

# Wildfire Smoke and Voting Behavior in the United States

Preliminary Results

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# Motivation

- Wildfire smoke is a very widespread *experiential* consequence of climate change in the U.S.
  - Unlike ambient air pollution, smoke events are visible, sudden, and directly attributable to wildfires — making them potentially more *salient* as climate signals
  - Tens of millions of Americans experience unhealthy air from wildfire smoke each year; awareness is high and rising
- Does smoke exposure change whether and how people vote?

# Prior Work

- **Fire proximity** → pro-environment voting in CA, but only among Democrats (Hazlett and Mildemberger, 2020, 2006–2012 CA elections)
- **Overall air pollution ( $PM_{10}$ )** → anti-incumbent voting in Germany (Bellani et al., 2024)
- **Rain on election day** → lower turnout (Gomez et al., 2007)

**Gap:** No study has linked wildfire-specific smoke  $PM_{2.5}$  to U.S. election outcomes

- Smoke differs from generic air pollution: episodic, visible, attributable to a specific cause
- Smoke differs from fire proximity: affects far more people, plausibly exogenous (wind-driven), isolates the experiential channel from property destruction

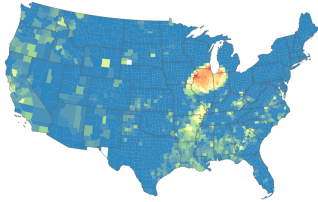
1. **Wildfire smoke  $\text{PM}_{2.5}$**  — Childs et al. (2022)
  - Daily, county-level, 2006–2023
  - Uses NOAA HMS satellite smoke plume classifications combined with machine learning to isolate wildfire-attributed  $\text{PM}_{2.5}$  from background pollution
2. **Election returns** — MIT Election Data Lab (MIT Election Data + Science Lab, 2024)
  - Presidential: county-level (2008, 2012, 2016, 2020)
  - House: precinct-level returns aggregated to county (2016, 2018, 2020, 2022)
3. **Analysis samples:**
  - Presidential: 12,428 county  $\times$  election obs (4 cycles)
  - House: 12,206 county  $\times$  election obs (4 cycles)

# Smoke Exposure Varies Dramatically Across Elections

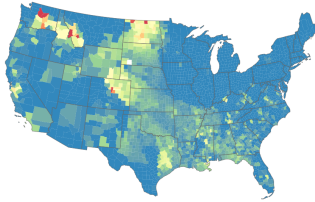
## Pre-Election Wildfire Smoke Exposure by County

Mean wildfire-attributed PM<sub>2.5</sub> in the 30 days before election day

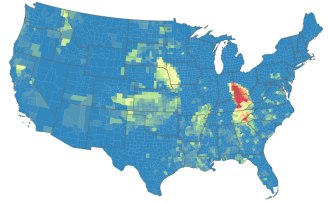
2008



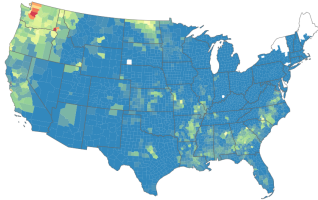
2012



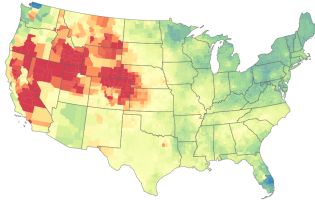
2016



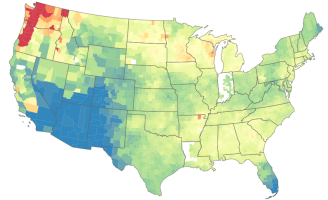
2018



2020



2022



# Empirical Strategy

Two-way fixed effects:

$$Y_{ct} = \alpha_c + \gamma_t + \beta \cdot \text{SmokePM}_{ct} + \varepsilon_{ct}$$

- $\alpha_c$ : County FE — absorb all time-invariant confounders
- $\gamma_t$ : Election year FE — absorb national swings
- SEs clustered by county
- Treatment: mean smoke  $\text{PM}_{2.5}$  in the 60 days before election

**Identifying assumption:** Conditional on county and year FE, variation in smoke exposure is uncorrelated with unobserved determinants of voting. This is plausible because smoke plume direction is determined by wind, not by county politics or demographics.

# Identification: Threats

## Potential threats:

- Spatially correlated shocks (e.g., drought affects both fires and local economy)
  - Mitigated: smoke travels hundreds of miles from fire origin
- Secular trends in fire-prone vs. non-fire-prone regions
  - Mitigated: county FE absorb levels; year FE absorb national trends

# TWFE with Continuous Treatment

Callaway et al. (2024) show TWFE with a continuous treatment can produce coefficients with ambiguous causal interpretation due to heterogeneous dose-response weighting.

## Why this is less concerning here:

- Treatment is atmospherically assigned — limiting selection into dose levels
- We estimate a linear slope corresponding to the average causal response (ACRT) decomposition, in which weights are non-negative
- Robustness: results are qualitatively similar when treatment is dichotomized or discretized into dose bins
- Note: the `contdid` estimator (Callaway et al., 2024) does not yet support time-varying continuous treatments (our setting), so the standard TWFE approach remains appropriate here



# Main Results: Presidential and House Elections

	(1)	(2)	(3)
	DEM Vote Share	Incumbent Share	Log Turnout
<i>Panel A: Presidential (2008–2020)</i>			
Smoke PM <sub>2.5</sub> (60d)	0.00077*** (0.00008)	−0.00379*** (0.00039)	0.00213*** (0.00016)
<i>N</i>	12,428	12,428	12,428
<i>Panel B: County-Level House (2016–2022)</i>			
Smoke PM <sub>2.5</sub> (60d)	0.00023** (0.00010)	−0.00116** (0.00047)	0.00114*** (0.00042)
<i>N</i>	11,155	11,155	12,197

County and year FE. SEs clustered by county. \*\*\*  $p < 0.01$ .

# Interpreting the Main Results

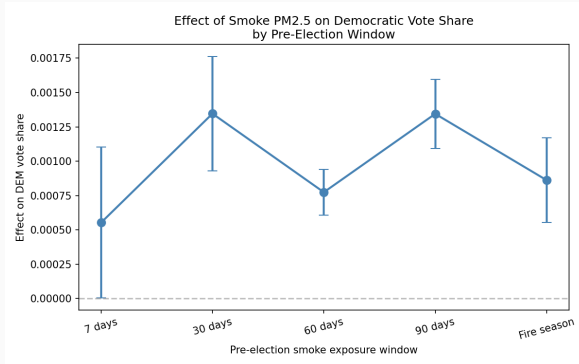
- **+10  $\mu\text{g}/\text{m}^3$  smoke  $\rightarrow$  +0.8 pp DEM vote share** (presidential)
- Anti-incumbent effect is  $\sim 5\times$  larger than pro-DEM effect
- No evidence of turnout suppression
- House effects are present and significant across 4 election cycles (2016–2022), but smaller in magnitude — consistent with the more candidate-driven nature of House races

## Effect Across the Partisan Spectrum

	<i>DEM Vote Share</i>		
	R-Leaning	Swing	D-Leaning
Smoke PM <sub>2.5</sub> (60d)	0.00039** (0.00017)	0.00049*** (0.00013)	0.00086*** (0.00013)
<i>N</i>	4,143	4,140	4,144

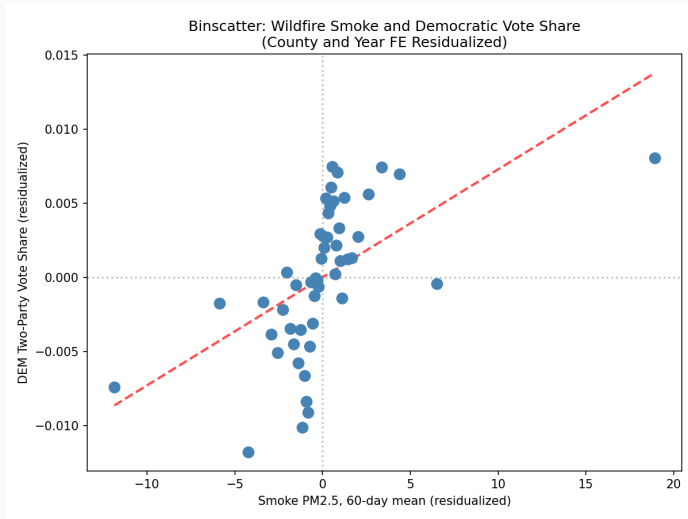
- Effect is **present in all terciles** of prior partisanship
- Somewhat larger in D-leaning counties
- Contrast with Hazlett and Mildenberger (2020, 2006–2012 CA): fire proximity affects *only* Democratic areas
- Smoke is a broader, less politically sorted treatment

# Temporal Dynamics



- Effect significant at all windows
- Strongest at 30 days
- Consistent with recency / salience mechanism
- Not just election-day disruption

# Binscatter: Smoke and Democratic Vote Share



County and year FE residualized. 50 equal-sized bins of smoke exposure.

## Robustness: Excluding 2020

	Presidential		County House	
	Full	Excl. 2020	Full	Excl. 2020
<i>Panel A: DEM Vote Share</i>				
Smoke PM <sub>2.5</sub> (60d)	0.00077*** (0.00008)	−0.00036* (0.00022)	0.00023** (0.00010)	0.00009 (0.00031)
<i>Panel B: Incumbent Vote Share</i>				
Smoke PM <sub>2.5</sub> (60d)	−0.00379*** (0.00039)	−0.01368*** (0.00141)	−0.00116** (0.00047)	0.00362*** (0.00124)
<i>Panel C: Log Total Votes</i>				
Smoke PM <sub>2.5</sub> (60d)	0.00213*** (0.00016)	0.00068*** (0.00025)	0.00114*** (0.00042)	0.00769*** (0.00159)

# Interpreting the 2020 Sensitivity

- The pro-DEM effect is not robust to excluding 2020 in either presidential or House races — the extreme Western fire season provides much of the identifying variation
- The **presidential anti-incumbent effect is robust and strengthens** without 2020 ( $-0.014$  vs.  $-0.004$ ), suggesting it is not an artifact of that single year
- The House anti-incumbent effect **flips sign** without 2020, indicating it is less robust than the presidential result
- Incumbent punishment in presidential races is the most robust finding across specifications

# Limitations and Next Steps

## Current limitations:

- Only 4 presidential elections; 4 House elections (smoke data: 2006–2023; elections through 2022)
- Pro-DEM shift is leveraged by the 2020 fire season
- County-level aggregation; no individual-level variation
- Turnout measure is crude (log total votes without population denominator)

## Planned extensions:

- Add 2024 elections as precinct data become available
- State legislative elections
- Wind direction as instrument for smoke exposure
- Conley spatial SEs for inference robust to spatial correlation



# Summary

1. Wildfire smoke **punishes incumbents** — the most robust finding
2. Smoke also **shifts votes toward Democrats**, but this is driven by the 2020 fire season
3. Effects are **nationally representative** and **cross the partisan spectrum**
4. Smoke is **plausibly exogenous** (wind-driven) and affects **far more people** than fire proximity

## References

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