Probabilities Recap

10-605 Machine Learning with Large Datasets

Fall 2022



Outline

- Setup
- Random variables
- Distribution function
- Expectation
- Multivariate Distributions
- Independence
- ROC curve
- Probability in Hashing (birthday paradox)



Setup

Sample Space

A set of all possible outcomes or realizations of some random trial.

Event

A subset of sample space

Probability Axioms

- \circ P(A) ≥ 0 for every A
- \circ P(Ω)=1;
- o If A1, A2, . . . are disjoint, then

$$P(\bigcup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} P(A_i)$$



Random variables

Definition

 \circ A random variable is a function that maps from the sample space to the reals (X : $\Omega \to R$), i.e., it assigns a real number X(ω) to each outcome ω.

Example

- X returns 1 if a coin is heads and 0 if a coin is tails. Y returns the number of heads after 3 flips of a fair coin.
- Random variables can take on many values, and we are often interested in the distribution over the values of a random variable, e.g., P(Y = 0)



Distribution function

Definition

- Suppose X is a random variable, x is a specific value that it can take,
- Cumulative distribution function (CDF) is the function $F: R \rightarrow [0, 1]$, where $F(x) = P(X \le x)$.
- If X is discrete \Rightarrow probability mass function: f(x) = P(X = x).



Distribution function (cont.)

 If X is continuous ⇒ probability density function for X if there exists a function f such that f(x) ≥ 0 for all x,

$$\int_{-\infty}^{\infty} f(x)dx = 1$$

and for every $a \le b$,

$$P(a \le X \le b) = \int_{a}^{b} f(x)dx$$

If F(x) is differentiable everywhere, f(x) = F'(x).



Example of distributions

Discrete variable	Probability function	Mean	Variance
Uniform $X \sim U[1, \ldots, N]$	1/ <i>N</i>	$\frac{N+1}{2}$	
Binomial $X \sim Bin(n, p)$	$\binom{n}{x}p^{x}(1-p)^{(n-x)}$	np	
Geometric $X \sim Geom(p)$	$(1-p)^{x-1}p$	1/p	
Poisson $X \sim Poisson(\lambda)$	$\frac{e^{-\lambda}\lambda^{x}}{x!}$	λ	
Continuous variable	Probability density function	Mean	Variance
Uniform $X \sim U(a, b)$	1/ (b-a)	(a + b)/2	
Gaussian $X \sim N(\mu, \sigma^2)$	$\frac{1}{\sqrt{2\pi}\sigma} \exp(-\frac{1}{2\sigma^2}(x-\mu)^2)$	μ	
Gamma $X \sim \Gamma(\alpha, \beta) \ (x \ge 0)$	$\frac{\frac{1}{\sqrt{2\pi}\sigma}\exp(-\frac{1}{2\sigma^2}(x-\mu)^2)}{\frac{1}{\Gamma(\alpha)\beta^a}x^{a-1}e^{-x/\beta}}$	$\alpha \beta$	
Exponential $X \sim exponen(\beta)$	$\frac{1}{\beta}e^{-\frac{x}{\beta}}$	β	

Expectation

Expected Values

Discrete random variable X

$$E[g(X)] = \sum_{x \in \chi} g(x)f(x)$$

Continuous random variable X

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx$$



Expectation (cont.)

Mean and variance

$$\mu = E(X)$$

$$var[X] = E[(X - \mu)^2]$$

We also have

$$var[X] = E[X^2] - \mu^2$$



Multivariate Distributions

Definition

$$F_{X,Y}(x,y) := P(X \le x, Y \le y)$$

and

$$f_{X,Y}(x,y) := \frac{\partial^2 F_{X,Y}(x,y)}{\partial x \partial y}$$

Marginal Distribution of X (discrete case)

$$f_X(x) = P(X = x) = \sum_y P(X = x, Y = y) = \sum_y f_{X,Y}(x, y)$$

What about continuous variable?



Independence

• Independent Variables

X and Y are independent if and only if

$$P(X = x, Y = y) = P(X = x)P(Y = y)$$

Or

$$f_{X,Y}(x,y) = f_X(x)f_Y(y)$$



Independence (cont.)

IID variable

 Independent and identically distributed (IID) random variables are drawn from the same distribution and are all mutually independent.

Linearity of Expectation

Even if the events are not independent, this property still holds

$$E[\sum_{x=1}^{n} X_i] = \sum_{i=1}^{n} E[X_i]$$



• Confusion matrix

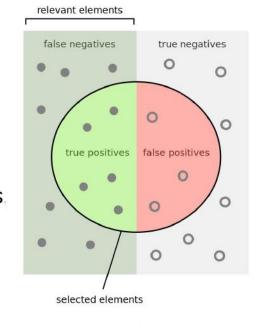
Actual Values

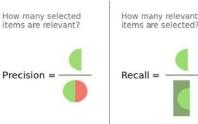
	,	Positive (1)	Negative (0)
d Values	Positive (1)	ТР	FP
Predicted	Negative (0)	FN	TN



- Statistics Computed from Confusion Matrix
 - Precision: Out of all the predicted positive instances how many were predicted correctly.

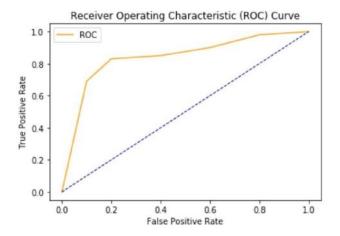
Recall: Out of all the positive classes
 how many instances were identified correctly.



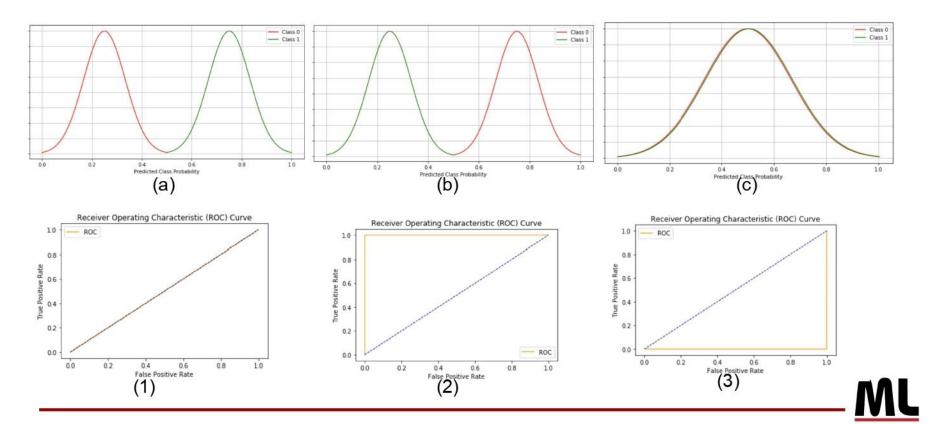




- Introduction to AUC ROC Curve
 - how good the model is for distinguishing the given classes, in terms of the predicted probability











Assumption

- n=number of people
- o k=365
- \circ P(person i is born on day j) = 1/k

We are interested in the event A that at least two people have the same

$$P(A) = 1 - P(\overline{A})$$

$$= 1 - \frac{k}{k} \cdot \frac{k-1}{k} \cdot \dots \cdot \frac{k-n+1}{k}$$

$$= 1 - \frac{k!}{(k-n)!k^n}.$$



Hashing

- Similar to assignments of birthdays
- n items mapped into k slots

Hashing problems dealing with probabilities

- the expected number of items mapping to same slot
- the expected number of empty slots
- the expected number of collisions



Empty slots

The probability that slot j remains empty after mapping all n items is

$$(1-\frac{1}{k})^n$$

The expected number of empty slots is

$$E(X) = \sum_{j=1}^{k} E(X_j) = k \left(1 - \frac{1}{k}\right)^n.$$

 \circ If k = n, we can get a max limitation of 0.367



KL Divergence

Question:

How different are two probability distributions from each other?

$$D_{KL}(P||Q) = \mathbb{E}_{x \sim P(\cdot)}[\log(rac{P(X)}{Q(X)})] = \sum_x P(x)\lograc{P(x)}{Q(x)}$$

It captures what is the expected "excess surprise" from using Q as a model for data when the actual distribution is P.

KL-divergence is NOT symmetric!

Concentration Inequalities

Can we figure out the probability that a random variable deviates from it's mean by a particular value; i.e. with how much probability does the following statement occur:

$$|\bar{X} - \mathbb{E}[X]| \le \delta$$

Concentration inequalities are a family of such statements that provide **exact** bounds on this probability.

Some common ones are; Markov's, Chebyshev's, Hoeffding's, Chernoff's Bounds etc.

Markov's Inequality

If X is non-negative, then for a positive value of a;

If
$$X \geq 0$$
, $P(X \geq a) \leq rac{\mathbb{E}[X]}{a}$

Chebyshev's Inequality

For a random variable X, with finite mean, and non-zero variance;

$$\Pr(|X - \mu| \geq k\sigma) \leq rac{1}{k^2}$$

Qualitatively, this statement tells us the probability that the value of a random variable deviates from it's mean by 'k' standard deviations is bounded by 1/k^2.

Johnson and Lindenstrauss Lemma

Lemma For any $0 < \epsilon < 1$ and any interger n let k be a possitive interger such that

$$k \ge \frac{24}{3\epsilon^2 - 2\epsilon^3} \log n \tag{2}$$

then for any set A of n points $\in \mathbb{R}^d$ there exists a map $f: \mathbb{R}^d \to \mathbb{R}^k$ such that for all $x_i, x_i \in A$

$$(1 - \epsilon)||x_i - x_j||^2 \le ||f(x_i) - f(x_j)||^2 \le (1 + \epsilon)||x_i - x_j||^2$$
(3)

Note: The proof involves Markov's inequality