Differences in traits predict phenological sensitivity to daylength more than temperature

Deirdre Loughnan[1], Faith A M Jones[1,2], and E M Wolkovich[1]

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[1] Department of Forest and Conservation, Faculty of Forestry, University of British Columbia, 2424 Main Mall Vancouver, BC Canada V6T 1Z4.
[2] Department of Wildlife, Fish and Environmental Studies, Swedish University of Agricultural Sciences, 901 83 Umeå, Sweden.

Corresponding Author: Deirdre Loughnan deirdre.loughnan@ubc.ca

Summary

Introduction

In plants, species variation can in part be explained by differences their strategies of growth, which are generally inferred from traits (Violle et al., 2007). Decades of research on plant traits have worked to build predictive models of species responses to the environment (Green et al., 2022), which could be promising to explain species-level variability in phenological responses. Phenology, however, has generally be excluded from plant trait research, making it difficult to leverage this important framework to explain phenological variation and improve predictions.

Drawing on the relationships between traits and species growth strategies can potentially further our understanding of how plant communities assemble. To date, many studies have focused on community-level responses, using metrics of functional diversity or community weighted means (such as classic work by Grime, 1974). But these methods fail to faaccount for the high degrees of intraspecific variation across traits (Albert et al., 2010) and are limited in their ability to account for the diversity of factors that lead to trait variation, like changes in species abundance across populations (Hawkins et al., 2017; Peres-Neto et al., 2017; Miller et al., 2019). To best understand the processes that define species strategies, we much account for variation at multiple scales—from proximate to ultimate drivers—and use analytical methods that can accurately partition variance across its many sources.

Variation in traits can arise from differences in environmental drivers at both proximate and ultimate scales. Proximate drivers include local short-term environmental factors, such as temperature, soil moisture, and daylength, that can alter observed plant traits, like the rate of plant growth, or leafout phenology (Menzel et al., 2006). But environmental factors at larger spatial and temporal scales may also act as ultimate drivers of variation. In temperate systems, temperature variation across the growing season can produce greater abiotic stress early in the season through high risks of frost events (Sakai and Larcher, 1987; Gotelli and Graves, 1996; Augspurger, 2009). If species avoid this risk by leafing out later, this variation could also produce greater biotic stress later in the season through a more competitive environment. These ultimate drivers may thus determine how species respond to

environmental cues and the effects of biotic interactions to shape species ecological niche and local species assemblages. Together, proximate and ultimate drivers shape a given species' shape their phenotype.

Research on trait ecology has illustrated the ways in which species phenotypes correlate with their growth strategies and responses to environments. The consistency in these relationships across communities allows trait values to provide inferences of growth strategies and community processes independent of species identity (McGill et al., 2006). In plants this has led to the development of frameworks, including the leaf economic spectrum (Wright et al., 2004) and wood economic spectrum (Chave et al., 2009). Under these frameworks, trait values follow distinct gradients that range from acquisitive strategies—fast growing plants that produce cheaper tissue —to conservative strategies— under which plants invest in long-lived tissue but grow at slower rates (Wright et al., 2004; Díaz et al., 2016). These traits serve as proxies for difficult to measure physiological processes and responses to biotic interactions (Shipley et al., 2016), but can still provide considerable insights into the mechanisms shaping plant communities globally.

Phenology itself is considered a trait with high inter- and intra-specific variability. Previous studies have found high variation in phenology for the same species when observed over different years or sites (Primack et al., 2009; Chuine et al., 2010). This variability is often given as justification for excluding phenology from broader functional trait frameworks. But given the importance of phenology in defining species temporal niches and the environmental conditions of growth periods, phenological variation may correlate with other functional traits along the axis of acquisitive to conservative growth strategies.

To understand whether phenology fits within major functional trait frameworks, we must understand the causes of variation within species. Phenological variation is generally observed in natural conditions where considerable differences in the major environmental cues that trigger many phenological events—temperature and daylength (Chuine, 2000; Körner and Basler, 2010)—vary across space and time. However, phenology is often highly predictable when decomposed into these underlying environmental cues (Laube et al., 2014). Within-species variation may also occur across other plant traits (e.g., leaf and wood structure traits), which may vary across latitudinal (Wiemann and Bruce, 2002) or environmental gradients (Pollock et al., 2012), though generally to a smaller scale compared to phenology. While our previous growth chamber study found no evidence that phenological cues varied spatially (Loughnan and Wolkovich, in prep), previous work in functional traits has found other traits to exhibit stronger spatial variation in response to environmental gradients (Gross et al., 2000; Wright et al., 2003). These results suggest understanding how phenology and other traits correlate across species will require methods that incorporate variation within species.

Here, we address this challenge using the phenological event of spring budburst alongside a suite of traits that can capture acquisitive to conservative growth strategies (Wright et al., 2004; Reich, 2014). Our study spans across species—including both shrubs and tree species—of temperate deciduous forests in North America. Seasonal differences in environmental conditions create strong gradients in frost risk, soil nutrients, and light availability, in addition to differences in biotic interactions, from herbivory or competition. These differences between early and late season conditions may select for variation in species growth strategies and as a result, correlate with woody plant traits. We predict that early spring species will budburst before canopy closure and have traits we associate with acquisitive growth—particularly shorter heights, small diameters at breast height (DBH), with lower investment in wood density and leaf tissue (low stem specific density and leaf mass area), but greater photosynthetic potential (high leaf nitrogen content, LNC). In contrast, we predict species with later budburst will predominately include canopy species that express more conservative growth strategies—exemplified by their greater investments in tissue, with taller heights, large DBH, greater SSD and LMA, but low LNC.

We decompose high phenological variation in budburst date by estimating phenological cues using ex-

periments, then combine multiple sources of variation using a joint modeling approach. We estimate the three major phenological cues for woody plant budburst: two temperature cues—chilling (associated with cool winter temperatures) and forcing (associated with warm spring temperatures) and daylength. These cues correlate with early to late phenology with early species (which we predict have acquisitive traits) showing weaker responses to temperature and daylength, while later species (which we predict have conservative traits) showing stronger responses. Using a powerful joint-modeling approach, we estimate the effects of other plant traits on phenological responses to cues, while partitioning the variance from species- and population-level differences. This approach represents a considerable advance in how we estimate trait-cue relationships, and may allow us to better estimate species phenological responses from available trait data.

Materials and Methods

Field sampling

We combined *in situ* trait data with budburst data from two growth chamber cutting experiments conducted across eastern and western temperate deciduous forests in North America. Both suites of data were collected from populations that span a latitudinal gradient of 4-6° for the eastern and western communities respectively. Our trait measurements were taken across eight populations, of which there were four eastern populations—Harvard Forest, Massachusetts, USA (42.55°N, 72.20°W), White Mountains, New Hampshire, USA (44.11°N, 52.14°W)), Second College Grant, New Hampshire, USA (44.79°N, 50.66°W), and St. Hippolyte, Quebec, Canada (45.98°N, 74.01°W), and four western population—E.C. Manning Park (49.06°N, 120.78°W), Sun Peaks (50.88°N, 119.89°W), Alex Fraser Research Forest (52.14°N, 122.14°W), and Smithers (54.78°N, 127.17°W), British Columbia (BC), Canada (Fig. 1). For the growth chamber studies on budburst phenology, cuttings were collected from the most southern and northern populations in each transect (n_{pop}=4).

Across all eight populations, we measured a diverse assemblage of species from the understory and canopy layers. We selected the angiosperm species that were most abundant, with an aim to maximize the number of closely related species and congeners between our eastern and western communities. We focused on angiosperm species only, as they are more likely to have similar environmental controls to their leafout phenology, and to standardize our leaf trait comparisons by excluding gymnosperms.

Functional traits

We measured all traits in the summer prior to each growth chamber study. For our eastern transect, traits were measured from 8-25 June 2015, and from 29 May to 30 July 2019 for our western transect, but before budset. At each population, we measured a total of five traits: height, diameter at breast height (DBH), stem specific density (SSD), leaf mass area (LMA), and the percent leaf nitrogen content (LNC). Each trait was measured for each species present at a population, resulting in us measuring 1-10 healthy adult individuals per species at each population. We also used the WSL xylem database (Schweingruber and Landolt, 2010) to collect data on species ring-type for the 72.3% of our species represented in the database.

We measured traits in accordance to the methods discussed by Pérez-Harguindeguy et al. (2013). We calculated tree height using trigonometric methods, using a TruePulse 200L rangefinder and measured DBH 1.37 m from the ground. For shrub heights, we measured the distance from the ground to the height of the top foliage and measured stem diameters at approximately 1 cm above ground-level. To measure stem specific density, we collected a 10 cm sample of branch wood, taken close to the base of the stem. All samples were kept cool during transport and measurements of stem volume taken

139 within 12 hours of sample collection using the water displacement method. We dried our stem samples
140 upon returning from the field at 105°C for 24h, and weighed their dry mass. Stem specific density was
141 calculated as the dry mass of the sample over its fresh volume.
142
143 For our two leaf traits, we haphazardly selected five, fully expanded, and hardened leaves, avoiding
144 leaves with considerable herbivory damage. We kept leaves cool during sampling and transport. For
145 each leaf, we took high resolution scans using a Canon flatbed scanner (CanoScan Lide 220) within 12
146 hours of collection, and estimated leaf area using the ImageJ software (version 2.0.0). Upon returning
147 from the field, we dried our leaves for 48 h at 70°C and weighed each leaf using a precision balance.
148 Leaf mass area was calculated as the ratio of the leaf dry mass over its area. To measure the percent
149 leaf nitrogen content, leaf samples were ground to a find powder and encapsulated for combustion
150 elemental analysis.
151

Growth chamber study

153 For our growth chamber study, we collected branch cuttings from our highest and lowest latitude pop-
154 ulations in each transect. In both our eastern and western controlled environment study, we included
155 two temperature treatments and daylength, for a total of eight distinct treatments. Our treatments
156 included two levels of chilling—with our eastern study having no additional chilling or 30 days at 4°C,
157 and our western study 30 days or 70 days of chilling at 4°C, with all non-field chilling occurring under
158 dark conditions. Our forcing treatments included either a cool regime of 15:5°C or a warm regime of
159 20:10°C, and our photoperiod treatments consisted of of either 8 or 12 hour daylengths. We recorded
160 budburst stages of each sample every 1-3 days for up to four months, defining the day of budburst
161 as the day of budbreak or shoot elongation (defined as code 07 by (Finn et al., 2007)). For a more
162 detailed discussion of study sample collection and methods see Loughnan and Wolkovich (in prep).
163

Statistical Anlaysis

165 Our analysis combined our *in situ* trait data with budburst data from the controlled environment. For
166 each trait, we developed a joint Bayesian model, in which the relationship between traits and cues
167 is used to estimate budburst. This statistical approach improves upon previous analyses of multiple
168 traits, as it allows us to carry through uncertainty between trait and phenology data—and is better
169 at partitioning the drivers of variation in species phenologies
170
171 Our joint model consists of two parts. The first is a hierarchical linear model, in which we have parti-
172 tioned the variation of individual observations (i) of a given trait value (Y_{trait}) to account for the effects
173 of species (j), population-level differences arising from transects (t), as well as the interaction between
174 transects and latitude (*transect × latitude*), and finally, residual variation or 'measurement error' (σ_{m}^2).
175

$$Y_{\text{trait}_{i,j}} \sim \mathcal{N}(\mu_{i,j}, \sigma_{\text{m}}^2) \tag{1}$$

$$\mu_{i,j} = \alpha_{\text{grand trait}} + \alpha_{\text{trait}_j} + \beta_{\text{transect}_t} \times transect + \beta_{\text{transect} \times \text{latitude}} \times (transect \cdot latitude) \tag{2}$$

$$\tag{3}$$

$$\boldsymbol{\alpha}_{\text{trait}} = \{\alpha_{\text{trait}_1}, ..., \alpha_{\text{trait}_n}\}^T \text{ such that } \boldsymbol{\alpha}_{\text{trait}} \sim \mathcal{N}(0, \sigma_{\text{trait}}^2) \tag{4}$$

$$\tag{5}$$

176 Transect was included as a categorical variable and latitude as a continuous variable in our model.
177 Traits were modeled using natural units, with the exception of LMA, which was rescaled by 100 for
178 numeric stability in the model. Each of our traits collected *in situ* were modeled individually, but
179 comparisons across species ring-porosity were made using the posterior estimates of our height model.
180 This allowed us to best account for inherent differences in wood anatomy across species and growth
181 form.

182

183 We used partial pooling to uniquely estimate species-level variance ($\alpha_{\text{sp,k}}$)—which controls for variation
184 in the number of trait estimates per species and trait variability—which are then used as predictors of
185 species-level estimates of each cue ($\beta_{\text{chill,k}}$, $\beta_{\text{force,k}}$, $\beta_{\text{photo,k}}$).

186

$$\beta_{\text{chill}_k} = \alpha_{\text{chill},k} + \beta_{\text{trait.chill}} \cdot \alpha_{\text{trait},k} \tag{6}$$
$$\beta_{\text{force}_k} = \alpha_{\text{force},k} + \beta_{\text{trait.force}} \cdot \alpha_{\text{trait},k}$$
$$\beta_{\text{photo}_k} = \alpha_{\text{photo},k} + \beta_{\text{trait.photo}} \cdot \alpha_{\text{trait},k}$$

187 In addition to the species-level estimates, the second part of our model estimates the overall effect of
188 each trait on each cue ($\beta_{\text{trait.chill}}$, $\beta_{\text{trait.force}}$, $\beta_{\text{trait.photo}}$). From this we can estimate how well traits
189 explain species-level differences—by estimating the the species-level cue variation not explained by
190 traits ($\alpha_{\text{chill,k}}$, $\alpha_{\text{force,k}}$, $\alpha_{\text{photo,k}}$) and individual species responses to cues (c_g, f_g, p_g, respectively).
191 Finally, our model estimates the residual budburst variation across species ($Y_{\text{pheno,k}}$), observations
192 (σ_{do}^2), as well as the variation in cues not attributed to the trait (using partial pooling).

$$Y_{\text{pheno}_{k,g}} \sim \mathcal{N}(\mu_{k,g}, \sigma_{\text{d}}^2) \tag{7}$$

193 with

$$\mu_{k,g} = \alpha_{\text{pheno},k} + \beta_{\text{chill}_k} \cdot c_g + \beta_{\text{force}_k} \cdot f_g + \beta_{\text{photo}_k} \cdot p_g \tag{8}$$

194 where α_{pheno_k}, α_{chill_k}, α_{force_k}, and α_{photo_k} are elements of the normal random vectors:

$$\boldsymbol{\alpha}_{\text{pheno}} = \{\alpha_{\text{pheno}_1}, ..., \alpha_{\text{pheno}_n}\}^T \text{ such that } \boldsymbol{\alpha}_{\text{pheno}} \sim \mathcal{N}(\mu_{\text{pheno}}, \sigma_{\text{pheno}}^2) \tag{9}$$
$$\boldsymbol{\alpha}_{\text{chill}} = \{\alpha_{\text{chill}_1}, ..., \alpha_{\text{chill}_n}\}^T \text{ such that } \boldsymbol{\alpha}_{\text{chill}} \sim \mathcal{N}(\mu_{\text{chill}}, \sigma_{\text{chill}}^2)$$
$$\boldsymbol{\alpha}_{\text{force}} = \{\alpha_{\text{force}_1}, ..., \alpha_{\text{force}_n}\}^T \text{ such that } \boldsymbol{\alpha}_{\text{force}} \sim \mathcal{N}(\mu_{\text{force}}, \sigma_{\text{force}}^2)$$
$$\boldsymbol{\alpha}_{\text{photo}} = \{\alpha_{\text{photo}_1}, ..., \alpha_{\text{photo}_n}\}^T \text{ such that } \boldsymbol{\alpha}_{\text{photo}} \sim \mathcal{N}(\mu_{\text{photo}}, \sigma_{\text{photo}}^2)$$

195 We included chilling as chill portions, and each environmental cue as continuous variables. We stan-
196 dardized all cues by z-scoring them, allowing us to make comparisons across cues (see Loughnan and
197 Wolkovich (in prep) for more details).

198

199 For our model, we used weakly informative priors unique to each trait. We validated our choice of priors
200 using prior predictive checks and confirmed model stability under wider priors. All models were coded
201 in the Stan programming language for Bayesian models using the rstan package (Stan Development
202 Team, 2018) in R version 4.3.1 (R Development Core Team, 2017). All models met basic diagnostic
203 checks, including no divergences, high effective sample sizes (n_{eff}) that exceeded 10% of the number
204 of iterations, and $\hat{R}$ values close to 1. Our model estimates are reported as the mean values, with the
205 90% uncertainty interval. Estimates are given to one decimal place, as this best represents the level of
206 precision we can infer from our data (Gelman et al., 2020).

207

Results

Across our eight populations, we measured 47 species of which 28 were in our eastern transect and 22 in our western transect. These include species dominant in both the understory and canopy layer, with our eastern community consisting of 13 shrubs and 15 trees, our western community consisting of 18 shrubs and 4 trees, and three species that occurred in both transects. In total we measured traits of 1428 unique individuals between the two transects across our five traits: height ($n = 1317$), DBH ($n = 1220$), SSD ($n = 1359$), LMA ($n = 1345$), LNC ($n = 1351$). Across our two growth chamber studies, we made observations of 4211 samples, with our observations of budburst spanning 82 and 113 days for our eastern and western studies respectfully.

Most of our traits showed negligible variation across populations, both in terms of differences between the two transects (main effect of transect only) or by latitude within each transect (an interactive effect between transect and latitude). Only diameter at breast height (DBH) (0.7, UI: 0.3,1, Table 2) and leaf mass area (LMA) differed by latitude within transects (0.3, UI: 0.2,0.4, Table 4)—with the DBH and LMA of eastern species increasing with latitude—but not western species (Fig. 2 d). Leaf nitrogen content (LNC) and structural traits, however, showed no differences across populations or transects (Table 1, 3, and 5). The differences we found across populations were small to negligible, especially in comparison to species-level differences, which varied considerably and up to 33 fold (Fig. 3).

Only a subset of our traits related to at least one budburst cue, but responses were generally weak, with the exception of trait-specific relationships to individual cues. We found similar cue relationships for our two structural traits, with tall plants with larger DBH exhibiting stronger photoperiod responses and producing earlier estimates of budburst under longer daylengths (-1.7, UI: -2.9, -0.5 for height and -2.2, UI: -3.3, -1 for DBH). But instead chilling and forcing cues were only related to SSD (Table 3). But we found no relationships between cues and the different types of wood porosity (Fig. 4). Of our two leaf traits, both LMA and LNC showed a relationship with photoperiod, with low LMA species advancing their budburst more in response to longer photoperiods compared to high LMA species (-7.0, UI: -10.5, -3.6). In contrast species with low LNC were delayed in their budburst under longer photoperiods relative to high LNC species (0, UI: -3.1, 3).

How our three cues shaped budburst timing also varied across our traits. For most of our trait models chilling, followed by forcing, were the strongest cues of budburst (Fig. 5). But in accounting for the correlations between LMA and cues, our model estimated similar responses to photoperiod (-7.0, UI: -10.5, -3.6) than chilling (-7.8, UI: -20.2, 4.1), and a relatively weak effect of forcing (90% UI crossing zero) (Table 4). In contrast, our LNC (Table 5) and SSD models (Table 3) both showed negligible relationships between photoperiod and budburst timing (90% UI crossing zero). These findings suggest that relationships with additional traits can alter the effects and relative importance of cues on budburst (Fig. 6). This—paired with strong species-level variation—may cause variation in the estimated order of species relative budburst dates across the different models.

In synthesizing the effects of multiple traits across species, our results can be used to make generalizations across ecologically important groups of species. But only some of our models estimated clear gradients in species timing between trees and shrubs (Fig. 7). In particular, we found height to have strong correlations between budburst timing and trait values, with earlier estimates of budbust for shrubs—especially under stronger cues—and later budburst estimates for trees (Fig. 7). Diameter at breast height showed similar trends as estimates from our height model (results not shown). But this was not the case for our two leaf traits. Leaf nitrogen content, for example, showed no distinct separation between shrub and tree functional groups (Fig. 7).

Discussion

Of the six traits we studied, we found most traits to relate to one or more phenological cue, but only two of those traits fit within a gradient of acquisitive to conservative growth. Using our joint modeling approach, we were able to estimate how these traits interact with cues to shape species budburst. We found that photoperiod—often the weakest cue to budburst (Laube et al., 2014; Zohner et al., 2016; Flynn and Wolkovich, 2018)—may be the most important cue in trait-phenology relationships. In general, these patterns were consistent across our populations and between our eastern and western transects, despite the differences in latitudes and community composition. Collectively our results provide new insights into the complexity of the mechanisms that underlie the relationships between phenotypes and environmental cues, while also challenging our existing understanding of these processes.

Across our eight forest populations, only two of our traits varied between our transects or with latitude. Western sites exhibited lower LMA and DBH, expressing a more acquisitive growth strategies on average. This may be result of the greater number of shrub species in these forest communities. This observed difference provides some support of our predictions that our two transects would differ in their resource partitioning and nutrient cycles based on the traits of their dominant species, with our shrub dominated Western community generally assumed to utilize resources early in the season prior to canopy closure.

While the strong differences in our community assemblages explained some of the trait variation we observed spatially, at the species-level the relationships between traits and budburst cues only partially supported our predictions for how phenology relates to species overall growth strategies. We predicted species with acquisitive traits—particularly small trees with low stem densities, and leaves with low LMA, and high LNC—to have early budburst via weak temperature and daylength responses. But we found most traits to only correlate with the effects of daylength on budburst. As predicted, conservative species with greater heights and DBH did have stronger photoperiod responses (associated with later budburst), but—contrary to our prediction—conservative species with high LMA and low LNC also showed weak responses to photoperiod. These results suggest that phenology is only partially aligns with trends found in established trait frameworks, but also offer new insight into potential tradeoffs in how varying physiological processes shape species temporal niches.

Despite our study spanning forest communities across North America, only specific traits varied spatially. This was surprising, as we expected the inherent variation in temperature and daylengths to impose population differences in traits. Other research using the similar plant traits have also found little variation in trait values across similar regional scales (Messier et al., 2010; Kang et al., 2013; Standen and Baltzer, 2023). This may be in part because most of our predictions for how traits pertain to species growth strategies are based on the synthesis of globally distributed species (Wright et al., 2004; Díaz et al., 2016). Our findings suggest the importance of studying ecological processes across spatial scales, as the ultimate and proximate drivers selecting species phenotypes appear to vary considerably.

In comparing our results with a global meta-analysis of tree trait relationships with budburst cues (Loughnan et al., in prep), we found similar trait-cue relationships for several of our traits. In addition to also finding relationships between tree height and photoperiod, we found similar responses in leaf traits, with high specific leaf area (the inverse of LMA) exhibiting strong responses to photoperiod (Loughnan et al., in prep). The consistency of this result, across these two spatial scales of study, suggest there are alternate underlying mechanisms shaping how species respond to photoperiod cues. It is possible that the unexpected trends we observed in our results are due to selection on other physiological processes, as many of our traits are associated with one or more ecological function (Wright et al., 2004; Pérez-Harguindeguy et al., 2013; Reich, 2014). The trends we observed in LMA, may be the result of other processes than light capture, such as tissue lifespan or decomposition rates (De La Riva

et al., 2016).

How traits shape species temporal niches

Individuals temporal niches are shaped by numerous—and complex—interactions with local environmental conditions and species assemblages. But studies that focus only on phenology fail to account for interactions between other traits and cues that also contribute to species temporal niches. In focusing solely on phenology, previous research has shown budburst timing is primarily driven by temperature and only weakly by photoperiod (Chuine et al., 2010; Basler and Körner, 2014; Laube et al., 2014). We, however, found only one trait to correlate with responses to temperature, suggesting other cues are impacting leaf and structural traits in temperate forests. Traits like leaf mass area can vary with soil moisture, with variation in leaf area allowing plants to reduce evaporation, selecting for high LMA under dry conditions (De La Riva et al., 2016). As a cue, soil moisture and changes in water use are known to shape other phenological events, including radial growth phenology and shoot elongation in woody plants (Cabon et al., 2020; Peters et al., 2021). If selection by soil moisture is shaping phenological responses, it may be contributing to the unexpected trends we observed in LMA and the absence of relationships with temperature. Thus, fully understanding how species growth strategies correlate with phenology may require a more holistic consideration of cues known to shape other plant traits.

Using functional traits to predict climate change responses

Our results offer novel insights into how broader correlations between plant trait syndromes and phenological cues can help predict phenological responses in plant communities with climate change. As temperatures rise, particularly at higher latitudes (Hoegh-Guldberg et al., 2018), these cues will become stronger and potentially select for earlier growth in some species, but photoperiod cues will remain fixed. This suggests the relationships between photoperiod and other traits have the potential to limit species abilities to track temperatures and constrain the extent to which woody plant phenologies will advance with climate change. Our results suggest that these effects will likely be stronger on larger species or canopy trees and species with relatively low LMA. These constraints could have cascading effects on forest communities, as variable species responses to increasing temperatures further alter species temporal niches and their interactions within communities.

Further, our findings of correlations between phenology and other commonly measured traits highlight that accurate forecasts of future changes in phenology may benefit from accounting for the response of other traits to climate change. Across temperature and precipitation gradients, leaf size and shape, for example, also change, as species shift to conserve water and mitigate effects of transpiration under higher temperatures (De La Riva et al., 2016). These changes could impact species photosynthetic potential and ultimately ecosystem services, such as carbon sequestration. While phenological research has focused on forecasting responses to temperature, the correlation of other traits with photoperiod suggest its importance as a cue. It is therefore critical that we continue to take a more holistic approach to studying the relationships between phenology and plant traits to accurately forecast the future impacts of climate change on communities.

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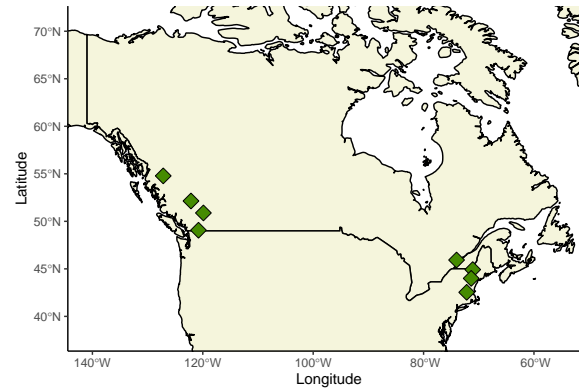
Figures

Figure 1: We measured leaf and structural traits in eight temperate deciduous forests, spanning four eastern populations and four western population, across a latitudinal gradients of 4-6°. The branch clippings used in our two growth chamber experiments were taken from the most northern and most southern populations in each transect.
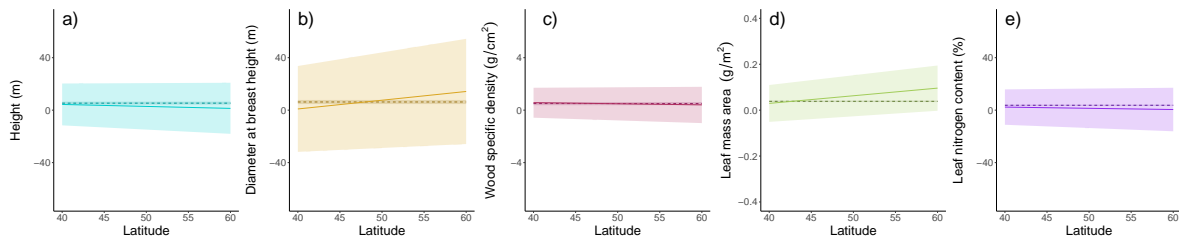

Figure 2: We found geographic differences for only one of our functional traits, the majority exhibiting no differences between latitudes or across transects. Depicted are the spatial trends for, a. height, b. diameter at base height, c. stem specific density, d. leaf mass area, and e. leaf nitrogen content. Dashed lines represent the western transect and solid lines the eastern transect.
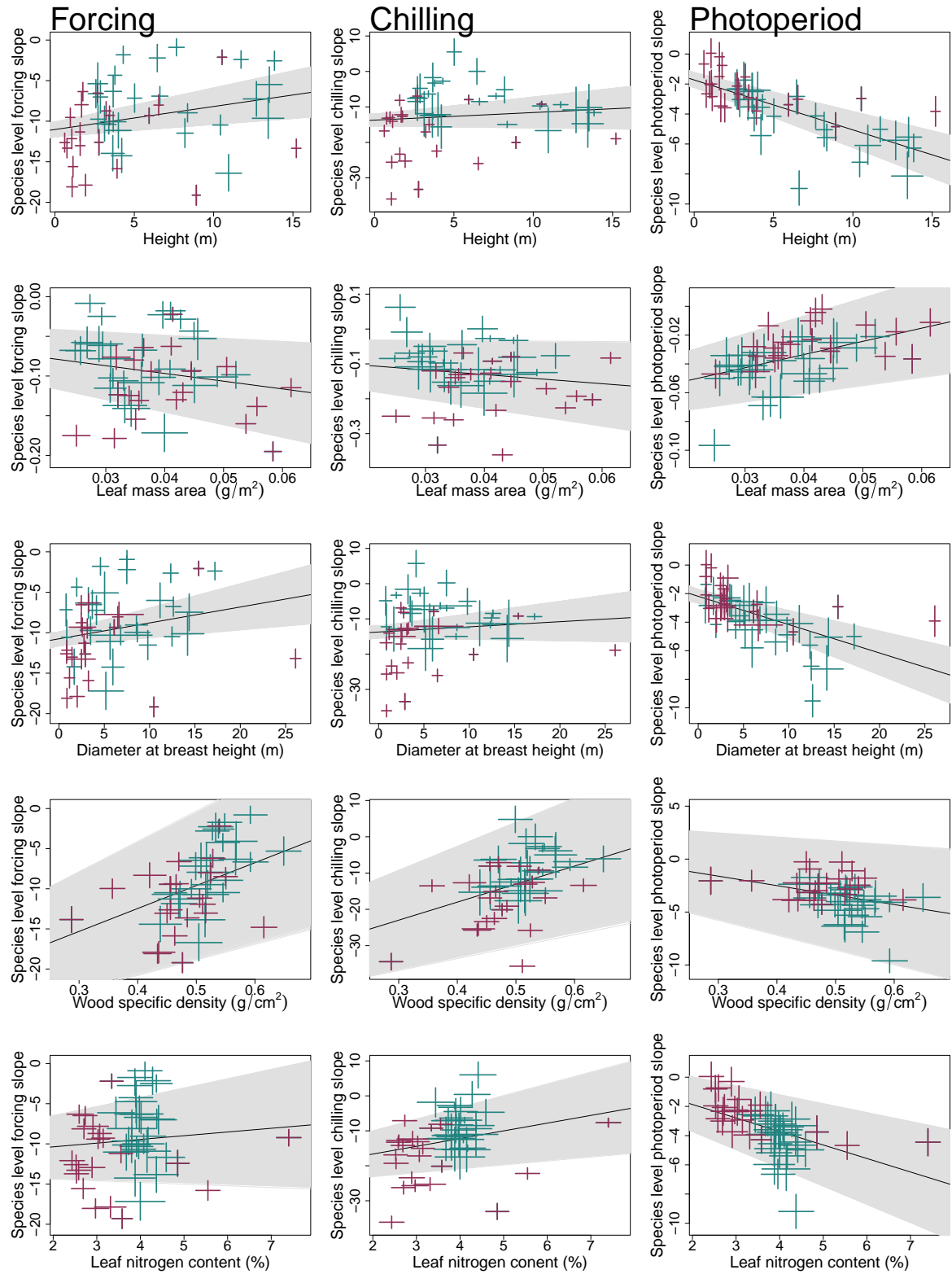
Figure 3: Relationships between species traits and cue responses showed considerable variation across a-c. height, d-f. leaf mass area, g-i. diameter at breast height, j-l. stem specific density, and m-o. the leaf nitrogen content. Point colours representing different species groups, with tree species are depicted in maroon and shrub species in teal. The blue band depicts the 90% uncertainty interval.
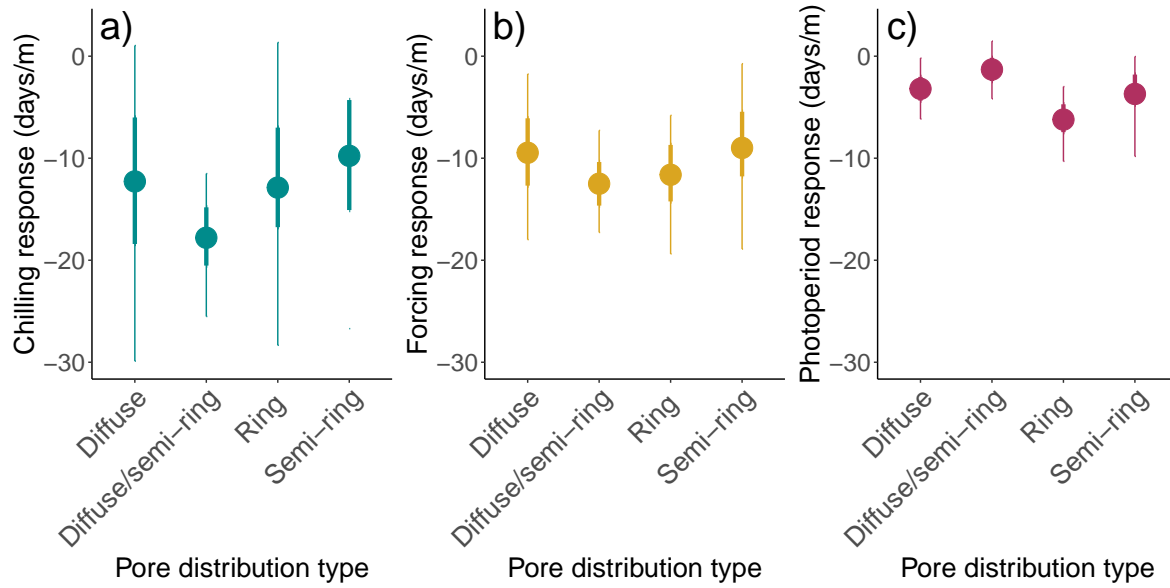
Figure 4: Despite species wood structures causing differing growth strategies in the spring, we did not find this trait to correlate with differences in cues across species. Thinner lines represent the 90% UI and thicker lines the 50% UI. Here we show the correlation using the posterior estimates from the height model only, but trends were consistent across all traits (Table 2-5).
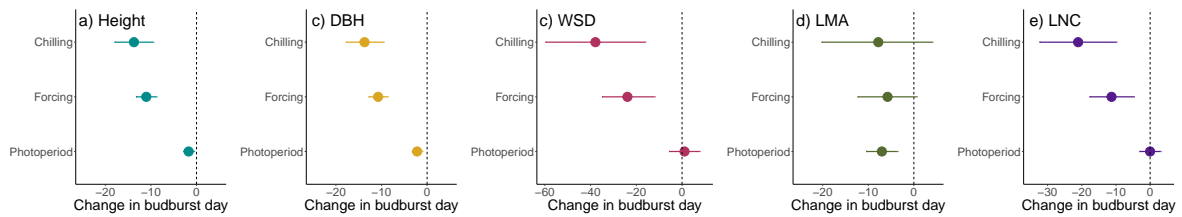

Figure 5: We found fairly consistent estimates for budburst cue responses to chilling, forcing, and photoperiod for each of our trait models: a. height, b. diameter at breast height, c. stem specific density, d. leaf mass area, and e. leaf nitrogen contnet. Lines represent 90% uncertainty intervals and points the mean estimate.
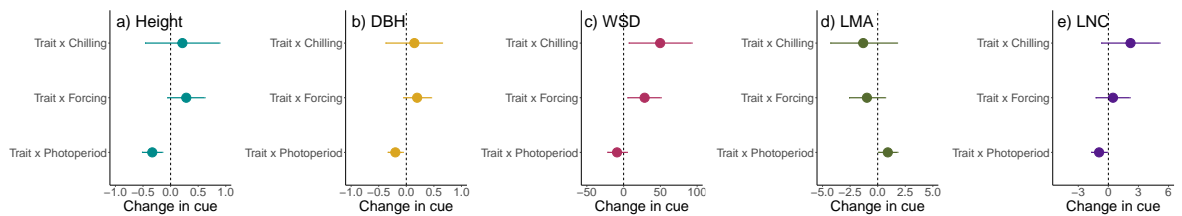

Figure 6: The relationships between traits and cue responses varied considerably across each of our trait models, a. height, b. diameter at breast height, c. stem specific density, d. leaf mass area, and e. leaf nitrogen content, and for individual cues. Lines represent 90% uncertainty intervals and points the mean estimate. Note the differences in the scale of the x-axis.
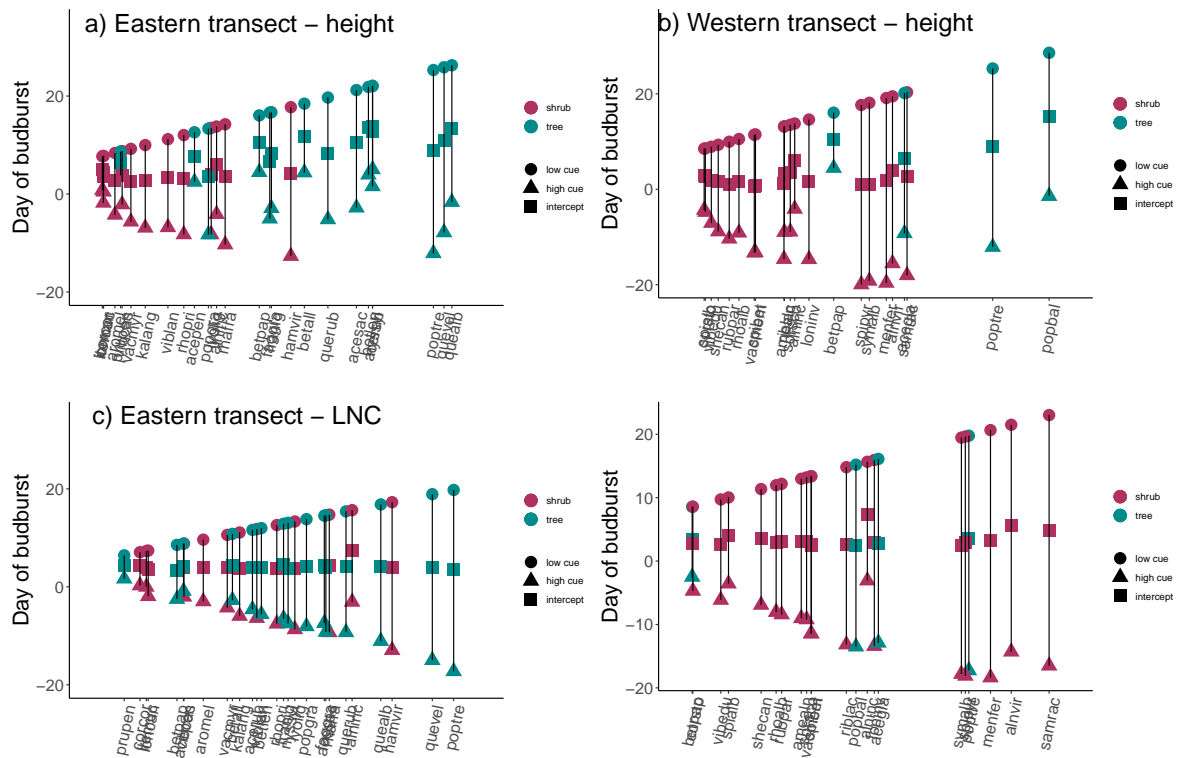
Figure 7: We found budburst estimates differed between our full model (intercept plus cues, depicted as triangles for high cues and as circles for low cues), versus the intercepts only model (without cues, shown as squares). Species are ordered in increasing budburst dates for both the eastern (a & c) and western (b & d) populations, spanning from early budbursting shrubs, in red, to late budbursting trees in blue. For traits such as height (a & b) we found distinct partitioning of budburst across shrub and tree species, but this was not the case for all traits, with our model of leaf nitrogen content showing highly mixed budburst order of shrub and tree species (c & d).

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Table 1: Summary output from a joint Bayesian model of height and budburst phenology in which species are partially pooled. The effect of transect is modeled as a categorical variable and latitude as continuous in its interaction term with transect. The model includes environmental cues as z-scored continuous variables, allowing comparisons to be made across cues.

	mean	5%	25%	75%	95%
Transect	5.00	-2.90	1.80	8.20	12.70
Transect x latitude	-0.10	-0.30	-0.20	-0.10	0.00
Forcing	-11.00	-13.30	-11.90	-10.10	-8.80
Chilling	-13.70	-18.00	-15.40	-12.00	-9.50
Photoperiod	-1.70	-2.90	-2.20	-1.20	-0.50
Trait x forcing	0.30	-0.10	0.10	0.40	0.60
Trait x chilling	0.20	-0.50	-0.10	0.50	0.90
Trait x photoperiod	-0.30	-0.50	-0.40	-0.30	-0.20

Table 2: Summary output from a joint Bayesian model of DBH and budburst phenology in which species are partially pooled. The effect of transect is modeled as a categorical variable and latitude as continuous in its interaction term with transect. The model includes environmental cues as z-scored continuous variables, allowing comparisons to be made across cues

	mean	5%	25%	75%	95%
Transect	-32.00	-48.40	-38.70	-25.30	-15.60
Transect x latitude	0.70	0.30	0.50	0.80	1.00
Forcing	-10.70	-12.80	-11.60	-9.90	-8.60
Chilling	-13.70	-17.80	-15.40	-12.10	-9.50
Photoperiod	-2.20	-3.30	-2.70	-1.70	-1.00
Trait x forcing	0.20	-0.10	0.10	0.30	0.50
Trait x chilling	0.10	-0.40	-0.10	0.40	0.60
Trait x photoperiod	-0.20	-0.30	-0.30	-0.10	-0.10

Table 3: Summary output from a joint Bayesian model of SSD and budburst phenology in which species are partially pooled. The effect of transect is modeled as a categorical variable and latitude as continuous in its interaction term with transect. The model includes environmental cues as z-scored continuous variables, allowing comparisons to be made across cues.

	mean	5%	25%	75%	95%
Transect	0.40	-0.10	0.20	0.60	1.00
Transect x latitude	-0.00	-0.00	-0.00	-0.00	0.00
Forcing	-23.00	-34.40	-27.70	-18.40	-11.50
Chilling	-36.20	-57.20	-45.30	-27.60	-13.80
Photoperiod	1.20	-5.50	-1.50	3.90	7.30
Trait x forcing	26.90	4.10	17.80	36.20	49.10
Trait x chilling	46.40	2.80	29.70	64.10	87.30
Trait x photoperiod	-9.10	-21.40	-14.50	-4.00	4.10

Table 4: Summary output from a joint Bayesian model of LMA and budburst phenology in which species are partially pooled. The effect of transect is modeled as a categorical variable and latitude as continuous in its interaction term with transect. The model includes environmental cues as z-scored continuous variables, allowing comparisons to be made across cues.

	mean	5%	25%	75%	95%
Transect	-14.20	-18.30	-15.90	-12.50	-10.20
Transect x latitude	0.30	0.20	0.30	0.40	0.40
Forcing	-5.80	-12.30	-8.30	-3.30	0.60
Chilling	-7.80	-20.20	-12.60	-2.80	4.10
Photoperiod	-7.00	-10.50	-8.40	-5.60	-3.60
Trait x forcing	-1.00	-2.60	-1.60	-0.30	0.70
Trait x chilling	-1.30	-4.30	-2.60	-0.10	1.80
Trait x photoperiod	0.90	0.10	0.60	1.30	1.80

Table 5: Summary output from a joint Bayesian model of LNC and budburst phenology in which species are partially pooled. The effect of transect is modeled as a categorical variable and latitude as continuous in its interaction term with transect. The model includes environmental cues as z-scored continuous variables, allowing comparisons to be made across cues.

	mean	5%	25%	75%	95%
Transect	2.30	-4.50	-0.60	5.10	9.20
Transect x latitude	-0.10	-0.20	-0.20	-0.00	0.10
Forcing	-11.30	-17.70	-14.00	-8.70	-4.80
Chilling	-21.10	-32.30	-25.40	-16.70	-9.90
Photoperiod	-0.00	-3.10	-1.30	1.30	3.00
Trait x forcing	0.50	-1.30	-0.20	1.20	2.10
Trait x chilling	2.20	-0.70	1.10	3.40	5.10
Trait x photoperiod	-0.90	-1.70	-1.30	-0.60	-0.10