

Advanced Topic: Derivation of Bayes Factors for Testing Two Proportions

July 19, 2016

Posterior Distributions and Predictive Distributions

Let's start by reviewing Bayes Theorem for the posterior under a Bernoulli sampling model and a conjugate prior Beta prior. We will start by considering the distributions with one group.

If $X_1, \dots, X_n \mid \theta \stackrel{iid}{\sim} \text{Ber}(\theta)$ where each of the X_i are independent and identically distributed with $P(X_i = 1 \mid \theta) = \theta$ then the conjugate prior for θ is Beta prior distribution,

$$p(\theta) = \frac{\theta^{a-1}(1-\theta)^{b-1}}{B(a, b)}$$

where $B(a, b)$ is known as the Beta function and is the normalizing constant of the Beta distribution¹

Bayes Theorem says that

$$p(\theta \mid X_1, \dots, X_n) = \frac{p(\theta)p(X_1, \dots, X_n \mid \theta)}{p(X_1, \dots, X_n)}$$

where the denominator is the marginal distribution or prior predictive distribution of the data. Using the definition of the Beta density we can find this without having to use calculus as long as we know the normalizing constants for the beta density. For those that are comfortable with calculus and integration see if you can confirm this using the definition in the footnote.

Starting with Bayes Theorem for the posterior density

$$p(\theta \mid X_1, \dots, X_n) = \frac{p(\theta)p(X_1, \dots, X_n \mid \theta)}{p(X_1, \dots, X_n)}$$

we begin by substituting the distributions of data and prior

$$\begin{aligned} &\propto \frac{\prod_i^n [\theta^{X_i}(1-\theta)^{1-X_i}] \theta^{a-1}(1-\theta)^{b-1}}{B(a, b)} \\ &= \frac{\theta^{\sum_{i=1}^n X_i} (1-\theta)^{\sum_{i=1}^n (1-X_i)} \theta^{a-1}(1-\theta)^{b-1}}{B(a, b)}. \end{aligned}$$

¹the Beta function is formally defined as $B(a, b) \equiv \int_0^1 \theta^{a-1}(1-\theta)^{b-1}$.

Letting $Y = \sum_{i=1}^n X_i$ and combining terms in the exponent, we have

$$\frac{\theta^{Y+a-1}(1-\theta)^{n-Y+b-1}}{B(a,b)}.$$

Recognizing that the numerator is the ‘kernel’ of a Beta density, we just need to multiply by $B(a,b)$ and divide by $B(Y+a, n-Y+b)$ so that the result is a Beta density:

$$p(\theta \mid X_1, \dots, X_n) = \frac{\theta^{Y+a-1}(1-\theta)^{n-Y+b-1}}{B(a,b)} \frac{B(a,b)}{B(Y+a, n-Y+b)}$$

where the $B(A,b)$ term cancels from numerator and denominator to obtain the normalized posterior density for θ .

Since the predictive distribution is the denominator in Bayes Theorem, we have that the marginal distribution is the inverse of the term in red:

$$p(X_1, \dots, X_n) = \frac{B(\sum X_i + a, n - \sum X_i + b)}{B(a,b)} \quad (1)$$

which is a ratio of Beta functions and the posterior density for θ simplifies to

$$p(\theta \mid X_1, \dots, X_n) = \frac{\theta^{\sum X_i + a - 1}(1-\theta)^{n - \sum X_i + b - 1}}{B(\sum X_i + a, n - \sum X_i + b)}.$$

Application to Bayes Factors

Recall that the Bayes Factor is defined as the ratio of prior predictive densities under two hypotheses H_1 and H_2 :

$$BF[H_1 : H_2] \equiv \frac{p(\text{data} \mid H_1)}{p(\text{data} \mid H_2)}$$

Let’s apply this to the case with two groups of Bernoulli observations $X_{A,i} \mid \theta_A \stackrel{iid}{\sim} \text{Ber}(\theta_A)$ for $i = 1, \dots, n_A$ and $X_{B,i} \mid \theta_B \stackrel{iid}{\sim} \text{Ber}(\theta_B)$ for $i = 1, \dots, n_B$ where we are interested in testing $H_1 : \theta_A = \theta_B$ versus $H_2 : \theta_A \neq \theta_B$

Under H_1 let’s denote the common value of the parameter as $\theta = \theta_A = \theta_B$. Our sampling model is that

$$X_{A,i} \mid \theta, H_1 \stackrel{iid}{\sim} \text{Ber}(\theta) \text{ for } i = 1, \dots, n_A$$

$$X_{B,i} \mid \theta, H_1 \stackrel{iid}{\sim} \text{Ber}(\theta) \text{ for } i = 1, \dots, n_B$$

If additionally the observations are independent across groups we may combine them into a single sample. Using a conjugate Beta prior

$$\theta \mid H_1 \sim B(a,b)$$

then the prior predictive distributions for the data $X_{A,1}, \dots, X_{A,n_A}, X_{B,1}, \dots, X_{B,n_B}$ will be

$$p(\text{data} \mid H_1) = \frac{B(Y_A + Y_B + a, n_A + n_B - Y_A - Y_B + b)}{B(a, b)}$$

where $Y_A = \sum_{i=1}^{n_A} X_{A,i}$ and $Y_B = \sum_{i=1}^{n_B} X_{B,i}$.

Under H_2 we assume that each group has its own probability of success:

$$X_{A,i} \mid \theta_A, H_2 \stackrel{iid}{\sim} \text{Ber}(\theta_A) \text{ for } i = 1, \dots, n_A$$

$$X_{B,i} \mid \theta_B, H_2 \stackrel{iid}{\sim} \text{Ber}(\theta_B) \text{ for } i = 1, \dots, n_B$$

and as before are independent. If we assign independent Beta priors to the θ 's for each group

$$\theta_A \sim \text{Beta}(a_A, b_A)$$

$$\theta_B \sim \text{Beta}(a_B, b_B)$$

then it is straightforward to show that we may apply the result about the predictive distribution to each group separately and that the joint predictive distribution is the product of the predictive distributions within each group:

$$p(\text{data} \mid H_1) = \frac{B(Y_A + a_A, n_A - Y_A + b_A)}{B(a_A, b_A)} \times \frac{B(Y_B + a_B, n_B - Y_B + b_B)}{B(a_B, b_B)}$$

The resulting Bayes factor is

$$BF[H_1 : H_2] = \frac{B(Y_A + Y_B + a, n_A + n_B - Y_A - Y_B + b)}{B(a, b)} \div \left[\frac{B(Y_A + a_A, n_A - Y_A + b_A)}{B(a_A, b_A)} \times \frac{B(Y_B + a_B, n_B - Y_B + b_B)}{B(a_B, b_B)} \right]$$

expressed as a function of the summary counts in the two groups and sample sizes. The beta function $B(\cdot)$ is available in most statistical/mathematical programming packages. When sample sizes are large, computing the log Bayes factor is recommended

$$\begin{aligned} \log(BF[H_1 : H_2]) = & \text{lbeta}(Y_A + Y_B + a, n_A + n_B - Y_A - Y_B + b) - \text{lbeta}(a, b) - \\ & [\text{lbeta}(Y_A + a_A, n_A - Y_A + b_A) - \text{lbeta}(a_A, b_A)] - \\ & [\text{lbeta}(Y_B + a_B, n_B - Y_B + b_B) - \text{lbeta}(a_B, b_B)] \end{aligned}$$

where lbeta is the log of the Beta function.

Hyperparameters

Günel and Dickey (1974) suggest that the prior hyperparameters under H_1 be obtained from the hyperparameters under H_2 by collapsing, so that $a = a_A + a_B$ and $b = b_A + b_B$. Their approach generalizes the special case here to contingency tables with Poisson or multinomial sampling.

For a default prior, we may use the Jeffrey's or reference prior within each group under H_2 , then $a_A = a_B = b_A = b_B = 1/2$ resulting in $a = 1, b = 1$ or a Uniform distribution for θ under H_1 .

References

Günel E. and Dickey, J. (1974) Bayes factors for independence in contingency tables *Biometrika* 61 (3): 545-557. doi: 10.1093/biomet/61.3.545 <http://biomet.oxfordjournals.org/content/61/3/545.short>