

# 36700 – Probability and Mathematical Statistics

## Spring 2019

### Homework 7

*Due Wednesday, April 10th at 12:40 PM*

- All homework assignments shall be uploaded using Gradescope through the Canvas portal). Late submissions are not allowed.

Q1. Let  $X, Y$  be two random variables. Let  $r(x) = \mathbb{E}(Y|X = x)$ , and  $f(x), g(x)$  be an arbitrary functions.

- (a) Use Q6 of HW2 to prove that (assuming all expectations are finite)

$$\mathbb{E}[(Y - r(X))g(X)] = 0.$$

- (b) Use the previous part to prove that (again, assuming all expectations are finite)

$$\mathbb{E}(Y - f(X))^2 = \mathbb{E}(Y - r(X))^2 + \mathbb{E}(r(X) - f(X))^2.$$

As a result,  $r(X)$  is the optimal predictor of  $Y$  based on  $X$  in terms of squared loss.

Q2. (Interpretation of least squares without linearity assumption) Let  $X, Y$  be two random variables with finite second moments. Let  $(\beta_0, \beta_1)$  be the minimizer of the following problem:

$$\min_{(b_0, b_1)} \mathbb{E}(Y - b_0 - b_1 X)^2.$$

Here, the interpretation of  $\beta_0 + \beta_1 X$  is no longer the conditional expectation since we do not assume  $r(x) = \mathbb{E}(Y|X = x)$  is linear. Instead, it is the best linear predictor of  $Y$  using  $X$ .

- (a) Express  $\beta_0$  and  $\beta_1$  using moments of  $(X, Y)$ .
- (b) Compare your result in part (a) with the least squares estimates  $(\hat{\beta}_0, \hat{\beta}_1)$ . What kind of estimates is it (MLE, MoM, Plug-in, Bayes)?
- Q3. In the simple regression problem, assume  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$ , where  $\epsilon_i$  are iid  $N(0, \sigma^2)$ . Here the unknown parameter is  $(\beta_0, \beta_1, \sigma^2)$ .

- (a) Find the MLE of  $(\beta_0, \beta_1, \sigma^2)$ .

- (b) Find the (exact) distributions of  $\hat{\beta}_0$  and  $\hat{\beta}_1$  conditioning on  $X_i$ 's (Hint: if  $Z_i \sim N(\mu_i, \sigma_i^2)$  are independent, then  $\sum_i a_i Z_i$  is normal with mean  $\sum_i a_i \mu_i$  and variance  $\sum_i a_i^2 \sigma_i^2$ . Moreover, in the lecture we already derived the mean and variance of  $\hat{\beta}_1$ , and the mean of  $\hat{\beta}_0$ , which you can directly use as known results).

(c) Find the MSE (mean squared error) of the MLE (for  $\beta_0, \beta_1$  only). Is the MLE consistent?

Q4. (Prediction intervals) In the same context as Q3, let's say we have a new independent data pair  $(X_0, Y_0)$  from the same model. But we only observe  $X_0$  and would like to predict  $Y_0$ . Our point estimate of  $Y_0$  is  $\hat{y}_0 = \hat{\beta}_0 + \hat{\beta}_1 X_0$ . Find a value  $\delta$  such that  $\mathbb{P}(\hat{y}_0 - \delta \leq Y_0 \leq \hat{y}_0 + \delta) = 1 - \alpha$ . [Hint: Let  $Y_0 = \beta_0 + \beta_1 X_0 + \epsilon_0$ . It is equivalent to have  $\mathbb{P}(-\delta \leq \beta_0 + \beta_1 X_0 + \epsilon_0 - \hat{y}_0 \leq \delta) = 1 - \alpha$ . Use Q3 to show that  $\beta_0 + \beta_1 X_0 + \epsilon_0 - \hat{y}_0$  is a normal variable with mean zero, and the desired  $\delta$  can be determined by its variance and  $\alpha$ .]

Q5. If we write

$$\mathbf{X} = \begin{bmatrix} 1 & X_1 \\ 1 & X_2 \\ \cdot & \cdot \\ \cdot & \cdot \\ \cdot & \cdot \\ 1 & X_n \end{bmatrix}, \quad \mathbf{Y} = \begin{bmatrix} Y_1 \\ Y_2 \\ \cdot \\ \cdot \\ \cdot \\ Y_n \end{bmatrix}, \quad \boldsymbol{\epsilon} = \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \cdot \\ \cdot \\ \cdot \\ \epsilon_n \end{bmatrix}, \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix},$$

then the regression model can be written as

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}.$$

Verify that the LS estimate is

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y},$$

where  $\cdot^T$  denotes the matrix transpose. In other words, you need to show that the matrix form solution  $\hat{\boldsymbol{\beta}}$  agrees with the LS solution given on page 3 of Lecture Note 16.

**Optional problem.** Even without normality assumption as in Q3 or the linearity assumption, one can still establish asymptotic normality of the least squares estimate  $(\hat{\beta}_0, \hat{\beta}_1)$  using the CLT and delta method. In this exercise, we will use the setting of Q2, assuming a random design (consider  $X_1, \dots, X_n$  as iid random variables).

- Prove that

$$\sqrt{n}(\hat{\beta}_0 - \beta_0) \rightsquigarrow N(0, \sigma_0^2), \quad \sqrt{n}(\hat{\beta}_1 - \beta_1) \rightsquigarrow N(0, \sigma_1^2)$$

and find the asymptotic variances  $\sigma_0^2, \sigma_1^2$  (expressed as functions of moments of  $X, Y$ ).

- Indeed we can establish joint asymptotic normality of  $(\hat{\beta}_0, \hat{\beta}_1)$ . Show that

$$\sqrt{n} \begin{bmatrix} \hat{\beta}_0 - \beta_0 \\ \hat{\beta}_1 - \beta_1 \end{bmatrix} \rightsquigarrow N(\mathbf{0}, \Sigma)$$

where  $\mathbf{0}$  is a zero vector of length 2, and find the asymptotic covariance matrix  $\Sigma$ .

- Using the previous part to construct a  $(1 - \alpha)$  confidence set for the two-dimensional vector  $(\beta_0, \beta_1)$ .