

## Homework 6

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**Problem 1** Let  $X_1, \dots, X_m$  i.i.d  $\sim F$ ,  $Y_1, \dots, Y_n$  i.i.d  $\sim G$  and i.i.d. within each sample, then

- (1) get the  $U$ -statistic  $U_n$  with kernel  $h(x_1, x_2, y_1, y_2) = I(x_1 < y_1, x_2 < y_2)$ ,
- (2) get the limit distribution of  $U$ -statistic  $U_n$  with  $m + n \rightarrow \infty, \frac{m}{n+m} \rightarrow p \in (0, 1)$ ,
- (3) get the limit distribution of  $U$ -statistic  $U_n$  under null hypothesis  $H_0 : F = G$ .

**Solution**

(1) The kernel is  $h(x_1, x_2, y_1, y_2) = I(x_1 < y_1, x_2 < y_2)$ , which is of order 2 in both  $x$  and  $y$ . The corresponding  $U$ -statistic is

$$U_n = \frac{1}{\binom{m}{2} \binom{n}{2}} \sum_{i < k}^m \sum_{j < l}^n I(X_i < Y_j, X_k < Y_l)$$

(2) The projection of  $U - \theta$  onto the set of all functions of the form  $\sum_{i=1}^m k_i(X_i) + \sum_{j=1}^n l_j(Y_j)$  is given by

$$\hat{U} = \frac{2}{m} \sum_{i=1}^m h_{1,0}(X_i) + \frac{2}{n} \sum_{j=1}^n h_{0,1}(Y_j),$$

where the functions  $h_{1,0}$  and  $h_{0,1}$  are defined by

$$\begin{aligned} h_{1,0}(x) &= Eh(x, X_2, Y_1, Y_2) - \theta, \\ h_{0,1}(y) &= Eh(X_1, X_2, y, Y_2) - \theta. \end{aligned}$$

The sequence  $\hat{U}$  is asymptotically normal by the central limit theorem. Then difference between  $\hat{U}$  and  $U_n - \theta$  is asymptotically negligible.

$$\begin{aligned} \theta &= Eh(X_1, X_2, Y_1, Y_2) \\ &= P(X_1 < Y_1)P(X_2 < Y_2) \\ &= (E[P(X_1 < y | Y_1 = y)])^2 \\ &= (E[F(Y_1)])^2 \\ &= \left( \int F(y) dG(y) \right)^2 \end{aligned}$$

Then calculate  $\zeta_{1,0}$  and  $\zeta_{0,1}$ .

$$\begin{aligned}
\zeta_{1,0} &= \text{cov}(h(X_1, X_2, Y_1, Y_2), h(X_1, X_3, Y_3, Y_4)) \\
&= \text{cov}(I(X_1 < Y_1, X_2 < Y_2), I(X_1 < Y_3, X_3 < Y_4)) \\
&= E[I(X_1 < Y_1, X_2 < Y_2)I(X_1 < Y_3, X_3 < Y_4)] \\
&\quad - E[I(X_1 < Y_1, X_2 < Y_2)]E[I(X_1 < Y_3, X_3 < Y_4)] \\
&= E[I(X_1 < \min(Y_1, Y_3))I(X_2 < Y_2)I(X_3 < Y_4)] \\
&\quad - E[I(X_1 < Y_1)I(X_2 < Y_2)]E[I(X_1 < Y_3)I(X_3 < Y_4)] \\
&= P(X_1 < \min(Y_1, Y_3))(P(X_1 < Y_1))^2 - (P(X_1 < Y_1))^4 \\
&= \left( \int F(y)dG(y) \right)^2 \int F(z)d(-G^2(z) + 2G(z)) - \left( \int F(y)dG(y) \right)^4
\end{aligned}$$

and

$$\begin{aligned}
\zeta_{0,1} &= \text{cov}(h(X_1, X_2, Y_1, Y_2), h(X_3, X_4, Y_1, Y_3)) \\
&= \text{cov}(I(X_1 < Y_1, X_2 < Y_2), I(X_3 < Y_1, X_4 < Y_3)) \\
&= E[I(X_1 < Y_1, X_2 < Y_2)I(X_3 < Y_1, X_4 < Y_3)] \\
&\quad - E[I(X_1 < Y_1, X_2 < Y_2)]E[I(X_3 < Y_1, X_4 < Y_3)] \\
&= E[I(Y_1 > \max(X_1, X_3))I(Y_2 > X_2)I(Y_3 > X_4)] \\
&\quad - E[I(X_1 < Y_1)I(X_2 < Y_2)]E[I(X_3 < Y_1)I(X_4 < Y_3)] \\
&= P(Y_1 > \max(X_1, X_3))(P(X_1 < Y_1))^2 - (P(X_1 < Y_1))^4 \\
&= \left( \int F(y)dG(y) \right)^2 \int F^2(z)dG(z) - \left( \int F(y)dG(y) \right)^4
\end{aligned}$$

Thus,

$$\sqrt{m+n}(U_n - \theta) \rightsquigarrow N\left(0, \frac{4\zeta_{1,0}}{p} + \frac{4\zeta_{0,1}}{1-p}\right)$$

where  $\theta = \left( \int F(y)dG(y) \right)^2$ ,  $\zeta_{1,0} = \left( \int F(y)dG(y) \right)^2 \int F(z)d(-G^2(z) + 2G(z)) - \left( \int F(y)dG(y) \right)^4$  and  $\zeta_{0,1} = \left( \int F(y)dG(y) \right)^2 \int F^2(z)dG(z) - \left( \int F(y)dG(y) \right)^4$ .

(3) Under  $H_0 : F = G$ ,  $\theta = 1/4$ ,  $\zeta_{1,0} = \frac{1}{4*12}$  and  $\zeta_{0,1} = \frac{1}{4*12}$ .

$$\sqrt{m+n}\left(U_n - \frac{1}{2}\right) \rightsquigarrow N\left(0, \frac{1}{12p} + \frac{1}{12(1-p)}\right)$$

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**Problem 2** Suppose the distribution of  $X$  is symmetric about zero with variance  $\sigma^2 > 0$  and  $EX^4 < \infty$ , consider kernel  $h(x, y) = xy + (x^2 - \sigma^2)(y^2 - \sigma^2)$ , then

- (1) prove that the  $U$ -statistic  $U_n$  with kernel  $h$  is degenerated of order 1,
- (2) get  $\lambda_1, \lambda_2$  and orthogonal functions  $\Phi_1(x), \Phi_2(x)$ , such that  $h(x, y) = \lambda_1\varphi_1(x)\varphi_1(y) + \lambda_2\varphi_2(x)\varphi_2(y)$ ,
- (3) get the limit distribution of  $nU_n$ .

**Solution**



**Problem 3** *Prove the Hoeffding decomposition in page 13.*

**Solution**



**Problem 4** *Prove the  $T$  decomposition in page 12.*

**Solution**

