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Topic 07: Regression1

Part 01: Regression - Overview

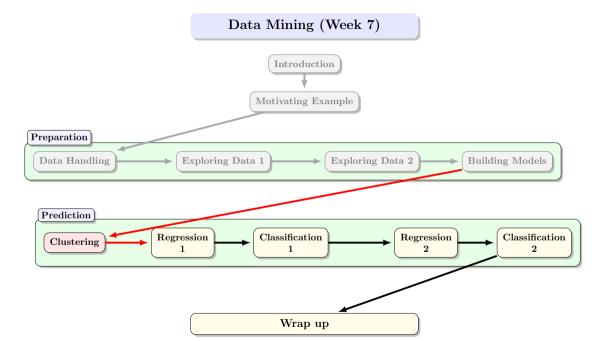
Dr Bernard Butler

Department of Computing and Mathematics, WIT. (bernard.butler@setu.ie)

Autumn Semester, 2024

Outline

- Regression as a means of minimising sum of the squared errors
- Regression assumptions what they mean, how they can be used for validation and model building
- Case studies from Diamond sales



Regression - Overview — Summary

1. Introduction

2. Linear regression assumptions

3. Reviewing regression results

- 4. Case Study 2: Diamond
- 4.1 Review

This week's aim is to give an overview of the linear regression: fitting linear models to data, to predict a numeric value.

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 - Diamond dataset: predicting diamond prices given their weights
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 - Advertising dataset: predicting widgets sold based on spending in different advertising channels
 - Credit dataset: predicting credit balance using income, status, etc.

Simple Linear Regression: Background

- Linear regression was discovered by Gauss and others around 1800. The "name" came later!
- With small data sets, calculations can be done by hand, but they are tedious and error-prone.
- The goal is simple: Given a training set of (x, y) data where y is assumed to have a linear relationship with x
 - Find the line that is the "best fit" to that data
 - Use the specification of that line to *predict* y for the test x values
- Note that the "linear relationship" of y upon x is just one of the underlying assumptions
- In practice, the data does not have an exact linear relationship, but it should be "close enough"—but what does that mean?

Definition 1 (Linear Combination of two vectors)

Given two vectors \mathbf{a} and \mathbf{b} , each with n elements, the linear combination (c) of \mathbf{a} and \mathbf{b} is

$$c \equiv a_1b_1 + a_2b_2 + \ldots + a_nb_n = \sum_{i=1}^n a_ib_i \equiv |\mathbf{a}||\mathbf{b}|\cos(\mathbf{a},\mathbf{b})$$

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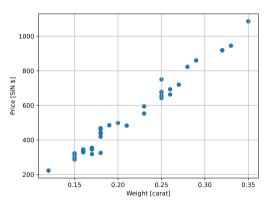
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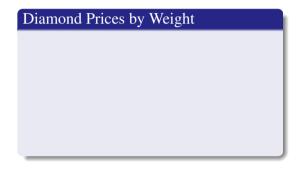
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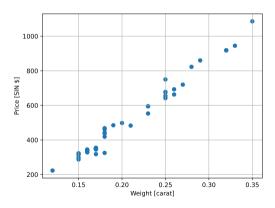
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- Linear combinations are used for prediction from linear (regression) models.







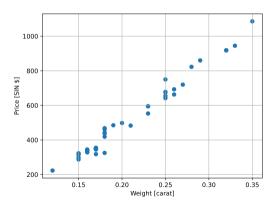
Relation between diamonds' price and weight



Diamond Prices by Weight

• Given the data on the left, can we use it to predict the price of a diamond that weighs 0.22 carat?

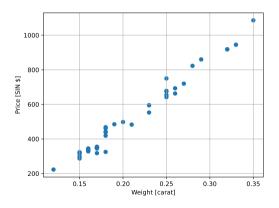
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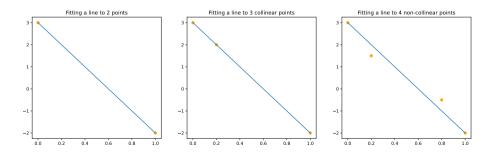
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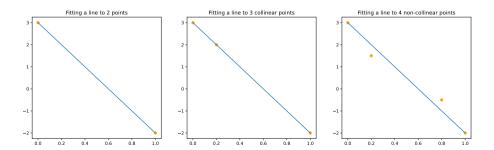
Diamond Prices by Weight

- Given the data on the left, can we use it to predict the price of a diamond that weighs 0.22 carat?
- NB we have not seen a diamond with that weight before in the data
- Can you think of at least 4 other factors that might affect the price?

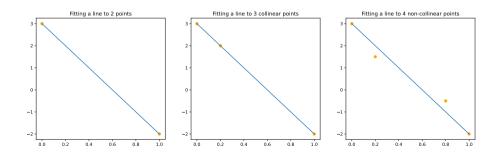
• Given data $\{x_i, y_i\}$ where i = 2, 3, ..., n and β_0, β_1 as the (unknown, but to be determined) *intercept* and *slope* of the regression line for this data.



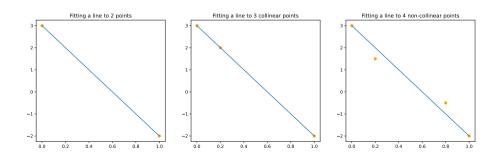
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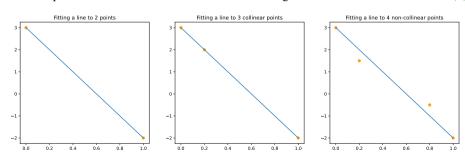
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- Otherwise the problem is overdetermined so need a more general formulation to solve for β_0, β_1 .



Definition 2 (Matrix formulation)

• General equation is $y_i = \beta_0 + \beta_1 x_i + \epsilon_i = \hat{y}_i + \epsilon_i$ (data = model + error), where \hat{y} is the predicted y for these values of β_0, β_1 .

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- So: our task is to solve the *overdetermined* (number of rows exceeds the number of columns) system of equations $y = X\beta$ for β
- Our geometric intuition is that the errors should be "balanced": no benefit to changing intercept (sliding up or down) or slope (tilting the line).

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$$\mathbf{y} = X\beta + \epsilon$$

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because $X^{\mathsf{T}} \epsilon \equiv 0$ implies the fitted line gives balanced errors and so is 'best'. Swapping sides, we have

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Note that everything on the right is a set of operations on the data.

For more info, and an alternative construction of the Normal equations, see https://goo.gl/TbLru3.

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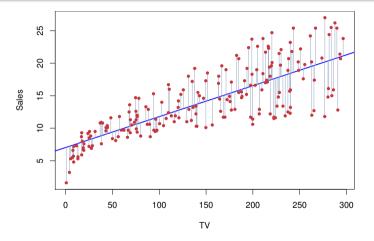
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Remember: after *learning* the β parameters using the training data $\{\mathbf{x}_i, y_i\}$, with the model encoded in the feature matrix X, it is then possible to predict \hat{y}_k for "new" (test) \mathbf{x}_k values, using separate *prediction* function calls.

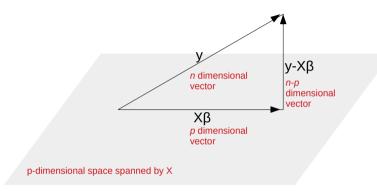
SLR: Residual Plot for the model



Source: ISLR, Fig 3.1: Advertising data with the model "Sales $\sim TV$ ".

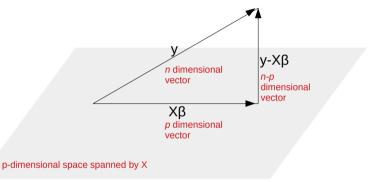
Note the vertical distance between the red dots (data points) y and the corresponding \hat{y} on the regression line, which is termed the *error* ϵ .

Geometrical interpretation of regression: n rows, p features, n > p



- Analogy: achieving photorealism with a limited palette of colours.
- Grey plane represents all the colours mixable from those colours.
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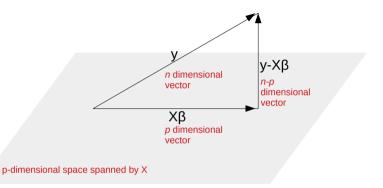
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This decomposition of n data dimensions (observations) into p model parameters and n residuals with rank n-p is helpful when interpreting regression diagnostics.

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According to the Gauss-Markov theorem, *Ordinary Least Squares* (OLS), which uses the Normal equations to minimise the sum of the squares of the errors ($\|\epsilon\|_2 \equiv \sqrt{e_1^2 + e_2^2 + e_3^2 + \ldots + e_n^2}$), is the *Best, Linear, Unbiased, Estimator* of that model that can be derived from the training data, provided some reasonably loose assumptions hold.

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In the rest of this lecture, we will generalise from Simple to Multiple Linear Regression, where $\beta = (\beta_0, \beta_1, \beta_2, \dots, \beta_p)$ and $2 \le p \le n$, so instead of fitting lines, we fit (hyper)planes to data.

Definition 4 (Linear Regression Assumptions)

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Because these assumptions depend both on the data and on the model fitted to that data, it is meaningless to say that "Data set A does not satisfy the linear regression assumptions", because this observation might not apply to all formulations of all models applied to that data.

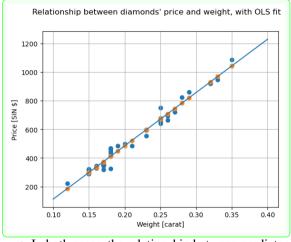
Definition 4 (Linear Regression Assumptions)

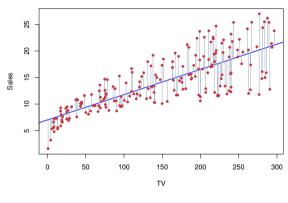
- The underlying relationship between the predictors and the response is linear in the regression parameters β .
- **②** The residual errors ϵ are drawn from a (multivariate) Normal distribution $N(\mu, \sigma^2)$ where $\mu = 0$.
- **1** The predictors are not pairwise collinear, i.e., each pair of predictors β_{j_1} and β_{j_2} (associated with columns $X(:,j_1)$ and $X(:,j_2)$) have low correlation (equivalently, the inner product of $X(:,j_1)$ and $X(:,j_2)$ is far from zero).
- There is no auto-correlation in y: each observation is independent of its "neighbours".
- **1** The errors are *homoscedastic* (i.e., $Var(\epsilon)$ is constant over the range of x or y).

Because these assumptions depend both on the data and on the model fitted to that data, it is meaningless to say that "Data set A does not satisfy the linear regression assumptions", because this observation might not apply to all formulations of all models applied to that data.

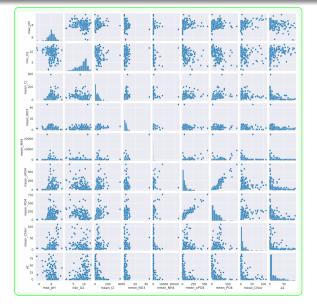
Consequently, these assumptions can be used constructively, when model building, or as checks, when validating models.

Linear relationship

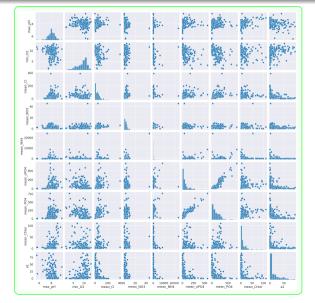




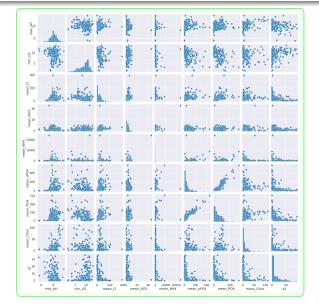
- In both cases, the relationship between predictor (feature) and target is approximately linear.
- Given a feature value, we can predict the target value using a simple linear formula.
- The predicted parameters are the *intercept* β_0 and *slope* β_1 of the line.
- Usually the vertical distance between a data point x, and its predicted value \hat{y}_i is $\epsilon_i \neq 0$



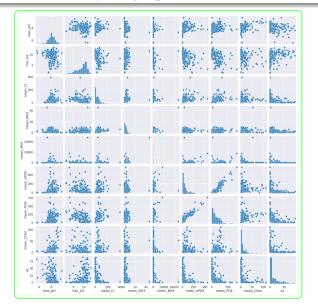
• The pairplot confirms what we saw in the corresponding correlation matrix: mean_PO4 and mean_OPO4 are highly correlated with each other (indeed, the relevant scatterplots indicate a strong linear relationship).



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- Either mean_PO4 or mean_OPO4 can be included in the model, but not both of them.



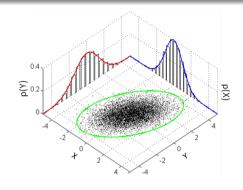
- The pairplot confirms what we saw in the corresponding correlation matrix: mean_PO4 and mean_OPO4 are highly correlated with each other (indeed, the relevant scatterplots indicate a strong linear relationship).
- Either mean_PO4 or mean_OPO4 can be included in the model, but not both of them.
- Also, the individual predictors do not have a strong linear relationship with a1 (look at the scatterplots in the last row and column) so, on their own, they are not likely to predict a1 well with a linear model.

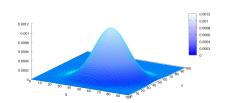


- The pairplot confirms what we saw in the corresponding correlation matrix: mean_PO4 and mean_OPO4 are highly correlated with each other (indeed, the relevant scatterplots indicate a strong linear relationship).
- Either mean_PO4 or mean_OPO4 can be included in the model, but not both of them.
- Also, the individual predictors do not have a strong linear relationship with a1 (look at the scatterplots in the last row and column) so, on their own, they are not likely to predict a1 well with a linear model.
- However, it is still possible that a combination of predictors might predict a1 well.

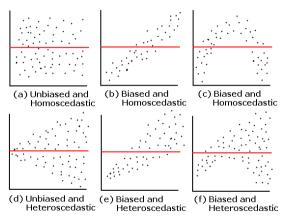
Errors are normally distributed

- Centred on zero so small errors are more common
- Symmetric so positive and negative balance out
- Normal distribution is a also called the Gaussian distribution and is "bell-shaped".





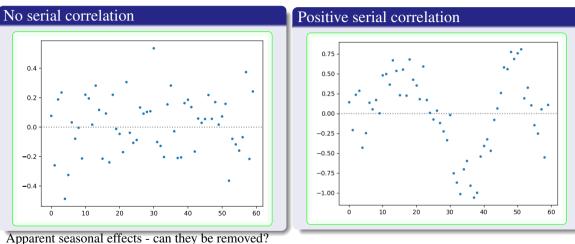
Bias and variance in regression



Source: https://bit.ly/3vC9zK7

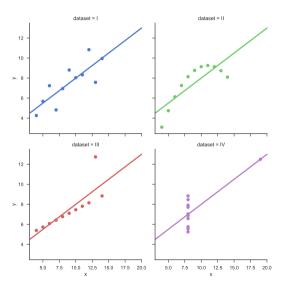
- Bias is caused by underfitting.
- Fix bias by adding suitable predictors.
- Overfitting causes large variance.
- If variance changes over the range, some errors get undue attention.
- Fix this by weighting the errors so they are balanced.

Errors should not be serially correlated



- Add feature to the model
- ② Include autoregressive terms (but thenr it is no longer Ordinary Least Squares (OLS)!)

Anscombe's quartet (1973)



Francis Anscombe devised 4 data sets to show different forms of misalignment between data and models. Sets I,II,III share the same *x* values. All 4 sets share approximately the same descriptive statistics (mean and variance), but little else is common to all 4!

Only I appears suited as it stands. The other data sets require some work, particularly IV.

What do you think needs to be done for each data set?

Common Cost Functions in Regression Models

Remember: we are trying to minimise a loss function based on the error, which we approximate with the residuals of the training set.

Common Cost Functions in Regression Models

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Measure	Definition	Purpose
Mean square error (MSE)	$\frac{(p_1-a_1)^2+\cdots+(p_m-a_m)^2}{m}$	Mathematically tractable but places greater emphasise on observations with large error
Root mean square error (RMSE)	$\sqrt{\frac{(p_1-a_1)^2+\cdots+(p_m-a_m)^2}{m}}$	Has same units as data
Mean absolute error (RMAE)	$\frac{ p_1-a_1 +\cdots+ p_m-a_m }{m}$	Does not overemphasise observa- tions with large error (like MSE
		does)
Relative square error (RSE)	$\frac{(p_1-a_1)^2+\cdots+(p_m-a_m)^2}{(p_1-\bar{a})^2+\cdots+(p_m-\bar{a})^2}$	Relative metric compares the
Root Relative square error (RRSE)	$\sqrt{\frac{(p_1-a_1)^2+\cdots+(p_m-a_m)^2}{(p_1-\bar{a})^2+\cdots+(p_m-\bar{a})^2}}$	error in the predictions with er- rors in the simplest model pos- sible (a model just always pre-
Relative absolute error (RAE)	$\frac{ p_1 - a_1 + \dots + p_m - a_m }{ p_1 - \bar{a} + \dots + p_m - \bar{a} }$	dicting the average value of y)

where a_j is the actual value, p_j is the predicted value, m is the number of observations, and \bar{a} represents the mean of the a_i .

Choices of Vector norms

Definition 5 (Manhattan norm)

 $\ell_1(\ldots) = \|\ldots\|_1$ is the *Manhattan* norm (length) of a vector. Let $\mathbf{x} = (x_1, x_2, \ldots, x_m)$. Then $\ell_1(\ldots) = \|\ldots\|_1 = |x_1| + |x_2| + \ldots + |x_m|$ is the *Manhattan* distance of \mathbf{x} from the origin. Think of having to *walk* from one junction in Manhattan to another, the distance is the difference in the Street numbers plus the difference in the Avenue numbers.

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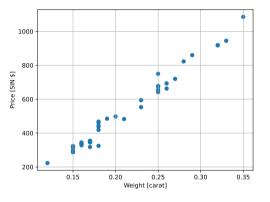
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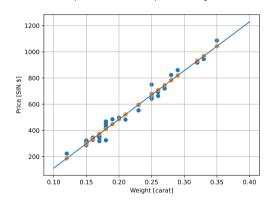
The Euclidean norm is very common, but the Manhattan norm is gaining popularity, because it is robust to outliers and computers are becoming powerful enough. However we generally use Euclidean norm in this module.

Case Study 2: Diamonds - Check relationship





Relationship between diamonds' price and weight, with OLS fit



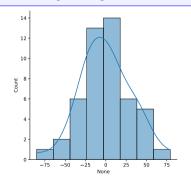
Clearly there is a linear relationship between a diamond's weight (in carats) and its price (in Singapore dollars, as here). So that is one assumption satisfied!

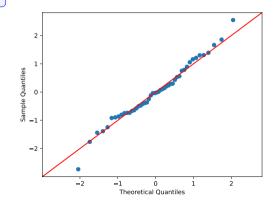
Sometimes the dependent variable has a linear dependence on a function of an attribute. Example functions include log, exp, sqrt, polynomial, etc. Even if the function is nonlinear in the attribute, that does not matter, as long as the model is linear in the regression parameters β .

Case Study 2: Diamonds - Check residual distribution

import seaborn as sns
resFig = "res/residHist.pdf"
sns_plot = sns.displot(x = residuals, kde=True)
sns_plot.savefig(resFig)

```
# Q-Q plot to verify the residuals distribution
resFig = "res/residualsqq.pdf"
fig = sm.qqplot(residuals, fit=True, line = '45')
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```

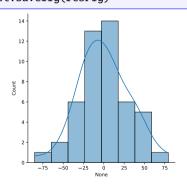


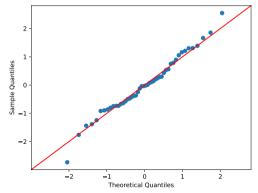


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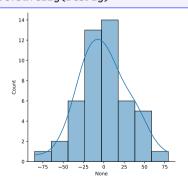


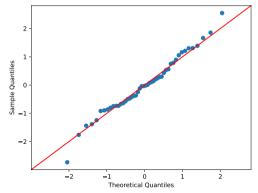
Both diagnostic plots indicate the residuals are reasonably close to Normal distribution centred on 0. The qqplot is perhaps more informative. Looking good so far!

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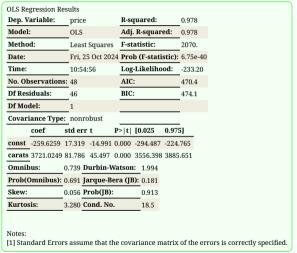




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OLS Reg	gression Re	sults					
Dep. Va	ariable:	_pric	e	R-squ	ared:		0.978
Model:		OLS		Adj. I	R-squar	ed:	0.978
Method	d:	Lea	st Squares	F-sta	tistic:		2070.
Date:		Fri,	25 Oct 2024	Prob	(F-stati	stic):	6.75e-
Time:		10:5	64:56	Log-I	ikeliho	od:	-233.2
No. Ob	servations	: 48		AIC:			470.4
Df Resi	duals:	46		BIC:			474.1
Df Mod	lel:	1					
Covari	ance Type:	non	robust				
	coef	std ei	ert P	> t [0.025	0.975	5]
const	-259.6259	17.31	9 -14.991 0.	.000 -:	294.487	-224.	765
carats	3721.0249	81.78	6 45.497 0.	.000 3	556.398	3885	.651
Omnib	us: (0.739	Durbin-Wa	tson:	1.994		
Prob(O	mnibus):	0.691	Jarque-Ber	a (JB):	0.181		
Skew:	(0.056	Prob(JB):		0.913		
Kurtos	is:	3.280	Cond. No.		18.5		

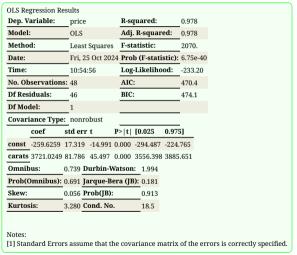
simpleModel.summary()



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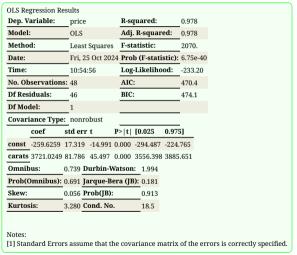
The output from Python's statsmodels.summary() call has lots of information!

How much of the variability of the data is explained by the model?



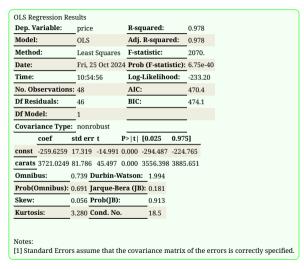
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- What is the probability that such data arose if price does not increase with weight?



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- How much of the variability of the data is explained by the model?
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- What score(s) indicate that the residuals are not serially correlated?
- What scores indicate that the distribution of the residuals is Normal?

OLS Re	gression Re	esults					
Dep. V	ariable:	_pric	e	R-s	quared:		0.978
Model:		OLS	5	Adj	j. R-squar	ed:	0.978
Metho	d:	Lea	st Squares	F-s	tatistic:		2070.
Date:		Fri,	25 Oct 202	24 Pro	ob (F-stati	stic):	6.75e-40
Time:		10:5	54:56	Log	g-Likeliho	od:	-233.20
No. Ob	servation	s: 48		AIC	C:		470.4
Df Resi	iduals:	46		BIC	: :		474.1
Df Mod	lel:	1					
Covari	ance Type	non	robust				
	coef	std e	er t	P> t	[0.025	0.975	5]
const	-259.6259	17.31	9 -14.991	0.000	-294.487	-224.	765
carats	3721.0249	81.78	6 45.497	0.000	3556.398	3885	.651
Omnib	us:	0.739	Durbin-V	Vatso	n: 1.994		
Prob(C	mnibus):	0.691	Jarque-B	era (J	B): 0.181		
Skew:		0.056	Prob(JB):	:	0.913		
Kurtos	is:	3.280	Cond. No		18.5		

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Definition 7 (Dep. variable)

This is synonymous with the *target*, which is price (in this dataset).

Definition 8 (Model)

statsmodels uses it here in the sense of *problem* formulation. We wish to solve an Ordinary Least Squares problem (assumes all the regression assumptions are met, so no special treatment was applied).

Definition 9 (No. Observations)

This is the number of rows (also known as instances or cases) in the training set.

Definition 10 (Df Model)

The model has one named feature (carats (weight of the diamond)) and one unnamed feature (constant, independent of carats). df, the number of degrees of freedom counts the named features.

Definition 11 (Df Residuals)

The number of degrees of freedom in the residuals is the number of residuals minus the number of features. A higher value tends to go with smaller model variance.

Definition 12 (Covariance Type)

If residuals have the same variance (homoscedastic), nonrobust covariance (the default) can be used.

OLS Regression Re	esults	
Dep. Variable:	price R-squared:	0.978
Model:	OLS Adj. R-squared:	0.978
Method:	Least Squares F-statistic:	2070.
Date:	Fri, 25 Oct 2024 Prob (F-statistic	6.75e-40
Time:	10:54:56 Log-Likelihood:	-233.20
No. Observation	s: 48 AIC:	470.4
Df Residuals:	46 BIC:	474.1
Df Model:	1	
Covariance Type	nonrobust	
coef	std err t P> t [0.025 0.9	75]
const -259.6259	17.319 -14.991 0.000 -294.487 -22	4.765
carats 3721.0249	81.786 45.497 0.000 3556.398 388	35.651
Omnibus:	0.739 Durbin-Watson: 1.994	
Prob(Omnibus):	0.691 Jarque-Bera (JB): 0.181	
Skew:	0.056 Prob(JB): 0.913	
Kurtosis:	3.280 Cond. No. 18.5	

Notes:

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ep. v	ariable:	_pric	e	K-sq	[uared:		0.97	3
Model:	:	OLS	;	Adj.	R-squar	ed:	0.97	3
Metho	d:	Lea	st Squares	F-st	atistic:		2070	١.
Date:		Fri,	25 Oct 2024	Prol	b (F-stati	stic):	6.75	e-40
Time:		10:5	64:56	Log	-Likeliho	od:	-233	.20
No. Ob	servations	s: 48		AIC	:		470.	1
Df Res	iduals:	46		BIC:	:		474.	1
Df Mod	iel:	1						
Covari	ance Type	: non	robust					
	coef	std e	rr t P	> t	[0.025	0.975	5]	
const	-259.6259	17.31	9 -14.991 0.	.000	-294.487	-224.	765	
carats	3721.0249	81.78	6 45.497 0	.000	3556.398	3885	.651	
Omnib	us:	0.739	Durbin-Wa	tson	1.994			
Prob(()mnibus):	0.691	Jarque-Ber	a (JB	0.181			
1100((0.056	Prob(JB):		0.913			
Skew:					18.5			

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Definition 13 (R-squared)

This is the ratio of the data variance explained by the model, to the variance of the data. It ranges from zero (model explains none of the data variance) to one (model explains all the data variance). A higher value is better, but be careful of overfitting the training set!

Definition 14 (Adj. R-squared)

Similar to R-squared, but it takes account of the number of features. Adding a feature generally increases R-squared, but if the feature did not help as much as its peers, adjusted R-squared shows a smaller increase than "normal" R-squared.

Definition 15 (F-statistic)

Ratio of the variance of a model with just the constant (intercept) feature to the variance of this model. Generally, large values of F are preferred.

Definition 16 (Prob (F statistic))

The value is assumed to follow the F distribution for given *dof*, so can lookup its probability. Small probability indicates that it is highly *unlikely* that the model is doing well purely by chance.

Definition 17 (log likelihood)

OLS is a special case of *maximum likelihood estimation*. Larger likelihood model fits the training data better.

OLS Reg	ression Re	esults						
Dep. Va	riable:	pric	e	R	-squ	ared:		0.978
Model:		OLS		A	dj. B	-squar	ed:	0.978
Method	l:	Lea	st Square	s F	stat	istic:		2070.
Date:		Fri,	25 Oct 20	24 P	rob	(F-stati	stic):	6.75e-40
Time:		10:5	4:56	L	og-L	ikeliho	od:	-233.20
No. Ob	servation	s: 48		A	IC:			470.4
Df Resi	duals:	46		В	IC:			474.1
Df Mod	el:	1						
Covaria	ance Type	non	robust					
	coef	std e	r t	P>	t [0	0.025	0.975	6]
const	-259.6259	17.31	9 -14.991	0.00	00 -2	94.487	-224.	765
carats	3721.0249	81.78	6 45.497	0.00	00 3	556.398	3885	.651
Omnib	us:	0.739	Durbin-	Nats	on:	1.994		
Prob(O	mnibus):	0.691	Jarque-B	era	(JB):	0.181		
Skew:		0.056	Prob(JB)	:		0.913		
Kurtos	is:	3.280	Cond. No). 		18.5		

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

OLS Reg	gression Re	sults					
Dep. Va	ariable:	_pric	e	R-squ	ared:		0.978
Model:		OLS	;	Adj. R	k-squar	ed:	0.978
Metho	d:	Lea	st Squares	F-stat	istic:		2070.
Date:		Fri,	25 Oct 2024	Prob ((F-stat	istic):	6.75e-40
Time:		10:5	64:56	Log-L	ikeliho	od:	-233.20
No. Ob	servations	: 48		AIC:			470.4
Df Resi	iduals:	46		BIC:			474.1
Df Mod	lel:	1					
Covari	ance Type	non	robust				
	coef	std e	rr t P	> t [0	0.025	0.975	5]
const	-259.6259	17.31	9 -14.991 0	.000 -2	94.487	-224.	765
carats	3721.0249	81.78	6 45.497 0	.000 35	556.398	3885	.651
Omnib	us:	0.739	Durbin-Wa	tson:	1.994		
Prob(C	mnibus):	0.691	Jarque-Ber	a (JB):	0.181		
Skew:		0.056	Prob(JB):		0.913		
	is:	2 200	Cond. No.		18.5		

[1] Standard Errors assume that the covariance matrix of the errors is correctly specifie

Definition 18 (AIC and BIC)

Akaike and Bayesian Information Criterion. These are calculated from the residuals and are derived from *information theory*. They allow for the number of features. Lower values are better.

Definition 19 (Features table: const, carats in this example)

coef is the parameter value for that feature, e.g., const=-259.6 here. P > |t| = 0 so it is highly unlikely the coef is zero, given the training data. We also have the 2.5% and 97.5% quantiles, giving the expected range of coef.

Definition 20 (Skew, Kurtosis)

Measures of asymmetry and of peak shape of the residual distribution. Ideal values are 0 (skew) and 3 (kurtosis).

Definition 21 (Durbin-Watson)

Measures the serial correlation of the residuals. Ideal value is 2 (no serial correlation).

Definition 22 (Cond. no)

OLS implementation solves a linear system of equations. Condition number measures column (hence feature independence). Large values mean the features are not independent (they are correlated), making the system more difficult to solve.

OLS Regression Results Dep. Variable: price R-squared: 0.978 Model: OLS Adj. R-squared: 0.978 Method: Least Squares F-statistic: 2070. Date: Fri, 25 Oct 2024 Prob (F-statistic): -233.20 No. Observations: 48 AIC: 470.4 Df Residuals: 46 BIC: 471.1 Df Model: 1
Model: OLS Adj. R-squarest 0.98 Method: Least Squarest F-statistic: 2070. Date: Fri, 25 Oct 2024 Prob (F-statistic) 6.75e-40 Time: 10:54:56 Log-Likelinost -233.20 No. Observations: 48 AIC: 470.4 Df Residuals: 46 BIC: 474.1 Df Model: 1 Polition of the properties of the propertie
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Df Residuals: 46 s BIC s 47.1 Df Model: 1 2 2 Covariance Type: nonr-bust p> t [0.025 s) 0.975 s conditions 259.6250 s 73.19 s 1.499 s 0.00 s 294.487 s 224.765 s
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carats 3721 0240 91 796 45 497 0 000 3556 399 3995 651
Circle 5721.0249 61.760 45.497 0.000 5550.598 5865.051
Omnibus: 0.739 Durbin-Watson: 1.994
Prob(Omnibus): 0.691 Jarque-Bera (JB): 0.181
Skew: 0.056 Prob(JB): 0.913
Kurtosis: 3.280 Cond. No. 18.5

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

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 - ② They can also be used constructively, to help identify promising candidate models.

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- Next time we return to classification...