

Chapter 9: Inference for Variances

DSCC 462

Computational Introduction to Statistics

Anson Kahng

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- Primarily, inference tends to be focused on means (and proportions, which we will cover next)
- However, variances are essential to inference of means
- We may be interested in determining:
 - Whether a population variance is equal to a predetermined value
 - Whether two population variances are equal

Sampling Distribution of Variance


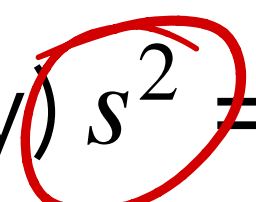
Sampling Distribution of Variance

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- Now, examine s^2/σ^2

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- We have $\underline{\frac{s^2}{\sigma^2}} = \frac{1}{n-1} \cdot \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma} \right)^2$

χ_n^2

$$\frac{x_i - \mu}{\sigma} \sim N(0, 1)$$

χ_1^2

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- Rewrite the sample variance as (approximately) $s^2 = \frac{1}{n-1} \cdot \sum_{i=1}^n (x_i - \mu)^2$
- Now, examine s^2/σ^2
- We have $\frac{s^2}{\sigma^2} = \frac{1}{n-1} \cdot \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma} \right)^2$
- Therefore, $\frac{s^2}{\sigma^2} = \frac{1}{n-1} \cdot \sum_{i=1}^n Z_i^2$, where Z_i is a standard normal random variable

Chi-squared (χ^2) distribution

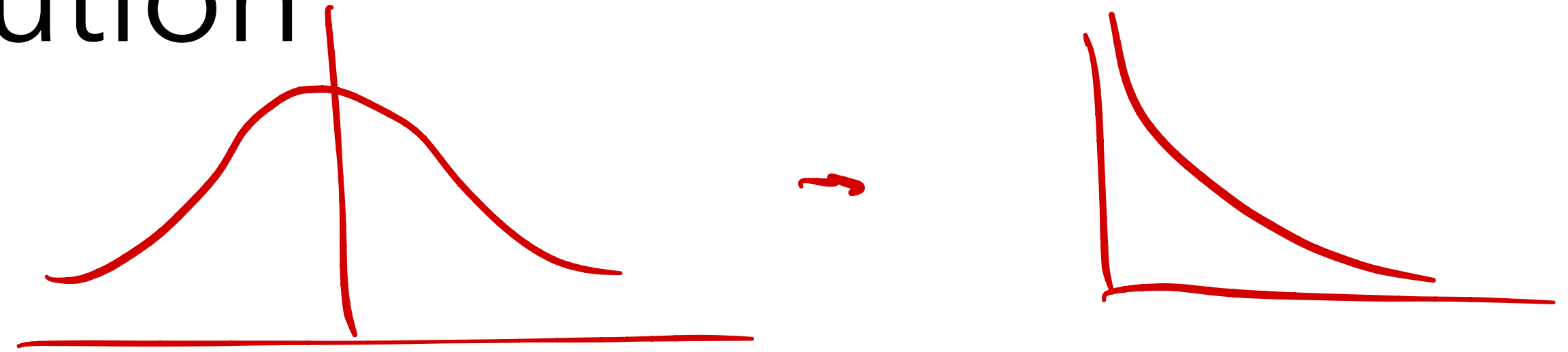
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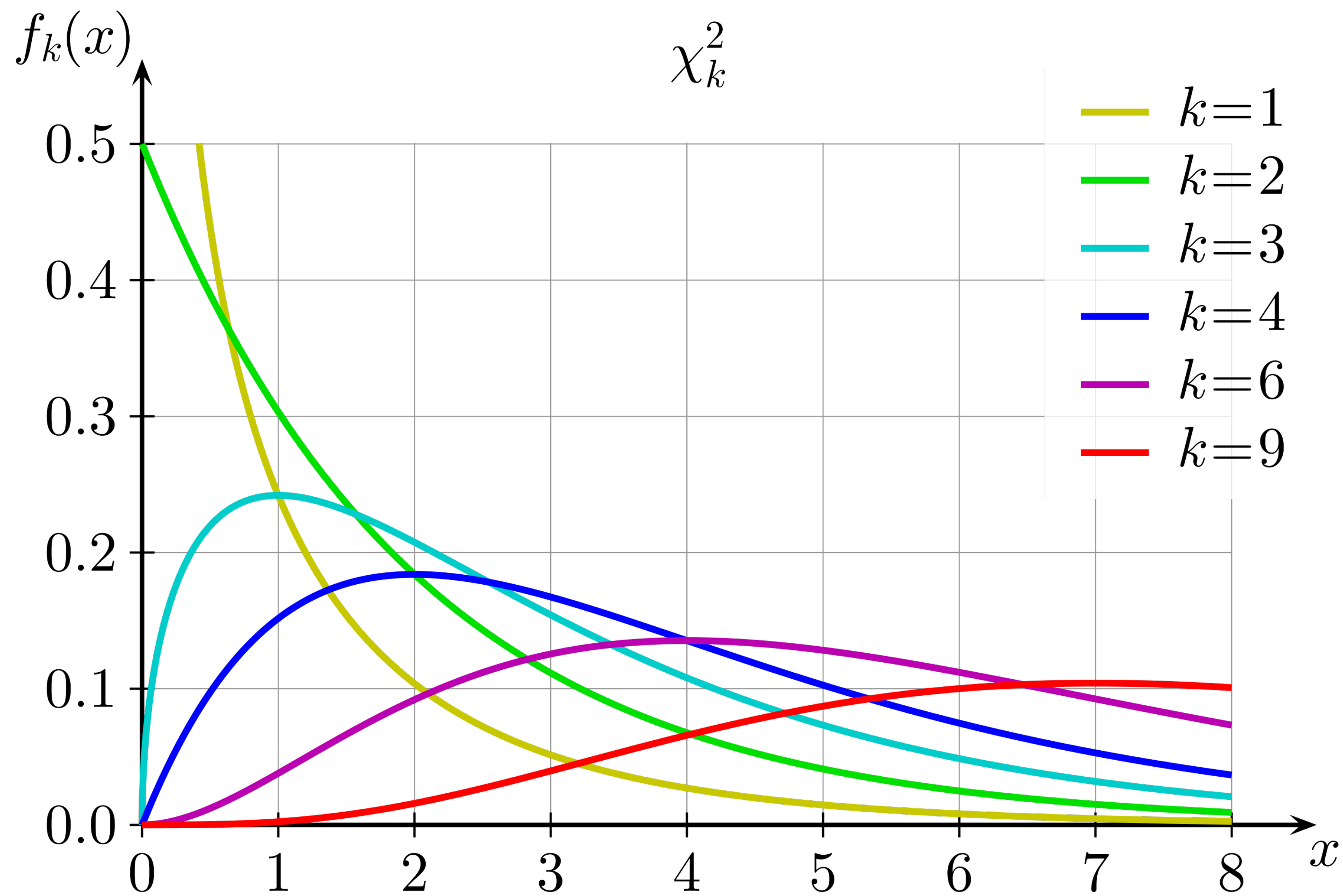


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- Mean of χ_k^2 is k (degrees of freedom)
- Variance of χ_k^2 is $2k$ (twice the degrees of freedom)

Chi-squared (χ^2) distribution



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- Recall that we assumed we knew σ^2 ; we don't actually know this and have to use s^2 as an estimate
- This means that $(n-1) \cdot \frac{s^2}{\sigma^2} \sim \chi_{n-1}^2$ (chi-squared distribution with $n-1$ dof)

Hypothesis Tests for One Population Variance

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- We have considered the case of comparing the *mean* of a population to a predetermined value
- We can also test whether the *variance* of a population is a specified value

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- Recall that we know $(n - 1) \cdot \frac{s^2}{\sigma^2} \sim \chi_{n-1}^2$, so it makes sense to define

$$T = (n - 1) \cdot \frac{s^2}{\sigma^2} \text{ based on our null hypothesis}$$

$$\text{null } H_0 : \sigma^2 = \sigma_0^2$$

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- Once we have our test statistic, we can compare to the chi-square distribution to calculate a p-value

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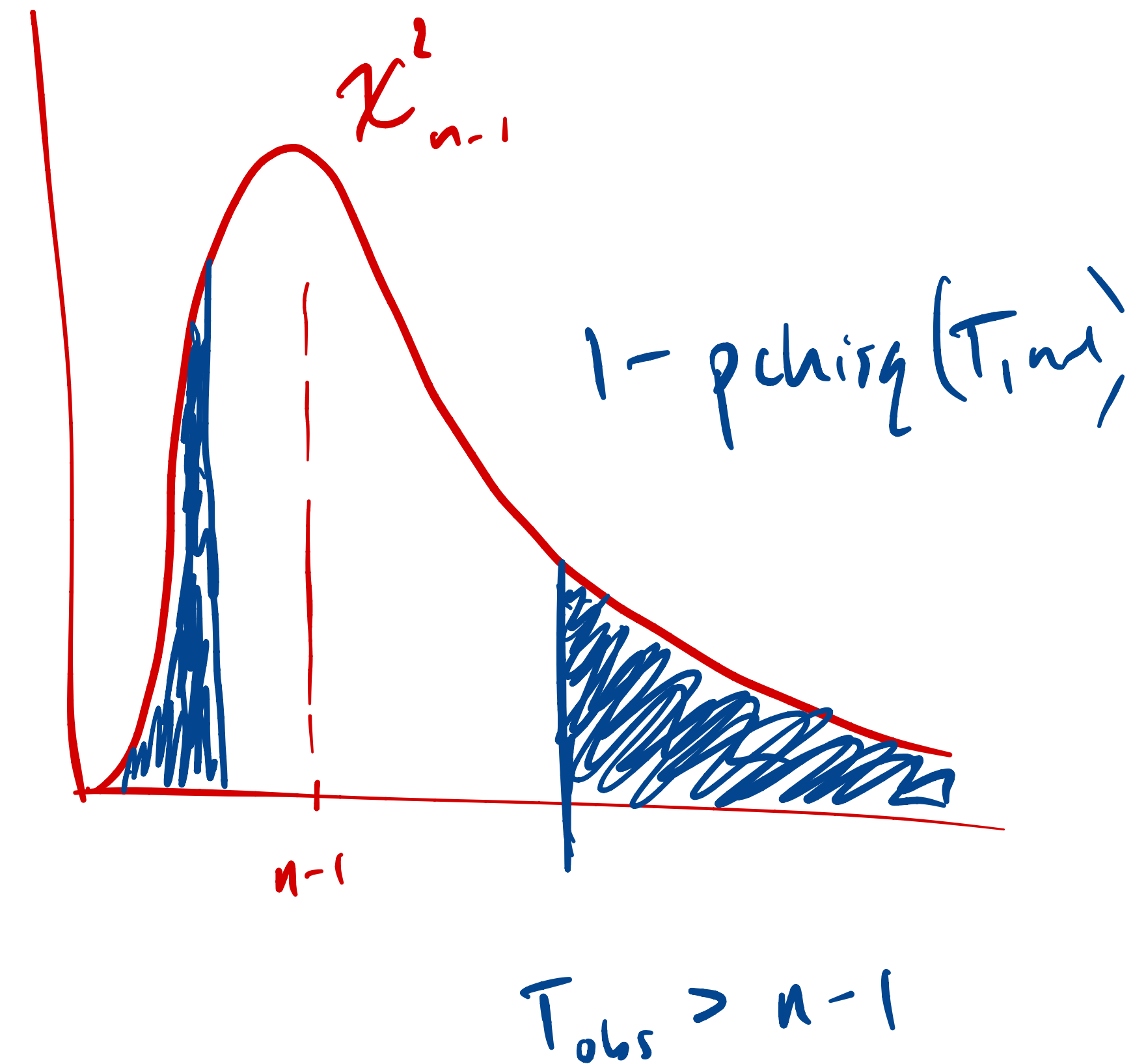
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 - If $T_{obs} > n - 1$, we have `2*(1-pchisq(Tobs, n-1))`



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- If the p-value is less than α , reject the null hypothesis and conclude that the population variances are unequal to each other
- If the p-value is greater than α , fail to reject the null hypothesis and conclude that the population variances are equal to each other

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- Hypotheses: $H_0 : \sigma^2 = 121$ v. $H_1 : \sigma^2 \neq 121$

- Calculate the T statistic: $T = (n-1) \cdot \frac{s^2}{\sigma^2} = 44 \cdot \frac{196}{121} = 71.3 > 44$

- Calculate the p-value: $2 \times (1 - \text{pchisq}(71.3, 44)) = 0.0115$

- Conclusion: $p > \alpha$, fail to reject.

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- Hypotheses:
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- Calculate the p-value:

$$s^2 = \frac{1}{n-1} \cdot \sum_{i=1}^n (x_i - \bar{x})^2 \sim \chi_{n-1}^2$$

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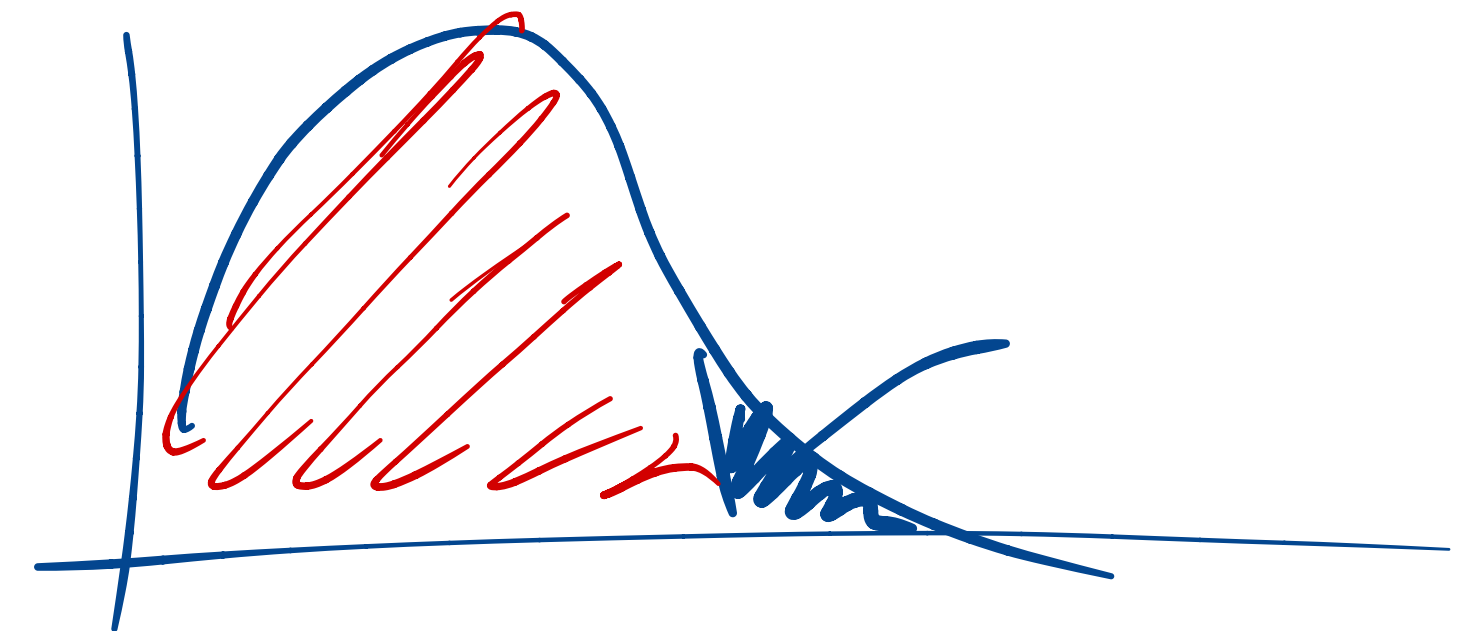
- Hypotheses: $H_0 : \sigma^2 \leq 121$ $H_1 : \sigma^2 > 121$

- Calculate the T statistic: $T = (n-1) \frac{s^2}{\sigma^2} = 71.3$

- Calculate the p-value:

$$1 - \text{pchisq}(71.3, 44) \approx 0.0055$$

- Conclusion: $p < \alpha \rightarrow$ Reject σ^2 increased since 2010



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- Because we have that $(n - 1) \cdot \frac{s^2}{\sigma^2} \sim \chi_{n-1}^2$, we have

$q^{\text{chisq}}(\alpha/2, n-1)$

$$\Pr\left(\chi_{\alpha/2, n-1}^2 \leq (n - 1) \cdot \frac{s^2}{\sigma^2} \leq \chi_{1-\alpha/2, n-1}^2\right) = \alpha$$

$$\Rightarrow \Pr\left(\frac{1}{\chi_{\alpha/2, n-1}^2} \geq \frac{\sigma^2}{(n - 1) \cdot s^2} \geq \frac{1}{\chi_{1-\alpha/2, n-1}^2}\right) = \alpha$$

$$\Rightarrow \Pr\left(\frac{(n - 1) \cdot s^2}{\chi_{1-\alpha/2, n-1}^2} \leq \sigma^2 \leq \frac{(n - 1) \cdot s^2}{\chi_{\alpha/2, n-1}^2}\right) = \alpha$$

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- Because we have that $(n - 1) \cdot \frac{s^2}{\sigma^2} \sim \chi_{n-1}^2$, we have

$$\Pr\left(\chi_{\alpha/2, n-1}^2 \leq (n - 1) \cdot \frac{s^2}{\sigma^2} \leq \chi_{1-\alpha/2, n-1}^2\right) = 1 - \alpha$$

$$\Rightarrow \Pr\left(\frac{1}{\chi_{\alpha/2, n-1}^2} \geq \frac{\sigma^2}{(n - 1) \cdot s^2} \geq \frac{1}{\chi_{1-\alpha/2, n-1}^2}\right) = 1 - \alpha$$

$$\Rightarrow \Pr\left(\frac{(n - 1) \cdot s^2}{\chi_{1-\alpha/2, n-1}^2} \leq \sigma^2 \leq \frac{(n - 1) \cdot s^2}{\chi_{\alpha/2, n-1}^2}\right) = 1 - \alpha$$

- Therefore, the interval $\left(\frac{(n - 1) \cdot s^2}{\chi_{1-\alpha/2, n-1}^2}, \frac{(n - 1) \cdot s^2}{\chi_{\alpha/2, n-1}^2}\right)$ contains σ^2 with probability $1 - \alpha$

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- One-sided, upper:

- $\left(0, \frac{(n - 1) \cdot s^2}{\chi^2_{\alpha, n-1}}\right)$

$$\chi^2_{\alpha, n-1}$$

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- One-sided, lower:

- $\left(\frac{(n - 1) \cdot s^2}{\chi^2_{1-\alpha, n-1}}, \infty \right)$

Confidence Intervals for One Population Variance: Example

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- Let's say we are interested in the amount of water the average person drinks per day. We sample a group of $n = 40$ people and find that the sample variance is $s^2 = 60 \text{ oz}^2$. What is a two-tailed 95% confidence interval for the variance of this distribution?

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- What is α ?

Confidence Intervals for One Population Variance: Example

$$(n-1) \cdot \frac{s^2}{r^2}$$

$$\frac{60}{40.26} \qquad \frac{98.92}{60}$$

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- What is α ? $\alpha = 0.05$

$$\text{CI: } \left(\frac{(n-1) \cdot s^2}{\chi^2_{1-\alpha/2, n-1}}, \frac{(n-1) \cdot s^2}{\chi^2_{\alpha/2, n-1}} \right) = \left(\frac{39 \cdot 60}{\chi^2_{0.975, 39}}, \frac{39 \cdot 60}{\chi^2_{0.025, 39}} \right)$$

$(40.26, 98.92)$

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- We have talked about testing whether the means of the two independent populations are equal to each other
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Hypothesis Tests for Two Population Variances

- Null hypothesis: Population variances are equal
- We attempt to see if there is a difference in variances
- One-tailed, lower hypothesis:
 - $H_0 : \sigma_1^2 \geq \sigma_2^2$ vs. $H_1 : \sigma_1^2 < \sigma_2^2$
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$$(n-1) \cdot \frac{s^2}{\sigma^2} \sim \chi^2_{n-1}$$

- Recall that for a single population, we have $\frac{s^2}{\sigma^2} \sim \frac{1}{n-1} \cdot \chi^2_{n-1}$

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F distribution

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- It turns out that this type of distribution has a name: F distribution

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Handwritten red annotations:

- A red circle around the fraction $\frac{s_1^2 / \sigma_1^2}{s_2^2 / \sigma_2^2}$.
- A red arrow pointing from the word "num" to the numerator s_1^2 / σ_1^2 .
- A red arrow pointing from the word "denom" to the denominator s_2^2 / σ_2^2 .

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- Once we have our test statistic, we can compare to the F distribution to calculate a p-value

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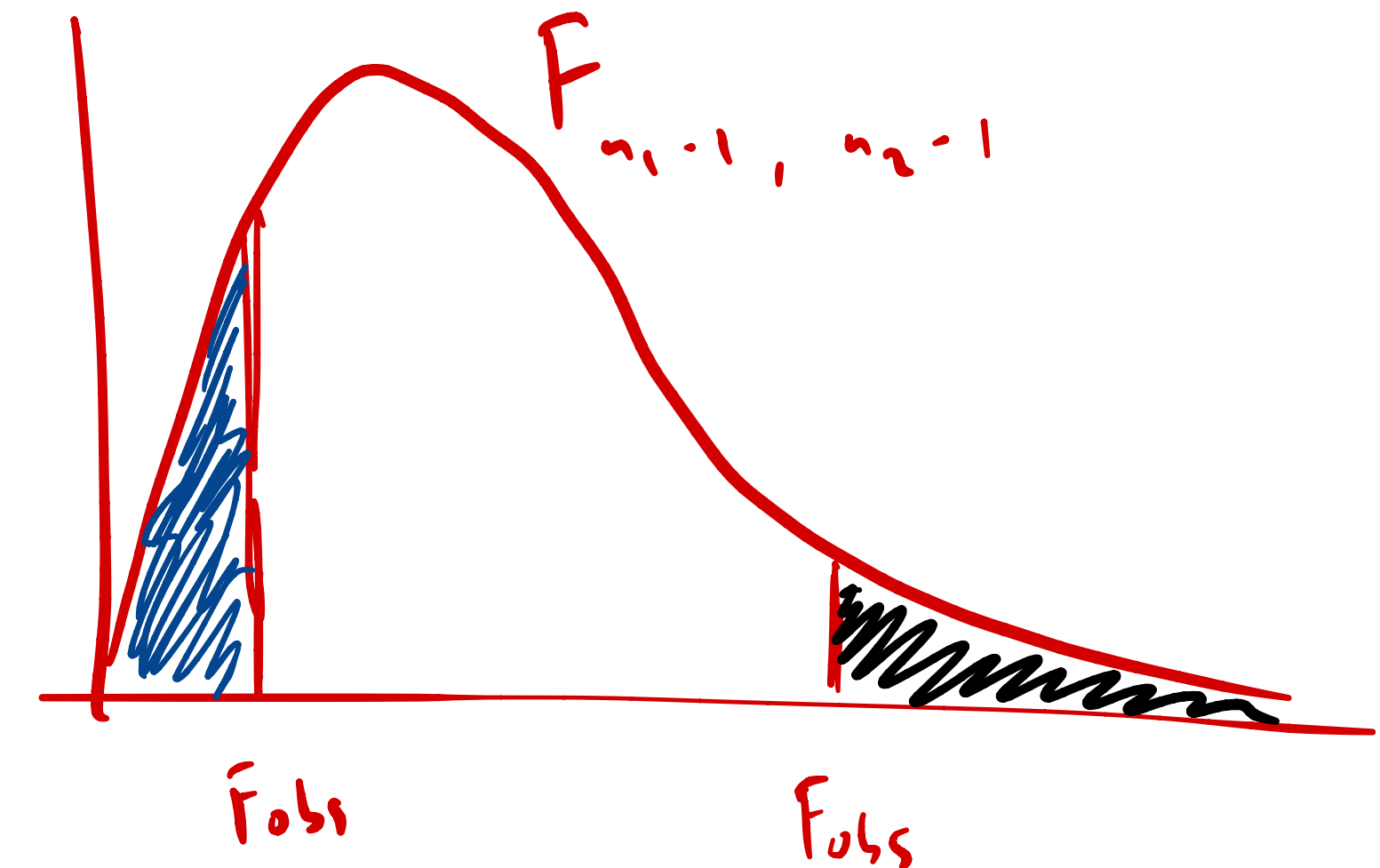
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 - Two-tailed hypothesis: $p = \Pr(|T| \geq |T_{obs}|)$
 - If $F_{obs} \leq 1$, we have $2 * pf(F_{obs}, n_1-1, n_2-1)$
 - If $F_{obs} > 1$, we have $2 * (1 - pf(F_{obs}, n_1-1, n_2-1))$



$$\bar{F} = \frac{\chi^2_{n_1-1} / (n_1-1)}{\chi^2_{n_2-1} / (n_2-1)}$$

$\mu = 1$ (pointing to the numerator)
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- If the p-value is greater than α , fail to reject the null hypothesis and conclude that the population variances are equal to each other

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$$qf($$

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- Two-tailed hypothesis:
 - Reject if $F_{obs} \leq F_{n_1-1, n_2-1, \alpha/2}$ or $F_{obs} \geq F_{n_1-1, n_2-1, 1-\alpha/2}$

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- Evaluate at $\alpha = 0.05$

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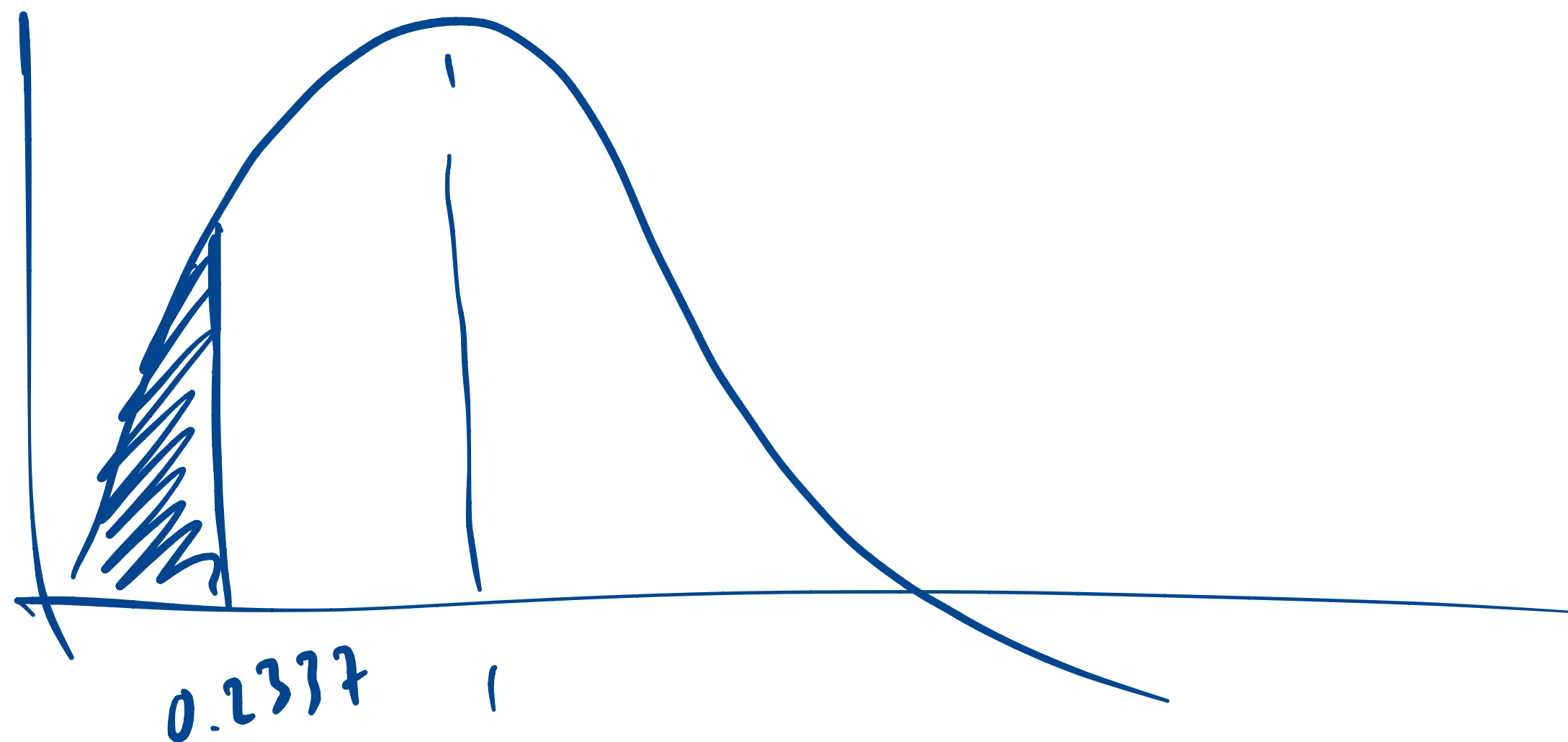
Hypothesis Tests for Two Population Variances

- Calculate the test statistic:

$$F = \frac{s_1^2}{s_2^2} = \frac{2.33}{9.97} \approx 0.2337$$

$$p = 2 \times pf(0.2337, 9, 19) \\ = \underline{0.03}$$

- Conclusion:



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- Similarly, we could have calculated the critical values for the F-test statistic and used them to complete the test:

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$$F = 0.1337$$

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$$\left. \begin{aligned} F_{n_1-1, n_2-1, \alpha/2} &= \text{qf}(0.025, 9, 19) = 0.2715 \\ F_{n_1-1, n_2-1, 1-\alpha/2} &= \text{qf}(0.975, 9, 19) = 2.880 \end{aligned} \right\}$$

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- Since $F = 0.2337 < 0.2715$, we reject the null hypothesis and conclude $\sigma_1^2 \neq \sigma_2^2$

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$$\frac{1}{\chi_{n-1, 1-\alpha/2}^2} \cdot \dots \quad , \quad \frac{1}{\chi_{n-1, \alpha/2}^2} \cdot \dots$$

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- A one-sided upper $(1 - \alpha) \cdot 100\%$ confidence interval for σ_1^2/σ_2^2 is given by
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$$\left(0, \frac{1}{F_{\alpha}} \cdot \frac{s_1^2}{s_2^2} \right)$$
- A one-sided lower $(1 - \alpha) \cdot 100\%$ confidence interval for σ_1^2/σ_2^2 is given by
$$\left(\frac{1}{F_{1-\alpha}} \cdot \frac{s_1^2}{s_2^2}, \infty \right)$$

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- What is α ? $\alpha = 0.05$ $(0, 0.689)$

- CI: $\left(0, \frac{1}{F_\alpha} \cdot \frac{s_1^2}{s_2^2} \right) = \left(0, \frac{1}{0.379} \cdot \frac{2.33}{9.97} \right)$
 $\uparrow qf(0.05, 9, 19) = 0.379$