Network analysis and simulation
Homework 1

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Exercise 1

In the first exercise two datasets are given and on those, a bunch of figures have been plotted, in which the data are showed and different measures of confidence are calculated on them.

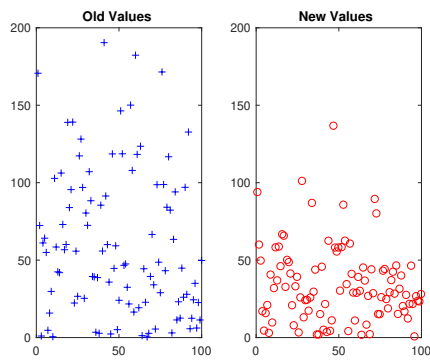In the follow the said figures are reported.
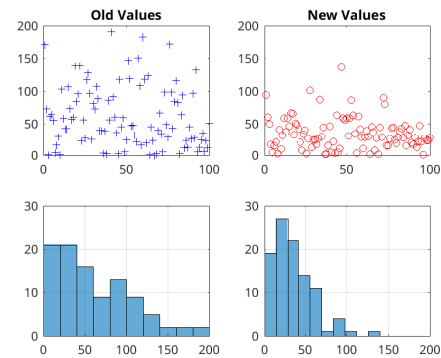

Figure 1: Plot of the data

Figure 2: Data plotted also in histograms divided in 10 bins (Figure 2.1)
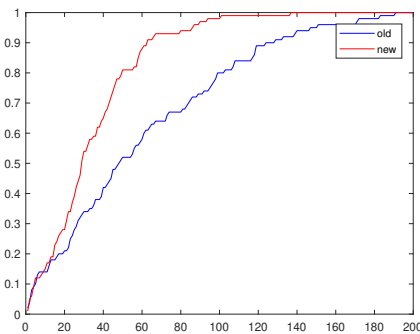

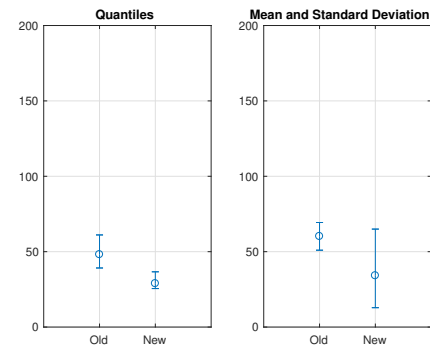Figure 3: Empirical distribution function of the data (Figure 2.2)

Figure 4: Box Plots of the data with Confidence Interval (CI) for median and mean (Figure 2.3)

Exercise 2

Executing the script correspondent to the second exercise, we found that in 56 experiments the CI does not contain the true value of the mean.

In Figure 7 and Figure 8 are reported the value of the sample mean and its CI (with $\gamma = 0.95$) for each experiment, sorted based on the lower extreme of the CI and using
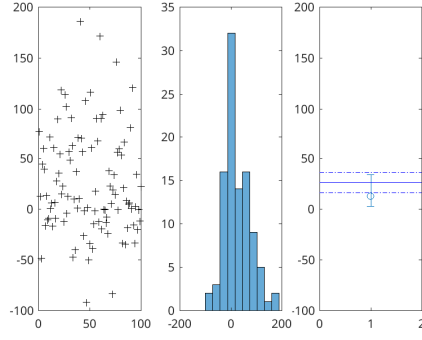
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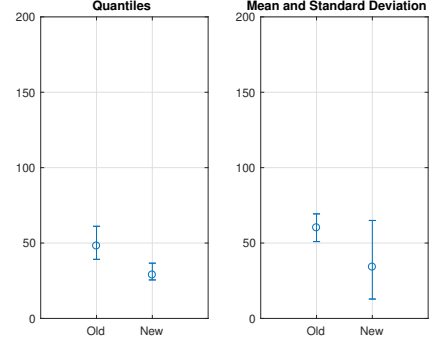Figure 5: Difference between old and new data (Figure 2.7)

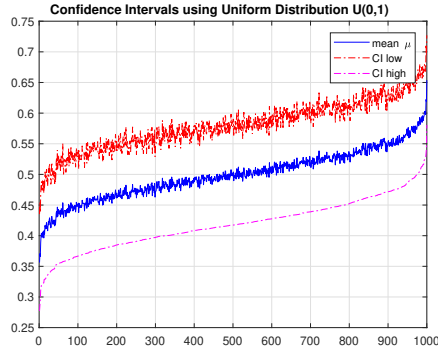Figure 6: Box Plots of the data with CI for median and mean (Figure 2.3)

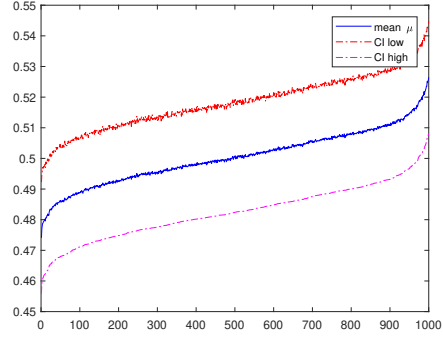Figure 7: Results of the experiment with $n = 48$

Figure 8: Results of the experiment with $n = 1000$

a different number of random variables in each experiment. In both cases, the sample mean is distributed around the true mean (that for this Uniform Distribution is 0.5).

Note how, increasing the number of random variables per experiment, the mean width of the CI get lower. Furthermore, the width of the CI is not constant.

Exercise 3

To calculate $\mathbf{E}(U_{(j)})$ we have firstly to compute the probability that at least an order statistic $U_{(j)}$ falls in $[u, u + du]$, that is $P_{U_{(j)}}(u)$. Actually the probability that more than one element falls in this interval is negligigble since is an $O(du^2)$, so we can take in account only the probability that exactly one element falls in $[u, u + du]$.

This probability is given by the following expression:

$$P_{U_{(j)}}(u) = P\left[j - 1 \text{ order statistics are in } [0, u]\right] \cdot P\left[1 \text{ order statistic is in } [u, u + du]\right]$$
$$\cdot P\left[n - j \text{ order statistics are in } [u + du, 1]\right] \tag{1}$$

The probability that a realization of a random variable $\mathbf{U}(0, 1)$ is in $[0, u]$ is u, we want $j - 1$ realization over n to be in this interval, so the resulting probability is:

$$P\left[j - 1 \text{ order statistics are in } [0, u]\right] = \binom{n}{j - 1} u^{j - 1}$$

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Now we want the next order statistic to be in $[u, u + du]$, so we can choose from the remaining $n - j + 1$ realization an element with probability du, that is:

$$P\left[1 \text{ order statistic is in } [u, u + du]\right] = (n - j + 1)du$$

Finally, the last $n - j$ elements have to be in $[u + du, 1]$, so:

$$P\left[n - j \text{ order statistics are in } [u + du, 1]\right] = (1 - u - du)^{n-j}$$

but since du is very small by definition, we can ignore it rewriting this last equation as $(1 - u)^{n-j}$.

Multiplying these three terms, as specified in Equation 1, the final probability is:

$$P_{U_{(j)}}(u) = \frac{n!}{(j - 1)!(n - j)!}u^{j-1} \cdot du \cdot (1 - u)^{n-j} \tag{2}$$

The keypoint of the proof is to note that Equation 2 is the PDF of a Beta distribution. Indeed a general Beta distribution is defined as

$$f(x) = \frac{1}{\mathbf{B}(\alpha, \beta)}x^{\alpha-1}(1 - x)^{\beta-1} \tag{3}$$

where $\mathbf{B}(\alpha, \beta)$ is called *Beta function* and is defined as

$$\mathbf{B}(\alpha, \beta) = \frac{(\alpha - 1)!(\beta - 1)!}{(\alpha + \beta - 1)!} \tag{4}$$

Given this definition, we can derive that $P_{U_{(j)}}(u)$ is distributed as a $Beta(j, n + 1 - j)$ and since the expected value of a *Beta* distribution is defined as $\mathbf{E}[x] = \frac{\alpha}{\alpha+\beta}$, we conclude that

$$\mathbf{E}\left[U_{(j)}\right] = \frac{j}{n + 1} \tag{5}$$

Exercise 4

In Figure 9 is plotted the accuracy of the sample mean versus the number of random variables in each experiment. This measure have been made based on the following formula:

$$A_i = |\overline{x_i} - \overline{x}| \tag{6}$$

where A_i is the accuracy of the experiment i and x_i is the sample mean of the same experiment. As can be seen in the said figure, the accuracy get lower as the number of random variables in each experiment increase, this happens since the higher is the number of random variables, the higher is the precision of the experiment.

In Figure 10 is reported the variance of each experiment and the relative confidence interval computed using bootstrap method. Finally, in **??** and **??** are reported the prediction intervals at level $\gamma = 0.95$ using theory and using bootstrap.
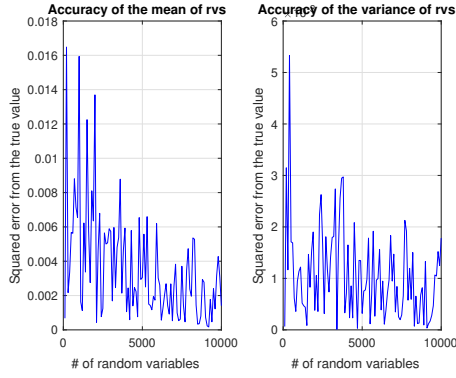
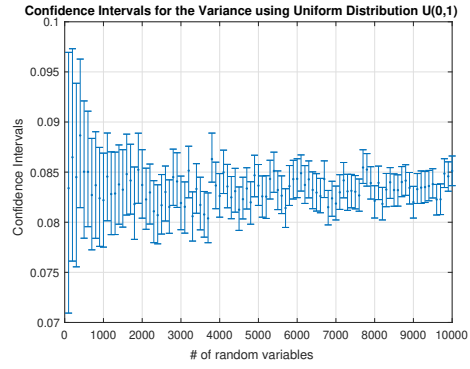Figure 9: Accuracy of the estimation versus n

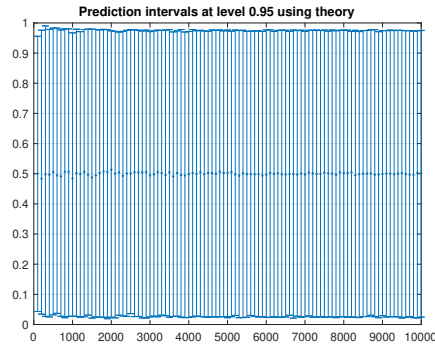Figure 10: Confidence intervals for the variance using rvs $U(0, 1)$

Figure 11: Prediction interval for a $U(0, 1)$ at level $\gamma = 0.95$ using theory
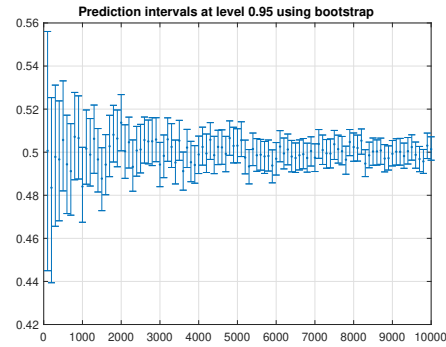

Figure 12: Prediction interval for a $U(0, 1)$ at level $\gamma = 0.95$ using bootstrap

Exercise 5

Redoing Exercise 2 the plot in Figure 11 and Figure 12 have obtained **(DP says: il verbo non so se sia giusto)**. Note that the sample mean now is distributed around the new true value, that is 0.

For the Exercise 4 the results are reported in Figure 13 and Figure 14. As in exercise 4, is visible how, the accuracy get higher (that is, the distance between the sample mean and the true value get lower) and the precision of the variance get higher (the CI becomes smaller) as n increases. Now the sample variance oscillates around 1 that is in fact the new true value.

For what regards the Prediction Interval (PI), we can see that using the relative theory, the PI is (as expected) between [-2,2], that is the γ quantile of the Normal Distribution $N(0, 1)$. In the bootstrap case, the distribution is assumed unknown, and increasing the number of realizations for rv, the size of the prediction interval get lower.
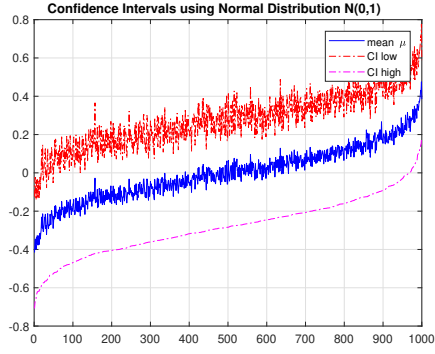
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Figure 13: Results of the experiment with $n = 48$ rvs $N(0,1)$
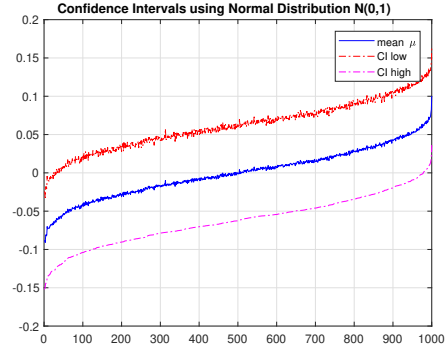

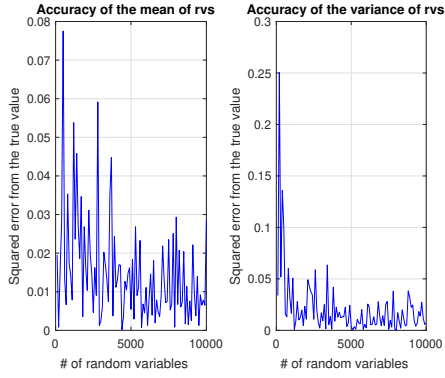Figure 14: Results of the experiment with $n = 1000$ rvs $N(0,1)$

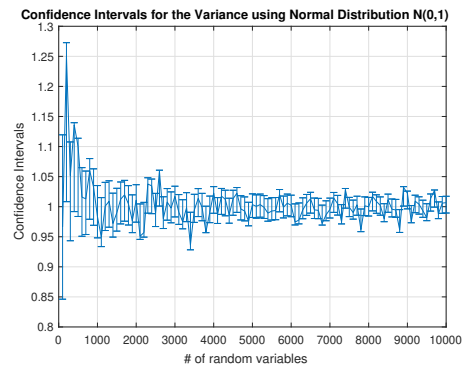Figure 15: Accuracy of the estimation versus n using rvs $N(0,1)$

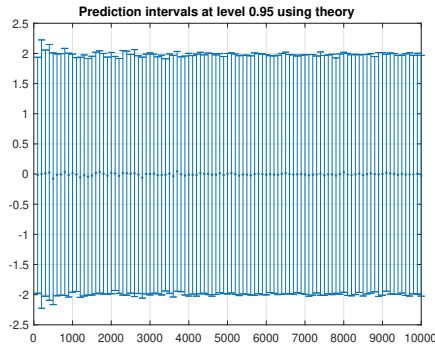Figure 16: Confidence intervals for the variance using rvs $N(0,1)$

Figure 17: Prediction interval for a $N(0,1)$ at level $\gamma = 0.95$ using theory
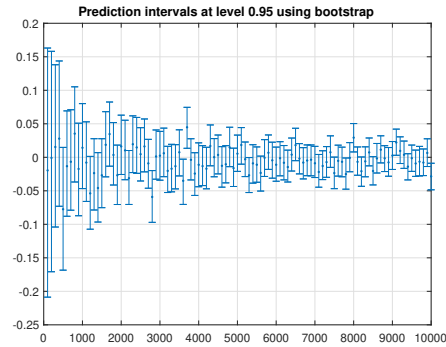

Figure 18: Prediction interval for a $N(0,1)$ at level $\gamma = 0.95$ using bootstrap