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Statistical Modeling

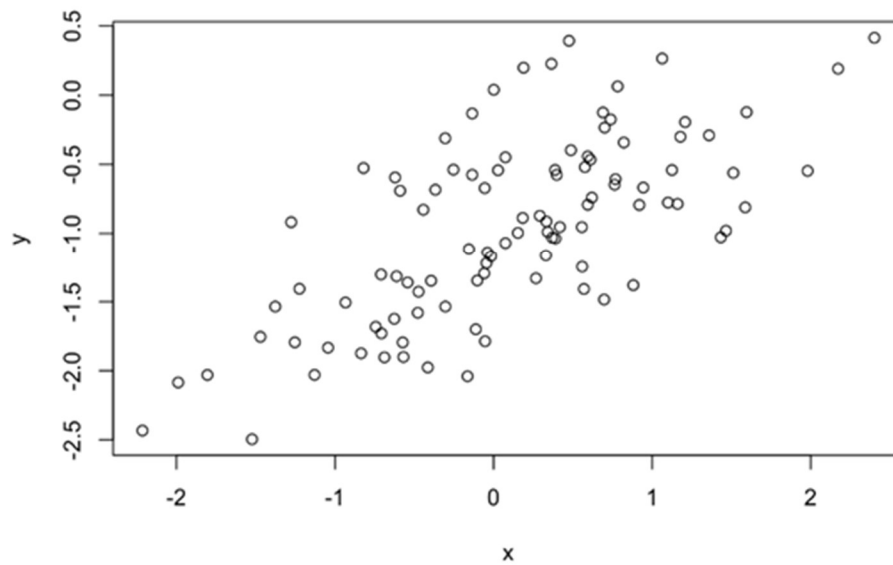
September 21, 2022

### Teamwork Formal Presentation and Submission Problems 1

#### Problem 13:

c)  $y$  is of length 100.  $\beta_0$  is -1 and  $\beta_1$  is 0.5

d) The plot of  $x$  and  $y$  is below



There is a positive relationship between  $x$  and  $y$ . The estimates for both betas seem to be very close to the actual values.

e) The summary of the fit is below:

Call:

`lm(formula = y ~ x)`

Residuals:

	Min	1Q	Median
	-0.93842	-0.30688	-0.06975
	3Q	Max	
	0.26970	1.17309	

Coefficients:

	Estimate	Std. Error	
(Intercept)	-1.01885	0.04849	
x	0.49947	0.05386	
	t value	Pr(> t )	
(Intercept)	-21.010	< 2e-16	***
x	9.273	4.58e-15	***

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Signif. codes:

0 '\*\*\*' 0.001 '\*\*' 0.01 '\*'  
0.05 '.' 0.1 ' ' 1

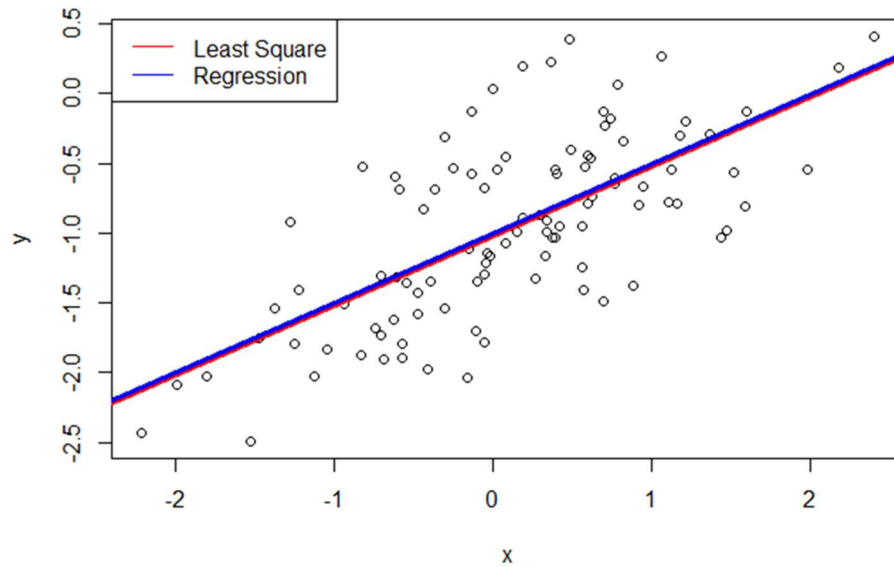
Residual standard error: 0.4814 on 98 degrees of freedom

Multiple R-squared: 0.4674, Adjusted R-squared: 0.4619

F-statistic: 85.99 on 1 and 98 DF, p-value: 4.583e-15

The linear regression fits a model close to the true value of the coefficients as was constructed. The model has a large F-statistic with a p-value close to 0 so the  $H_0$  can be rejected.

f) The plot with the 2 model lines is shown below:



g) There is evidence that model fit has increased over the training data given the slight increase in  $R^2$  and RSE. However, the p-value of the t-statistic suggests that there isn't a relationship between  $y$  and  $x^2$ . The summary of fit\_sq is shown below:

Call:

```
lm(formula = y ~ x + I(x^2))
```

Residuals:

Min	1Q	Median	3Q	Max
-0.98252	-0.31270	-0.06441	0.29014	1.13500

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.97164	0.05883	-16.517	< 2e-16 ***
x	0.50858	0.05399	9.420	2.4e-15 ***
I(x^2)	-0.05946	0.04238	-1.403	0.164

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.479 on 97 degrees of freedom

Multiple R-squared: 0.4779, Adjusted R-squared: 0.4672

F-statistic: 44.4 on 2 and 97 DF, p-value: 2.038e-14

h) The error seen in  $R^2$  and the RSE both decrease significantly, which is expected. The summary and the plot for `lm.fit` are shown below:

Call:

`lm(formula = y1 ~ x1)`

Residuals:

Min	1Q	Median	3Q	Max
-0.136567	-0.028264	0.001012	0.031550	0.131670

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.998814	0.005173	-193.09	<2e-16 ***
x1	0.505777	0.005235	96.61	<2e-16 ***

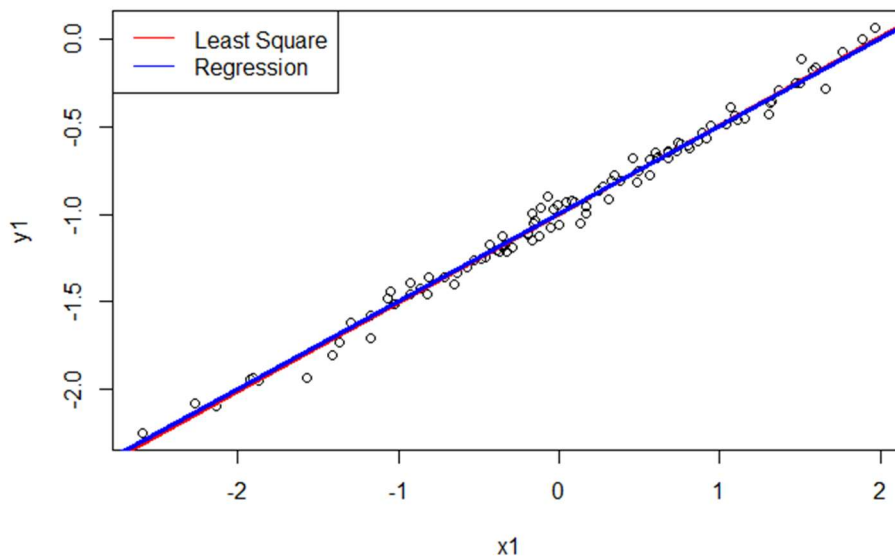
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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.05166 on 98 degrees of freedom

Multiple R-squared: 0.9896, Adjusted R-squared: 0.9895

F-statistic: 9333 on 1 and 98 DF, p-value: < 2.2e-16



i) The error seen in  $R^2$  and the RSE both increase significantly from part h), which is expected. The summary and the plot for `lm.fit2` are shown below:

Call:

```
lm(formula = y2 ~ x2)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.16208	-0.30181	0.00268	0.29152	1.14658

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.94557	0.04517	-20.93	<2e-16 ***
x2	0.49953	0.04736	10.55	<2e-16 ***

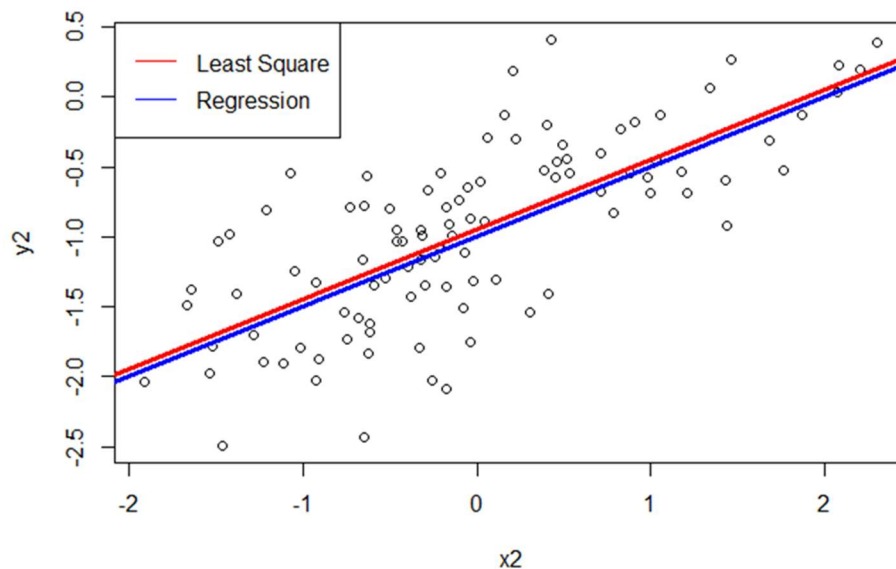
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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.4514 on 98 degrees of freedom

Multiple R-squared: 0.5317, Adjusted R-squared: 0.5269

F-statistic: 111.2 on 1 and 98 DF, p-value: 2.2e-16



j) All 3 intervals seem to be centered on about 0.5, with the second fit's interval being narrowest and the last fit's interval being widest. All three intervals are printed below in order.

2.5 % 97.5 %

(Intercept) -1.1150804 -0.9226122

x 0.3925794 0.6063602

> confint(lm.fit)

2.5 % 97.5 %

(Intercept) -1.0090795 -0.9885493

x1 0.4953877 0.5161661

> confint(lm.fit2)

2.5 % 97.5 %

(Intercept) -1.0352203 -0.8559276

x2 0.4055479 0.5935197

Conclusion Paragraph: The above data showcases that all three of the models have similar performances, and we can see that in the confidence intervals above that are all centered on approximately 0.5. The narrowest interval is the second model's and the widest interval is the first model's, with the third model only slightly more narrow than the first one. This leads us to the conclusion that as the noise increases the interval widens and the model becomes less predictable, and as the noise decreases the interval becomes narrower and the model becomes more predictable.

## Supplemental:

To analyze the data, we used a linear model with all variables involved and a correlation of all variables. Some predictors that look important are:

- zn (proportion of residential land zoned for lots over 25,000 sq.ft.),
- Dis (weighted mean of distances to five Boston employment centres.),
- Rad (index of accessibility to radial highways.),
- and medv (median value of owner-occupied homes in \$1000s.).

	crim	zn	indus	chas	nox	rm	age	dis
crim	1.00000000	-0.20046922	0.40658341	-0.055891582	0.42097171	-0.21924670	0.35273425	-0.37967009
zn	-0.20046922	1.00000000	-0.53382819	-0.042696719	-0.51660371	0.31199059	-0.56953734	0.66440822
indus	0.40658341	-0.53382819	1.00000000	0.062938027	0.76365145	-0.39167585	0.64477851	-0.70802699
chas	-0.05589158	-0.04269672	0.06293803	1.00000000	0.09120281	0.09125123	0.08651777	-0.09917578
nox	0.42097171	-0.51660371	0.76365145	0.091202807	1.00000000	-0.30218819	0.73147010	-0.76923011
rm	-0.21924670	0.31199059	-0.39167585	0.091251225	-0.30218819	1.00000000	-0.24026493	0.20524621
age	0.35273425	-0.56953734	0.64477851	0.086517774	0.73147010	-0.24026493	1.00000000	-0.74788054
dis	-0.37967009	0.66440822	-0.70802699	-0.099175780	-0.76923011	0.20524621	-0.74788054	1.00000000
rad	0.62550515	-0.31194783	0.59512927	-0.007368241	0.61144056	-0.20984667	0.45602245	-0.49458793
tax	0.58276431	-0.31456332	0.72076018	-0.035586518	0.66802320	-0.29204783	0.50645559	-0.53443158
ptratio	0.28994558	-0.39167855	0.38324756	-0.121515174	0.18893268	-0.35550149	0.26151501	-0.23247054
lstat	0.45562148	-0.41299457	0.60379972	-0.053929298	0.59087892	-0.61380827	0.60233853	-0.49699583
medv	-0.38830461	0.36044534	-0.48372516	0.175260177	-0.42732077	0.69535995	-0.37695457	0.24992873
	rad	tax	ptratio	lstat	medv			
crim	0.625505145	0.58276431	0.2899456	0.4556215	-0.3883046			
zn	-0.311947826	-0.31456332	-0.3916785	-0.4129946	0.3604453			
indus	0.595129275	0.72076018	0.3832476	0.6037997	-0.4837252			
chas	-0.007368241	-0.03558652	-0.1215152	-0.0539293	0.1752602			
nox	0.611440563	0.66802320	0.1889327	0.5908789	-0.4273208			
rm	-0.209846668	-0.29204783	-0.3555015	-0.6138083	0.6953599			
age	0.456022452	0.50645559	0.2615150	0.6023385	-0.3769546			
dis	-0.494587930	-0.53443158	-0.2324705	-0.4969958	0.2499287			
rad	1.000000000	0.91022819	0.4647412	0.4886763	-0.3816262			
tax	0.910228189	1.00000000	0.4608530	0.5439934	-0.4685359			
ptratio	0.464741179	0.46085304	1.0000000	0.3740443	-0.5077867			
lstat	0.488676335	0.54399341	0.3740443	1.0000000	-0.7376627			
medv	-0.381626231	-0.46853593	-0.5077867	-0.7376627	1.0000000			

The following summary is a result of all the predictors being used.

Call:

```
lm(formula = crim ~ ., data = Boston)
```

Residuals:

Min	1Q	Median	3Q	Max
-8.534	-2.248	-0.348	1.087	73.923

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	13.7783938	7.0818258	1.946	0.052271	.
zn	0.0457100	0.0187903	2.433	0.015344	*
indus	-0.0583501	0.0836351	-0.698	0.485709	
chas	-0.8253776	1.1833963	-0.697	0.485841	
nox	-9.9575865	5.2898242	-1.882	0.060370	.
rm	0.6289107	0.6070924	1.036	0.300738	
age	-0.0008483	0.0179482	-0.047	0.962323	
dis	-1.0122467	0.2824676	-3.584	0.000373	***
rad	0.6124653	0.0875358	6.997	8.59e-12	***
tax	-0.0037756	0.0051723	-0.730	0.465757	
ptratio	-0.3040728	0.1863598	-1.632	0.103393	
lstat	0.1388006	0.0757213	1.833	0.067398	.
medv	-0.2200564	0.0598240	-3.678	0.000261	***

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6.46 on 493 degrees of freedom

Multiple R-squared: 0.4493, Adjusted R-squared: 0.4359

F-statistic: 33.52 on 12 and 493 DF, p-value: < 2.2e-16



With just the 4 mentioned predictors, the F-statistic increases from 33.52 to 95.84.

Call:

```
lm(formula = crim ~ zn + dis + rad + medv, data = Boston)
```

Residuals:

```
      Min       1Q   Median       3Q      Max
-8.459 -1.960 -0.331  0.857  74.718
```

Coefficients:

```
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  5.26548     1.34674   3.910 0.000105 ***
zn           0.05487     0.01735   3.163 0.001658 **
dis          -0.72291     0.20254  -3.569 0.000393 ***
rad           0.50021     0.04044  12.370 < 2e-16 ***
medv         -0.19122     0.03566  -5.362 1.26e-07 ***
```

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6.5 on 501 degrees of freedom

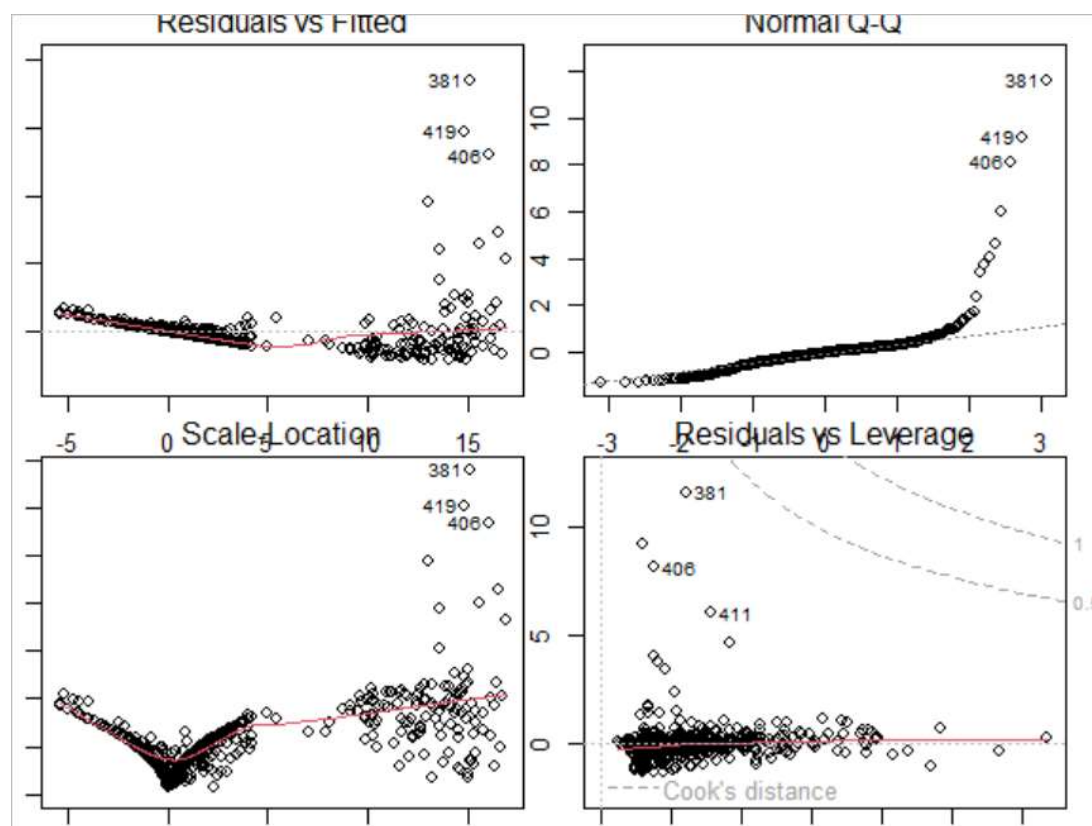
Multiple R-squared: 0.4335, Adjusted R-squared: 0.429

F-statistic: 95.84 on 4 and 501 DF, p-value: < 2.2e-16

We ran many different models on the data:

Name	RSE	Adjusted R^2	F-Statistic
All predictors	6.46	.4359	33.52
Medv,dis,rad,zn	6.5	.429	95.84
Medv,dis^0.5,rad,zn	6.465	.4351	98.24
Medv,dis,rad,zn^0.5	6.505	.4281	95.52
Crim^0.5~all predictors	.6936	.7716	143.2
Ln(Crim)~medv,dis,zn,rad	.878	.8349	639.4
Ln(crim)~all predictors	.781	.8694	281
Ln(crim)~all predictors on Boston1	.768	.8704	282

We analyzed the diagnostic plots for linear model with all predictors and found that the residual values increased significantly at higher fitted values. We tried to square the response but found that the natural log was better. The 4 predictors we picked out were a worse overall model than using all predictors. Finally, we removed 3 outliers we believed were skewing out data and renamed the dataset "Boston1" and see the model is only slightly better.



The best model we found was the take the natural log of the response and to include all the predictors.

Call:

```
lm(formula = I(log(crim, base = 2.72)) ~ ., data = Boston)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-2.58529	-0.56856	-0.04957	0.47295	2.66877

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	-4.3784836	0.8561394	-5.114	4.52e-07	***
zn	-0.0115074	0.0022716	-5.066	5.76e-07	***
indus	0.0208393	0.0101109	2.061	0.03982	*
chas	-0.0632406	0.1430637	-0.442	0.65865	
nox	3.9152451	0.6394999	6.122	1.88e-09	***
rm	-0.0093813	0.0733929	-0.128	0.89834	
age	0.0055267	0.0021698	2.547	0.01117	*
dis	-0.0104253	0.0341482	-0.305	0.76027	
rad	0.1475944	0.0105824	13.947	< 2e-16	***
tax	-0.0001312	0.0006253	-0.210	0.83394	
ptratio	-0.0476792	0.0225295	-2.116	0.03482	*
lstat	0.0341910	0.0091541	3.735	0.00021	***
medv	0.0062483	0.0072323	0.864	0.38803	

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.781 on 493 degrees of freedom

Multiple R-squared: 0.8725, Adjusted R-squared: 0.8694

F-statistic: 281 on 12 and 493 DF, p-value: < 2.2e-16

In this model, we used all other 12 predictors to help predict the Boston suburbs' crime rate. In the model created, we transformed the response variable by taking the natural log of each which resulted in the adjusted R<sup>2</sup> being 86.94%. After removing 3 outliers and renaming the dataset "Boston1" we got our R<sup>2</sup> up to 87.04%. We chose to use this model because out of all the ones we tested included other transforming functions of the response variable, using only the predictor variable with the lowest p-scores, and transforming the predictor variables as well, the model with just taking the natural log of the response variable explained the most variability of the response and had a high F-Statistic.