

## Model Based Statistics in Biology.

### Part IV. The General Linear Model. Multiple Explanatory Variables.

#### Chapter 13.7 Hierarchical ANOVA Random within Fixed

To be completed

ReCap. Part I (Chapters 1,2,3,4), Part II (Ch 5, 6, 7)

ReCap Part III (Ch 9, 10, 11)

ReCap Multiple Regression (Ch 12)

Ch13.xls

13.1 Fixed Effects ANOVA - no interactive effects)

13.2 Fixed Effects ANOVA - interactive effects)

13.3 Fixed\*Random Effects - Paired t-test)

13.4 Fixed\*Random Effects - Randomized Block)

13.5 Repeated Measures, within subject

13.6 Nested Random Effects - Hierarchical ANOVA)

13.7 Random within Fixed (problem of confounding)

13.8 More Than Two Factors (to be written)

on chalk board

**ReCap** Part I (Chapters 1,2,3,4) Quantitative reasoning is based on models, including statistical analysis based on models.

**ReCap** Part II (Chapters 5,6,7)

Hypothesis testing uses the logic of the null hypothesis to declare a decision.

Estimation is concerned with the specific value of an unknown population parameter.

**ReCap** (Ch 9, 10,11) The General Linear Model with a single explanatory variable.

**ReCap** (Ch 12) GLM with more than one regression variable (multiple regression)

**ReCap** (Ch 13) GLM with more than one categorical variable (ANOVA).

Two fixed factors (Ch 13.1, Ch13.2)

Today: Special case of two factor ANOVA: Repeated Measures

One fixed and one random factor (Paired t-test, Randomized block)

## Introduction.

A mixed model consists of at least one fixed factor of interest, and at least one random factor (defined unit or random block).

For nested designs, the main effect likelihood ratio is formed relative to the nested term  
Here is an example of a mixed model nested design.

Mixed model Drug is fixed, Subj is random

Design is Random within Fixed

### Mixed Model - Correct F-ratios

Source	df	SS	MS	F	---->	p
Drug	2	665.68	332.84	1.74		0.23
Subj(Drug)	9	1720.68	191.19	146.88		$>10^3$
Error	<u>12</u>	<u>15.62</u>	1.3017			
Total	23	2401.97				

### Mixed Model - Incorrect F-ratios (all ratios relative to residual)

Source	df	SS	MS	F	---->	p
Drug	2	665.68	332.84	255.7		$>10^3$
Subj(Drug)	9	1720.68	191.19	146.55		$>10^3$
Error	<u>12</u>	<u>15.62</u>	1.3017			
Total	23	2401.97				

The simple rule is test over the random within fixed term.

What about more complex designs?

This is addressed by writing out the expected mean squares.

Forming the likelihood ratios (and  $F$ -ratios) relative to the residual SS is not always correct. Here is a procedure for identifying the correct  $F$ -ratio. To illustrate it, we will use the Extra Sleep example.

Here is a step by step procedure to write out the correct LRs and  $F$ -ratios.

List the terms in the model vertically, as in the ANOVA table

List the same terms horizontally.

In each row, show the row term.

		Drug	Subj%in% Drug	Error
EMS	Drug	Drug		
EMS	Subj%in%Drug		Subj%in% Drug	
EMS	Error			Error

Each EMS includes itself

		Drug	Subj%in% Drug	Error
EMS	Drug	Drug		Error
EMS	Subj%in%Drug		Subj%in% Drug	Error
EMS	Error			Error

Each EMS includes the fixed error term

		Drug	Subj%in% Drug	Error
EMS	Drug	Drug	Subj%in% Drug	Error
EMS	Subj%in%Drug		Subj%in% Drug	Error
EMS	Error			Error

Each EMS includes crossed (or nested) random terms

		Drug	Subj%in% Drug	Error
EMS	Drug	Drug	Subj%in% Drug	Error
EMS	Subj%in%Drug		Subj%in% Drug	Error
EMS	Error			Error

Correct

Denominator MS  
Subj%in% Drug  
Error

Identify the denominator MS for the  $F$ -ratio

The denominator MS cancels all but the term of interest

Incorrect

Denominator MS

Drug MS / Error results in \*two\* uncancelled terms.

The  $F$ -test is confounded.

		Drug	Subj%in% Drug	Error
EMS	Drug	Drug	Subj%in% Drug	Error
EMS	Subj%in%Drug			Error
EMS	Error			Error

In this example each ratio is formed relative to the term below it in the ANOVA table. This will not always be true in more complex analyses. In these, the EMS table must be written out to identify the correct ratio.

## Distinguishing crossed from nested designs.

If the random units are nested within the fixed factor, the interaction terms cannot be estimated. A two-way table of the random and fixed factors will be incomplete. As a result, the interactive effect cannot be estimated. Instead, it is folded into a new factor, a random factor within another factor.

Here is an example of data that appears to be crossed, but is in fact nested.

Table 1 from Gutsell 1951. The effect of sulfamerazine on the erythrocyte and hemoglobin content of trout blood. *Biometrics* 7:171-179.

	Treatment 0		Treatment 5		Treatment 10		Treatment 15		
Fish	A	B	A	B	A	B	A	B	
1	6.7	7	9.9	9.9	10.4	9.9	9.3	11	
2	7.8	7.8	8.4	9.6	8.1	9.6	9.3	9.3	
3	5.5	6.8	10.4	10.2	10.6	10.2	7.2	11	
4	8.4	7	9.3	10.4	8.7	10.4	7.8	9	
5	7	7.5	10.7	11.3	10.7	11.3	9.3	8.4	
6	7.8	6.5	11.9	9.1	9.1	10.9	10.2	8.4	
7	8.6	5.8	7.1	9	8.8	8	8.7	6.8	
8	7.4	7.1	6.4	10.6	8.1	10.2	8.6	7.2	
9	5.8	6.5	8.6	11.7	7.8	6.1	9.3	8.1	
10	7	5.5	10.6	9.6	8	10.7	7.2	11	
Mean	7.2	6.75	9.33	10.14	9.03	9.73	8.69	9.02	

Sulfamerazine is sulfa antibiotic found to be successful in treating furunculosis a serious, septicemic, bacterial disease caused by *Aeromonas salmonicida*. The experiment investigated toxic effects as measured by reduced erythrocyte and hemoglobin content in fish blood. To the author's surprise the drug increased hemoglobin in brown trout, as evident in Table 1.

The table layout suggests that treatment, fish ID, and trough (A or B) are crossed, with repeated measurement of 10 fish.

However the layout was as follows:

*Five pounds of brown trout from one large pool were placed in each of 8 adjacent upper troughs in the hatchery. Each of the 4 treatments was assigned to 2 troughs by random selection, so that the test was run in duplicate and was fully randomized.*

Thus there was 8 troughs x 10 trout/trough = 80 trout.

Consequently none of the two-way interactive effects can be estimated. For example the sample number in the two way table for Treatment x Trough is:

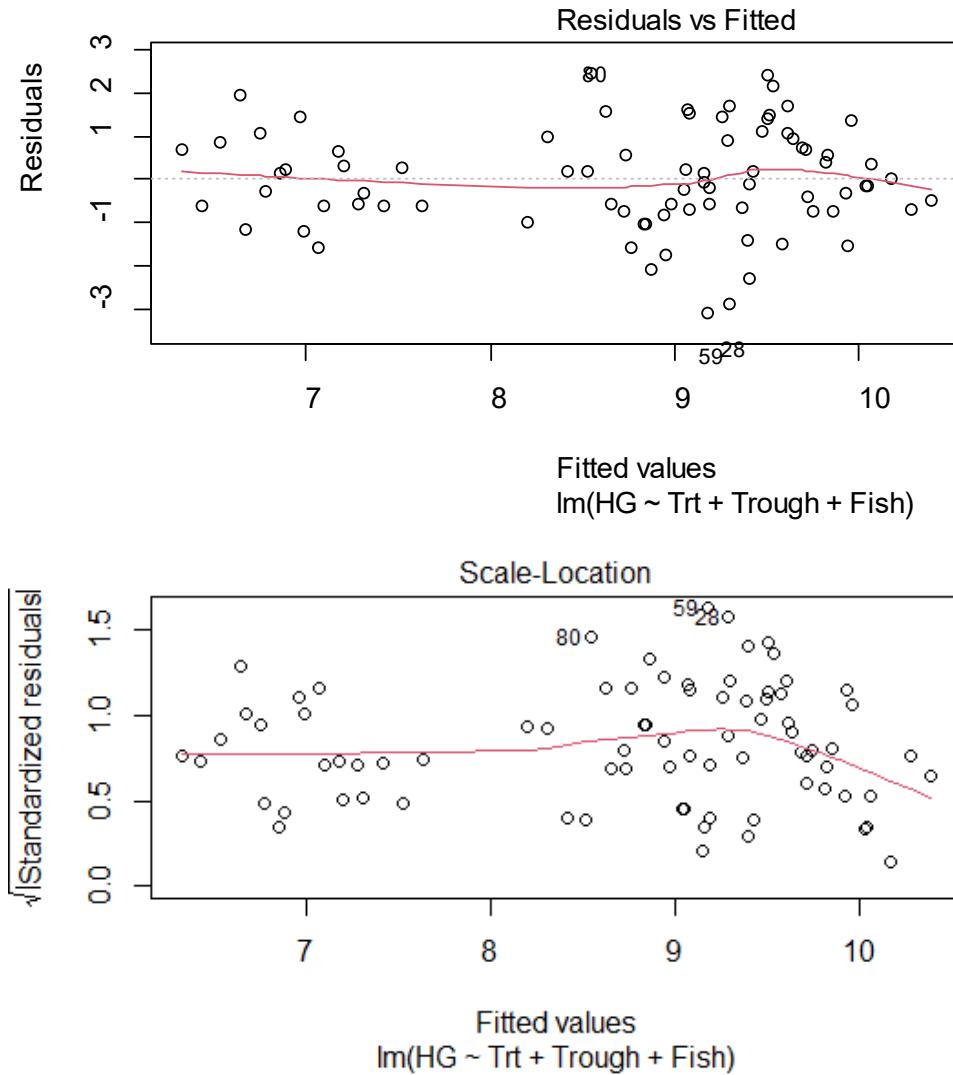
		Trough							
		A	B	C	D	E	F	G	H
Trt 0	10	10							
Trt 5			10	10					
Trt 10				10	10				
Trt 15						10	10		

Similarly sparse displays appear in the two way tables for Fish x Trough and Fish x Treatment.

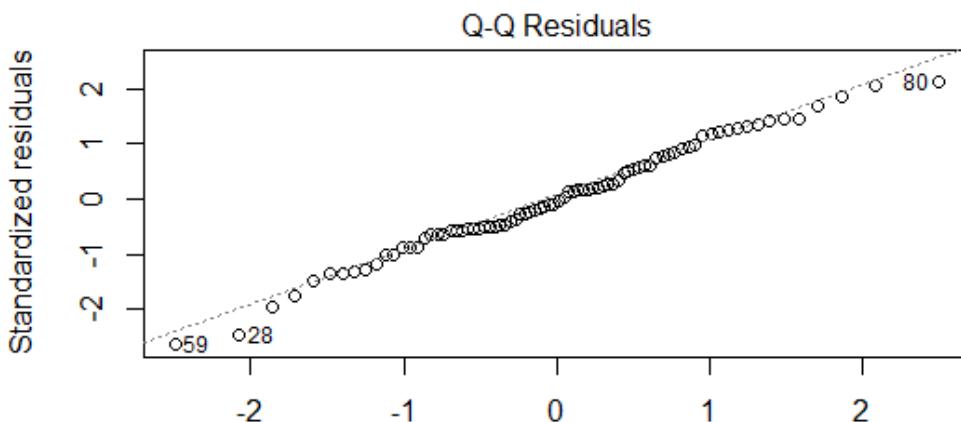
For nested design the interactive effect variance that cannot be estimated appears as the sum or the variance due to main and interactive effects.

Thus  $\text{var}(\text{Trt}) + \text{Var}(\text{Trt} \times \text{Trough})$  become  $\text{var}(\text{Trough}(\text{Treatment}))$ , where the nested term  $\text{Trough}(\text{Treatment})$  is read left to right, as “Trough within Treatment.”

1. Write the model.
2. Execute, obtain fitted values (means) and residuals.
3. Use residuals to check model  
Straight line? Not applicable.  
Homogeneous errors? Only slightly heterogeneous in first plot



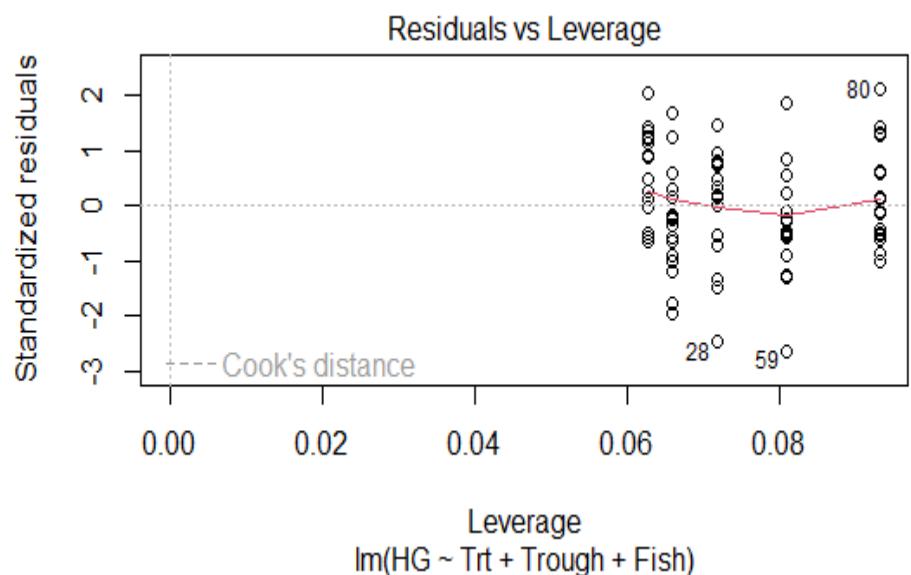
Second plot shows reduction in heterogeneity at large values, suggesting an upper limit to the response variable. This is consistent with lack of increase in hemoglobin with increase in dose beyond the recommended level.



Normal errors?  
Yes

Influential outliers with high leverage?

No.



Conclusion. Normal error model acceptable

Anova Table

Response: HG						
	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
Trt	3	90.560	30.1868	20.4958	9.066e-10	***
Trough	1	2.415	2.4151	1.6398	0.20435	
Fish	1	7.520	7.5200	5.1058	0.02679	*
Residuals	74	108.989	1.4728			

Df correct? Yes  
Df correct? No  
Df correct? No

Model has been executed as crossed design, with 2 troughs instead of 8.  
Fish is regression variable. Somewhat disturbingly Hg depends on fish order.  
 $LR = \{(7.53+109)/109\}^{80/2} = 14.5$  Adequate evidence.

Run model as crossed design to show allocation of interaction term to nested terms.

o.