

$$\textcircled{1} \quad \text{a) } J(\theta) = (\vec{x}\theta - \vec{y})^T (\vec{x}\theta - \vec{y})$$

$$= \theta^T x^T x \theta - 2\vec{y}^T x \theta + \vec{y}^2$$

$$\frac{\partial J}{\partial \theta} = (x^T x + (x^T x)^T) \theta - 2\vec{y}^T x = 0$$

$$x^T x \theta = \vec{y}^T x$$

$$\theta = (x^T x)^{-1} \vec{y}^T x$$

$\begin{matrix} x & \vec{v} \\ \vdots & \vdots \\ m \times (n+2) \end{matrix}$    
  $\begin{matrix} -y^T \\ -\vec{v}^T \\ \vdots \\ n+2 \times m \end{matrix}$    
  $\theta = (x^T x)^{-1} \vec{y}^T x$    
  $x^T x \in \mathbb{R}^{(n+1) \times (n+1)}$   
 $x^T x \in \mathbb{R}^{(n+1) \times (n+1)}$    
  $\vec{y}^T \in \mathbb{R}^{m \times (n+1)}$   
 $\vec{y}^T \in \mathbb{R}^{(n+1) \times 1}$

$$\hat{\theta} = \underline{x}^T \vec{y}$$

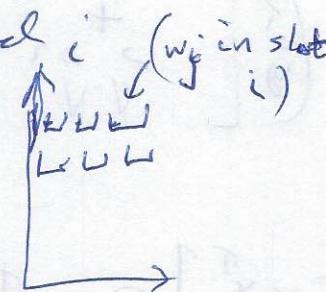
$$\Rightarrow J(\hat{\theta}) = (\underline{x}^T \vec{y} - \vec{y})^T (\underline{x}^T \vec{y} - \vec{y})$$

$$\text{b) } \tilde{J}(\hat{\theta}_{\text{new}}) = (\tilde{x}^T \vec{y} - \vec{y})^T (\tilde{x}^T \vec{y} - \vec{y})$$

$$= ([x \vec{v}] \begin{bmatrix} x^T \\ \vec{v}^T \end{bmatrix} \vec{y} - \vec{y})^T ([x \vec{v}] \begin{bmatrix} x^T \\ \vec{v}^T \end{bmatrix} \vec{y} - \vec{y})$$

$$\hat{\theta}_{\text{new}} = \cancel{\left( \tilde{x}^T \vec{y} \right)}^T \tilde{x}^T \vec{y} - \vec{y} \cancel{\left( \tilde{x}^T \vec{y} \right)}^T \left( \tilde{x}^T \vec{y} - \vec{y} \right)$$

② a)  $c_1, \dots, c_i, c_v$  = number of times word  $i$  ( $w_i$  in slot  $i$ ) appears in document



$$p(x|y) = p(c_i=m | y=x)$$

~~$$L = \left( \prod_{i=1}^n p(w_i=m | y=x) \right)^{\phi_y^{(i)}} (1-\phi_y^{(i)})^{1-\phi_y^{(i)}}$$~~

↑ probability 1

(~~(c)~~) ~~# of possible combination~~

~~$$L = \left( \prod_{i=1}^n p(w_i=m | y=x) \right)^{\phi_y^{(i)}} (1-\phi_y^{(i)})^{1-\phi_y^{(i)}}$$~~

bc measure  
x are II, conditioned  
on y.

$$L = \left( \prod_{i=1}^n p(w_i=m | y=y) \right)^{\phi_y^{(i)}} (1-\phi_y^{(i)})^{1-\phi_y^{(i)}}$$

$$L = \left( \prod_{i=1}^n p(x_i=m | y=y) \right)^{\phi_y^{(i)}} (1-\phi_y^{(i)})^{1-\phi_y^{(i)}}$$

$$p(w_i|y) = \frac{1}{n} \sum_{i=1}^n \log(p(x_i=m | y=y)) + ny \log \phi$$

$$\ln(L) = \sum_{i=1}^n \log(p(x_i=m | y=y)) + n(1-y) \log(1-\phi)$$

$$= \sum_{i=1}^n \log(p(x_i=m | y=y)) + ny \log \phi + n(1-y) \log(1-\phi)$$

$p(w_i|y)$ :  
probability word  $i$  is  
in doc, given  $y$

$$\begin{aligned}
 J(\hat{\theta}_{\text{new}}) &= \left( \vec{y} \left[ \begin{array}{c|cc} x & x^T \\ \hline v & v^T \end{array} \right] \vec{y} - \vec{y}^T \right) \left( \left[ x x^T + \vec{v} \vec{v}^T \right] \vec{y} - \vec{y} \right) \\
 &\quad - \vec{y}^T \left[ \begin{array}{c|cc} x & x^T \\ \hline v & v^T \end{array} \right] \vec{y} - \vec{y}^T \left[ \begin{array}{c|cc} x & x^T \\ \hline v & v^T \end{array} \right] \vec{y} \\
 &\quad - \vec{y}^T \left[ x x^T + \vec{v} \vec{v}^T \right] \vec{y} + \vec{y}^T \vec{y} \\
 \tilde{J}(\hat{\theta}_{\text{new}}) &= \vec{y}^T \left( x x^T \right) \vec{y} + 2 \vec{y}^T \vec{v} \vec{v}^T \vec{y} + (\vec{v} \vec{v}^T) (\vec{v} \vec{v}^T) \vec{y} - 2 \vec{y}^T (x x^T + \vec{v} \vec{v}^T) \vec{y} \\
 &= \vec{y}^T \left( x x^T \right) \vec{y} - 2 \vec{y}^T \left( x x^T \vec{y} - \vec{y} \right) + 2 \vec{y}^T \vec{v} \vec{v}^T (\vec{v} \vec{v}^T - \vec{I}) \vec{y} \\
 &= \cancel{\vec{y}^T \left( x x^T \right) \vec{y}} - 2 \vec{y}^T \left( \vec{v} \vec{v}^T \right) \vec{y} + \cancel{2 \vec{y}^T \left( x x^T \vec{y} - \vec{y} \right)} + \cancel{2 \vec{y}^T \vec{v} \vec{v}^T (\vec{v} \vec{v}^T - \vec{I}) \vec{y}} \\
 f(x, \vec{v}, \vec{y}) &= 2 \vec{y}^T x x^T \vec{y} + \vec{y}^T (\vec{v} \vec{v}^T) (\vec{v} \vec{v}^T) \vec{y} \\
 &= 2 \vec{y}^T \vec{v} \vec{v}^T \vec{y} + \vec{y}^T (\vec{v} \vec{v}^T) (\vec{v} \vec{v}^T) \vec{y} \\
 &= - \vec{y}^T (\vec{v} \vec{v}^T) \vec{y} - (\vec{v} \vec{v}^T)^2 \\
 \tilde{J}(\hat{\theta}_{\text{new}}) &= J(\hat{\theta}) + f(x, \vec{v}, \vec{y})
 \end{aligned}$$

show  $f(x, \vec{v}, \vec{y}) \leq 0$

if no, b/c 1 only describes bias  
 reduces to extra feature,  
 don't account for variance

$$P(y=1/x) = \frac{1}{\left[ 1 + \prod_{j=1}^V \frac{p(x_j | y=0)^{c_j}}{p(x_j | y=1)} \frac{p(y=0)}{p(y=1)} \right]}$$

$\Rightarrow F$

$$F \leq 1 \Rightarrow \hat{y} = 1$$

$$= \prod_{j=1}^V \left[ p(x_j | y=0)^{c_j} / p(x_j | y=1)^{c_j} \frac{p(y=0)}{p(y=1)} \right]$$

$$\log(L) = \sum_{j=1}^V \left[ \log(p(x_j | y=1)) - \log(p(x_j | y=0)) + \log(p(y=0)) - \log(p(y=1)) \right] \leq \log(1) = 0$$

$\log(L) \leq 0$

$$\log(L) = \sum_{j=1}^V c_j \left[ \log(p(x_j | y=1)) - \log(p(x_j | y=0)) + \sqrt{\log \left[ \frac{(1-t)}{t} \right] - \log(t)} \right]$$

$$= \sum_{j=1}^V (w_j + b) \geq 0$$

$$P(y=1|x) \geq P(y=0|x) \Rightarrow$$

$$\Rightarrow \frac{P(x|y=1) P(y=1)}{P(x)} \geq \frac{P(x|y=0) P(y=0)}{P(x)}$$

$$\Rightarrow N(\mu_0, \Sigma_0)\phi \geq N(\mu_1, \Sigma_1)(1-\phi)$$

$$\ln(N(\mu_0, \Sigma_0)) + \ln(\phi) \geq \ln(N(\mu_1, \Sigma_1)) + \ln(1-\phi)$$

$$\ln\left(\frac{1}{2\pi^{n/2}}\right) + \ln\left(\frac{1}{|\Sigma_0|^{1/2}}\right) + \left(-\frac{1}{2}(x-\mu_0)^T \Sigma_0^{-1} (x-\mu_0)\right)$$

$$+ \ln\phi \geq \ln\left(\frac{1}{2\pi^{n/2}}\right) + \ln\left(\frac{1}{|\Sigma_1|^{1/2}}\right)$$

$$+ \left(-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1} (x-\mu_1)\right)$$

$$\ln(1-\phi) - \frac{1}{2}\frac{(x-\mu_1)^T \Sigma_1^{-1} (x-\mu_1)}{(x-\mu_0)^T \Sigma_0^{-1} (x-\mu_0)}$$

$$\Rightarrow \left[ \frac{1}{2}\ln|\Sigma_0| + \frac{1}{2}\ln|\Sigma_1| \right] - \frac{(x-\mu_1)^T \Sigma_1^{-1} (x-\mu_1)}{(x-\mu_0)^T \Sigma_0^{-1} (x-\mu_0)} + \ln\phi - \ln(1-\phi) \geq 0$$

$$C \triangleq \left[ \frac{1}{2} \ln(\bar{\Sigma}_0) + \frac{1}{2} \ln(\bar{\Sigma}_1) \right]$$

$$2b) + \ln \phi - \ln(1-\phi) + \frac{1}{2} \left[ \mu_0^T \bar{\Sigma}_0^{-1} \mu_0 + \mu_1^T \bar{\Sigma}_1^{-1} \mu_1 \right] \\ \cancel{- \frac{1}{2} (x - \mu_0)^T [\bar{\Sigma}_0^{-1} \bar{\Sigma}_1^{-1}] (x - \mu_0) \mu_1^T \bar{\Sigma}_1^{-1} \mu_1}$$

~~$$-\frac{1}{2} (x - \mu_i)^T \bar{\Sigma}_i^{-1} (x - \mu_i)$$~~

$$= -\frac{1}{2} \left[ x^T \bar{\Sigma}_i^{-1} x - 2x^T \bar{\Sigma}_i^{-1} \mu_i + \mu_i^T \bar{\Sigma}_i^{-1} \mu_i \right]$$

~~$$\partial \bar{\Sigma}_i = -\frac{1}{2} x^T (\bar{\Sigma}_0^{-1} + \bar{\Sigma}_1^{-1}) x$$~~

$$= 2x^T (\bar{\Sigma}_0^{-1} + \bar{\Sigma}_1^{-1}) x + \mu_i$$

~~$$= 2x^T \bar{\Sigma}_0^{-1} \mu_0 - 2x^T \bar{\Sigma}_1^{-1} \mu_1 +$$~~

$$-2x^T \left[ \bar{\Sigma}_0^{-1} \mu_0 + \bar{\Sigma}_1^{-1} \mu_1 \right]$$

$$+ \mu_0^T \bar{\Sigma}_0^{-1} \mu_1 + \mu_1^T \bar{\Sigma}_1^{-1} \mu_1$$

2b, cont'd)

$$\beta = -2 \left[ \frac{1}{\Sigma_0} u_0 - \frac{1}{\Sigma_1} u_1 \right]$$

$$A = \frac{1}{2} \left[ \Sigma_0^{-1} \Sigma_1^{-1} \right]$$

③ i)  $p(y|n) = b(y) \exp(n^T \theta - \alpha(n))$

$$n^T = \log(1-\phi)$$

$$\alpha(n) = -\log(\phi)$$

$$T(y) = \cancel{y+1}$$

$$b(y) = y-1$$

$$\cancel{e^{n^T}} = 1-\phi$$

$$\phi = 1 - e^{n^T}$$

$$3 \text{aii}) \log(p_{xy}^{(i)} | x; \phi) = (y-1) \log \left( \frac{\phi}{1-\phi} \right) (1-\phi) \phi$$

$$= (y-1) \log(1-\phi) + \log(\phi)$$

$$= y \log(1-\phi) + \log \phi - \log(1-\phi)$$

$$\phi = 1 - e^{-\frac{x}{\delta^T x}}$$

$$= y \log(1 - (1 - e^{-\frac{x}{\delta^T x}})) + \log(1 - e^{-\frac{x}{\delta^T x}})$$

$$- (\log(1 - (1 - e^{-\frac{x}{\delta^T x}})))$$

$$= y \left( \log(e^{-\frac{x}{\delta^T x}}) \right) + \log(1 - e^{-\frac{x}{\delta^T x}})$$

$$- \log(e^{-\frac{x}{\delta^T x}})$$

$$\log p_{xy}^{(i)} = \sum_{j=1}^n y_j e^{\frac{x}{\delta^T x}} (y_j - 1) + \log(1 - e^{\frac{x}{\delta^T x}})$$

3# b) (cont'd)

$$\frac{\partial^2 l}{\partial \theta^2} = \frac{\sum_{i=1}^n y_i x^T - \underline{x^T}}{2\theta}$$

$$= \frac{n + x^T (-x e^{\theta x})}{(1 - e^{\theta x})^2} = \frac{\sum_{i=1}^n x^T e^{\theta x}}{(1 - e^{\theta x})^2} \quad \text{OK}$$

$$\theta := \theta - \frac{\nabla l(\theta)}{H}$$

$$\tilde{x}^T \left( \frac{-x^T e^{\theta x}}{(1 - e^{\theta x})^2} \right) \tilde{x} \quad \text{OK}$$

$$= \frac{-(\tilde{x}^T x)(x^T \tilde{x})e^{\theta x}}{(1 - e^{\theta x})^2}$$

$$= \frac{-(\tilde{x}^T)^2 e^{\theta x}}{(1 - e^{\theta x})} \quad \in V$$

$$4) \quad \mathcal{L} = \sum_{i=1}^n \frac{1}{2} \|w\|^2 + \alpha_i (y^{(i)} - w^T x^{(i)} - b - \epsilon)$$

$$+ \alpha_i^* (w^T x^{(i)} + b - y^{(i)} - \epsilon)$$

$$\max_{\alpha, b} \left( \min_{w, b} (\mathcal{L}) \right) = 0$$

$$\frac{\partial \mathcal{L}}{\partial w} = \sum_{i=1}^n \left( \|w\| + \alpha_i x^{(i)} + \alpha_i^* x^{(i)} \right) = 0$$

$$\|w\| = \sqrt{\sum_{i=1}^n \alpha_i^* x^{(i)} (\alpha_i - \alpha_i^*)}$$

$$\frac{\partial \mathcal{L}}{\partial b} = -\alpha_1 b + \alpha_1^* b = 0 \quad \text{and} \quad \alpha_1^* - \alpha_1 = 0$$

$$\frac{\partial \mathcal{L}}{\partial \alpha_i} = \alpha_i^* - \alpha_i = 0 \quad \text{and} \quad \alpha_i^* - \alpha_i = 0$$

$$\frac{\partial \mathcal{L}}{\partial \epsilon} = y^{(i)} - w^T x^{(i)} - b - \epsilon = 0 = y^{(i)} - b - \epsilon = 0$$

$$\frac{\partial \mathcal{L}}{\partial \alpha_i^*} = \left( w^T x^{(i)} + b - y^{(i)} - \epsilon \right) = 0$$

$$\frac{\partial \mathcal{L}}{\partial \alpha_i^*} = \left( w^T x^{(i)} (\alpha_i - \alpha_i^*) + b - y^{(i)} - \epsilon \right) = 0$$

$$+ \alpha_i^* x^{(i)} (\alpha_i - \alpha_i^*) + y^{(i)} - w^T x^{(i)} (\alpha_i - \alpha_i^*) - \epsilon = 0$$

$$+ \alpha_i^* x^{(i)} (\alpha_i - \alpha_i^*) - y^{(i)} - w^T x^{(i)} (\alpha_i - \alpha_i^*) - \epsilon = 0$$

$$- 2\epsilon = 0$$

4a ii)

$$L = \log \prod_{i=1}^m (1 - e^{-\theta^T x_i})$$

$$\theta = \theta_0$$

$\uparrow$   
cat poisson distribution

$$L = \log \sum_{F=1}^m \left[ (y-1) \frac{\log(\theta^T x)}{e^{\theta^T x}} + \log \left[ \frac{-x^T \theta}{(1-e^{\theta^T x})} \right] \right]$$

?

$$= \mathbb{E} \left[ (y-1) \theta^T x - \log \theta^T x \right. \\ \left. + \log [1 - e^{\theta^T x}] \right]$$

4b)

$$\frac{\partial L}{\partial \theta} = \sum \left[ (y-1)x^T - \cancel{x^T} \cancel{+ \frac{x e^{\theta^T x}}{(1-e^{\theta^T x})}} \right]$$

$$= \sum \left[ y x^T + \cancel{\frac{x e^{\theta^T x}}{(1-e^{\theta^T x})}} \right]$$

$$= \sum \left[ y x^T + \cancel{\frac{x e^{\theta^T x}}{(1-e^{\theta^T x})}} + \cancel{x e^{\theta^T x}} \right]$$

$$= \sum \left[ (y-1)x^T (1 - e^{\theta^T x}) - x e^{\theta^T x} \right]$$

$$\cancel{\sum (y-1)x^T} = \sum y x^T - \cancel{\frac{x^T (1 - e^{\theta^T x})}{(1 - e^{\theta^T x})} + x e^{\theta^T x}}$$

$$= \sum y x^T - \cancel{\frac{x^T}{(1 - e^{\theta^T x})}}$$

$$\begin{aligned}
 4(c) \quad L = & \frac{1}{2} \sum_{i=1}^n x^{(i)} \times ((\alpha_i - \alpha_i^*)^2 \\
 & + \sum_{i=1}^n \alpha_i \left[ -y^{(i)} + \frac{1}{2} \sum_{j=1}^n (\alpha_j - \alpha_i) \times \right. \\
 & \quad \left. + b + \epsilon \right] \\
 & + \sum_{i=1}^n \alpha_i^* \left[ -y^{(i)} + \frac{1}{2} \sum_{j=1}^n (\alpha_j - \alpha_i) \times \right. \\
 & \quad \left. + b + \epsilon \right] \\
 & + \sum_{i \neq j}^n (\alpha_i + \alpha_i^*) \times \cancel{\phi(x_j) \phi(x_i)} \\
 - f(w, x) = & \sum_{i=1}^n x^{(i)} \cancel{\alpha_i} (\alpha_i - \vec{\alpha}) \times + b \\
 & \text{just change } n \text{ to } \\
 & (n+1), \text{ repeat compute } w \\
 f(w, x) = & \sum_{i=1}^n (\alpha_i - \vec{\alpha}) k(x^{(i)}, x) + b
 \end{aligned}$$

⑤

(2015 Practice Midterm)  
at least  
w/p prob  
 $\frac{1}{1-g}$

a)  $\hat{\mathbb{E}}(h_0) \leq \mathbb{E}(h_0) + \gamma$

$\hat{\mathbb{E}}(h^*) \leq \mathbb{E}(h^*) + n + \gamma$

$\leq \mathbb{E}(\hat{h}) + n + \gamma$

$\leq \hat{\mathbb{E}}(\hat{h}) + n + 2\gamma$

$\Rightarrow \hat{\mathbb{E}}(h_0) > \hat{\mathbb{E}}(\hat{h}) + n + \gamma$

w/ prob at most  $\delta$

b)

$\mathbb{E}(h_0) > \mathbb{E}(h^*) + n$

Show prob of yes is at most  $\delta$

$\mathbb{E}(h_0) - \mathbb{E}(h^*) > n$

~~$\mathbb{E}(h_0) \rightarrow \mathbb{E}(k^{h_0}) + n$~~

5 b

$$\text{if } \hat{\varepsilon}(h_0) > \varepsilon(h^*) + n$$

show that probability  
of YES is at most  $\delta$   
with prob  $(1-\delta)$

$$\begin{aligned} \hat{\varepsilon}(h_0) &\geq \varepsilon(h_0) - 2\gamma \\ \varepsilon(h_0) &> \varepsilon(h^*) + n \\ &> \varepsilon(h^*) + n - 2\gamma \\ &> \cancel{\left( \varepsilon(h^*) + n - 2\gamma \right)} \\ &> \varepsilon(h^*) + n - 2\gamma \\ &> \varepsilon(h^*) + n - 2\gamma \\ &> \varepsilon(h^*) + n - 2\gamma \end{aligned}$$

$$\begin{aligned} \hat{\varepsilon}(h_0) &\geq \varepsilon(h_0) - \gamma \\ &\geq \varepsilon(h_0) - 2\gamma \\ &\geq \cancel{\varepsilon(h_0) - 2\gamma} \\ &\geq \varepsilon(h^*) + n - \gamma \\ &\geq \varepsilon(h^*) + n - 2\gamma \end{aligned}$$

5c)

If  $m \rightarrow \infty, \gamma \rightarrow 0$ , so

$$\hat{\varepsilon}(h_0) = \hat{\varepsilon}(h^*) \leq \varepsilon(h^*) + \alpha + \gamma \\ \leq \varepsilon(h^*) + n + \gamma$$

w/ prob  $\geq 1 - \delta$

if  $\gamma \rightarrow 0$

$$\hat{\varepsilon} \leftarrow \hat{\varepsilon}(h^*)$$

~~we're done~~ show

~~is  $\hat{\varepsilon}(h_0)$  an optimal~~

~~if  $m \rightarrow \infty$  w/ prob  $\geq 1 - \delta$~~

$$\hat{\varepsilon}(h^*) - \hat{\varepsilon}(h_0) \leq \tau \quad \text{w/ prob } \geq 1 - \delta$$

$$\text{or } [\hat{\varepsilon}(h^*) - \hat{\varepsilon}(h_0)] < (n - 2\gamma)$$

w/ prob  $\geq 1 - \delta$

$\Rightarrow$  true w/ prob  
at least  $1 - \delta$

$$\begin{aligned}
 \hat{\varepsilon}(h^*) &\leq \varepsilon(h^*) + \gamma \\
 &\leq \varepsilon(h^*) + \gamma \\
 &\leq \hat{\varepsilon}(\hat{h}) + \cancel{\gamma} + \gamma \\
 &\leq \hat{\varepsilon}(\hat{h}) + \cancel{\gamma} + 2\gamma \\
 \text{w/ prob } &\geq 1 - \delta \\
 \Rightarrow & \hat{\varepsilon}(h_0) - \hat{\varepsilon}(\hat{h}) \leq 2\gamma \quad (1) \\
 \text{w/ prob } &\geq 1 - \delta
 \end{aligned}$$

$$\hat{\varepsilon}(h_0) < \hat{\varepsilon}(\hat{h}) + \cancel{\gamma} - 2\gamma$$

$$\hat{\varepsilon}(h_0) - \hat{\varepsilon}(\hat{h}) < \gamma - 2\gamma \quad (2)$$

given (1), if ~~(1) holds~~,

$$2\gamma < \gamma - 2\gamma$$

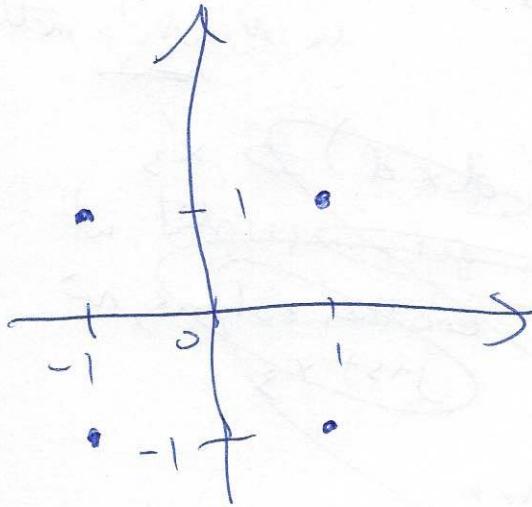
$$\Rightarrow (2) \text{ holds w/ prob } \geq 1 - \delta$$

⑥  
 a) false, in general steps  
 under Newton's gradient step or  
 take longer, but converge to  
 optimum in fewer steps,  
 can't infer which method will  
 converge in less time

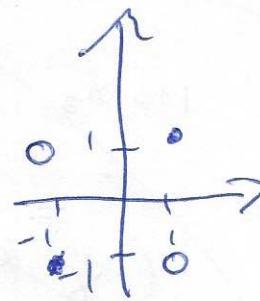
b) false, b/c stochastic  
 gradient descent is  
stochastic, we randomly  
 sample ~~train~~ data for training  
 dataset, ~~some data can be~~  
 different training data will produce  
 different optimal parents

c) if  $K(x_i, y)$  is valid kernel  $\Rightarrow$   
~~if~~  $K$  is symmetric, PSD  $\Rightarrow$  ~~Σ~~  
~~Σ~~  
 $K_{norm} = \frac{K(x, y)}{\sqrt{K(x, x)K(y, y)}}$   $\hat{z}^T K z = \sum_k \left( z^T \phi(x^{(k)}) \right)^2$   
 $\hat{z}^T K_{norm} z = \sum_k \left( \sum_i \hat{z}_i \phi(x^{(i)}) \right)^2$   
 $\hat{z}_i \triangleq \underbrace{z}_{SKLEFT} \geq 0$

(d)



No, counter example



(e) True

if  $h \geq 0$ ,  $\Rightarrow$  ~~dist~~  $h$  is distance, or margin from separating hyperplane

if  $h(x_0) = 0$ , and ~~if~~  $h(x_1) = c$

$$c = h(x_1) - h(x_0) = w(x_1 - x_0) + b$$

(f) We can't tell, b/c VC dimension just tells us size of largest dataset that classifier shatters (can classify all points), but ~~but~~ showing datasets of VC-dimension 3 that are not shattered by classifier

6.9) ~~True~~ ~~Fibonacci~~  
 i)  $y \leftarrow \theta = (x^T x)^{-1} x^T y$   $\theta = (x^T x)^{-1} x^T y = \text{Wentz's method}$   
~~(onestep)~~  
~~linear iteration~~  
~~to Newton's method~~  
 $x_3' = 100x_3$   $(d \times d) \Rightarrow x_3'$   
 $\Rightarrow (x^T x')^{-1}$  gets multiplied w/  
~~other columns, not~~  
~~just  $x_3'$~~   
 ~~$(d \times n) \otimes (n \times d) (d \times n)$~~   
 ~~$(n \times d) (n \times d) (n \times 1)$~~   
 ~~$(d \times n) (n \times d) (d \times n) (n \times 1)$~~   
 ~~$(d \times 1)$~~

6.10) ii)  ~~$\nabla J = *(\theta^T x - y)^*$~~   
 ~~$\theta_0 = [0] \Rightarrow \nabla J_0 = -y^*$~~   
~~false~~  ~~$\nabla(J_0(x)) = -y^*$~~   
 $\Rightarrow \theta_1 = \theta_0 - \alpha \nabla J_0$   
 $\Rightarrow \theta_1^{(+)}) = \theta_0^{(+)}) - 2\nabla J_0^{(+)})$   
 $= +y^*$   
 $\Rightarrow \theta_2^{(+)}) = y^*$   
 $\Rightarrow \theta_2 = 100\theta_2$