

## WEAK Supervision Nuggets

- INCOMPLETE CASE  $\rightarrow$  Simple Estimation task
- Correlations  $\rightarrow$  Inverse Covariance & Graph

GIVEN:  $x^{(1)} \dots x^{(n)} \in \mathbb{R}^d$  (Data Points)  
 $\lambda_1 \dots \lambda_n : \lambda_i : \mathbb{R}^d \rightarrow \{-1, 1\} \cup \{\text{ABSTAIN}\}$   
find  $P(y^{(i)} | \bar{x}, x^{(i)}) \quad y \in \{-1, 1\}$

IDEA:  $\lambda_i$  is a noisy voter/ function (INACCURATE)  
 $\lambda_i$ : "the classifier says yes"  
 $\lambda_i$ : "NAME IN DS"  
... Programmatic labels ...

Model 0: No ABSTAINS, INDEPENDENT CLASS.

Each classifier  $\lambda_i$  has hidden accuracy  $P(x \neq y \rightarrow \text{observed label})$

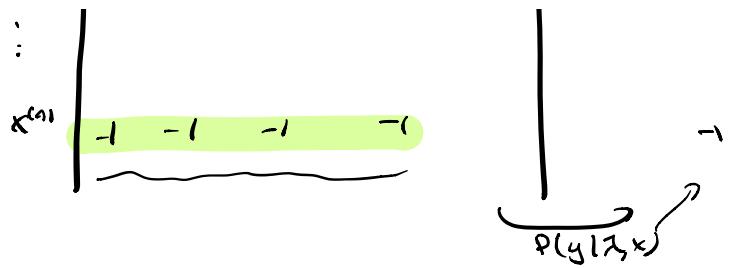
w/ Prob  $P_i(\lambda_i(x) = y)$  "  $\lambda_i$  is right"

$1 - P_i(\lambda_i(x) = -y)$  "  $\lambda_i$  is wrong"

Sorry, we don't see  $y$ .

$$P(\lambda_i(x) = 1 | y=1) = P(\lambda_i(x) = -1 | y=-1) = P_i$$

<u>Given</u>				<u>Undeserved</u>
Data		$\lambda_1 \lambda_2 \lambda_3 \dots \lambda_n$		$y$
$x^{(1)}$		1 1 1 1		1
$x^{(2)}$		1 -1 1		-1



$$\textcircled{1} \quad \mathbb{E}[\lambda_i \lambda_j] = P_i \cdot \underbrace{1}_{\in \{-1, 1\}} + (1-P_i) \cdot (-1) = 2P_i - 1 \quad \det a_i \\ a_i \stackrel{!}{=} 2P_i - 1 \\ Q_i \in [-1, 1]$$

$$\textcircled{2} \quad \mathbb{E}[\lambda_i \lambda_j] = 1 \text{ if } i=j \quad (\mathbb{E}[\lambda^2] = \mathbb{E}[1])$$

$$\textcircled{3} \quad \mathbb{E}[\lambda_i \lambda_j] \quad i \neq j \\ = P_i P_j \cdot 1 + (1-P_i)(1-P_j) \cdot (-1) \quad \text{is true} \\ P_i(1-P_j)(-1) + (1-P_i)P_j \cdot (-1) \\ = a_i \cdot a_j$$

form a MATRIX  $m \in \mathbb{R}^{m \times m}$   $M_{ij} = \mathbb{E}[\lambda_i \lambda_j]$

NB: we can estimate  $m$  from data (observed)

Unlike  $y$ :

"Agreements  $\downarrow$  disagreements" no need for  $y$ .

### Simple Algorithm

$$M_{ij} M_{jk} = a_i a_j a_k$$

$$\frac{M_{ij} M_{jk}}{M_{ik}} \approx a_j^2 \quad \text{solve "up to the sign of } a_i\text{"}$$

we know magnitude, not sign

$M_{ij}$  is observed  $M_{ij} = a_i a_j$  if I know  $\text{Sign}(a_i)$

$$\text{Sign}(M_{ii}) = \text{Sign}(a_i) \text{ Sign}(a_i) \Rightarrow \text{Sign}(a_i)$$

$\Rightarrow \alpha_i, -\alpha_i$  both are solutions.

$\sum \alpha_i > 0$ , breaks symmetry "Labels are not all  
WEAKLY PREDICTIVE  
OR Adversarial"

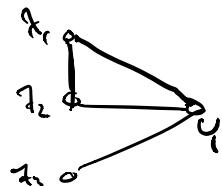
• WHAT if  $m_{ij} = 0 \Rightarrow \alpha_i = 0$  or  $\alpha_j = 0$

$$\Rightarrow \alpha_i = 0 \Rightarrow 2\alpha_i - 1 = 0 \Rightarrow \rho_i = \frac{1}{2} \quad \rho_i \text{ are located}$$

$|P_i - \frac{1}{2}|$  distance away from  $\frac{1}{2}$

Remark: Simple Solution to "Em-like"  
Symmetry  $\Rightarrow$  has to be in model

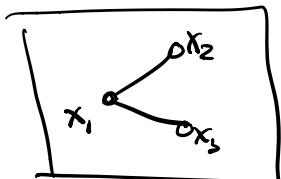
WHAT IF WE Labels are correlated



$$\mathbb{E}[\lambda_1, \lambda_2 | \gamma] = \mathbb{E}[\lambda_1 | \gamma] \mathbb{E}[\lambda_2 | \gamma]$$

"if there edge between two r.v. (nodes)"  
 $\Rightarrow$  then they are not INDEPENDENT  
 $(i,j) \notin E$  holds for any  $i,j$ .

Nugget Structure of INVERSE COVARIANCE MATRICES



$$x_1 \sim N(0, 1)$$

$$x_2 = x_1 + \epsilon_2 \quad \epsilon_2 \sim N(0, 1)$$

$$x_3 = x_1 + \epsilon_3 \quad \epsilon_3 \sim N(0, 1)$$

$$x_2 \sim N(\mu_1, 1)$$

$$\mathbb{E}(x_2) = 0 \quad \mathbb{E}(x_2) = \mathbb{E}(x_1) + \mathbb{E}(\epsilon_2) = 0$$

$$\mathbb{E}(x_3) = \mathbb{E}(x_1) + \mathbb{E}(\epsilon_3) = 0$$

$$2. \mathbb{E}[x_i^2] = 1 \quad \mathbb{E}[x_2^2] = \mathbb{E}[(x_1 + \epsilon_2)^2] = \mathbb{E}[x_1^2] + \mathbb{E}[\epsilon_2^2] + \cancel{\mathbb{E}[x_1 \epsilon_2]} = 2$$

$$3. \mathbb{E}[x_1 x_2] = \mathbb{E}[x_1^2] + \mathbb{E}[x_1 \epsilon_2] = 1$$

$$\Sigma = \begin{bmatrix} 1 & 1 & 1 \\ 1 & 2 & 1 \\ 1 & 1 & 2 \end{bmatrix}$$

NO EDGE

$$\Sigma^{-1} = \begin{bmatrix} 3 & -1 & -1 \\ -1 & 1 & 0 \\ -1 & 0 & 1 \end{bmatrix}$$

$V = \{x_1, x_2, x_3\}$   
 $E = \{(x_1, x_2), (x_1, x_3)\}$

We say a probability distribution  $p: \mathbb{R}^d \rightarrow [0, 1]$

factorizes or agrees with a graph  $G = (V, E)$

if  $p(x) = c_0 \cdot \prod_{(x_i, x_j) \in E} p_e(x_i, x_j) \cdot \prod_{x_i \in V} p_v(x_i)$

$c_0$  is the normalization constant

GAUSSIANS

$$\log p(x) = \log \exp \left\{ x^T \Sigma^{-1} x \right\} + C$$

$$\text{"IR factors"} = \log c_0 + \sum_{(x_i, x_j) \in E} \log p_e(x_i, x_j) + \sum_{v \in V} \log p_v(x_i)$$

$$A = \Sigma^{-1} = \sum_{i,j} A_{ij} x_i x_j$$

$$\text{For } (i, j) \notin E \quad 2x_i x_j \sum_{k, l} A_{kl} x_k x_l = (A_{ij} + A_{ji})$$

BECUSE  $\Sigma^{-1}$  IS symmetric (covariance matrix)

$$A_{ij} = A_{ji} \Rightarrow 2A_{ij}.$$

If we differentiate factorized expression?

$(i,j) \notin E \Rightarrow$  factors term must be 0.

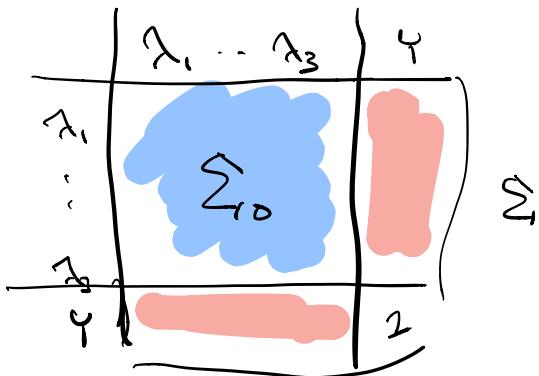
$$\Rightarrow A_{ij} = \sum_{\cdot i}^{-1} = 0$$

So Gaussians have this structure if  $a_{ij}$  are not connected

$$\text{Graph} \Rightarrow \sum_{ij}^{-1} = 0$$

More complex for Gaussians, (Finsler 2018) (Wainwright & Yu 2019)

## Back to our Problem



WE SEE

ASSUME WE KNOW Graph

structure

$\Rightarrow$  zero in  $\Sigma^{-1}$

$$\text{let } \Omega = \{1, 2, 3\}^2$$

$$(\Sigma^{-1})_{ij} = (\Sigma_D^{-1} - UU^T)^{-1} \quad \begin{matrix} \text{OBSERVE} \\ \text{SOME RANK} \\ \text{ONE diff row} \end{matrix} \quad \text{(MATRIX INVERSE)}$$

$$(B - \Omega\Omega^T)^T = B^T + \frac{B^{-1}U\Omega^T B^{-1}}{1 - \Omega^T B^{-1}\Omega} \quad z = \frac{B^{-1}U}{\sqrt{1 - \Omega^T B^{-1}\Omega}}$$

$$= \Sigma_D^{-1} + \underbrace{zz^T}_{\text{RANK ONE}}$$

$$(i,j) \notin E \Rightarrow \Sigma_{ij}^{-1} = 0$$

$$\Omega = (\Sigma_D^{-1})_{ij} + zz^T \quad \text{so for every missing edge we get an equation.}$$

$$(\hat{\Sigma}^{-1})_{ij} = B_{ij}$$

$B_{ij}^2 = z_i^T z_j$

 $\log B_{ij}^2 = \log z_i^T + \log z_j^T$ 

LINEAR

STRUCTURAL ABOUT INVERSE COVARIANCE TO CREATE A SEQUENCE  
of linear Equations  $\Rightarrow$  solved those.

$\Rightarrow$  Solution to the considered case

### IN THE NOTES

- Higher Rank Correlations
- How can we learn (structure REARR)

### REVIEW:

- WEAK Supervision  $\nrightarrow$  Incremental
- formal theory.
- Nugget Graphs  $\nrightarrow$  Prob distribution
- "METHODS OF MOMENTS" STYLE of Algorithms