

Stochastic and combinatorial optimization

- Unconstrained optimization:
 - Input variables are continuous
 - Response is differentiable

We could apply Steepest descent, Newton, BFGS, CG

Now:

- Variables can be discrete (scheduling problem, traveling salesman)
- Outcome can be discrete, noisy (typical in statistics) or having multiple local minima

Stochastic and combinatorial optimization

Given set of states S, objective to minimize *f(s)*, typically S is large.

- Sometimes, exhaustive search is possible (shortest path algorithm)
- Often exhaustive search is computationally expensive, sometimes NP-hard (traveling salesman)
- Alternative a solution (sometimes exact) can be obtained by stochastic methods (ex: simulated annealing, genetic algorithms)

Idea comes from physics (melted metal is being cooled)

- Parameters:
 - Energy of the metal (decreasing, but not monotonic)
 - Temperature (decreasing)

How to find minimum energy (global one)?

- 0. Set k = 1 and initialize state s.
- 1. Compute the temperature T(k).
- 2. Set i = 0 and j = 0.
- 3. Generate a new state r and compute $\delta f = f(r) f(s)$.
- 4. Based on δf , decide whether to move from state s to state r. If $\delta f < 0$,

accept state r;

otherwise,

accept state r with a probability $P(\delta f, T(k))$.

If state r is accepted, set s = r and i = i + 1.

- If i is equal to the limit for the number of successes at a given temperature, go to step 1.
- 6. Set j = j + 1. If j is less than the limit for the number of iterations at given temperature, go to step 3.
- 7. If i = 0, deliver s as the optimum; otherwise, if $k < k_{\text{max}}$, set k = k + 1 and go to step 1; otherwise, issue message that 'algorithm did not converge in k_{max} iterations'.

Comments

- Check https://www.youtube.com/watch?v=iaq Fpr4KZc
- How to generate new state?
 - Continuous inputs: choose new point at some distance r from the current point (r can be a random variable..)
 - Discrete inputs: same idea or rearrangements
- How to choose selection probability
 - Sometimes chosen as exp(-δf/T)
- How to choose temperature function?
 - continuous or noisy functions, taken constant
 - Another choice

$$T(k+1) = b(k)T(k), \ b(k) = (\log(k))^{-1}$$

Example: Traveling salesman

- Assume constant temperature
 - 1. Choose initial configuration $(A_1...A_n)$
- 2. Generate new configuration by 2-rearrangement:

$$(1, \underline{2, 3, 4, 5, 6}, 7, 8, 9) \rightarrow (1, \underline{6, 5, 4, 3, 2}, 7, 8, 9).$$

 $1, \underline{2, 3, 4, 5, 6}, 7, 8, \uparrow 9) \rightarrow (1, 7, 8, \underline{2, 3, 4, 5, 6}, 9)$

- 3. Measure difference in path length δf between new and old configuration
- If shorter path found, accept it otherwise accept it with probability P(δf)
- Repeat until maximum iteration condition fulfilled

- Idea comes from biology (the fittest survives)
- Variables=genotypes
- Observation=organism, characterized by genetic code
- State space= Population of organisms
- Objective function=fitness of organism

New points are obtained from old points by crossover and mutation, the population retains only fittest organisms (with better objective function)

- How to code the points
 - Enumerate each point in S
 - Code for observation i is presented by binary representation of i
 - Other encodings are possible
- Mutation and recombination rules

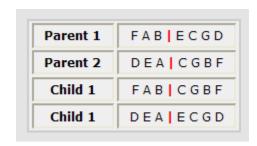
Generation k	Generation $k+1$				
G.					
Crossover					
$x_i^{(k)}$ 11001001	(7- 1-4)				
→	$x_i^{(k+1)}$ 11011010				
$x_j^{(k)}$ 00111010					
	rsion				
$x_i^{(k)}$ 11101011 \rightarrow	$x_i^{(k+1)}$ 11010111				
Mutation					
$x_i^{(k)}$ 11101011 \rightarrow	$x_i^{(k+1)}$ 10111011				
	one				
$x_i^{(k)}$ 11101011 \rightarrow	$x_i^{(k+1)}$ 11101011				
732A38					

- 0. Determine a representation of the problem, and define an initial population, $x_1^{(0)}, x_2^{(0)}, \ldots, x_n^{(0)}$. Set k = 0.
- 1. Compute the objective function (the "fitness") for each member of the population, $f(x_i^{(k)})$ and assign probabilities p_i to each item in the population, perhaps proportional to its fitness.
- 2. Choose (with replacement) a probability sample of size $m \leq n$. This is the reproducing population.
- 3. Randomly form a new population $x_1^{(k+1)}, x_2^{(k+1)}, \ldots, x_n^{(k+1)}$ from the reproducing population, using various mutation and recombination rules (see Table 6.2). This may be done using random selection of the rule for each individual of pair of individuals.
- 4. If convergence criteria are met, stop, and deliver $\arg\min_{x_i^{(k+1)}} f(x_i^{(k+1)})$ as the optimum; otherwise, set k = k + 1 and go to step 1.

Traveling salesman problem

Encoding and crossover

First idea - encode tours as $A_1...A_n$. Problem:



- 1. Instead: Remove FAB from DEACGBF -> DECG. Obtain first child by appending: FABDECG
- 2. Second child is obtained by taking prefix from parent 2

Traveling salesman problem

Mutation

- If taking small population and using only crossover the input domain becomes limited, may converge to local solution
- Taking large initial population may be computationally heavy
- Mutation allows to investigate entire input domain
- In traveling salesman, mutation =moving a city in the tour to another position

Traveling salesman problem

Other issues

- Reproduction: Among m tours selected at step 2, two best are selected for reproduction, two worst replaced by children
- Neighborhood size: large m some tours are never parents, global solution may not be attained
- Mutation probalility should be fixed

Stochastic gradient descent

Machine learning models minimize empirical risk

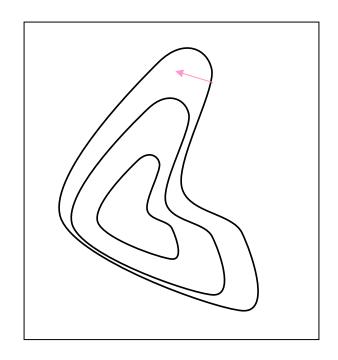
$$R(\theta, \hat{\theta}) = \frac{1}{N} \sum_{i=1}^{N} L(\theta, \hat{\theta})$$

- Regression models: $R(Y, \hat{Y}) = (Y \hat{Y})^2$
- Usual gradient descent

$$\theta_{k+1} = \theta_k - \alpha_k \nabla R(\theta, \theta_k)$$

$$\theta_{k+1} = \theta_k - \alpha_k \frac{1}{N} \sum_{i=1}^{N} \nabla L(\theta, \theta_k)$$

Takes a lot of time per iteration for large N



Stochastic gradient descent

- Main idea: $L(\theta_i, \hat{\theta}) \approx \frac{1}{N} \sum_{i=1}^{N} \nabla L(\theta, \theta_k)$ for any i
 - Update formula $\theta_{k+1} = \theta_k \alpha_k \nabla L(\theta, \theta_k)$
 - Second order gradient descent $\theta_{k+1} = \theta_k \alpha_k \Gamma_k \nabla L(\theta, \theta_k)$
 - Converges to a local minimum but may require a lot more iterations than gradient descent
- Less time per iteration (by factor N)
- Converges to the same kind of optimal point as the gradient descent
- Can be used in online learning

Stochastic gradient descent

- Use package sgd in R:
 - Linear model, GLM, Cox model,...

			- U		
r.dioxide [‡]	density ‡	pH [‡]	sulphates [‡]	alcohol [‡]	quality ‡
170.0	1.0010	3.00	0.45	8.8	1
132.0	0.9940	3.30	0.49	9.5	1
186.0	0.9956	3.19	0.40	9.9	1
186.0	0.9956	3.19	0.40	9.9	1
97.0	0.9951	3.26	0.44	10.1	1
136.0	0.9949	3.18	0.47	9.6	1
170.0	1.0010	3.00	0.45	8.8	1

> sgd.theta

[,1]
[1,] -0.067246995
[2,] -0.498257484
[3,] -0.170646610
[4,] 0.079520738
[5,] -0.007496517
[6,] -0.010871955
[7,] 0.080419024
[8,] -0.055379999
[9,] -0.067940438
[10,] -0.104186087
[11,] 0.052921863
[12,] 0.766583940
> sgd.theta\$converged
[1] TRUE

EM algorithm

- Model depends on the data which are known Y and data that can not be observed Z (latent).
- Data depend on some parameters $oldsymbol{ heta}$

How to compute the ML of model parameters θ ?

- All data known: Apply unconstrained optimization (for ex. Gradient descent)
 - Can not be used because gradient contains unknown data (latent variables)
- Now: Use EM algorithm

EM algorithm

Let

$$Q(\theta, \theta^{t}) = \int \log p(Y, Z \mid \theta) p(Z \mid \theta^{t}, Y) dz = E_{Z \mid \theta^{t}, Y} \log \operatorname{lik}(\theta \mid Y, Z)$$

EM algorithm

- 1. Choose starting point θ^t ,
- 2. E-step: Derive $Q(\theta^t, \theta)$
- 3. M-step: $\theta^{t+1} = \arg \max_{\theta} Q(\theta^t, \theta)$, set t = t + 1
- 4. Repeat until convergence

Example: Normal data with missing values

R: EM Algorithm

Example: Normal data with some missing observations

```
em.norm <- function(Y) {
Yobs <- Y[!is.na(Y)]
Ymiss <- Y[is.na(Y)]
n <- length(c(Yobs, Ymiss))
r <- length(Yobs)
# Initial values
mut <- 1
sit <- 0.1
# Define log-likelihood function
11 <- function(y, mu, sigma2, n) {
-0.5*n*log(2*pi*sigma2)-0.5*sum((y-mu)^2)/sigma2
}
# Compute the log-likelihood for the initial values
1ltm1 <- ll(Yobs, mut, sit, n)</pre>
```

R: EM Algorithm

Example: Normal data with some missing observations

```
repeat{
# E-step
EY <- sum(Yobs) + (n-r)*mut
EY2 \leftarrow sum(Yobs^2) + (n-r)*(mut^2 + sit)
# M-step
mut1 <- EY / n
sit1 \leftarrow EY2 / n - mut1^2
# Update parameter values
mut. <- mut.1
sit <- sit1
# Compute log-likelihood using current estimates
llt <- ll(Yobs, mut, sit, n)</pre>
# Print current parameter values and likelihood
cat(mut, sit, llt, "\n")
# Stop if converged
if (abs(lltm1 - llt) < 0.001) break
lltm1 <- llt
```

R: EM Algorithm

Example: Normal data with some missing observations

```
# Generate complete data N(5,1) and set 5 last as missing values x <- rnorm(20,5)
x[16:20] <- NA
# Run the EM-algorithm function
em.norm(x)
3.856777 3.279395 -33.95757
4.570972 1.523877 -26.37452
4.74952 0.9255991 -23.4036
4.794157 0.7660673 -22.6868
4.805317 0.7255617 -22.53068
4.808106 0.7153964 -22.49403
4.808804 0.7128526 -22.48504
4.808978 0.7122165 -22.4828
4.809022 0.7120575 -22.48225
```

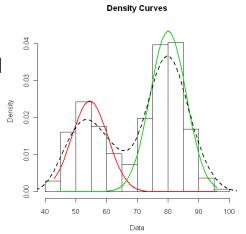
Note that the function only generates parameter values

EM algorithm

- Applications
 - Mixture models (π_k is a latent variable)
 - In regression and classification
 - Mixed data comes from different sources
 - Clustering
 - Density in each cluster is normally distributed
 - Cluster label is latent

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

Maximizing likelihood directly leads to numerical problems → latent class variables are introduced and EM is used



EM algorithm for gaussian mixtures

- 1. Initialize the means μ_k , covariances Σ_k and mixing coefficients π_k , and evaluate the initial value of the log likelihood.
- 2. **E step**. Evaluate the responsibilities using the current parameter values

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}.$$
(9.23)

3. **M step**. Re-estimate the parameters using the current responsibilities

$$\boldsymbol{\mu}_{k}^{\text{new}} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{x}_{n}$$
 (9.24)

$$\Sigma_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \left(\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}} \right) \left(\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}} \right)^{\text{T}}$$
(9.25)

$$\pi_k^{\text{new}} = \frac{N_k}{N} \tag{9.26}$$

where

$$N_k = \sum_{n=1}^{N} \gamma(z_{nk}). {(9.27)}$$

4. Evaluate the log likelihood

$$\ln p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$
(9.28)

and check for convergence of either the parameters or the log likelihood. If the convergence criterion is not satisfied return to step 2.

$$Ez_{nk} = \gamma(z_{nk})$$

Source: Pattern recognition by Bishop