

## Kalman Filter

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- If the observation function  $h(\cdot)$  is linear, then

- $$\underline{y} = \underline{H}\underline{x} + \underline{n}$$

where  $\underline{y}$  is a linear function of  $h$  but with some added noise.

$$\text{Then } \hat{\underline{x}} = \underline{L}\underline{y}$$

where 
$$\underline{L} = \underline{H}^{-1}$$

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## Kalman Filter

$\underline{y}$ : observation

$\underline{x}$ : state

$\underline{n}$ : random noise (measurement noise)

$$\underline{n} \sim \mathcal{N}(0, \underline{R})$$

$$\underline{y} = h(\underline{x}) + \underline{n}$$

The estimation problem  $\hat{\underline{x}} = \mathcal{L}(\underline{y})$

$\hat{\underline{x}}$ : estimate

$\mathcal{L}$ : estimator

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## Over constrained System

- If  $\underline{H}$  has more rows than the number of columns then there is no inverse of  $\underline{H}$ .
- But we can find the best  $\hat{\underline{x}}$  such that the residue  $|\underline{y} - \underline{H}\hat{\underline{x}}|$  is as small as possible.
- The best estimator for a linear observation function happens to be a linear function.
- We denote the residue as  $\underline{n} = \underline{y} - \underline{H}\hat{\underline{x}}$

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- Different equations can have noise terms of different variance.

So some equations may be more reliable.

- We formulate minimisation of the norm of  $\underline{Wn}$  where  $\underline{W}$  is a matrix as

$$\|\underline{Wn}\|^2 = \underline{n}^T \underline{W}^T \underline{Wn}$$

which is equivalent to solving the system

$$\underline{W}\underline{y} = \underline{W}\underline{H}\underline{x}$$

$$\text{because } \underline{Wn} = \underline{W}(\underline{y} - \underline{H}\underline{x}) = \underline{W}\underline{y} - \underline{W}\underline{H}\underline{x}$$

- That is we want to solve  $\underline{W}\underline{y} = \underline{W}\underline{H}\underline{x}$  subject to minimising  $\|\underline{W}\underline{y} - \underline{W}\underline{H}\underline{x}\|^2$

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$$\bullet \quad \frac{df}{d\underline{x}} = 2\underline{A}^T \underline{A} * \underline{x} - 2\underline{A}^T \underline{b} = 0$$

$$\bullet \quad \text{Therefore } \underline{A}^T \underline{A}\underline{x} = \underline{A}^T \underline{b} \quad \text{or} \quad \hat{\underline{x}} = (\underline{A}^T \underline{A})^{-1} \underline{A}^T \underline{b}$$

$$\bullet \quad \text{The given system of equations } \underline{W}\underline{H}\underline{x} = \underline{W}\underline{y}$$

$$\text{Substituting } \underline{A} \equiv \underline{W}\underline{H} \text{ and } \underline{b} \equiv \underline{W}\underline{y} \text{ in } \hat{\underline{x}} = (\underline{A}^T \underline{A})^{-1} \underline{A}^T \underline{b}$$

$$\begin{aligned} \bullet \quad \hat{\underline{x}} &= ((\underline{W}\underline{H})^T \underline{W}\underline{H})^{-1} (\underline{W}\underline{H})^T \underline{W}\underline{y} = (\underline{H}^T \underline{W}^T \underline{W}\underline{H})^{-1} \underline{H}^T \underline{W}^T \underline{W}\underline{y} \\ &= (\underline{H}^T \underline{R}^{-1} \underline{H})^{-1} (\underline{H}^T \underline{R}^{-1}) \underline{y} \end{aligned}$$

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- Consider the standard notation for a system of equations

$$\underline{A}\underline{x} = \underline{b} \text{ to be solved for } \hat{\underline{x}} \text{ so as to minimise } \|\underline{b} - \underline{A}\underline{x}\|^2$$

- The objective to be minimised can be re-written as  $\|\underline{A}\underline{x} - \underline{b}\|^2$  or

$$(\underline{A}\underline{x} - \underline{b})^T (\underline{A}\underline{x} - \underline{b}) \quad \text{or} \quad (\underline{x}^T \underline{A}^T - \underline{b}^T) (\underline{A}\underline{x} - \underline{b})$$

- Or  $f \equiv (\underline{x}^T \underline{A}^T \underline{A}\underline{x} - \underline{x}^T \underline{A}^T \underline{b} - \underline{b}^T \underline{A}\underline{x} + \underline{b}^T \underline{b})$

$$\frac{df}{d\underline{x}} = 2\underline{A}^T \underline{A}\underline{x} - \underline{A}^T \underline{b} - \underline{A}^T \underline{b} + 0$$

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- $\underline{R}$  is the covariance matrix of the measurement noise.

- Entries in  $\underline{R}$  which correspond to a large variance or covariance will translate to smaller entries in the  $\underline{W}$  matrix.

$$\underline{R}^{-1} = \underline{W}^T \underline{W}$$

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### Quality of the estimator

- The size of the covariance matrix should be small.
- Covariance matrix  $\underline{P} = \mathbb{E}[(x - \hat{x})(x - \hat{x})^T]$
- A smaller norm of  $\underline{P}$  implies reduced uncertainty or fluctuations
- The estimate  $\hat{x}$  should be unbiased

$$\mathbb{E}[\underline{x} - \hat{\underline{x}}] = 0$$

- For a linear estimator  $\hat{\underline{x}} = \underline{L}\underline{y}$  corresponding to the system  $\underline{y} = \underline{H}\underline{x} + \underline{n}$

$$\mathbb{E}[\underline{x} - \hat{\underline{x}}] = \mathbb{E}[\underline{x} - \underline{L}\underline{y}] = \mathbb{E}[\underline{x} - \underline{L}(\underline{H}\underline{x} + \underline{n})] = \mathbb{E}[(\underline{I} - \underline{L}\underline{H})\underline{x}] - \mathbb{E}[\underline{L}\underline{n}] = 0$$

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- The covariance matrix of our estimate

$$\begin{aligned}\underline{P} &= \mathbb{E}[(\underline{x} - \hat{\underline{x}})(\underline{x} - \hat{\underline{x}})^T] \\ &= \mathbb{E}[(\underline{x} - \underline{L}\underline{y})(\underline{x} - \underline{L}\underline{y})^T]\end{aligned}$$

- Substituting  $\underline{y} = \underline{H}\underline{x} + \underline{n}$  we have

$$\begin{aligned}&= \mathbb{E}[(\underline{x} - \underline{L}\underline{H}\underline{x} - \underline{L}\underline{n})(\underline{x} - \underline{L}\underline{H}\underline{x} - \underline{L}\underline{n})^T] \\ &= \mathbb{E}[(\underline{I} - \underline{L}\underline{H})\underline{x} - \underline{L}\underline{n}][(\underline{I} - \underline{L}\underline{H})\underline{x} - \underline{L}\underline{n}]^T = \mathbb{E}[(\underline{L}\underline{n})(\underline{L}\underline{n})^T]\end{aligned}$$

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$$\bullet \mathbb{E}[(\underline{I} - \underline{L}\underline{H})\underline{x}] = 0$$

$$\text{or } (\underline{I} - \underline{L}\underline{H})\mathbb{E}[\underline{x}] = 0$$

$$\therefore \underline{I} = \underline{L}\underline{H} \text{ or } \underline{L}\underline{H} = \underline{I}$$

$$\bullet \mathbb{E}[\hat{\underline{x}}] = \underline{x}$$

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$$\begin{aligned}\bullet &= \mathbb{E}[\underline{L}\underline{n}\underline{n}^T\underline{L}] \\ &= \underline{L}\mathbb{E}[\underline{n}\underline{n}^T]\underline{L} \\ &= \underline{L}\underline{R}\underline{L}^T\end{aligned}$$

$$\text{We define } \underline{P} \equiv \underline{L}\underline{R}\underline{L}^T$$

- Here  $\underline{R}$  is the covariance matrix of the measurement noise.
- Substituting for the estimator  $\underline{L} = (\underline{H}^T \underline{R}^{-1} \underline{H})^{-1} \underline{H}^T \underline{R}^{-1}$  we get

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- $\underline{P} \equiv \underline{L}\underline{R}\underline{L}^\top$
- $\underline{L} = (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \underline{H}^\top \underline{R}^{-1}$

$$\begin{aligned} \underline{P} &= (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \underline{H}^\top \underline{R}^{-1} \underline{R} \underline{R}^{-1} \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \\ &= (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \underline{H}^\top \underline{R}^{-1} \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \\ &= (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \end{aligned}$$

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### Claim

- The estimator  $\underline{L} = (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \underline{H}^\top \underline{R}^{-1}$  has the minimum norm covariance, given by  $\underline{P} \equiv \underline{L}\underline{R}\underline{L}^\top$ . We denote this optimal estimator as  $\underline{L}_0$

$$\underline{L}_0 = \underline{P}\underline{H}^\top \underline{R}^{-1}$$

Proof:

We know that  $\underline{L}_0 = (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \underline{H}^\top \underline{R}^{-1}$  is the solution to the normal equations  $\underline{W}\underline{H}\underline{x} = \underline{W}\underline{y}$

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- Here we have used the result that the inverse of a symmetric and invertible matrix will also be symmetric.

That is, given  $A^\top = A$  and  $A^{-1}$  exists we have

$$A^{-1}A = I$$

Taking transpose of both sides  $(A^{-1}A)^\top = A^\top(A^{-1})^\top = I^\top = I$

Premultiplying both sides by  $A^{-1}$

$$A^{-1}A(A^{-1})^\top = A^{-1}I$$

$I(A^{-1})^\top = A^{-1}I$  or  $(A^{-1})^\top = A^{-1}$  Therefore  $A^{-1}$  is symmetric.

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- The covariance matrix for estimator  $\underline{L}_0$  is  $\underline{L}_0 \underline{R} \underline{L}_0^\top \equiv \underline{P}$

- $\underline{L}_0$  is an unbiased estimator.

$$\text{So } \underline{L}_0 \underline{H} = \underline{I}$$

- Now consider an alternative unbiased estimator  $\underline{L}$

- We trivially write  $\underline{L} = \underline{L}_0 + (\underline{L} - \underline{L}_0)$

- The covariance matrix for  $\underline{L}$  is given by

$$\underline{P} = \underline{L}\underline{R}\underline{L}^\top$$

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- $\underline{P} = \underline{L} \underline{R} \underline{L}^\top$ 

$$= (\underline{L}_0 + (\underline{L} - \underline{L}_0)) \underline{R} (\underline{L}_0 + (\underline{L} - \underline{L}_0))^\top$$

$$= \underline{L}_0 \underline{R} \underline{L}_0^\top + (\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top + \underline{L}_0 \underline{R} (\underline{L} - \underline{L}_0)^\top + (\underline{L} - \underline{L}_0) \underline{R} (\underline{L} - \underline{L}_0)^\top$$
- $\underline{L}_0^\top = (\underline{H}^\top \underline{R}^{-1})^\top (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}^\top$ 

$$= \underline{R}^{-1} \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}$$

$$= \underline{R}^{-1} \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}$$

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- Since  $(\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top = 0$ 

$$\left( (\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top \right)^\top = 0$$
or  $\left( (\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top \right)^\top = 0$ 

$$\underline{L}_0 \underline{R} (\underline{L} - \underline{L}_0)^\top = 0$$
- $\underline{P} = \underline{L}_0 \underline{R} \underline{L}_0^\top + (\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top + \underline{L}_0 \underline{R} (\underline{L} - \underline{L}_0)^\top + (\underline{L} - \underline{L}_0) \underline{R} (\underline{L} - \underline{L}_0)^\top$ 

$$= \underline{L}_0 \underline{R} \underline{L}_0^\top + (\underline{L} - \underline{L}_0) \underline{R} (\underline{L} - \underline{L}_0)^\top$$

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- $\underline{R} \underline{L}_0^\top = \underline{R} \underline{R}^{-1} \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} = \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}$
- $(\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top = (\underline{L} - \underline{L}_0) \underline{H} (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}$ 

$$= (\underline{L} \underline{H} - \underline{L}_0 \underline{H}) (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}$$
- Since  $\underline{L}$  and  $\underline{L}_0$  are unbiased estimators
$$\underline{L} \underline{H} - \underline{L}_0 \underline{H} = \underline{I}$$
- $\therefore (\underline{L} - \underline{L}_0) \underline{R} \underline{L}_0^\top = 0$

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- $\underline{P} = \underline{L}_0 \underline{R} \underline{L}_0^\top + (\underline{L} - \underline{L}_0) \underline{R} (\underline{L} - \underline{L}_0)^\top$
- This is the sum of two positive semi-definite matrices.
- For such matrices, the norm of the sum is greater than or equal to either norm.
- The norm of  $\underline{P}$  will be minimised when the 2nd term on the RHS vanishes when  $\underline{L} = \underline{L}_0$
- Therefore, the estimator  $\underline{L}_0 = (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1} \underline{H}^\top \underline{R}^{-1}$  has the minimum covariance, given by  $\underline{P} = (\underline{H}^\top \underline{R}^{-1} \underline{H})^{-1}$

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- In fact,  $\underline{L}_0 = \underline{P}\underline{H}^\top \underline{R}^{-1}$  in terms of  $\underline{P}$
- So far we have seen how to estimate the state  $\hat{\underline{x}}$  given the measurement equation.

$$\underline{y} = \underline{H}\underline{x} + \underline{n} \quad \underline{n} \sim \mathcal{N}(0, \mathbb{R})$$

$$\begin{aligned} \hat{\underline{x}} &= \underline{P}\underline{H}^\top \underline{R}^{-1} \underline{y} \\ &= \underline{L}\underline{y} \end{aligned}$$

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## Dynamic System

- In a linear dynamic system, the states update from one time tick to another.

$$\text{State Update:} \quad \underline{x}_{k+1} = \underline{F}_k \underline{x}_k + \underline{G}_k \underline{u}_k + \underline{\eta}_k$$

$$\text{Measurement:} \quad \underline{y}_k = \underline{H}_k \underline{x}_k + \underline{\xi}_k$$

- $\underline{\eta}_k \sim \mathcal{N}(0, \underline{Q}_k)$  is the system noise
- $\underline{\xi}_k \sim \mathcal{N}(0, \underline{R}_k)$  is the measurement noise

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- At every time tick there are 2 evidences which determine the current state.
  - Contribution (evidence) from the previous state, through state update
  - Evidence from the measurement made at the current time tick.  
(state update from observing the evidence)

- Notation

$\hat{\underline{x}}_{k|k-1}$  state estimate at time  $k$  given the measurements up to time  $k - 1$

$\hat{\underline{x}}_{k|k}$  state estimate after taking account the measurement up to time tick  $k$

$\hat{\underline{y}}_k$  predicted observation at time tick  $k$

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## Update Step

- There are two sources that provide estimate to the state  $\hat{\underline{x}}_k$

### 2. Evidence from the new measurement

$$\underline{y}_k = \underline{H}_k \underline{x}_k + \underline{\xi}_k$$

$$\underline{R}_k = \mathbb{E}[\underline{\xi}_k \underline{\xi}_k^T]$$

- There are two sources that provide estimate to the state  $\hat{\underline{x}}_k$

### 1. Estimate from the state propagation $\hat{\underline{x}}_{k|k-1}$

$$\hat{\underline{x}}_{k|k-1} = \underline{x}_k + \underline{e}_k$$

The estimate  $\hat{\underline{x}}_{k|k-1}$  is considered to have a random deviation or error from the true state  $\underline{x}_k$

The covariance  $P_{k|k-1}$  of the error term summarises the uncertainty of prediction  $\hat{\underline{x}}_{k|k-1}$  given the past history of measurements.

$$P_{k|k-1} = \mathbb{E}[\underline{e}_k \underline{e}_k^T]$$

- Given the evidences about the unknown true state  $\underline{x}_k$  our task is to compute an estimate  $\hat{\underline{x}}_k$  of  $\underline{x}_k$

- We collect the two evidences

$$\hat{\underline{x}}_{k|k-1} = \underline{x}_k + \underline{e}_k$$

$$\underline{y}_k = \underline{H}_k \underline{x}_k + \underline{\xi}_k$$

to form a system of equations

$$\underline{y} = \underline{H} \underline{x}_k + \underline{n}$$

- $\underline{y} = \underline{H}x_k + \underline{n}$

$$\underline{y} = \begin{bmatrix} \hat{x}_{k|k-1} \\ y_k \end{bmatrix} \quad \underline{H} = \begin{bmatrix} I \\ \underline{H}_k \end{bmatrix} \quad \underline{n} = \begin{bmatrix} \underline{e}_k \\ \underline{n}_k \end{bmatrix}$$

- $\hat{x}_{k|k-1} = \underline{I}x_k + \underline{e}_k$   
 $y_k = \underline{H}_k x_k + \underline{n}_k$

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- The covariance matrix of the vector  $\underline{n} = \begin{bmatrix} \underline{e}_k \\ \underline{n}_k \end{bmatrix}$  is formulated as

$$R = \begin{bmatrix} \underline{P}_{k|k-1} & 0 \\ 0 & \underline{R}_k \end{bmatrix}$$

Here we assume that the two noise vectors are independent.

- The formulation  $\underline{y} = \underline{H}x_k + \underline{n}$  is a classical estimation problem.

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- The solution to the estimation problem  $\underline{y} = \underline{H}_k x_k + \underline{n}$  is

$$\hat{x}_{k|k} = \underline{P}_{k|k} \underline{H}^T \underline{R}^{-1} \underline{y}$$

$$\underline{P}_{k|k} = (\underline{H}^T \underline{R}^{-1} \underline{H})^{-1}$$

This is the update stage of the Kalman Filter

- Simplification

$$\underline{P}_{k|k}^{-1} = \underline{H}^T \underline{R}^{-1} \underline{H}$$

$$= \begin{bmatrix} I & \underline{H}_k^T \end{bmatrix} \begin{bmatrix} \underline{P}_{k|k-1}^{-1} & 0 \\ 0 & \underline{R}_k^{-1} \end{bmatrix} \begin{bmatrix} I \\ \underline{H}_k \end{bmatrix}$$

$$= \underline{P}_{k|k-1}^{-1} + \underline{H}_k^T \underline{R}_k^{-1} \underline{H}_k$$

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### Simplifying the posterior estimate

- $\hat{x}_{k|k} = \underline{P}_{k|k} \underline{H}^T \underline{R}^{-1} \underline{y}$

$$= \underline{P}_{k|k} \begin{bmatrix} I & \underline{H}_k^T \end{bmatrix} \begin{bmatrix} \underline{P}_{k|k-1}^{-1} & 0 \\ 0 & \underline{R}_k^{-1} \end{bmatrix} \begin{bmatrix} \hat{x}_{k|k-1} \\ y_k \end{bmatrix}$$

$$= \underline{P}_{k|k} \begin{bmatrix} \underline{P}_{k|k-1}^{-1} & \underline{H}_k^T \underline{R}_k^{-1} \end{bmatrix} \begin{bmatrix} \hat{x}_{k|k-1} \\ y_k \end{bmatrix}$$

$$= \underline{P}_{k|k} \left( \underline{P}_{k|k-1}^{-1} \hat{x}_{k|k-1} + \underline{H}_k^T \underline{R}_k^{-1} y_k \right)$$

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$$\begin{aligned}
\bullet \quad \hat{\underline{x}}_{k|k} &= \underline{P}_{k|k} \underline{H}^\top \underline{R}^{-1} \underline{y} \\
&= \underline{P}_{k|k} \left( \left( \underline{P}_{k|k}^{-1} - \underline{H}_k^\top \underline{R}_k^{-1} \underline{H}_k \right) \hat{\underline{x}}_{k|k-1} + \underline{H}_k^\top \underline{R}_k^{-1} \underline{y}_k \right) \\
&= \left( \underline{I} - \underline{P}_{k|k} \underline{H}_k^\top \underline{R}_k^{-1} \underline{H}_k \right) \hat{\underline{x}}_{k|k-1} + \underline{P}_{k|k} \underline{H}_k^\top \underline{R}_k^{-1} \underline{y}_k \\
&= \hat{\underline{x}}_{k|k-1} - \underline{P}_{k|k} \underline{H}_k^\top \underline{R}_k^{-1} \underline{H}_k \hat{\underline{x}}_{k|k-1} + \underline{P}_{k|k} \underline{H}_k^\top \underline{R}_k^{-1} \underline{y}_k \\
&= \hat{\underline{x}}_{k|k-1} + \left( \underline{P}_{k|k} \underline{H}_k^\top \underline{R}_k^{-1} \right) \left( \underline{y}_k - \underline{H}_k \hat{\underline{x}}_{k|k-1} \right)
\end{aligned}$$