

Chapter 12: Nonparametric Inference

DSCC 462

Computational Introduction to Statistics

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Fall 2022

Plan for Today

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- Introduce nonparametric analogues to hypothesis tests

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- *Wilcoxon Signed-Rank Test*

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 - Nonparametric analog to the one-sample or paired t-test

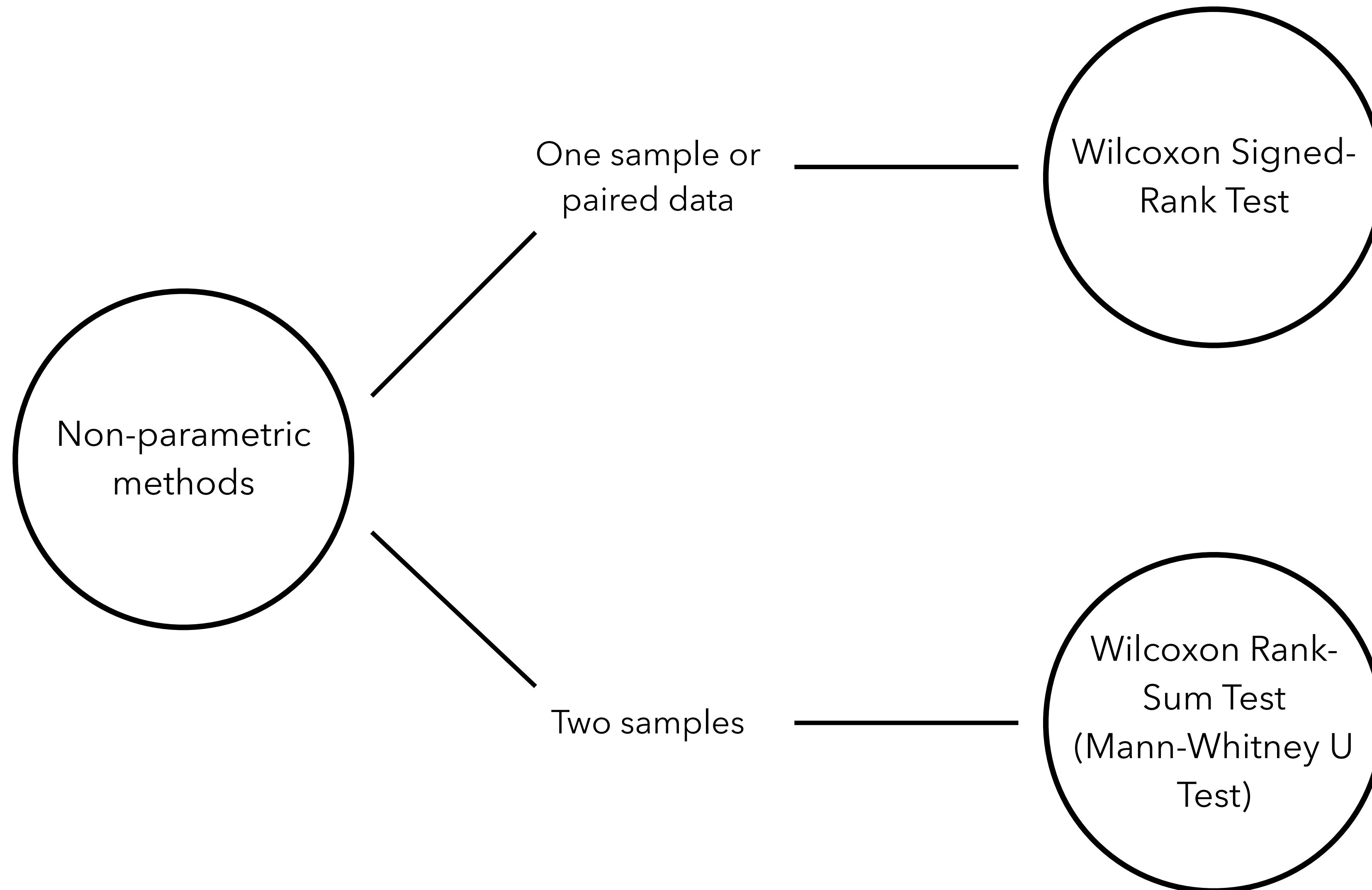
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- Introduce nonparametric analogues to hypothesis tests
- *Wilcoxon Signed-Rank Test*
 - Nonparametric analog to the one-sample or paired t-test
- *Wilcoxon Rank-Sum Test (Mann-Whitney U Test)*

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Plan for Today, Visualized



Nonparametric Methods

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- In this case, we use *nonparametric methods*, which make fewer assumptions regarding the underlying distribution
 - Also known as *distribution-free methods*

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 - Make a claim, develop hypotheses, state significance level
 - Calculate a test statistic based on a random sample of data
 - Determine whether to reject or fail to reject the null hypothesis based on the test statistic and significance level

Motivating Example #1

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 - Note that we consider medians for nonparametric tests as opposed to means

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- H_0 : The median difference in tumor size equals 0
- H_1 : The median difference in tumor size is different from 0
- Test at the $\alpha = 0.05$ significance level

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d_i	\rightarrow	<u>0.5</u>	<u>1</u>	-1	<u>1</u>	<u>1</u>	2
Ranks	\rightarrow	1	3	-3	3	3	5

Wilcoxon Signed-Rank Test: Steps

- Next, take the difference for each pair of observations
- Ignoring the sign of these observations, rank their absolute values from smallest to largest
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 - Remove pair from data set and reduce number of pairs by 1
- Tied observations are assigned an average rank
- Finally separate the ranks by sign to either $+$ or $-$

Wilcoxon Signed-Rank Test: Data Table

↓

Subject	Tumor Size (mm)		Difference	Rank	Signed Rank	
	Before	After			+	-
1	36.3	27.1	9.2			
2	21.7	17.4	4.3	4.5		
3	45.1	33.1	12.0			
4	27.8	32.1	-4.3	4.5		
5	5.1	8.3	→ -2.2	2		
6	23.4	22.1	→ 1.3	1		
7	25.0	31.2	-6.2			
8	12.6	16.4	→ -3.8	3		
9	19.9	12.5	7.4			
10	22.1	22.1	→ 0	—	—	—
11	18.6	4.8	13.8			
12	8.9	22.6	-13.7			
13	12.7	6.4	6.3			
14	29.3	18.3	9.0			
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6	23.4	22.1	1.3	1		
7	25.0	31.2	-6.2	7		
8	12.6	16.4	-3.8	3		
9	19.9	12.5	7.4	9		
10	22.1	22.1	0	-		
11	18.6	4.8	13.8	14		
12	8.9	22.6	-13.7	13		
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6	23.4	22.1	1.3	1	1	
7	25.0	31.2	-6.2	7		7
8	12.6	16.4	-3.8	3		3
9	19.9	12.5	7.4	9	9	
10	22.1	22.1	0	-		
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$\Sigma \rightarrow T^+$

$\bar{\Sigma} \rightarrow T^-$

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- Under the null hypothesis, the median of the underlying population differences is equal to 0

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$$T^+ = \sum \text{ranks all } +$$

- Calculate the sum of the positive ranks, T^+ , and the sum of the negative ranks, T^-
- Calculate $T = T^+ - T^-$
- Under the null hypothesis, the median of the underlying population differences is equal to 0
- Thus, we expect approximately equal numbers of positive and negative ranks

$$- \frac{n(n+1)}{2} \quad / \quad \frac{n(n+1)}{2}$$

Wilcoxon Signed-Rank Test

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$\begin{array}{ccccccc} -3 & -2 & -1 & 0 & 1 & 2 & 3 \\ \hline \end{array}$
- Under the null hypothesis, the median of the underlying population differences is equal to 0
- Thus, we expect approximately equal numbers of positive and negative ranks
- Additionally, the sum of the positive ranks should be approximately equal to the sum of the negative ranks, so T should be approximately 0

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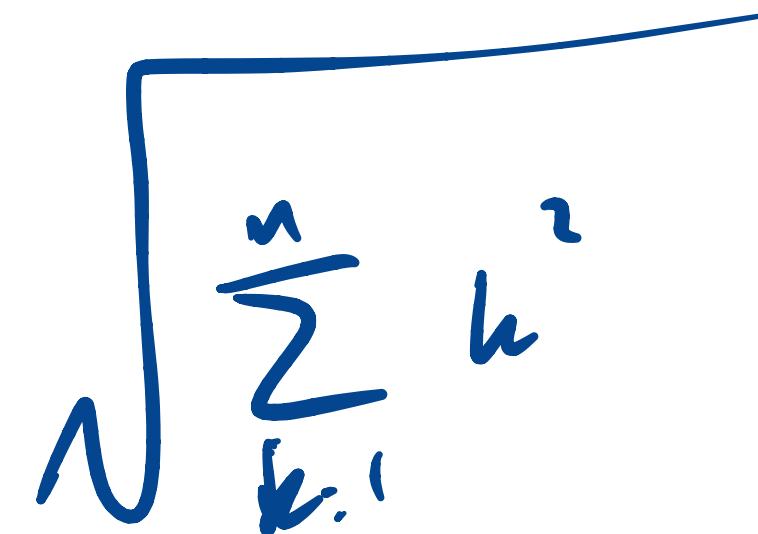
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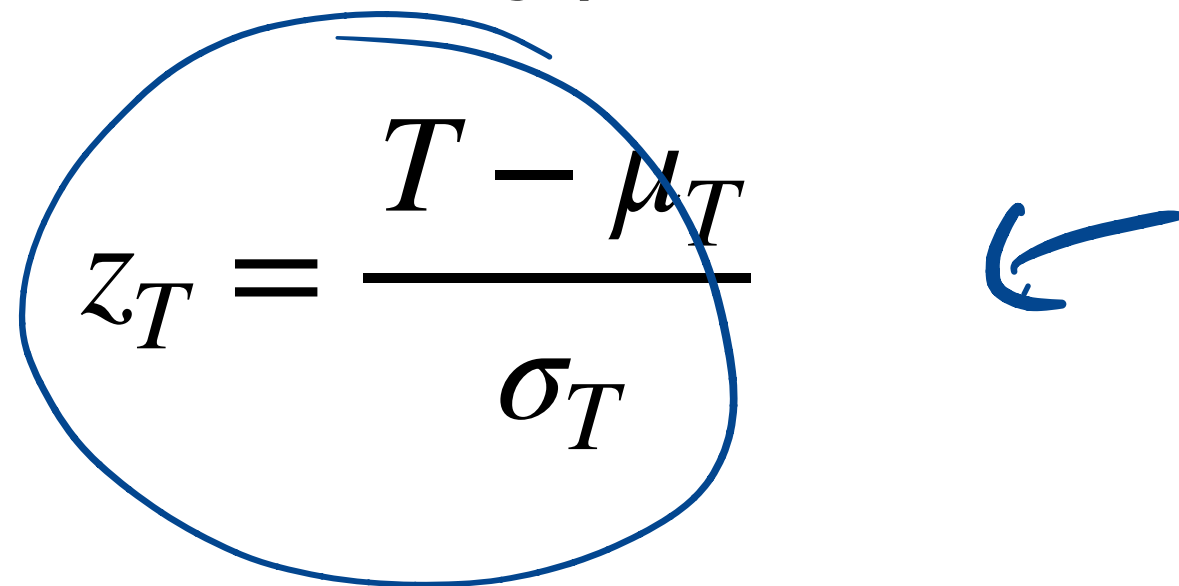
$$\sigma_T = \sqrt{\frac{n(n+1)(2n+1)}{6}}$$



A handwritten blue formula for the standard deviation σ_T , which is $\sqrt{\sum_{k=1}^n k^2}$. A blue arrow points from this handwritten formula towards the printed formula for σ_T in the previous block.

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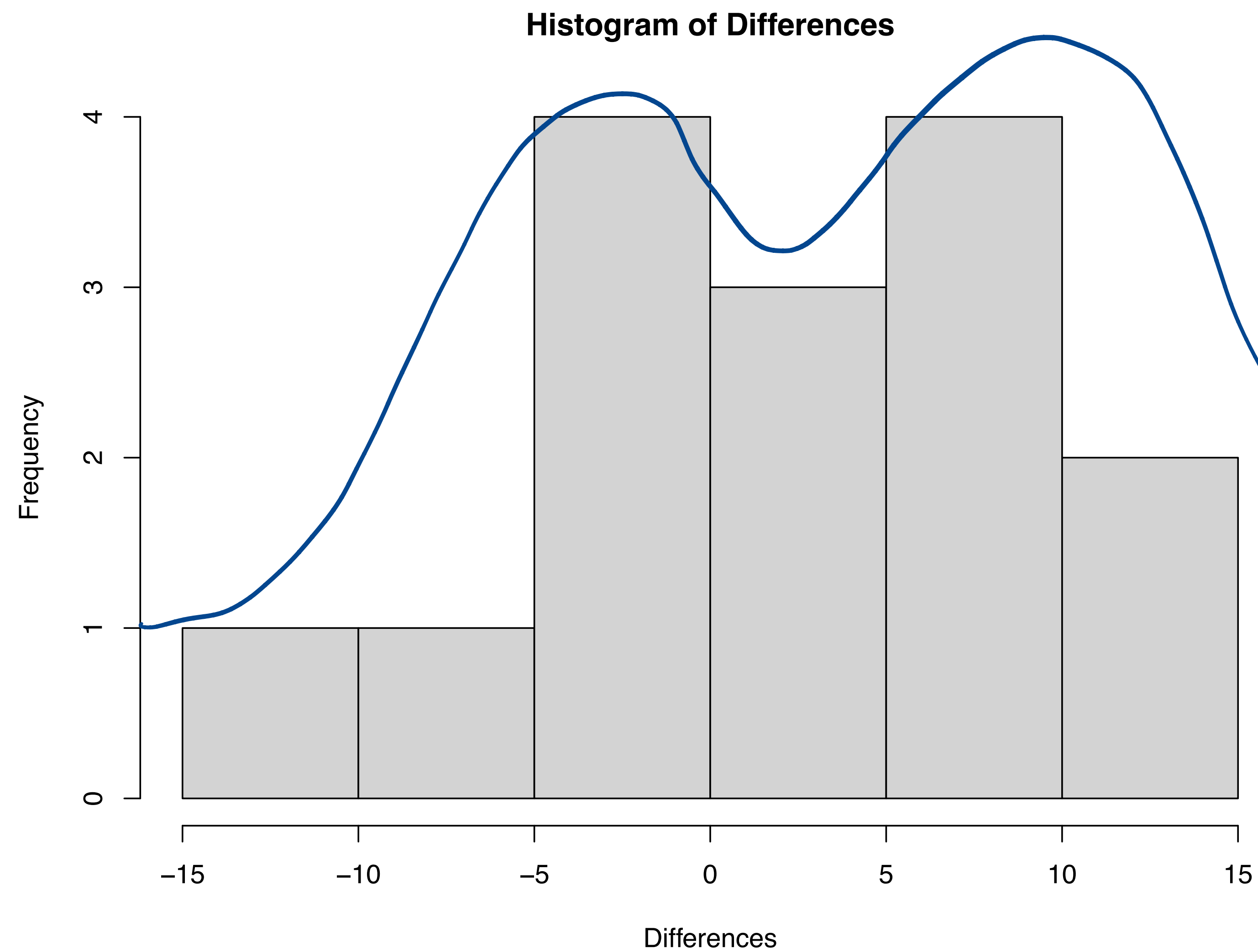
- Note that

$$\mu_T = 0$$

$$\sigma_T = \sqrt{\frac{n(n+1)(2n+1)}{6}}$$

- $Z_T \sim N(0,1)$ given that n is large enough (typically $n > 12$)

Histogram of Differences



Wilcoxon Signed-Rank Test

Subject	Signed Rank	
	+	-
1	11	
2	4.5	
3	12	
4		4.5
5		2
6	1	
7		7
8		3
9	9	
10		
11	14	
12		13
13	8	
14	10	
15	6	

Wilcoxon Signed-Rank Test

- Based on the histogram, the difference in tumor size does not appear to be normally distributed, so we want to use the Wilcoxon signed-rank test

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- $T^+ =$

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Wilcoxon Signed-Rank Test

- Based on the histogram, the difference in tumor size does not appear to be normally distributed, so we want to use the Wilcoxon signed-rank test
- $T^+ = 75.5$
- $T^- = 29.5$
- $T = 46$
- $n = 14 > 12$

Subject	Signed Rank	
	+	-
1	11	
2	4.5	
3	12	
4		4.5
5		2
6	1	
7		7
8		3
9	9	
10		
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75.5 29.5

Wilcoxon Signed-Rank Test

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- Given $T = 46$, we then have the following:

Wilcoxon Signed-Rank Test

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$$\mu_T = 0$$

Wilcoxon Signed-Rank Test

$$n = 14$$

- Given $T = 46$, we then have the following:

$$\mu_T = 0$$

$$\sigma_T = \sqrt{\frac{n(n+1)(2n+1)}{6}} = 31.86$$

Wilcoxon Signed-Rank Test

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- Thus,

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- Given $T = 46$, we then have the following:

$$\mu_T = 0$$

$$\sigma_T = \sqrt{\frac{n(n+1)(2n+1)}{6}} = 31.86$$

- Thus,

$$z_T = \frac{T - \mu_T}{\sigma_T} = \frac{46 - 0}{31.86} = \underline{1.44}$$

Wilcoxon Signed-Rank Test

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- Calculating the p-value, we have

Wilcoxon Signed-Rank Test

$$Z_T = \frac{T - \mu_T}{\sigma_T} \sim N(0,1)$$

- Calculating the p-value, we have

$$p = 2 \cdot \Pr(Z > 1.44) = 2 \cdot (1 - \text{pnorm}(1.44)) = 0.149$$

- Conclusion: $p = 0.149 > \alpha = 0.05$

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- In that case, we can use `psignrank(T, n)` in R to calculate the exact p-value

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- If the sample size is $n \leq 12$, we cannot use the normal approximation
- In that case, we can use `psignrank(T, n)` in R to calculate the exact p-value
 - $2 * (1 - \text{psignrank}(\overset{75.5}{\cancel{74}}, n=14)) = \cancel{0.1213} \quad 0.135$

***** R requires $T = T^+$ for this to work correctly!

Wilcoxon Signed-Rank Test: R Code

$$d_i = x_i - y_i$$

$$Z \sim N(0,1)$$

```
> wilcox.test(before, after, paired=T, exact=F, correct=F)
```

Wilcoxon signed rank test

data: before and after

V = 64, p-value = 0.1961

alternative hypothesis: true location shift is not equal to 0

\uparrow $p_{norm}(T, n)$

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Wilcoxon signed rank test

data: before and after

V = 64, p-value = 0.2163

alternative hypothesis: true location shift is not equal to 0

\uparrow $p_{signrank}(T, n)$

Motivating Example #2

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- Cannot assume that body temperature is normally distributed
- Take a sample of $n_1 = 12$ people who do not have the disease
- Take a sample of $n_2 = 15$ people who do have the disease
- How can we compare the median body temperature for these two populations?

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- Take a sample of $n_2 = 15$ people who do have the disease
- How can we compare the median body temperature for these two populations?
 - *Wilcoxon Rank-Sum Test (Mann-Whitney U Test)*

Wilcoxon Rank-Sum Test

Wilcoxon Rank-Sum Test

- Used to compare samples from independent populations

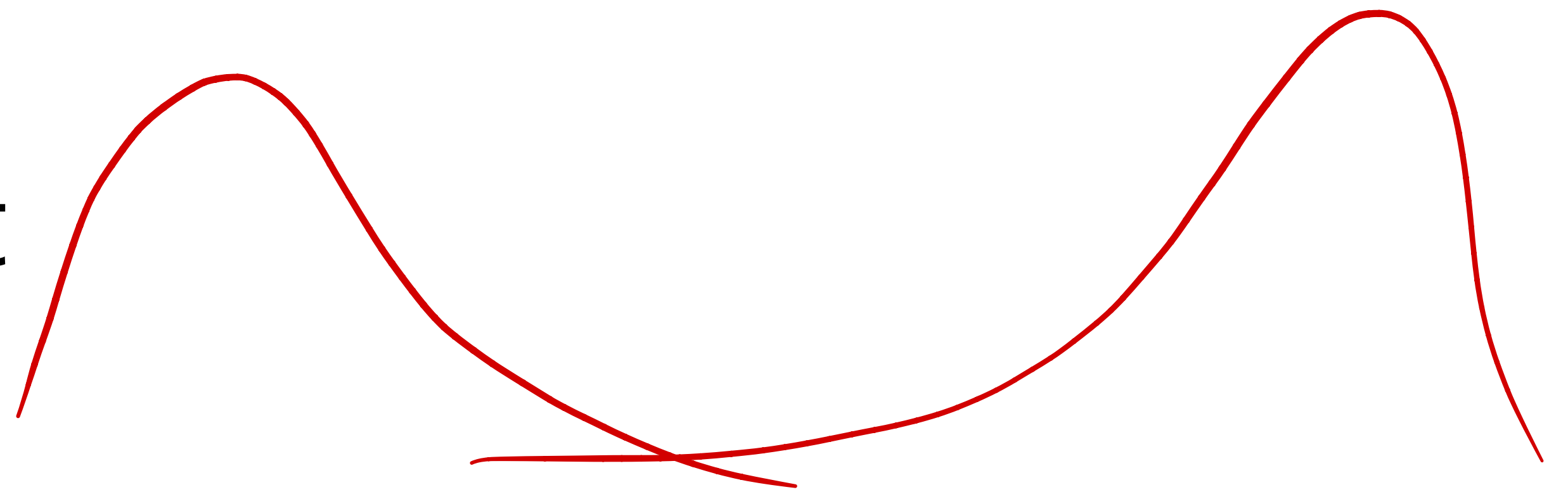
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- H_0 : The medians of the two populations are identical

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- Test at the $\alpha = 0.05$ significance level

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- Thus, the average ranks for the two samples (i.e., W_1/n_1 and W_2/n_2) should be approximately equal

Data Table

No Disease		Disease	
Temp	Rank	Temp	Rank
98.1	1	99.3	8
98.5	2	99.4	[4.5 4.5
98.6	3	99.4	
98.8	4	99.5	
98.9	5	99.5	
99.0	6	99.6	
99.2	7	99.7	
99.5		99.7	
99.6		100.0	
99.7		100.0	
100.5		100.1	
101.0		100.1	
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99.5	12	99.7	17
99.6	14.5	100.0	19.5
99.7	17	100.0	19.5
100.5	24	100.1	22
101.0	25	100.1	22
		100.1	22
		101.1	26
		101.9	27

$w_1 =$

$w_2 =$

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$$W - \binom{n_1}{2}$$

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• Then,

$$\mu_W = \frac{n_1(n_1 + n_2 + 1)}{2} \text{ and } \sigma_W = \sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}}$$

$$\mu = n_1 \left(\frac{n_1 + n_2}{2} \right)$$

avg rank

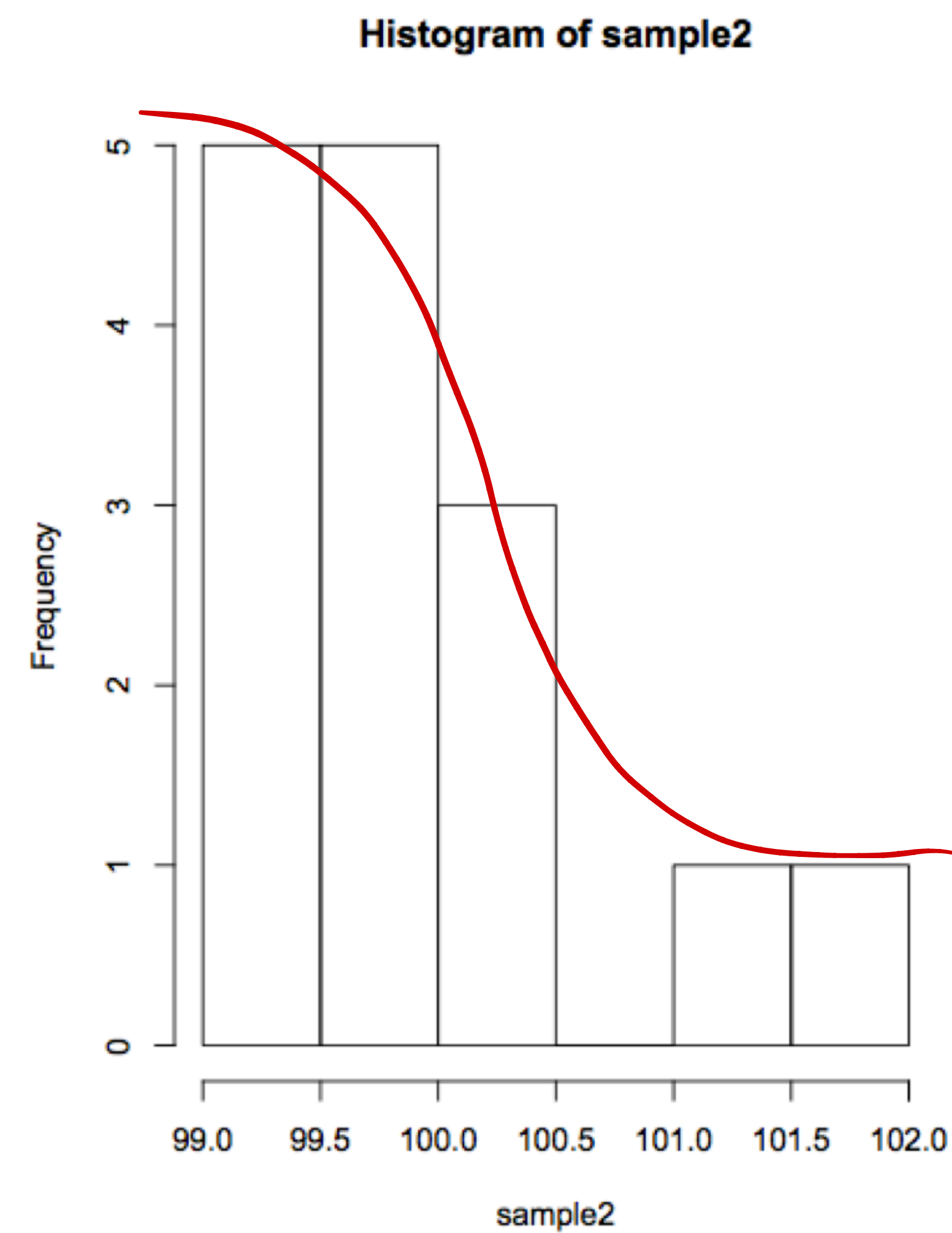
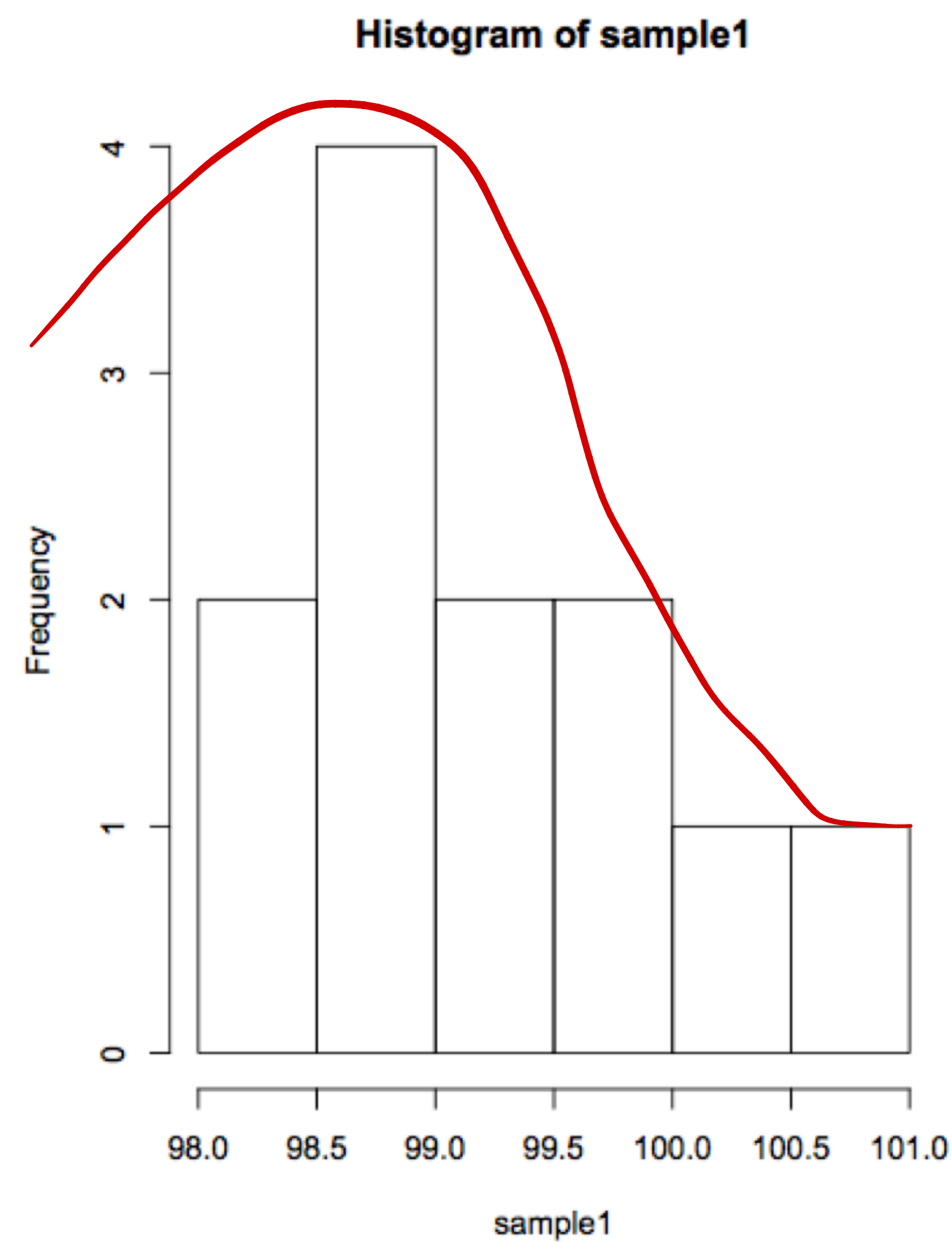
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- $\mu_W = \frac{n_1(n_1 + n_2 + 1)}{2}$ and $\sigma_W = \sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}}$

- $z_W \sim N(0,1)$ when n_1 and n_2 are large enough ($n_1, n_2 > 10$)

Histograms of Samples



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- The sum of ranks for sample 1 is 120.5 $= W_1$
- The sum of ranks for sample 2 is 257.5 $= W_2$

$$W = \min(W_1, W_2) = \underline{120.5}$$

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- Wilcoxon rank sum test is appropriate since both populations have similar shapes
- The sum of ranks for sample 1 is 120.5
- The sum of ranks for sample 2 is 257.5
- Thus, $W = 120.5$, $n_1 = 12$, and $n_2 = 15$

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Wilcoxon Rank-Sum Test

$$\frac{1}{n_1 + n_2} \cdot 1 + \frac{1}{n_1 + n_2} \cdot 2 + \dots + \frac{1}{n_1 + n_2} \cdot (n_1 + n_2)$$

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- $\mu_W = \frac{n_1(n_1 + n_2 + 1)}{2} =$

- $\sigma_W = \sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}} =$

	A	B
x_1	21	y_1
\vdots		\vdots
x_{n_1}		y_{n_2}

Ranks

1
:
 $n_1 + n_2$

$$\frac{(n_1 + n_2) + 1}{2}$$

Wilcoxon Rank-Sum Test

- Given $W = 120.5$, $n_1 = 12$, and $n_2 = 15$, we then have

- $\mu_W = \frac{n_1(n_1 + n_2 + 1)}{2} = \frac{12(12+15+1)}{2} = 168$

- $\sigma_W = \sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}} = 20.49$

- Thus, we have $z_W = \frac{W - \mu_W}{\sigma_W} = \frac{120.5 - 168}{20.49} = -2.318$

Wilcoxon Rank-Sum Test

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Wilcoxon Rank-Sum Test

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$$p = \Pr(Z < -2.318) = \text{pnorm}(-2.318) = 0.01$$

- Conclusion:

$$p < \alpha \rightarrow \text{Reject } H_0.$$

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$$W = \min(W_1, W_2)$$

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 - In this case, $W_{\text{obs}} = W - \frac{n_1(n_1+1)}{2}$ 
 - `pwilcox(120.5-78, 12, 15) = 0.0093`

Wilcoxon Rank-Sum Test: R Code

$F \rightarrow N(0,1)$
 $T \rightarrow \text{Wilcoxon distr.}$

```
> wilcox.test(sample1, sample2, exact=F, correct=F, alt="less")
```

Wilcoxon rank sum test

data: sample1 and sample2

W = 42.5, p-value = 0.01009

alternative hypothesis: true location shift is less than 0

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