## Bayesian statistics with R 3. Analyses by hand

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# Back to Bayes

#### A simple example

- Let us take a simple example to fix ideas.
- 120 deer were radio-tracked over winter.
- 61 close to a plant, 59 far from any human activity.
- Question: is there a treatment effect on survival?

	Released	Alive	Dead	Other
treatment	61	19	38	4
control	59	21	38	0

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- How would the classical statistician justify this estimate?

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- K the number of alive individuals at the end of the winter, so that  $P(K = k) = \binom{n}{k} \theta^k (1 \theta)^{n-k}$ .
- The classical approach is to maximise the corresponding likelihood with respect to  $\theta$  to obtain the entirely plausible MLE:

$$\hat{\theta} = k/n = 19/57$$

4

• The Bayesian starts off with a prior.

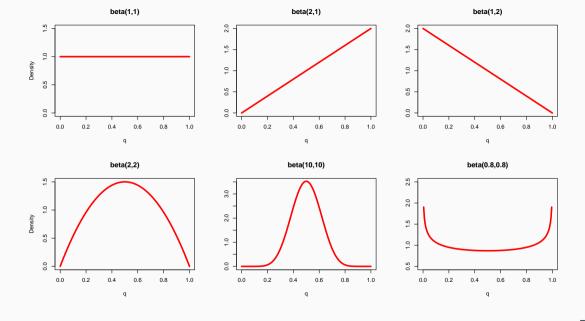
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$$q(\theta \mid \alpha,\beta) = \frac{1}{\mathsf{Beta}(\alpha,\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$
 with  $\mathsf{Beta}(\alpha,\beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$  and  $\Gamma(n) = (n-1)!$ 



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Take a Beta prior with a Binomial likelihood, you get a Beta posterior (conjugacy)

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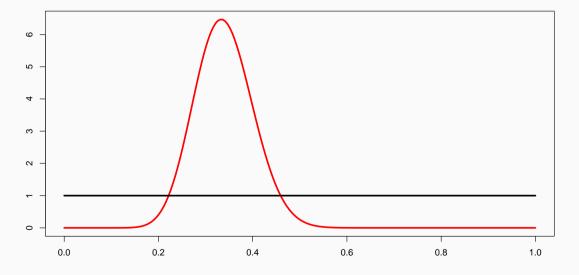
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- Note that in this specific situation, the posterior has an explicit expression, easy to manipulate.
- In particular,  $E(Beta(a,b)) = \frac{a}{a+b} = 20/59$  to be compared with the MLE 19/57.

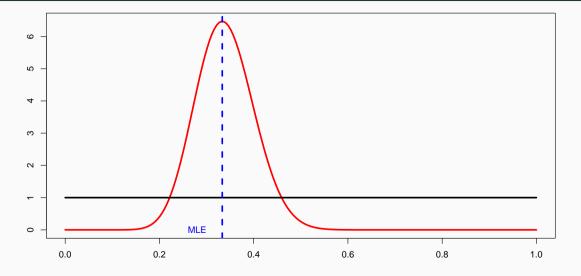
#### A general result

This is a general result, the Bayesian and frequentist estimates will always agree if there is sufficient data, so long as the likelihood is not explicitly ruled out by the prior.

## Prior Beta(1,1) and posterior survival Beta(20,39)



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#### **Notation**

Our model so far

$$y \sim \mathsf{Binomial}(N, \theta)$$

 $heta \sim \mathsf{Beta}(1,1)$ 

[likelihood]

[prior for  $\theta$ ]

