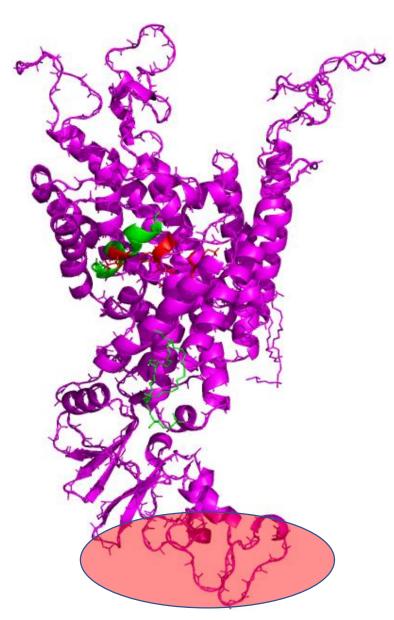
Data analysis challenge

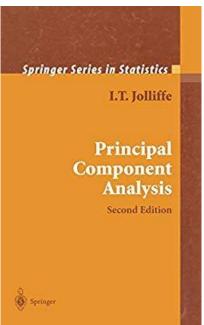
- · Aim: to understand what is «wrong» with the mutated protein
- Why? We need a way of comparing movements among the different molecular simulations.
- How?
 - Because this analysis requires a lot of processing, you will have access to Power9 at BSC.
 - 3 main steps: generate simplified trajectories, perform analysis, and visualize the data.
 - · Example scripts are provided for 2 analyses:
 - · Principal Component Analysis (PCA) to see the most dominant motions (with most variance)
 - · Autocorrelation Map to see correlated or anti-correlated motions
- Questions:
 - Is it possible to parallelize the calculations?
 - · How can PCA be used to:
 - · compare the movement of different chlorine and protonation states («healthy» or mutant protein)?
 - · compare the movements of the «healthy» with the mutated protein?
 - Can you find a graphical way to easily compare 2 autocorrelation maps? e.g.
 - · Comparison of the same chlorine state but a different protonation state for the «healthy» protein
 - · Comparison of the same chlorine state but a different protonation state for the mutant protein
 - · Comparison of the same chlorine and protonation state of «healthy» protein vs. mutant

Could PCA be improved in this analysis of simulations?

The loop which has larger contribution in the PCA is also located farthest from the center of coordinates (in the center of protein) and PCA analysis doesn't show other regions significantly contributing.



PCA textbooks recommend scaling & centering when variances are different between variables



For instance:

$$x_i' = \frac{(x_i - \bar{x})}{std. \, dev_i}$$

A User's Guide to **Principal Components**

WILEY

J. EDWARD JACKSON

In Section 1.6, the matter of scaling that context, scaling referred to the sc obtain the pc's. In particular, we had

Scaling of Data

CHAPTER 3

3.1 INTRODUCTION

WILEY SERIES IN PROBABILITY AND STATISTICS

1. U-vectors that were orthonormal and that produced pc's, z, with variances equal to the corresponding characteristic root

2. V-vectors whose coefficients were in the same units as the data and whose pc's, L1/2z, had variances equal to the square of the characteristic roots

3. W-vectors whose pc's, y, had unit variance

It was pointed out that each of these three sets of pc's conveyed exactly the same amount of information and the choice of which one to use was strictly a

matter of the use to which it was This chapter is also concerned data which are being scaled and employed. Once the method of proceed, for the most part, as des some modifications unique to eac be on the matrix from which the the purpose of this chapter to desc why each might be employed, and in Chapters 1 and 2 that might b Specificially, the following thre

2. Even if the original variables are in the same units, the variances may differ widely, often because they are related to their means. If this gives undue weight to certain variables, the correlation matrix should be employed here also unless, possibly, taking logs of the variables or the use of some other variance-stabilizing transformation will suffice). Harris (1985), in discussing some tests for the homogeneity of variance of correlated variables, recommended the use of these procedures to decide whether or not a covariance matrix may be employed.

1. No scaling at all. The final variate vector is x.

2. Scaling the data such that each variable has a zero mean (i.e., in terms

of deviation from the mean). The final variate vector is $\mathbf{x} - \hat{\mathbf{x}}$

3. Scaling the data such that each variable is in standard units. (i.e., has zero mean and unit standard deviation). Each variable is expressed as

2. Properties of Population Principal Components

the original variables rearranged in decreasing order of the size of their variances. Also, the first few PCs account for little of the off-diagonal elements of Σ in this case (see Property A3) above. In most circumstances, such a transformation to PCs is of little value, and it will not occur if the relation rather than coverience matrix

The example has shown that it is unwise to use PCs on a covariance matrix when x consists of measurements of different types, unless there is a strong conviction that the units of measurements chosen for each element of x are the only ones that make sense. Even when this condition holds, using the covariance matrix will not provide very informative PCs if the variables have widely differing variances. I irthermore, with covariance matrices and variables the PC scores are difficult to interpret—what does it mean to add a temperature to a weight? For correlation matrices, the standardized variates are all dimensionless and can be happily combined to give PC scores (Legendre and Legendre, 1983, p. 129).

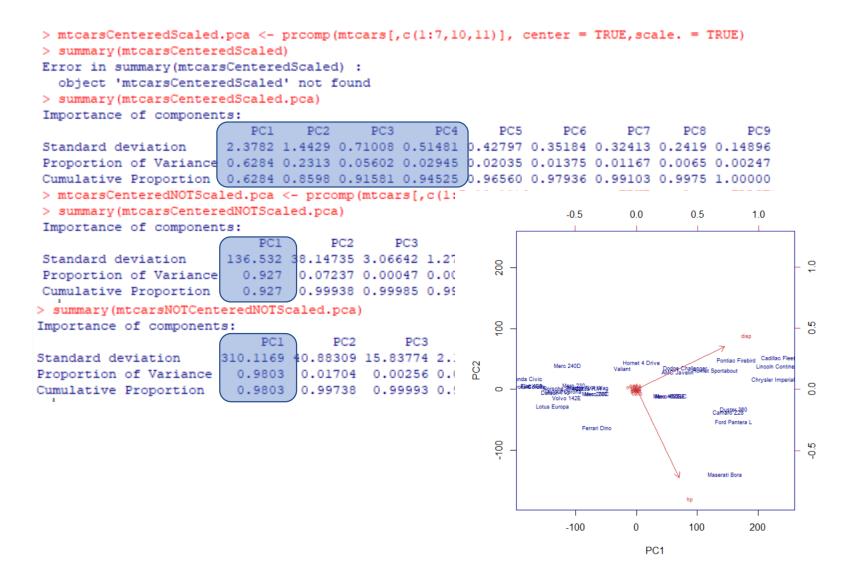
When they are not centered and scaled, fewer PC dominate and mask the effect of other PCs

Dataset of cars (included by default in R package)

> summary(mtcars)]	
mpg	cyl	disp	hp	mpg	Miles/(US) gallon
Min. :10.40	Min. :4.000	Min. : 71.1	Min. : 52.0	cyl	Number of cylinders
1st Qu.:15.43	lst Qu.:4.000	1st Qu.:120.8	1st Qu.: 96.5	""	Number of cylinaers
Median :19.20	Median :6.000	Median :196.3	Median :123.0	disp	Displacement
Mean :20.09	Mean :6.188	Mean :230.7	Mean :146.7	uisp	(cu.in.)
3rd Qu.:22.80	3rd Qu.:8.000	3rd Qu.:326.0	3rd Qu.:180.0		
Max. :33.90	Max. :8.000	Max. :472.0	Max. :335.0	hp	Gross horsepower
drat	wt	qsec	VS	drat	Rear axle ratio
Min. :2.760	Min. :1.513	Min. :14.50	Min. :0.0000	urat	iteal axic ratio
lst Qu.:3.080	lst Qu.:2.581	1st Qu.:16.89	lst Qu.:0.0000	wt	Weight (1000 lbs)
Median :3.695	Median :3.325	Median :17.71	Median :0.0000		1/4 mile time
Mean :3.597	Mean :3.217	Mean :17.85	Mean :0.4375	qsec	1/4 mile time
3rd Qu.:3.920	3rd Qu.:3.610	3rd Qu.:18.90	3rd Qu.:1.0000		Engine (0 = V-
Max. :4.930	Max. :5.424	Max. :22.90	Max. :1.0000	vs	shaped, 1 =
am	gear	carb	1	•	straight)
Min. :0.0000	Min. :3.000	Min. :1.000			
1st Qu.:0.0000	1st Qu.:3.000	1st Qu.:2.000			Transmission (0 =
Median :0.0000	Median :4.000	Median :2.000		am	automatic, 1 =
Mean :0.4062	Mean :3.688	Mean :2.812			manual)
3rd Qu.:1.0000	3rd Qu.:4.000	3rd Qu.:4.000			Number of forward
Max. :1.0000	Max. :5.000	Max. :8.000		gear	gears
					Number of
				carb	carburetors

Removed (binary/categoric data)

When they are not centered and scaled, fewer PC dominate and mask the effect of other PCs



Cpptraj can do it?

We couldn't find a way of transforming the coordinates like this:

$$x_i' = \frac{(x_i - \bar{x})}{std. \, dev_i}$$

Nor could find a way of calling diagmatrix in cpptraj so that it does it automatically. Neither we could find a way to do it using GROMACS

Our approach: DIY

We did export each "frame" of the simulation in text-readable format:

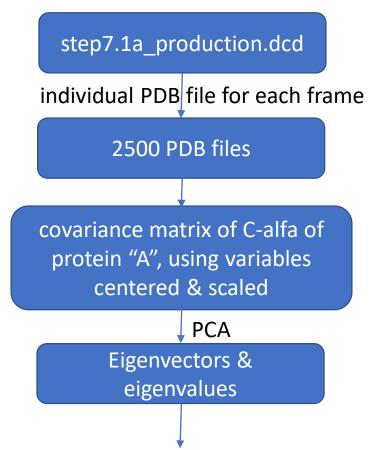
- 2500 PDB files (23 Mb), each with two proteins and many chlorine, sodium, lipids, Water molecules...

We removed chlorine, sodium, lipids, water, the protein "B", and only kept Carbon-alpha of protein A.

Prepared a table where each cell's value was centered & scaled

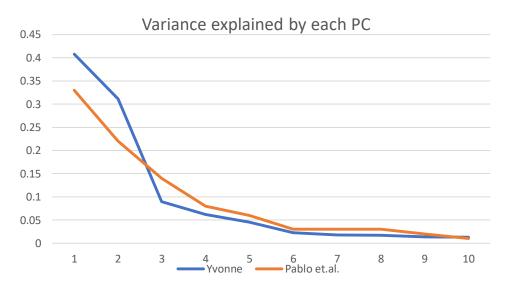
	x1	y1	z1	x2	y2	z2
T1						
T2						
•••						
T2500						

Our approach: DIY

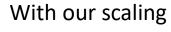


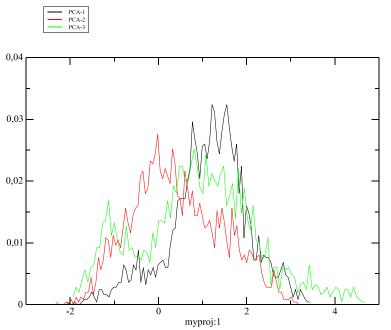
Convert eigenvector and eigenvalues into cpptraj format and load them into Cpptraj, and do the same analysis after PCA

With centering & scaling, more principal components are needed to describe the C-alfa movements in the simulation (more information is captured by PCs)



Our results: histogram of projection using PCA eigenvectors





Original (without scaling)

