Exercise 4

Machine Learning I

4A-1.

Prerequisites

Leibnitz rule for constants $-\infty < a, b < +\infty$:

$$\frac{d}{dx}\left(\int_a^b f(x,t)dt\right) = \int_a^b \frac{\partial}{\partial x} f(x,t)dt$$

Additionally, let at least one of the following conditions hold:

(i) $f(x,t)$ is measurable and nonnegative

(ii) $\int_a^b |f(x,t)|dt$ is finite

then we can switch the order of integration according to Tonelli/Fubini, respectively.

Note: In our case, at least one of (i), (ii) is nearly always satisfied. Think about why.

Lastly, we need $(*)$ *Theorem 1* concerning uniform convergence from these notes:

http://www.math.ucla.edu/~tao/resource/general/131bh.1.03s/week45.pdf

Let D be the domain of $\mathbf{x}$.

$$\begin{aligned} E[y(\mathbf{x}) - t] &= \int_{t_1}^{t_2} \int_D (y(\mathbf{x}) - t)^2 p(\mathbf{x}, t) d\mathbf{x}\, dt \\ &= \underbrace{\int_D \int_{t_1}^{t_2} (y(\mathbf{x}) - t)^2 p(\mathbf{x}, t) dt\, d\mathbf{x}}_{Fubini/Tonelli} \\ &= \int_D p(\mathbf{x}) \int_{t_1}^{t_2} (y(\mathbf{x}) - t)^2 p(t|\mathbf{x}) dt\, d\mathbf{x} \end{aligned}$$

Minimizing the loss cumulative loss for all $\mathbf{t}$ equals minimizing the loss for each t_i separately$(**)$. Note: Inside the interior integral, $\mathbf{x}$ is constant. Let $y(\mathbf{x}) = z$:

$$\begin{aligned} \frac{\partial}{\partial z} \int_{t_1}^{t_2} (z - t)^2 p(t|\mathbf{x}) dt &= \underbrace{\int_{t_1}^{t_2} \frac{\partial}{\partial z} (z - t)^2 p(t|\mathbf{x}) dt}_{Leibnitz\ rule} \\ &= 2 \int_{t_1}^{t_2} (z - t) p(t|\mathbf{x}) dt \end{aligned}$$

We can now solve for z:

$$
\begin{aligned}
& 2\int_{t_1}^{t_2}(z-t)p(t|\mathbf{x})dt && = 0 \\
\Leftrightarrow \quad & \underbrace{z\int_{t_1}^{t_2}p(t|\mathbf{x})dt}_{=1} && = \int_{t_1}^{t_2}tp(t|\mathbf{x})dt \\
\Leftrightarrow \quad & y(\mathbf{x}) && = E[t|\mathbf{x}]
\end{aligned}
$$

According to the Leibnitz rule, this only holds for finite limits t_1, t_2. To extend this proof to the infinite domain, we construct the sequence:

$$
f_n' = \frac{\partial}{\partial z}\int_{-n}^{n}(z-t)^2 p(t|\mathbf{x})dt
$$

Because probabilities sum to one, if n tends to infinity, $p(t|\mathbf{x})$ vanishes for most t_i. The $(z-t)^2$ will not compensate that, as we required $E[y(\mathbf{x})-t]$ to be finite earlier.

Accordingly, $\lim\limits_{n\to\infty} f_n{}'$ converges to g uniformly. Due to $(*)$, the functions f_n converge uniformly to f, with $f' = g$.

Spoken plainly, this means if we have an infinite domain $\mathbb{R}$, we can approximate the solution arbitrarily close by increasing $[t_1, t_2]$.

$(**)$ Only because the x_i are independent. In our situation this is the case, otherwise we would also have to integrate over all possibilities $p(x|x_i, \dots x_0)$.

4A-2.

Auxiliary calculation

Using the Jacobian integral substitution, the area an infinitesimal d-dimensional volume element is

$$d^nV = \left|\det \frac{\partial(x_i)}{\partial(r,\phi_j)}\right| drd\theta_1 \dots d\theta_{n-1}$$

$$\begin{aligned} d^nV &= \left|\det \frac{\partial(x_i)}{\partial(r,\phi_j)}\right| drd\theta_1 \dots d\theta_{n-1} \\ &= r^{n-1} \underbrace{\sin^{n-2}\theta_1 \sin^{n-3}\theta_2 \cdot \dots \cdot \sin\theta_{n-2}}_{g(\theta_1,\dots,\theta_{n-1})} drd\theta_1 \dots d\theta_{n-1} \\ &= r^{n-1} g(\boldsymbol{\theta}) drd\theta_1 \dots d\theta_{n-1} \end{aligned}$$

Let this be called

Above can be seen here:

https://en.wikipedia.org/wiki/N-sphere#Spherical_coordinates

Furthermore: We have $\|\mathbf{x}\|^2 = r^2$ for all $\boldsymbol{\theta}$, because the squared length of a coordinate point is r^2.

The surface area $S_D = S_{n-1}$ of an n dimensional sphere with radius r is denoted by

$$S_{n-1} = \frac{dV_n(R)}{dR} = nC_nR^{n-1}$$

where $V_n(R)$ denotes the volume of a sphere of radius R.
This can be seen here:
http://scipp.ucsc.edu/~haber/ph116A/volume_11.pdf

There we can also find the identity (Eq.7):

$$\begin{aligned} nC_n &= \int \dots \int d\Omega_{n-1} \\ (*) \quad &= \int \dots \int g(\boldsymbol{\theta}) d\theta_1 d\theta_2 \dots d\theta_{n-1} \end{aligned}$$

$$p(r,\theta_1,\theta_2,\dots,\theta_{n-1}) = \int_0^r\int_0^{2\pi}\dots\int_0^{\pi}\underbrace{(2\pi)^{-\frac{D}{2}}e^{-\frac{1}{2}r^2}}_{pdf\ normal\ dist.}\underbrace{r^{n-1}\sin^{n-2}\theta_1\sin^{n-3}\theta_2\cdot\dots\cdot\sin\theta_{n-2}\,d\theta_1\dots d\theta_{n-1}dr}_{infinitismal\ area}$$

$$= \int_0^r(2\pi)^{-\frac{D}{2}}e^{-\frac{1}{2}r^2}r^{n-1}\int_0^{2\pi}\dots\int_0^{\pi}\underbrace{g(\boldsymbol{\theta})d\theta_1\dots d\theta_{n-1}}_{=nC_n\cdot 1, acc.to.(*)}dr$$

$$= \int_0^r(2\pi)^{-\frac{D}{2}}e^{-\frac{1}{2}r^2}r^{n-1}S_D\,dr$$

Ergo $p(r)dr = S_D r^{d-1}(2\pi)^{-\frac{D}{2}}e^{-\frac{1}{2}r^2}$.

Now we are looking for the maximum density $\max_r p(r)$:

$$\frac{d}{dr}\left[\log r^{d-1} + \log e^{-\frac{1}{2}r^2}\right] = \frac{(d-1)r^{d-2}}{r^{d-1}} - r = 0.$$

$\Leftrightarrow (d-1) = r^2$.
Because radii are non-negative, we have a maximum at $\sqrt{d-1}$.

Now if we set $\|\mathbf{x}\| = \sqrt{d-1}$, we get

$$\frac{p(\mathbf{x})}{p(\mathbf{0})} = \frac{(2\pi)^{-\frac{D}{2}}e^{-\frac{1}{2}\|\mathbf{x}\|^2}}{(2\pi)^{-\frac{D}{2}}e^{-\frac{1}{2}\|\mathbf{0}\|^2}} = e^{-\frac{1}{2}(d-1)}.$$

4A-3.

Let $L(\mathbf{w})$:

$$L(\mathbf{w}) = \frac{1}{2}\sum_{n=1}^{N}\left(t_n - \mathbf{w}^T\boldsymbol{\phi}(x_n)\right)^2 + \frac{\lambda}{2}\mathbf{w}^T\mathbf{w}$$

Derivative of $\ln L(\mathbf{w})$ with respect to w_i:

$$\frac{\partial}{\partial w_i}\ln L(\mathbf{w}) = \sum_{n=1}^{N}\phi_i(x_n)\left(t_n - \mathbf{w}^T\boldsymbol{\phi}(x_n)\right) + \lambda w_i$$

Conversion to matrix/vector operations:

$$\sum_{n=1}^{N}\phi_i(x_n)\left(t_n - \mathbf{w}^T\boldsymbol{\phi}(x_n)\right) + \lambda w_i = \mathrm{col_i}(\boldsymbol{\Phi})^T(\mathbf{t} - \boldsymbol{\Phi}\mathbf{w}) + \lambda w_i$$

Generalized for all w:

$$\frac{\partial}{\partial \mathbf{w}}\ln L(\mathbf{w}) = \boldsymbol{\Phi}^T(\mathbf{t} - \boldsymbol{\Phi}\mathbf{w}) + \lambda\mathbf{w}$$

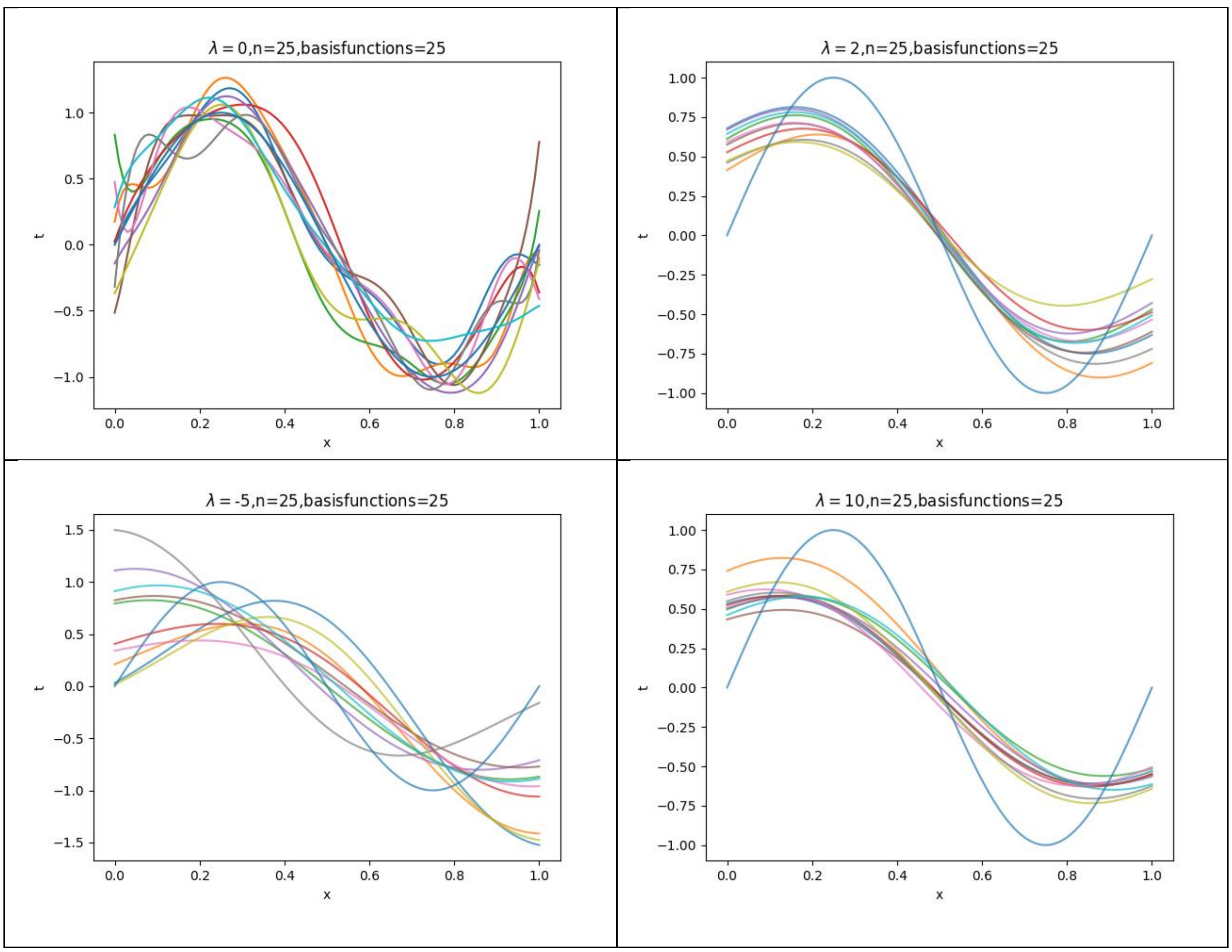
Setting zero and solving for $\mathbf{w}$:

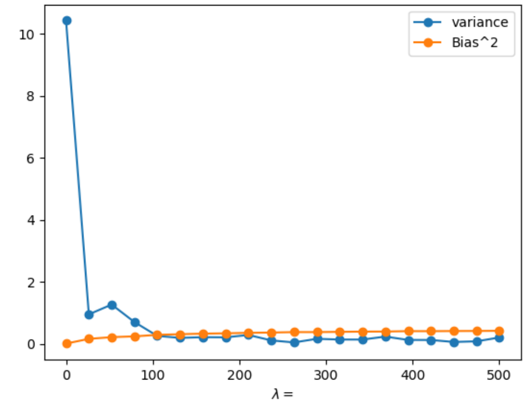
$$\begin{aligned} & \mathbf{\Phi}^T(\mathbf{t} - \mathbf{\Phi}\mathbf{w}) + \lambda\mathbf{w} && = 0 \\ \Leftrightarrow \quad & \mathbf{\Phi}^T\mathbf{t} + (-\mathbf{\Phi}^T\mathbf{\Phi} + \lambda\mathbf{I})\mathbf{w} && = 0 \\ \Leftrightarrow \quad & \mathbf{w} && = (-\lambda\mathbf{I} + \mathbf{\Phi}^T\mathbf{\Phi})^+\mathbf{\Phi}^T\mathbf{t} \end{aligned}$$

As usual, $\mathbf{A}^+$ denotes the Penrose pseudo inverse.
Because λ can be an arbitrary normative factor, it is also possible to write:

$$\mathbf{w} = (\lambda\mathbf{I} + \mathbf{\Phi}^T\mathbf{\Phi})^+\mathbf{\Phi}^T\mathbf{t}$$

Pictures from Python:




variance
Bias^2
0
2
4
6
8
10
0
100
200
300
400
500
$\lambda =$