

Homework 2

DATA130021 Financial Econometrics

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Problem 1

The unit of yearly log returns is *year*. We have the definition as

$$r_t = \log(1 + R_t) = \log P_t - \log P_{t-1} \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mu, \sigma^2),$$

where $\mu = 0.08$ and $\sigma = 0.15$. From the assumption, the price P_0 is \$80 today. Then we have

$$\begin{aligned} P(P_1 \geq 90) &= P\left(\log\left(\frac{P_1}{P_0}\right) \geq \log\left(\frac{90}{80}\right)\right) \\ &= P(r_1 \geq 0.11778) \\ &= P\left(\frac{r_1 - 0.08}{0.15} \geq \frac{0.11778 - 0.08}{0.15}\right) \\ &= 1 - \Phi(0.2519) \\ &\approx 0.4006. \end{aligned}$$

Hence, the probability that the price one year from now is no less than \$90 is approximately 40.06%.

Problem 2

Definitions of simple returns and log returns are

$$R_t = \frac{P_t + D_t}{P_{t-1}} - 1,$$

$$r_t = \log(1 + R_t) = \log\left(\frac{P_t + D_t}{P_{t-1}}\right),$$

and multiple-period gross returns are products of single-period gross returns so that

$$1 + R_t(k) = \prod_{i=t-k+1}^t (1 + R_i) = \prod_{i=t-k+1}^t \frac{P_i + D_i}{P_{i-1}}.$$

Then we have the k -period returns and log returns as

$$R_t(k) = -1 + \prod_{i=t-k+1}^t (1 + R_i) = -1 + \prod_{i=t-k+1}^t \frac{P_i + D_i}{P_{i-1}},$$

$$r_t(k) = \log(1 + R_t(k)) = \sum_{i=t-k+1}^t \log\left(\frac{P_i + D_i}{P_{i-1}}\right) = \sum_{i=t-k+1}^t r_i.$$

a. $R_3(2)$

$$\begin{aligned}
R_3(2) &= (1 + R_2)(1 + R_3) - 1 \\
&= \frac{P_2 + D_2}{P_1} \cdot \frac{P_3 + D_3}{P_2} - 1 \\
&= \frac{85 + 0.1}{82} \cdot \frac{83 + 0.1}{85} - 1 \\
&\approx 0.0146.
\end{aligned}$$

b. $r_4(3)$

$$\begin{aligned}
r_4(3) &= r_2 + r_3 + r_4 \\
&= \log(1 + R_2) + \log(1 + R_3) + \log(1 + R_4) \\
&= \log\left(\frac{P_2 + D_2}{P_1}\right) + \log\left(\frac{P_3 + D_3}{P_2}\right) + \log\left(\frac{P_4 + D_4}{P_3}\right) \\
&= \log\left(\frac{85 + 0.1}{82}\right) + \log\left(\frac{83 + 0.1}{85}\right) + \log\left(\frac{87 + 0.125}{83}\right) \\
&\approx 0.0630.
\end{aligned}$$

Problem 3

a. From the assumption, variables $r_1, r_2, \dots \stackrel{\text{i.i.d}}{\sim} \mathcal{N}(\mu, \sigma^2)$. Then

$$\log\left(\frac{X_k}{X_0}\right) = \log X_k - \log X_0 = \sum_{i=1}^k r_i \sim \mathcal{N}(k\mu, k\sigma^2)$$

Hence, we have

$$\begin{aligned}
P(X_2 > 1.3X_0) &= P\left(\log\left(\frac{X_2}{X_0}\right) > \log(1.3)\right) \\
&= P(r_1 + r_2 > 0.26236) \\
&= P\left(\frac{r_1 + r_2 - 2\mu}{\sqrt{2}\sigma} > \frac{0.26236 - 2\mu}{\sqrt{2}\sigma}\right) \\
&= 1 - \Phi\left(\frac{0.26236 - 2\mu}{\sqrt{2}\sigma}\right).
\end{aligned}$$

b. The relationship between variables X_1 and r_1 is

$$X_1 = X_0 e^{r_1} \Leftrightarrow h(X_1) = r_1 = \log X_1 - \log X_0 \sim \mathcal{N}(\mu, \sigma^2),$$

and the density of r_1 is

$$f_{r_1}(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}, \quad x \in (-\infty, +\infty).$$

Apply formula (A.4) and we get that the density of X_1 is

$$\begin{aligned}
f_{X_1}(y) &= f_{r_1}(h(y)) |h'(y)| \\
&= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(h(y)-\mu)^2} \cdot \frac{1}{y} \\
&= \frac{1}{\sqrt{2\pi}\sigma y} e^{-\frac{1}{2\sigma^2}(\log y - \log X_0 - \mu)^2}, \quad y \in (X_0, +\infty).
\end{aligned}$$

c. From problem 3.a, we know that

$$\log\left(\frac{X_k}{X_0}\right) = \log X_k - \log X_0 = \sum_{i=1}^k r_i \sim \mathcal{N}(k\mu, k\sigma^2).$$

Then

$$h(X_k) = Z = \frac{\log\left(\frac{X_k}{X_0}\right) - k\mu}{\sqrt{k}\sigma} \sim \mathcal{N}(0, 1),$$

Since $h(\cdot)$ is strictly increasing, we can get its inverse function $g(\cdot)$

$$g(Z) = X_k = X_0 e^{\sqrt{k}\sigma Z + k\mu}.$$

Then we have

$$\begin{aligned} F_{X_k}^{-1}(p) &= g(F_Z^{-1}(p)) \\ &= g(\Phi^{-1}(p)) \\ &= X_0 e^{\sqrt{k}\sigma\Phi^{-1}(p) + k\mu}. \end{aligned}$$

For the 0.9 quantile of X_k for all k , we have the following formula

$$\begin{aligned} F_{X_k}^{-1}(0.9) &= X_0 e^{\sqrt{k}\sigma\Phi^{-1}(0.9) + k\mu} \\ &= X_0 e^{1.28155\sqrt{k}\sigma + k\mu}. \end{aligned}$$

d. From problem 3.c, we know that

$$X_k = X_0 e^{\sqrt{k}\sigma Z + k\mu},$$

where $Z \sim \mathcal{N}(0, 1)$. Recall the moment-generating function of normal distribution

$$M_Z(t) = \mathbb{E}(e^{tZ}) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}.$$

Then the expected value of X_K^2 for any k is

$$\begin{aligned} \mathbb{E}(X_k^2) &= \mathbb{E}\left(X_0^2 e^{2\sqrt{k}\sigma Z + 2k\mu}\right) \\ &= X_0^2 e^{2k\mu} \mathbb{E}\left(e^{2\sqrt{k}\sigma Z}\right) \\ &= X_0^2 e^{2k\mu} M_Z(2\sqrt{k}\sigma) \\ &= X_0^2 e^{2k\mu} \cdot e^{2\sqrt{k}\mu\sigma + 2k\sigma^4} \\ &= X_0^2 e^{2k\mu + 2\sqrt{k}\mu\sigma + 2k\sigma^4}. \end{aligned}$$

e. From problem 3.d, we know that

$$\mathbb{E}(X_k^2) = X_0^2 e^{2k\mu + 2\sqrt{k}\mu\sigma + 2k\sigma^4}.$$

Similarly, we can get that

$$\begin{aligned} \mathbb{E}(X_k) &= \mathbb{E}\left(X_0 e^{\sqrt{k}\sigma Z + k\mu}\right) \\ &= X_0 e^{k\mu} \mathbb{E}\left(e^{\sqrt{k}\sigma Z}\right) \\ &= X_0 e^{k\mu} M_Z(\sqrt{k}\sigma) \\ &= X_0 e^{k\mu} \cdot e^{\sqrt{k}\mu\sigma + \frac{1}{2}k\sigma^4} \\ &= X_0 e^{k\mu + \sqrt{k}\mu\sigma + \frac{1}{2}k\sigma^4}. \end{aligned}$$

Then the variance of X_k for any k is

$$\begin{aligned}
\text{Var}(X_k) &= \mathbb{E}(X_k^2) - \mathbb{E}^2(X_k) \\
&= X_0^2 e^{2k\mu + 2\sqrt{k}\mu\sigma + 2k\sigma^4} - \left(X_0 e^{k\mu + \sqrt{k}\mu\sigma + \frac{1}{2}k\sigma^4}\right)^2 \\
&= X_0^2 e^{2k\mu + 2\sqrt{k}\mu\sigma + 2k\sigma^4} - X_0^2 e^{2k\mu + 2\sqrt{k}\mu\sigma + k\sigma^4} \\
&= X_0^2 (e^{k\sigma^4} - 1) e^{2k\mu + 2\sqrt{k}\mu\sigma + k\sigma^4}.
\end{aligned}$$

Problem 4

From the assumption, daily log returns are normally distributed, that is, $r_1, r_2, \dots \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mu, \sigma^2)$, where $\mu = 0.0005$ and $\sigma^2 = 0.012$. From the definition, we have

$$r_t(t) = \log\left(\frac{P_t}{P_0}\right) \sim \mathcal{N}(t\mu, t\sigma^2).$$

Then

$$\begin{aligned}
\mathbb{P}\left(\frac{P_t}{P_0} \geq 2\right) &\geq 0.9 \Leftrightarrow \mathbb{P}\left(\log\left(\frac{P_t}{P_0}\right) \geq \log(2)\right) \geq 0.9 \\
&\Leftrightarrow \mathbb{P}\left(r_t(t) \geq \log(2)\right) \geq 0.9 \\
&\Leftrightarrow \mathbb{P}\left(\frac{r_t(t) - t\mu}{\sqrt{t}\sigma} \geq \frac{\log(2) - t\mu}{\sqrt{t}\sigma}\right) \geq 0.9 \\
&\Leftrightarrow 1 - \Phi\left(\frac{\log(2) - t\mu}{\sqrt{t}\sigma}\right) \geq 0.9 \\
&\Leftrightarrow \Phi\left(\frac{\log(2) - t\mu}{\sqrt{t}\sigma}\right) \leq 0.1 \\
&\Leftrightarrow \mu t + \sigma \Phi^{-1}(0.1)\sqrt{t} - \log(2) \geq 0 \\
&\Leftrightarrow \mu^2 t^2 - \left[\sigma^2 \Phi^{-1}(0.1)^2 + 2\log(2)\mu\right]t + \log^2(2) \geq 0 \quad (*).
\end{aligned}$$

Solve the quadratic inequality (*) and we get

$$t \geq 81584 \text{ days.}$$

In conclusion, the probability that the price has doubled is at least 90% after 81584 days (about 223.5 years), which means it is almost impossible in reality.

Problem 5

Codes for simulations of log-normal geometric random walk models with $P_0 = 1$, $\mu = \pm 0.01$ and $\sigma = 0.01, 0.05$:

```

library(ggplot2)
set.seed(999) # Set random seed

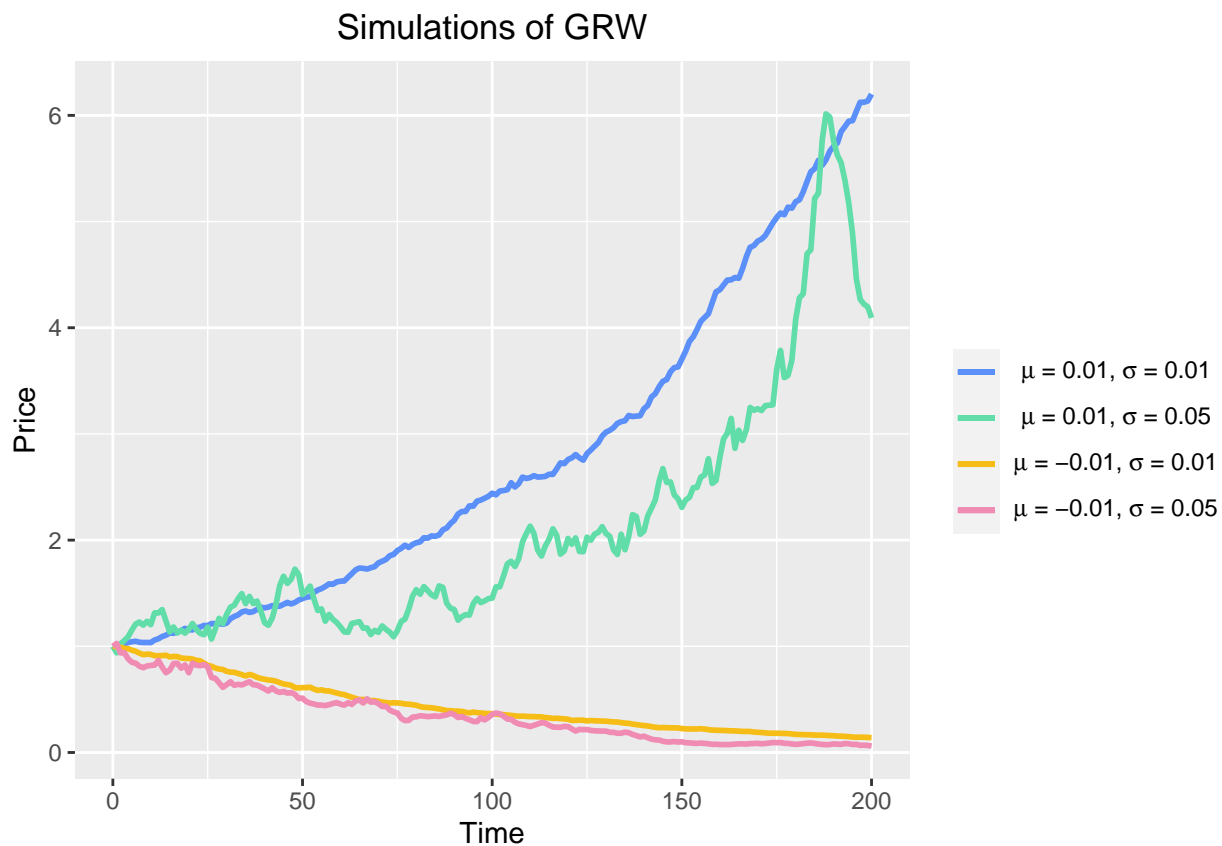
Time <- c(0:200) # Simulates 200 times
Price_1 <- c(1, exp(cumsum(rnorm(200, 0.01, 0.01))))
Price_2 <- c(1, exp(cumsum(rnorm(200, 0.01, 0.05))))
Price_3 <- c(1, exp(cumsum(rnorm(200, -0.01, 0.01))))
Price_4 <- c(1, exp(cumsum(rnorm(200, -0.01, 0.05))))
df <- data.frame(Time, Price_1, Price_2, Price_3, Price_4)

```

```

ggplot() +
  geom_line(data=df, aes(x=Time, y=Price_1, colour="Price_1"), size=1) +
  geom_line(data=df, aes(x=Time, y=Price_2, colour="Price_2"), size=1) +
  geom_line(data=df, aes(x=Time, y=Price_3, colour="Price_3"), size=1) +
  geom_line(data=df, aes(x=Time, y=Price_4, colour="Price_4"), size=1) +
  labs(title="Simulations of GRW", x="Time", y="Price") +
  theme(plot.title=element_text(hjust=0.5)) +
  scale_colour_manual(
    name="",
    values=c(
      "Price_1"="#5B8FF9",
      "Price_2"="#61DDAA",
      "Price_3"="#F6BD16",
      "Price_4"="#F08BB4"
    ),
  ),
  labels=c(
    expression(paste(mu, " = 0.01, ", sigma, " = 0.01 ")),
    expression(paste(mu, " = 0.01, ", sigma, " = 0.05 ")),
    expression(paste(mu, " = -0.01, ", sigma, " = 0.01")),
    expression(paste(mu, " = -0.01, ", sigma, " = 0.05"))
  )
)

```



From the figure above, we find that the chains with $\sigma = 0.05$ have a larger volatility than the ones with $\sigma = 0.01$. Chains with positive drift $\mu = 0.01$ show the increasing trend and chains with negative drift $\mu = -0.01$ seem to converge to zero. These simulation results are consistent with common sense.