Bayesian statistics with RBayesian analyses in R with the Jags software

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Bayes in practice

Software implementation (R friendly)

Oldies but goodies:

- WinBUGS, OpenBUGS: Where it all began.
- Jags: What we will use in this course.

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If you're not into coding:

- brms: Bayesian regression models with Stan.
- MCMCglmm: Generalised Linear Mixed Models.
- Check out the CRAN Task View: Bayesian Inference for more.

Introduction to JAGS (Just Another Gibbs Sampler)

Martyn Plummer



Real example

Impact of climatic conditions on white stork breeding success



Assess effects of temperature and rainfall on productivity.

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- We need to specify priors for parameters.

Read in the data

```
nbchicks \leftarrow c(151,105,73,107,113,87,77,108,118,122,112,120,122,89,69,71,
              53.41.53.31.35.14.18)
nbpairs <- c(173,164,103,113,122,112,98,121,132,136,133,137,145,117,90,80,
            67,54,58,39,42,23,23)
```

```
temp \leftarrow c(15.1, 13.3, 15.3, 13.3, 14.6, 15.6, 13.1, 13.1, 15.0, 11.7, 15.3, 14.4, 14.4
          12.7.11.7.11.9,15.9,13.4,14.0,13.9,12.9,15.1,13.0)
```

rain $\leftarrow c(67,52,88,61,32,36,72,43,92,32,86,28,57,55,66,26,28,96,48,90,86,$

78,87)

datax <- list(N = 23, nbchicks = nbchicks, nbpairs = nbpairs,</pre> tomp = (tomp - moon(tomp))/cd(tomp)

Write down the model

[likelihood]	$nbchicks_i \sim Binomial(nbpairs_i, p_i)$
[linear model]	$logit(p_i) = a + b_{temp} \; temp_i + b_{rain} \; rain_i$
[prior for a]	$a \sim Normal(0, 1000)$
[prior for b_{temp}]	$b_{temp} \sim Normal(0, 1000)$
[prior for b_{rain}]	$b_{rain} \sim Normal(0, 1000)$

Build the model

```
{
# Likelihood
    for( i in 1 : N){
        nbchicks[i] ~ dbin(p[i],nbpairs[i])
        logit(p[i]) <- a + b.temp * temp[i] + b.rain * rain[i]
      }
# ...</pre>
```

Specify priors

```
# Priors
a ~ dnorm(0,0.001)
b.temp ~ dnorm(0,0.001)
b.rain ~ dnorm(0,0.001)
}
```

Warning: Jags uses precision for Normal distributions (1 / variance)

You need to write everything in a file

```
model <-
paste("
model
    for( i in 1 : N)
         nbchicks[i] ~ dbin(p[i],nbpairs[i])
         logit(p[i]) <- a + b.temp * temp[i] + b.rain * rain[i]</pre>
a \sim dnorm(0, 0.001)
b.temp \sim dnorm(0,0.001)
b.rain \sim dnorm(0.0.001)
```

Alternatively, you may write a R function

```
logistic <- function() {</pre>
    for( i in 1 : N)
        nbchicks[i] ~ dbin(p[i],nbpairs[i])
        logit(p[i]) <- a + b.temp * temp[i] + b.rain * rain[i]</pre>
# priors for regression parameters
a \sim dnorm(0, 0.001)
b.temp \sim dnorm(0,0.001)
b.rain ~ dnorm(0,0.001)
```

Let us specify a few additional things

```
# list of lists of initial values (one for each MCMC chain)
init1 \leftarrow list(a = -0.5, b.temp = -0.5, b.rain = -0.5)
init2 \leftarrow list(a = 0.5, b.temp = 0.5, b.rain = 0.5)
inits <- list(init1.init2)
# specify parameters that need to be estimated
parameters <- c("a","b.temp","b.rain")</pre>
# specify nb iterations for burn-in and final inference
nb.burnin <- 1000
nb. iterations <- 2000 # beware: nb. iterations includes nb. burnin!
```

Run Jags

```
# load R2jags
library(R2jags)
# run Jags
storks <- jags(data = datax,
               inits = inits,
               parameters.to.save = parameters,
               model.file = "code/logistic.txt",
               # model.file = logistic, # if a function was written
               n.chains = 2.
               n.iter = nb.iterations,
               n.burnin = nb.burnin)
storks
```

Inspect parameter estimates

```
#> Compiling model graph
#>
     Resolving undeclared variables
     Allocating nodes
#>
  Graph information:
#>
     Observed stochastic nodes: 23
#>
     Unobserved stochastic nodes: 3
#>
     Total graph size: 181
#>
  Initializing model
#> Inference for Bugs model at "code/logistic.txt", fit using jags,
   2 chains, each with 2000 iterations (first 1000 discarded)
\# n.sims = 2000 iterations saved
           mu.vect sd.vect 2.5%
                                     25%
                                             50%
                                                    75% 97.5% Rhat n.eff
#>
             1.545 0.085 1.429 1.509 1.550 1.587 1.659 1.173
#> a
                                                                      2000
#> b.rain -0.154 0.063 -0.278 -0.190 -0.156 -0.119 -0.024 1.061
                                                                        31
#> b.temp 0.022 0.060 -0.094 -0.017
                                          0.020
                                                  0.060
                                                         0.141 1.006 1300
#> deviance 206.430 29.947 201.764 202.698 203.843 205.630 212.695 1.066
                                                                      2000
```

Your turn: Practical 5

Assess convergence

Reminder - MCMC Algorithm

 MCMC algorithms can be used to construct a Markov chain with a given stationary distribution (set to be the posterior distribution).

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- For the MCMC algorithm, the posterior distribution is only needed to be known up to proportionality.
- Once the stationary distribution is reached we can regard the realisations of the chain as a (dependent) sample from the posterior distribution (and obtain Monte Carlo estimates).
- We consider some important implementation issues.

MCMC – Proposal Distribution

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- To implement a MCMC algorithm, we often need to specify a proposal distribution from which we generate candidate value then accept/reject.
- This typically involves
 - specifying a given distribution family (e.g. normal, uniform), and then,
 - setting the parameters of the given distribution.
- Although the exact distribution specified is essentially arbitrary it will have a significant effect on the performance of the MCMC algorithm.

• If only small moves can be proposed, the acceptance probability is high, but it will take a long time to explore the posterior distribution.

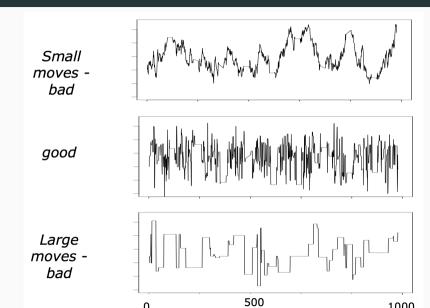
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- The movement around the parameter space is often referred to as mixing.

Good/Bad Traces



Autocorrelation functions

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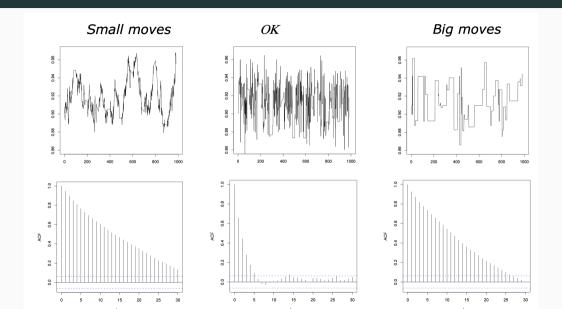
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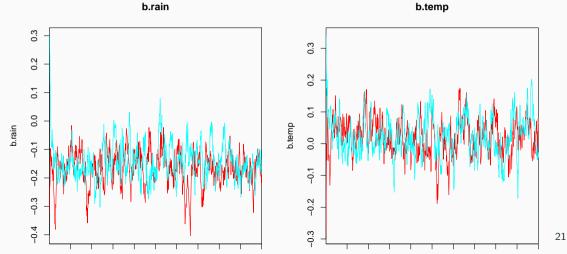
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- Autocorrelation function (ACF) plots are a convenient way of displaying the strength of autocorrelation in the given sample values.
- ACF plots provide the autocorrelation between successively sampled values separated by k iterations, referred to as lag, (i.e. $cor(\theta_t, \theta_{t+k})$) for increasing values of k.

ACFs



Traceplots for the storks

traceplot(storks,mfrow = c(1, 2), varname = c('b.rain','b.temp'), ask = FA



Autocorrelation for the storks

autocorr.plot(as.mcmc(storks),ask = FALSE) b.rain а Autocorrelation Autocorrelation -1.0 15 20 15 20 Lag Lag deviance b.temp Autocorrelation 0.5 Autocorrelation 0.5

How do good chains behave?

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- Once there, explore efficiently: The post-convergence sample size required for suitable numerical summaries.

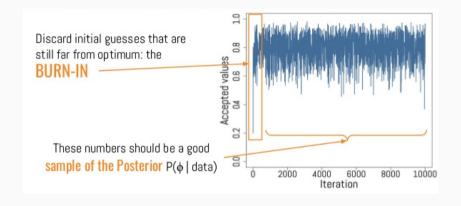
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- In practice, we must discard observations from the start of the chain and just use observations from the chain once it has converged.
- The initial observations that we discard are referred to as the **burn-in**.
- The simplest method to determine the length of the burn-in period is to look at trace plots.

Burn-in (if simulations cheap, be conservative)



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- The effective sample size (n.eff) measures chain length while taking into account the autocorrelation of the chain.
 - n.eff is less than the number of MCMC iterations.
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 - n.eff is less than the number of MCMC iterations.
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- We need n.eff \geq 100 independent steps.

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- Measures the ratio of the total variability combining multiple chains (between-chain plus within-chain) to the within-chain variability. Asks the question is there a chain effect? Very much alike the F test in an ANOVA.
- ullet Values near 1 indicates likely convergence, a value of ≤ 1.1 is considered acceptable.
- Necessary condition, not sufficient; In other words, these diagnostics cannot tell you that you have converged for sure, only that you have not.

n.eff and \hat{R} for the storks

storks

#>

- #> Inference for Bugs model at "code/logistic.txt", fit using jags,
- #> 2 chains, each with 2000 iterations (first 1000 discarded)
- # n.sims = 2000 iterations saved
- mu.vect sd.vect 2.5% 25% 50% 75% 97.5% Rhat #>

- #> a 1.545 0.085 1.429 1.509 1.550 1.587 1.659 1.173
- #> b.rain -0.154 0.063 -0.278 -0.190 -0.156 -0.119 -0.024 1.061
- #> b.temp 0.022 0.060 -0.094 -0.017 0.020 0.060 0.141 1.006
- #> deviance 206.430 29.947 201.764 202.698 203.843 205.630 212.695 1.066 #>
- #> For each parameter, n.eff is a crude measure of effective sample size,
- #> and Rhat is the potential scale reduction factor (at convergence, Rhat=

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- Check autocorrelation, effective sample size and \hat{R} .

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- Something wrong with your model?
 - Start with a simpler model (remove complexities).
 - Use simulations.
- Change your sampler. Upgrade to Nimble or Stan.

MCMC makes you queens and kings

of the stats world

Get all values sampled from posteriors

```
res <- as.mcmc(storks) # convert outputs in a list
res <- rbind(res[[1]],res[[2]]) # put two MCMC lists on top of each other
head(res)
                       b.rain b.temp
#>
                                             deviance
#> [1.] 0.07642911 -0.32041118 -0.2902827755 1205.9032
#> [2,] 0.42572884 -0.22453244 -0.1844443635 747.2488
#> [3,] 0.65598954 -0.23096623 -0.0621171663 520.0681
#> [4,] 0.80114761 -0.21617920 -0.0009756621 415.1433
#> [5,] 0.93302269 -0.12248125 -0.0523636288
                                             349.0541
#> [6,] 1.03283474 -0.09644411 -0.0494129797 305.5391
tail(res)
```

#> a b.rain b.temp deviance #> [1995,] 1.592979 -0.1785135 -0.01251340 202.9100 #> [1996] 1.61/32/ -0.159/6/ -0.050/7115 20/ 6339

Compute a posteriori Pr(rain < 0)

```
# probability that the effect of rainfall is negative
mean(res[,'b.rain'] < 0)
#> [1] 0.9915
```

Compute a posteriori Pr(temp < 0)

```
# probability that the effect of temperature is negative
mean(res[,'b.temp'] < 0)
#> [1] 0.366
```

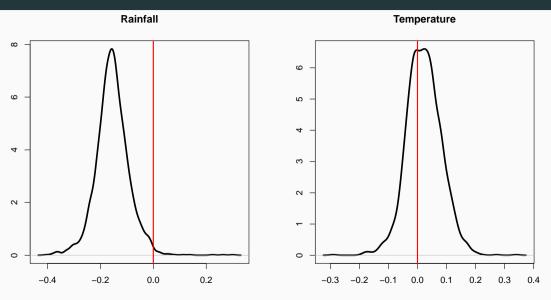
Get credible interval for the rain effect

```
quantile(res[,'b.rain'],c(0.025,0.975))
#> 2.5% 97.5%
#> -0.27838080 -0.02448805
```

Get credible interval for the temperature effect

```
quantile(res[,'b.temp'],c(0.025,0.975))
#> 2.5% 97.5%
#> -0.09373495 0.14065101
```

Graphical summaries



Your turn: Practical 6