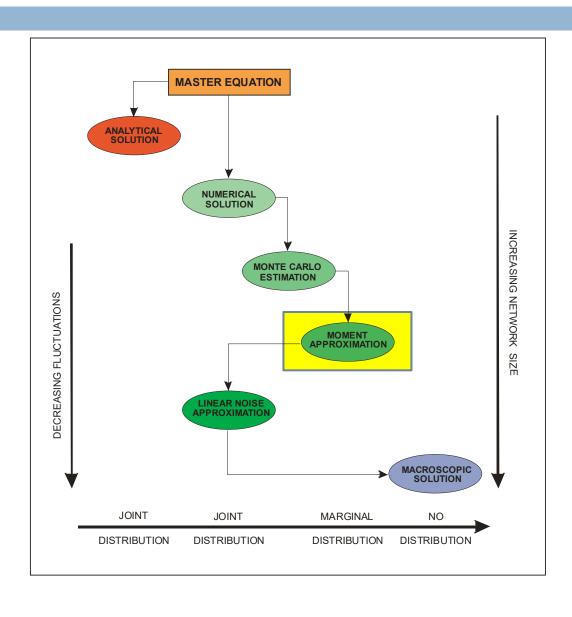
LECTURE #7

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Available Methods



- When a reaction network contains many species and many reactions, it may not be possible to accurately estimate, in <u>reasonable time</u>, the statistical behavior of the DA and population processes by Monte Carlo sampling.
- In this case, we may try an alternative technique known as moment approximation (MA).

□ Let

$$Z_m(t) = \mu_Z(m;t) + W_m(t), t > 0, m = 1, 2, ..., M$$

where

$$\mu_{Z}(m;t) \triangleq \mathrm{E}[Z_{m}(t)]$$

$$W_{m}(t) \triangleq Z_{m}(t) - \mu_{Z}(m;t)$$

This equation is <u>exact !!</u>

In most cases, we are interested in differential equations that govern the dynamic evolution of the means

$$\mu_Z(m;t) = \mathrm{E}[Z_m(t)]$$

and covariances

$$c_Z(m, m'; t) = \text{cov}[Z_m(t), Z_{m'}(t)] = \text{E}[W_m(t)W_{m'}(t)]$$

□ From the ME, we can show that:

$$\frac{d\,\mu_Z(m;t)}{dt} = \mathrm{E} \Big[\alpha_m(\mathbf{Z}(t))\Big], \quad t>0, m=1,2,...,M$$
 see supplement 5 for details
$$\frac{dc_Z(m,m';t)}{dt} = \mathrm{E} \Big[\alpha_m(\mathbf{Z}(t))\Big] \underbrace{\delta(m-m')}_{\text{Kronecker delta}} + \mathrm{E} \Big[[Z_{m'}(t) - \mu_Z(m';t)] \alpha_m(\mathbf{Z}(t)) \Big] + \mathrm{E} \Big[[Z_m(t) - \mu_Z(m;t)] \alpha_{m'}(\mathbf{Z}(t)) \Big],$$

$$t>0, m, m'=1,2,...,M$$

Can we solve this system of equations?

- This is <u>not possible</u> in general (e.g., we cannot evaluate the expectations).
- Since $\mathbf{Z}(t) = \boldsymbol{\mu}_z(t) + \mathbf{W}(t)$, we can use the following Taylor series expansion:

$$\alpha_{m}(\mathbf{Z}(t)) = \alpha_{m}(\mathbf{\mu}_{z}(t)) + \sum_{m_{1}} h_{m;m_{1}}(\mathbf{\mu}_{z}(t))W_{m_{1}}(t)$$

$$+ \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}}(\mathbf{\mu}_{z}(t))W_{m_{1}}(t)W_{m_{2}}(t)$$

$$+ \frac{1}{6} \sum_{m_{1}} \sum_{m_{2}} \sum_{m_{3}} h_{m;m_{1},m_{2},m_{3}}(\mathbf{\mu}_{z}(t))W_{m_{1}}(t)W_{m_{2}}(t)W_{m_{3}}(t)$$

$$+ \cdots$$

In this case, we obtain:

 $+\cdots$

$$\begin{split} \frac{d\,\mu_Z(m;t)}{dt} &= \alpha_m(\pmb{\mu}_Z(t)) + \frac{1}{2} \sum_{m_1} \sum_{m_2} h_{m;m_1,m_2}(\pmb{\mu}_Z(t)) c_Z(m_1,m_2;t) \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} h_{m;m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m_1,m_2,m_3;t) + \cdots \\ \frac{dc_Z(m,m';t)}{dt} &= \alpha_m(\pmb{\mu}_Z(t)) \delta(m-m') \\ &+ \sum_{m_1} \left[h_{m';m_1}(\pmb{\mu}_Z(t)) c_Z(m,m_1;t) + h_{m;m_1}(\pmb{\mu}_Z(t)) c_Z(m',m_1;t) \right] \\ &+ \frac{1}{2} \left[\sum_{m_1} \sum_{m_2} h_{m;m_1,m_2}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2;t) \right] \delta(m-m') \\ &+ \frac{1}{2} \sum_{m_1} \sum_{m_2} \left[h_{m';m_1,m_2}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2;t) + h_{m;m_1,m_2}(\pmb{\mu}_Z(t)) c_Z(m',m_1,m_2;t) \right] \\ &+ \frac{1}{6} \left[\sum_{m_1} \sum_{m_2} \sum_{m_3} h_{m;m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_3} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m_1,m_2,m_3;t) \right] \delta(m-m') \\ &+ \frac{1}{6} \sum_{m_1} \sum_{m_2} \sum_{m_2} \left[h_{m';m_1,m_2,m_3}(\pmb{\mu}_Z(t)) c_Z(m,m$$

- The previous equations show that the mean and covariance dynamics of the DA process are in general governed by a system of <u>coupled</u> first-order differential equations driven by the third-, fourth-, and higher-order central moments.
- The dependency of these equations on propensity function derivatives tells us how the mean and fluctuation dynamics are affected by the presence of nonlinearities in the propensity functions.

If the propensity functions are <u>linear</u>, then all second- and higherorder derivatives of these functions will be zero, which implies that

$$\frac{d\mu_Z(m;t)}{dt} = \alpha_m(\boldsymbol{\mu}_Z(t))$$

$$\frac{dc_Z(m,m';t)}{dt} = \alpha_m(\boldsymbol{\mu}_Z(t))\delta(m-m') \xrightarrow{\text{constants}}$$

$$+\sum_{m_1} \left[h_{m';m_1}c_Z(m,m_1;t) + h_{m;m_1}c_Z(m',m_1;t)\right]$$

- These equations are exact!!
- In this case, the means can be calculated <u>independently</u> from the covariances.

□ If the propensity functions are <u>quadratic</u>, then

$$\begin{split} \frac{d\,\mu_{Z}(m;t)}{dt} &= \alpha_{m}(\boldsymbol{\mu}_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}} c_{Z}(m_{1},m_{2};t) \\ \frac{dc_{Z}(m,m';t)}{dt} &= \left[\alpha_{m}(\boldsymbol{\mu}_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}} c_{Z}(m_{1},m_{2};t) \right] \delta(m-m') \\ &+ \sum_{m_{1}} \left[h_{m';m_{1}}(\boldsymbol{\mu}_{Z}) c_{Z}(m,m_{1};t) + h_{m;m_{1}}(\boldsymbol{\mu}_{Z}) c_{Z}(m',m_{1};t) \right] \\ &+ \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} \left[h_{m';m_{1},m_{2}} \frac{c_{Z}(m,m_{1},m_{2};t)}{c_{Z}(m,m_{1},m_{2};t)} + h_{m;m_{1},m_{2}} \frac{c_{Z}(m',m_{1},m_{2};t)}{c_{Z}(m',m_{1},m_{2};t)} \right] \end{split}$$

- □ These equations are <u>exact</u>!!
- In this case, the means <u>cannot</u> be calculated independently from the covariances.
- Calculation requires third-order central moments!!

□ Note that $\mathbf{X}(t) = \mathbf{x}_0 + \mathbf{S}\mathbf{Z}(t)$ implies

$$\boldsymbol{\mu}_{X}(t) = \mathbf{x}_{0} + \mathbf{S}\boldsymbol{\mu}_{Z}(t)$$
$$\mathbf{C}_{X}(t) = \mathbf{S}\mathbf{C}_{Z}(t)\mathbf{S}^{T}$$

- Therefore, the means and covariances of the population process can be easily calculated from the means and covariances of the DA process.
- We can also derive differential equations for the <u>central</u> moments of the population process.

- When the propensity functions are <u>nonlinear</u>, the moment equations always form an <u>infinite hierarchy</u>, with lower order moments depending on higher order moments, indicating that <u>exact</u> solutions <u>cannot</u> be obtained in practice.
- To address this problem, we can replace the moments at some stage of the hierarchy with appropriately chosen functions of lower-order moments.
- This approach results in a <u>moment closure</u> scheme that produces a <u>self-contained</u> system of differential equations whose solution provides approximate values for the moments of the DA and populations processes.

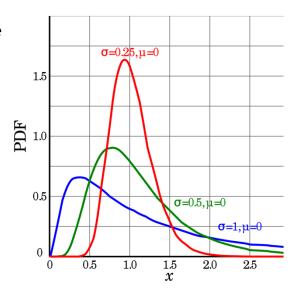
- A possible solution to the moment closure problem is to assume the form of the probability distribution of the DA or population processes.
- Example: We may assume a probability distribution for the DA process that can be uniquely defined from the means and covariances (the Gaussian is one such distribution).
 - This distribution will impose functional relationships between the higher-order central moments and the means and covariances.
 - This relationship can be used to close the system of moment equations.

- **Example:** We may assume that $p_Z(\mathbf{z};t)$ is approximately Gaussian.
- In this case, the third-order central moments are zero and we have (for at most quadratic propensities):

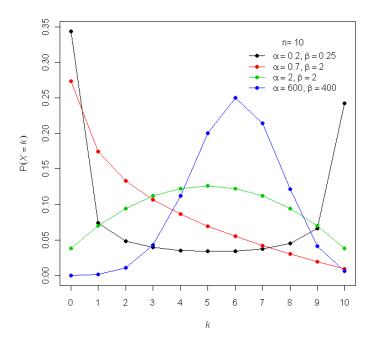
$$\frac{d\mu_{Z}(m;t)}{dt} = \alpha_{m}(\boldsymbol{\mu}_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}} c_{Z}(m_{1},m_{2};t)
\frac{dc_{Z}(m,m';t)}{dt} = \left[\alpha_{m}(\boldsymbol{\mu}_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}} c_{Z}(m_{1},m_{2};t)\right] \delta(m-m')
+ \sum_{m_{1}} \left[h_{m';m_{1}}(\boldsymbol{\mu}_{Z}) c_{Z}(m,m_{1};t) + h_{m;m_{1}}(\boldsymbol{\mu}_{Z}) c_{Z}(m',m_{1};t)\right]$$

MA techniques based on a Gaussian assumption cannot characterize bimodality or highly skewed probability distributions and may result in undesirable negative DA and population values.

- In addition to the normal distribution, several alternative approximating distributions have been suggested in the literature: log-normal, Poisson, binomial, beta binomial, and mixtures of distributions.
- Using the log-normal distribution has some advantages over using the Gaussian distribution.
 - The log-normal distribution has nonnegative support and exhibits nonzero skewness, two properties that are important in the context of certain nonlinear reaction networks, such as biochemical reaction networks.



On the other hand, the <u>beta binomial distribution</u> is a discrete distribution with a flexible shape that, in some cases, can capture the dynamic evolution of the true probability distribution of the population process better than other distributions.



https://en.wikipedia.org/wiki/Beta-binomial distribution

- It may be difficult to specify an appropriate probability distribution for the population process, since this distribution must assign <u>zero</u> probabilities to <u>stoichiometrically impossible</u> populations.
- It may be easier to specify a probability distribution for the DA process, since this process is usually confined within a welldefined subset of the positive orthant of the multidimensional DA state-space.
- Approximating the moments of the DA process by employing continuous distributions, such as log-normal, may be difficult to justify due to the discrete nature of this process.

Moment Approximation with MaxEnt

- Although the previous strategy may lead to a sufficiently accurate estimation of the low-order moments of the DA and population processes, it cannot provide an <u>analytical</u> expression for the probability distributions of these processes.
- However, we may be able to address this problem using MaxEnt.

Moment Approximation with MaxEnt

■ **Example:** If the propensity functions of the reaction network at hand are <u>at most quadratic</u> and if we use a <u>log-normal-based</u> moment closure scheme, then the MaxEnt approximation of the probability distribution of the DA process associated with the *m*-th reaction will be given by the following Gibbs distribution:

$$\hat{p}_{Z}(z_{m};t) = \left(\sum_{u \ge 0} \exp\left\{-\lambda_{1}(t)u - \lambda_{2}(t)u^{2} - \lambda_{3}(t)u^{3}\right\}\right)^{-1} \times \exp\left\{-\lambda_{1}(t)z_{m} - \lambda_{2}(t)z_{m}^{2} - \lambda_{3}(t)z_{m}^{3}\right\}$$

where the coefficients $\lambda_1(t), \lambda_2(t), \lambda_3(t)$ must be determined so that

$$\sum_{z_{m}\geq 0} z_{m} \hat{p}_{Z}(z_{m};t) = \mu_{Z}(m;t)$$

$$\sum_{z_{m}\geq 0} z_{m}^{2} \hat{p}_{Z}(z_{m};t) = c_{Z}(m,m;t) + \mu_{Z}^{2}(m;t)$$

$$\sum_{z_{m}\geq 0} z_{m}^{3} \hat{p}(z_{m};t) = \left[\frac{c_{Z}(m,m;t) + \mu_{Z}^{2}(m;t)}{\mu_{Z}(m;t)}\right]^{3}$$

Moment Approximation with MaxEnt

 In this case the mean and covariance dynamics can be calculated from

$$\begin{split} \frac{d\mu_{Z}(m;t)}{dt} &= \alpha_{m}(\boldsymbol{\mu}_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}} c_{Z}(m_{1},m_{2};t) \\ \frac{dc_{Z}(m,m';t)}{dt} &= \left[\alpha_{m}(\boldsymbol{\mu}_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}} c_{Z}(m_{1},m_{2};t) \right] \Delta(m-m') \\ &+ \sum_{m_{1}} \left[h_{m';m_{1}}(\boldsymbol{\mu}_{Z}) c_{Z}(m,m_{1};t) + h_{m;m_{1}}(\boldsymbol{\mu}_{Z}) c_{Z}(m',m_{1};t) \right] \\ &+ \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} \left[h_{m';m_{1},m_{2}} c_{Z}(m,m_{1},m_{2};t) + h_{m;m_{1},m_{2}} c_{Z}(m',m_{1},m_{2};t) \right] \end{split}$$

where the third-order moments can be related to the first- and second-order moments by

$$E[Z_1(t)Z_2(t)Z_3(t)] = \frac{E[Z_1(t)Z_2(t)]E[Z_1(t)Z_3(t)]E[Z_2(t)Z_3(t)]}{E[Z_1(t)]E[Z_2(t)]E[Z_3(t)]}$$

property of log-normal distribution

Moment Approximation by Truncation

- Several moment closure schemes have been proposed in the literature based on <u>truncating</u> high-order central moments or <u>cumulants</u>.
- The typical assumption behind these methods is that the solution of the ME has <u>negligible</u> high-order central moments or cumulants, which can be set equal to zero without affecting the mean and covariance dynamics.
- This assumption is not true in general, except for normal random variables whose central moments of odd order and cumulants of order larger than 2 are zero.

https://en.wikipedia.org/wiki/Cumulant

Moment Approximation by Truncation

- For non-normal random variables, there is an <u>infinite</u> number of non-vanishing central moments or cumulants in general.
- Example: all cumulants of a <u>Poisson</u> random variable are equal to the mean value.
- Truncation of high-order central moments or cumulants <u>cannot</u> be easily justified and may lead to a distribution that is a <u>non-valid</u> probability distribution.

Moment Approximation – Stabilization

Recall that

$$\frac{d\mu_{Z}(m;t)}{dt} = E[\alpha_{m}(\mathbf{Z}(t))] \qquad T_{m}(\mu_{Z}(t))$$

$$= \alpha_{m}(\mu_{Z}(t)) + \frac{1}{2} \sum_{m_{1}} \sum_{m_{2}} h_{m;m_{1},m_{2}}(\mu_{Z}(t)) c_{Z}(m_{1},m_{2};t)$$

$$+ \frac{1}{6} \sum_{m_{1}} \sum_{m_{2}} \sum_{m_{3}} h_{m;m_{1},m_{2},m_{3}}(\mu_{Z}(t)) c_{Z}(m_{1},m_{2},m_{3};t) + \cdots$$
The MA method is based on replacing the mean value $E[\alpha_{L}(\mathbf{Z}(t))]$

The MA method is based on replacing the mean value $E[\alpha_m(\mathbf{Z}(t))]$ by $\alpha_m(\boldsymbol{\mu}_Z(t)) + T_m(\boldsymbol{\mu}_Z(t))$, where $T_m(\boldsymbol{\mu}_Z(t))$ is a term calculated by solving differential equations for higher-order central moments.

Moment Approximation – Stabilization

When the propensity function is <u>convex</u>, then <u>Jensen's inequality</u> implies that

$$E[\alpha_m(\mathbf{Z}(t))] \ge \alpha_m(E[\mathbf{Z}(t)]) = \alpha_m(\mathbf{\mu}_Z(t))$$

which results in

$$E[\alpha_m(\mathbf{Z}(t))] = \alpha_m(\boldsymbol{\mu}_Z(t)) + T_m(\boldsymbol{\mu}_Z(t)) \ge \alpha_m(\boldsymbol{\mu}_Z(t))$$

Consequently, we must have

$$T_m(\boldsymbol{\mu}_Z(t)) \geq 0$$

https://en.wikipedia.org/wiki/Jensen inequality

Moment Approximation – Stabilization

- However, this may not be true, in which case the method may result in additional errors that can lead to instabilities.
- \square To address this problem, we can replace $\mathbb{E}[\alpha_m(\mathbf{Z}(t))]$ by

$$\alpha_m(\boldsymbol{\mu}_Z(t)) + \max\{0, T_m(\boldsymbol{\mu}_Z(t))\}$$

 In many instances, this simple modification results in a more accurate and more stable implementation of the MA method.

Moment Approximation – Consistency

- A moment closure scheme used in a particular application must produce moments that satisfy several necessary conditions.
- □ For example, all moments must be <u>nonnegative</u> and <u>invariant</u> under permutations and the resulting covariances must define a <u>symmetric positive semi-definite</u> matrix.
- If a necessary condition is violated by a particular moment
 closure scheme, then this is indicative of its inappropriateness.

Moment Approximation – Consistency

- The assumptions underlying a given moment closure scheme may be inconsistent with the statistics of a particular reaction network at hand.
- In this case, the moment equations may produce unrealistic solutions or even result in unstable and unbounded solutions.