

[SP-6]

I. Exponential Distribution and Memoryless

1.

(a). Let X be an r.v. representing offers. Given the problem,

$$X \sim \text{Expo}(\lambda)$$

Since

$$E[X] = \frac{1}{\lambda} = 10000$$

$$\Rightarrow \lambda = 1/10000$$

$$\Rightarrow X \sim \text{Expo}\left(\frac{1}{10000}\right)$$

Let Y be an r.v. representing # offers Fred will have. Then, by problem,

$$Y \sim \text{Geom}(p)$$

$$p = P(X \geq 15000)$$

$$= 1 - P(X < 15000)$$

$$= 1 - (1 - e^{-\frac{15000}{10000}}) = 0.223$$

$$\Rightarrow Y \sim \text{Geom}(0.223)$$

$$\Rightarrow E[Y] = (1 - 0.223) / 0.223 = 3.48.$$

$$\Rightarrow \text{Expected \# offers} = 3.48 + 1 = 4.48.$$

(b). Find the expected amount of money that Fred gets for the car.

We know that Fred accepts the offer only if the offer is larger than \$15000.

$$\text{WTF: } E[X | X \geq 15000]$$

By memoryless property,

$$\begin{aligned} E[X | X \geq 15000] &= E[X - 15000 + 15000 | X \geq 15000] \\ &= 15000 + E[X - 15000 | X \geq 15000] \\ &\geq 15000 + E[X] \geq 25000. \end{aligned}$$

2. Find $E[X^3]$ for $X \sim \text{Exp}(\lambda)$
By LOTUS

$$\begin{aligned} E[X^3] &= \int_0^{\infty} x^3 \lambda e^{-\lambda x} dx \\ &= x^3 (-e^{-\lambda x}) \Big|_0^{\infty} - \int_0^{\infty} 3x^2 (-e^{-\lambda x}) dx \\ &= 0 + \frac{3}{\lambda} \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx \\ &= \frac{3}{\lambda} E[X^2] \\ &= \frac{3}{\lambda} \{ \text{Var}(X) + (E[X])^2 \} \\ &= \frac{3}{\lambda} \left[\frac{1}{\lambda^2} + \left(\frac{1}{\lambda}\right)^2 \right] \\ &= \frac{6}{\lambda^3}. \quad \checkmark \end{aligned}$$

3. We find the cdf of M .

$$\begin{aligned} P(M \leq m) &= P(\min(X_1, \dots, X_n) \leq m) \\ &= 1 - P(\min(X_1, \dots, X_n) > m). \end{aligned}$$

Then,

$$\begin{aligned} &P(\min(X_1, \dots, X_n) > m) \\ &= P(X_1 > m, \dots, X_n > m) \quad \downarrow \text{by independence.} \\ &= P(X_1 > m) \cdots P(X_n > m) \\ &= e^{-\lambda_1 m} e^{-\lambda_2 m} \cdots e^{-\lambda_n m} \\ &= e^{-(\lambda_1 + \lambda_2 + \dots + \lambda_n) m} \end{aligned}$$

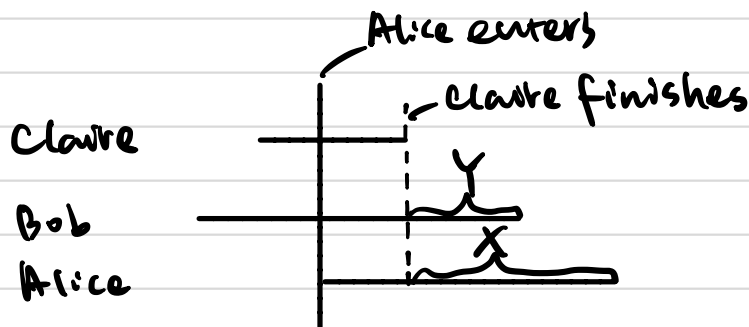
Therefore,

$$\begin{aligned} P(M \leq m) &= 1 - e^{-(\lambda_1 + \lambda_2 + \dots + \lambda_n) m} \\ \Rightarrow M &\sim \text{Exp}(\lambda_1 + \lambda_2 + \dots + \lambda_n). \quad \checkmark \end{aligned}$$

Intuition: $\{ \text{Check the end of the file} \}$.

4.

(a).

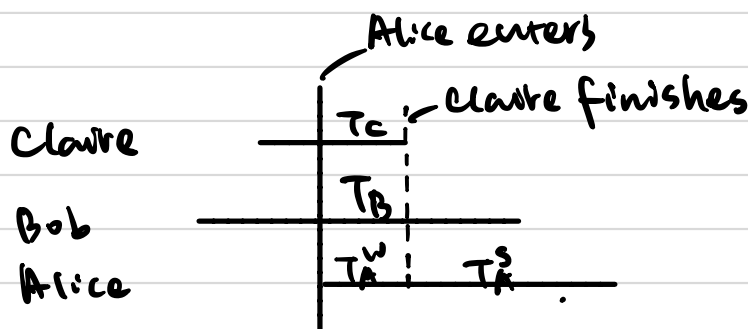


Without loss of generality, we assume Claire finishes first. By memoryless property, $X, Y \sim \text{Exp}(\lambda)$. By symmetry, $P(X > Y) = P(Y > X)$.

Since $P(X = Y) = 0$,

$$\Rightarrow P(X > Y) = \frac{1}{2}.$$

(b).



Let T_A be an r.v. representing the total time Alice spends. T_A consists of waiting time T_A^W and serving time T_A^S :

$$T_A = T_A^W + T_A^S.$$

Based on Question 3,

$$T_A^W = \min(T_B, T_C) \sim \text{Exp}(2\lambda)$$

where $T_B, T_C \sim \text{Exp}(\lambda)$ and are independent.

And,

$$T_A^S \sim \text{Exp}(\lambda)$$

$$E[T_A] = E[T_A^W] + E[T_A^S] = \frac{1}{2\lambda} + \frac{1}{\lambda} = \frac{3}{2\lambda}.$$

II. Moment Generating Functions (MGFs)

1. Find $E[X^3]$ for $X \sim \text{Expo}(\lambda)$ by using MGF.
In Lecture, we know

$$E[X^3] = \frac{3!}{\lambda^3} = \frac{6}{\lambda^3} \quad \checkmark$$

↳ comparing to 1.2.1

2. Suppose X has MGF $M_X(t)$.

The MGF of $-X$ is

$$\begin{aligned} M_{-X}(t) &= E[e^{t(-X)}] \\ &= E[e^{-tX}] \\ &= M_X(-t) \end{aligned} \quad \checkmark$$

The MGF of $Y = aX + b$ is → copied wrongly.

$$\begin{aligned} M_Y(t) &= E[e^{t(aX+b)}] \\ &= E[e^{atX + bt}] \\ &= e^{bt} E[e^{atX}] \\ &= e^{bt} M_X(at) \end{aligned} \quad \checkmark$$

3. Since U_1, U_2, \dots, U_{60} be i.i.d. $\text{Unif}(0,1)$,

$$M_X(t) = M_{U_1}(t) \dots M_{U_{60}}(t).$$

Let $U \sim \text{Unif}(0,1)$. Then,

$$\begin{aligned} M_U(t) &= E[e^{tU}] = \int_0^1 e^{tu} du \\ &= \frac{1}{t} e^{tu} \Big|_0^1 = \frac{1}{t} (e^t - 1) \quad \checkmark \quad \text{for } t \neq 0 \end{aligned}$$

Therefore,

$$M_X(t) = \frac{(e^t - 1)^{60}}{t^{60}} \quad \checkmark$$

for $t=0$,
 $M_X(t) = 1.$

4. In lecture, we have MGF of $X \sim \text{Pois}(\lambda)$

$$M(t) = e^{\lambda(e^t - 1)}$$

Then,

$$g(t) = \ln M(t)$$

$$= \lambda(e^t - 1)$$

$$= \lambda \sum_{j=1}^{\infty} \frac{t^j}{j!}$$

$$= \sum_{j=1}^{\infty} \frac{\lambda}{j!} t^j.$$

$\Rightarrow j^{\text{th}}$ cumulant of X is λ for all $j \geq 1$. ✓

1.3.

Intuition: $M = \min(X_1, \dots, X_n)$, it makes sense that M should have a continuous, memoryless distribution (\Rightarrow Exponential).

If we interpret λ_j as rate, it makes sense that M has a combined rate $\lambda_1 + \lambda_2 + \dots + \lambda_n$. X_1 : wait time for a green car to pass by, X_2 : ...
... blue car, ... M : wait time for a car with any of these colors.