

## PCA & ICA

+ RECAP PCA & SOLVE IT

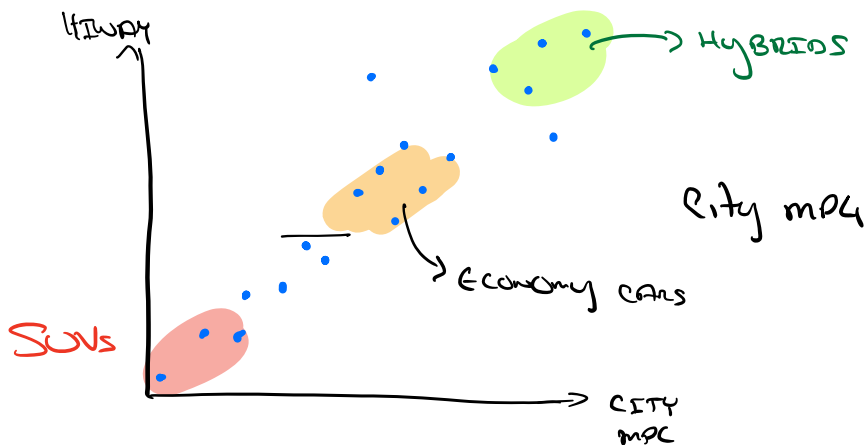
+ ICA & "the cocktail party" "LEARN UP TO SYMMETRY"

UP NEXT: Self-Supervised MACHINE LEARNING!

# PCA: Principal Component Analysis

STRUCTURE	Prob.	Non Prob.
"CLUSTER"	GMM	K-MEANS
"SUBSPACE"	Factor Analysis	PCA ← THIS SECTION

Ex: GIVEN PAIRS (Highway mpg, city mpg) of some cars

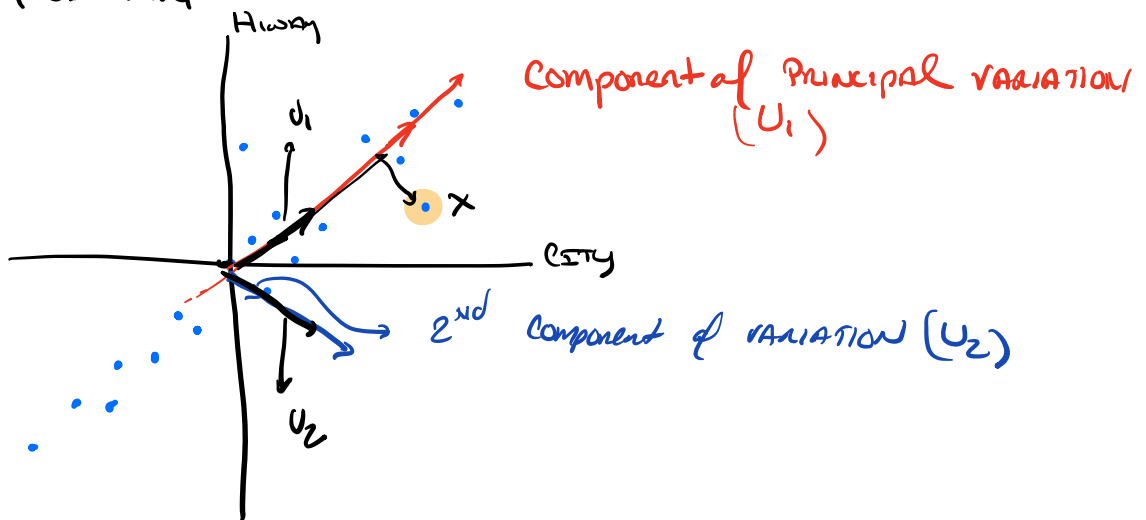


Question: "Good mpg"

① CENTER DATA

$$\mu = \frac{1}{n} \sum x^{(i)}$$

$$x^{(i)} \mapsto x^{(i)} - \mu$$



Now  $\|u_1\| = \|u_2\| = 1$  by convention.

- $u_1$  IS "How good is mpg"
- $u_2$  IS "difference between highway & city" (roughly)

WE CAN WRITE  $x = \alpha_1 u_1 + \alpha_2 u_2$

↳ WE MAY JUST KEEP THIS COMPONENT

"Explains more variation"

TODAY: How we find these directions, AND some caveats

- think about 1000s of dims  $\rightarrow$  10s of dims
- A dimensionality reduction method

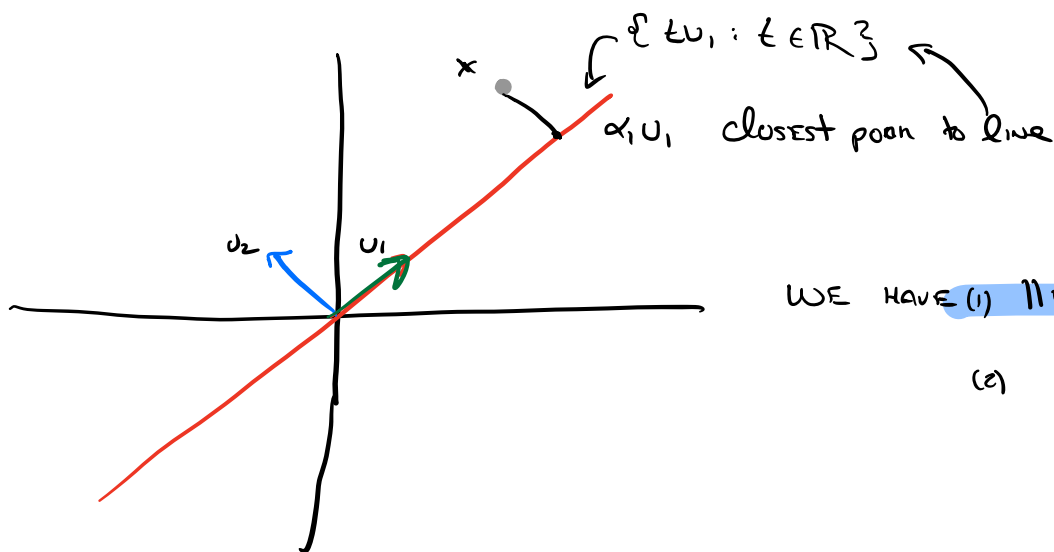
## Preprocessing

Given  $x^{(1)} \dots x^{(n)} \in \mathbb{R}^d$

1. CENTER the data  $x^{(i)} \mapsto x^{(i)} - \mu$  in which  $\mu = \frac{1}{n} \sum x^{(i)}$
2. MAY NEED TO RESCALE components e.g. "FEET PER gallon"  
? MPG

WE will assume data is preprocessed

## PCA AS OPTIMIZATION



WE HAVE (i)  $\|u_i\| = 1$  (unit vectors)

(ii)  $u_i \cdot u_j = \delta_{ij}$  (orthogonal)

How do you find closest point to the line?

$$\alpha_1 = \underset{\alpha}{\operatorname{argmin}} \|x - \alpha u_1\|^2$$

$$= \underset{\alpha}{\operatorname{argmin}} \|x\|^2 + \alpha^2 \|u_1\|^2 - 2\alpha (u_1 \cdot x)$$

Differentiate w.r.t  $\alpha$

$$2(\alpha - u_1 \cdot x) = 0 \Rightarrow \alpha = u_1 \cdot x$$

Generalize:  $U_1 \dots U_k \in \mathbb{R}^d$  AND  $x \in \mathbb{R}^d$  USE  $U_i \cdot U_j = \delta_{ij}$

$$\underset{\alpha_1, \dots, \alpha_k}{\text{Argmin}} \|x - \sum_{i=1}^k \alpha_i U_i\|^2 = \underset{\alpha}{\text{argmin}} \|x\|^2 + \sum_{i=1}^k \alpha_i^2 \|U_i\|^2 - 2\alpha_i \langle U_i, x \rangle$$

Hence  $\alpha_i = U_i \cdot x$

WE CALL  $\|x - \sum_{i=1}^k \alpha_i U_i\|^2$  THE RESIDUAL

WE CAN find PCA by either

- IN CLASS
- ① MAXIMIZE Projected Subspace
  - ② MINIMIZE Residual

$$\underset{\substack{U \in \mathbb{R}^d \\ \|U\|=1}}{\text{MAX}} \frac{1}{n} \sum_{i=1}^n (U \cdot x^{(i)})^2$$

WE NEED SOME facts  
to solve this

LET  $A$  BE SYMMETRIC & SQUARE, THEN

$$A = U \Lambda U^T \text{ in which}$$

$$\cdot U U^T = I \text{ (ORTHONORMAL)}$$

$\cdot \Lambda$  IS DIAGONAL

$$\Lambda_{ii} = \lambda_i \text{ AND } \lambda_1 \geq \dots \geq \lambda_n \text{ by convention eigenvalues}$$

Recall: IF  $x = \sum_{i=1}^n \alpha_i U_i$  where  $[U_1 \dots U_n] = U$

$$\begin{aligned} Ax &= U \Lambda U^T x = U \Lambda \sum_{i=1}^n \alpha_i \overset{\text{STANDARD BASIS vector}}{e_i} \quad (U_i \cdot U_j = \delta_{ij}) \\ &= U \sum_{i=1}^n \lambda_i \alpha_i e_i \quad \text{diagonal } \Lambda \\ &= \sum_i \lambda_i \alpha_i U_i \end{aligned}$$

IF  $x = c U_i$  THEN  $x$  IS AN EIGENVECTOR, AND  $Ax = \lambda_i x$

$$\max_{x: \|x\|^2=1} x^T A x = \max_{\alpha: \|\alpha\|^2=1} \sum_{i=1}^n \alpha_i^2 \lambda_i$$

Hence, we set  $\alpha_i = 1$ , the principal eigenvalue

Which  $x$  attains it? If  $\lambda_1 = \lambda_2$ ?

Now, back to PCA!

$$\max_{U: \|U\|_F^2=1} \frac{1}{n} \sum_{i=1}^n (U^T x^{(i)})^2$$

↑ THE PROJECTION ONTO  $U$

$$U \in \mathbb{R}^d = \frac{1}{n} \sum_{i=1}^n U^T x^{(i)} (x^{(i)})^T U = U^T \left( \frac{1}{n} \sum_{i=1}^n x^{(i)} (x^{(i)})^T \right) U$$

↙ COVARIANCE of DATA  
(WE SUBTRACTED MEAN)

$\therefore U$  is principal Eigenvector

WHAT IF WE WANT MORE DIMENSIONS? WE KEEP  $U_{1:k}$ !

How do we REPRESENT DATA?

$$x^{(i)} \mapsto \sum_{j=1}^k (x^{(i)} \cdot U_j) U_j$$

↙ WE KEEP THESE  $k$  SCALARS (the  $\alpha_k$  ABOVE)

A map from  $\mathbb{R}^d \rightarrow \mathbb{R}^k$

How do we CHOOSE  $k$ ?

ONE APPROACH "Amount of Explained VARIANCE"

$$\frac{\sum_{i=1}^k \lambda_i}{\sum_{i=1}^n \lambda_i} \geq 0.9$$

$$(\text{ASIDE } \text{tr}(A) = \sum_i A_{ii} = \sum_j \lambda_j)$$

NB: ONLY MAKES SENSE if  $\lambda_j \geq 0$ . Hence COVARIANCE IS important

LUCKING INSTABILITY: Suppose  $\lambda_k = \lambda_{k+1} \dots$  WHAT HAPPENS?

REP IS UNSTABLE HERE

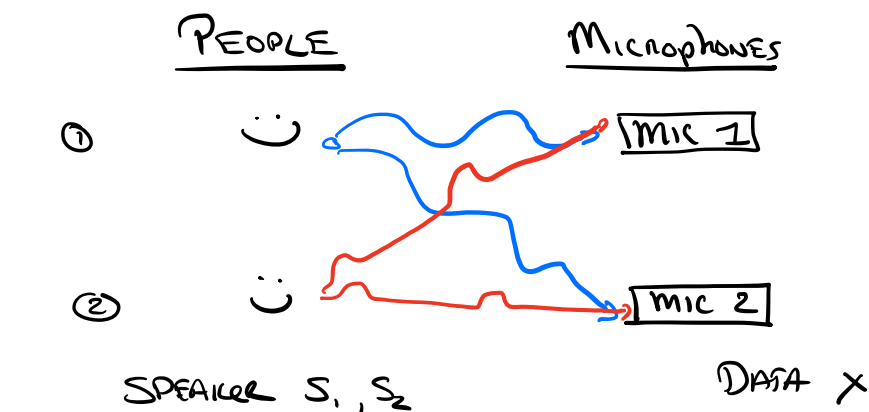
## RECAP of PCA

- Dimensionality Reduction technique (e.g. Visualization)
- MAIN IDEA IS TO project ON A SUBSPACE, nice theory.



## ICA INDEPENDENT Component Analysis

- high-level story
- Key facts  $\neq$  likelihood
- model

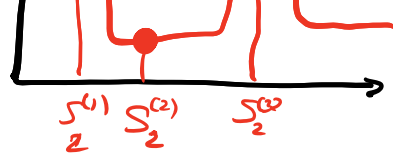
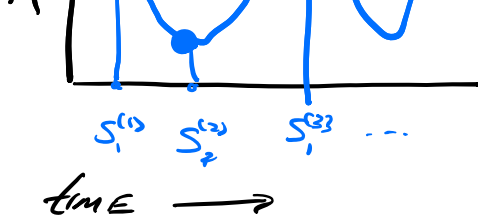
## Cocktail Party Problem (IN HW!)



NB: WE SEE A MIXTURE of

  $\neq$   AT EACH MIC





$S_j^{(t)}$  IS INTENSITY AT TIME  $t$  from SPEAKER  $j$

WE DO NOT OBSERVE  $S^{(t)}$  only  $x^{(t)}$  - the microphones  
ex model  $x_j^{(t)} = a_{j1} S_1^{(t)} + a_{j2} S_2^{(t)}$

"MICROPHONE  $j$  SEES A MIXTURE OF  $S_1^{(t)}$  &  $S_2^{(t)}$ "

OR

$$\text{OBSERVED } x^{(t)} = A \text{ LATENT } S^{(t)}$$

LATENT

FOR SIMPLICITY, ASSUME # of SPEAKERS = # of mics =  $d$

GIVEN:  $x^{(1)}, x^{(2)}, \dots, x^{(n)} \in \mathbb{R}^d$        $d$  IS # of microphones & SPEAKERS

DO: find  $s^{(1)}, \dots, s^{(n)} \in \mathbb{R}^d$   
 AND  $A \in \mathbb{R}^{d \times d}$  st.  $x^{(t)} = A s^{(t)}$

WE CALL  $A$  THE MIXING MATRIX AND  $W = A^{-1}$  UNMIXING MATRIX

WRITE  $W = \begin{bmatrix} w_1^T \\ \vdots \\ w_d^T \end{bmatrix}$  SO THAT  $S_j^{(t)} = w_j \cdot x^{(t)}$

### SOME CAVEATS

- WE ASSUME  $A$  DOES NOT VARY w/ TIME AND IS FULL RANK

• THERE ARE INHERENT Ambiguity

• WE CAN'T DETERMINE SPEAKER ID (could swap 1 & 2)

• CAN'T DETERMINE ABSOLUTE INTENSITY

$$(cA)(c^{-1}s^{(t)}) = As^{(t)} \text{ for any } c \neq 0$$

• Surprising Speakers CANNOT be Gaussian

Suppose  $x^{(t)} \sim N(\mu, AA^T)$  then if  $U^T U = I$   $AU$  generates the SAME data.

Nevertheless, we can recover something meaningful!

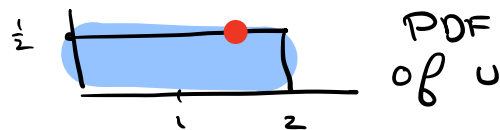
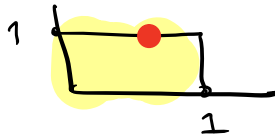
Algorithm: Joint MLE, solved by GRAD DESCENT

DETOUR: Density under linear transform (Key Confusion)

Ex:  $S \sim \text{Uniform}[0,1]$   $U = 2S$  what is PDF of  $U$ ?

TEMPTED TO WRITE  $P_U(\frac{x}{2}) = P_S(x)$

PDF of  $S$



$$P_S(x) = \begin{cases} 1 & \text{if } x \in [0,1] \\ 0 & \text{o.w} \end{cases}$$

$$P_U(x) = P_S(\frac{x}{2}) \cdot \frac{1}{2}$$

THE KEY ISSUE IS the NORMALIZATION CONSTANT

for INVERTIBLE MATRIX  $A$ ,  $U = AS$

$$P_U(x) = P_S(A^{-1}x) | \det(A^{-1}) |$$

$$= P_S(\frac{1}{W}x) | \det(W) | \quad (\frac{1}{\det(A)} = \det(A^{-1}))$$

CHANGE OF VAR  
formula for  
integrals



From HERE ICA is MLE:

$$P(S) = \prod_{j=1}^d P_S(s_j)$$

"SOURCES ARE INDEPENDENT,

AND HAVE SAME DISTRIBUTION"

$$P(X) = \prod_{j=1}^d P_S(W \cdot x)$$

$\cdot |\det(W)|$

(USE LINEAR TRANSFORM ABUSE)

Now written in terms of  $K$  and  $A$ .

Key technical bit: Use non-rotationally invariant distribution

SET  $P_S(x) \propto g'(x)$  for  $g(x) = (1 + e^{-x})^{-1}$

$$\text{Solve } \ell(W) = \sum_{t=1}^n \sum_{j=1}^d \log g'(w_j \cdot x^{(t)}) + \log |\det(W)|$$

- $\log |\det(W)|$
- USE GD & you're done!

RECAP: • SAW PCA. WORKAROUND dimensionality reduction

• ICA. Key ideas for HW. Introduce "up to symmetry".