 4(Y+2)}}{8}$$

$$= \frac{-8Y \pm \sqrt{16Y^2 - 16Y - 32}}{8}$$

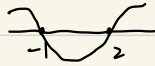
If the roots need to be real, then  $16Y^2 - 16Y - 32 \geq 0$

$$16Y^2 - 16Y - 32 \geq 0$$

$$Y^2 - Y - 2 \geq 0$$

$$(Y-2)(Y+1) \geq 0$$

$$\therefore Y \leq -1 \text{ or } Y \geq 2$$



But  $0 < Y < 5$ ,  $\therefore Y \geq 2$

$$P(Y \geq 2) = 1 - P(Y < 2)$$

$$= 1 - \int_0^2 \frac{1}{5} \cdot dx$$

$$= 1 - \frac{1}{5} \times 10^2$$

$$= 1 - \frac{2}{5}$$

$$= \frac{3}{5}$$

$\therefore$  the probability of having real roots is  $\frac{3}{5}$