

Global-local shrinkage prior for variable selection in graph-structured models.

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Dependence between variables may be induced by various factors in different applications:

- Genomic studies: dependence structure between genes obtained from biological pathways or inferred computationally (e.g., based on co-expression),
- Environmental studies: dependence structure between covariates collected over years,
- ...

In many domains high-dimensional data are generated: the number of variables p may be greater than the number of observations n :

- Genomic studies: high-throughput technologies provide genetic/genomic information on the whole genome,
- Environmental studies: high-throughput technologies provide regular and intense monitoring of phenotypic traits over time,
- ...



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Context

For example in linear model context:

↪ Spectrometric data:

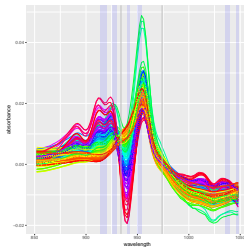
$$\begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix}$$

Fat content measured on n individuals

~

$$\begin{bmatrix} X_{11} & X_{1p} \\ \vdots & \vdots \\ X_{n1} & X_{np} \end{bmatrix}$$

Spectra sampled at p wavelengths



↪ Genomic data:

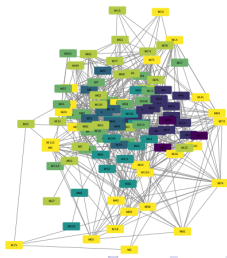
$$\begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix}$$

Production measured on n individuals

~

$$\begin{bmatrix} X_{11} & X_{1p} \\ \vdots & \vdots \\ X_{n1} & X_{np} \end{bmatrix}$$

Expression values of p genes



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Need to use statistical approaches incorporating such structures and dealing with $p \gg n$

↪ **Regularization methods**

- To help the model building process by reducing the model complexity of models
- To prevent ill-posed problems (non-invertible matrix, overfitting)
- To lead to parsimonious models

In the following we will focus on **Bayesian approaches**



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Bayesian approaches

In Bayesian framework additional information is integrated into models via prior distributions

↔ Regularization is done by specifying **specific priors**

Selection

To shrink towards zero small coefficients while leaving large signals large:
Shrinkage priors

Structure

Priors with a **variance-covariance matrix** related to structure information between variables

Objective

Taking advantage that most of the dependence between variables may be encoded by an undirected graph \mathcal{G}
↔ To propose **shrinkage** priors integrating **graph structure** information to select graph-structured variables

Shrinkage priors

Two classes of shrinkage priors:

- **Spike-and-slab priors:** Discrete mixture of two distributions (Mitchell and Beauchamp, 1988; George and McCulloch, 1997)
 - **Continuous shrinkage priors:** Unimodal continuous distributions (Bayesian Lasso prior, Horseshoe prior, Elastic-Net prior, ...) (Kyung et al., 2010; Carvalho et al., 2008)
- ↪ the class of global-local priors (Carvalho et al., 2010; Polson and Scott, 2010): a scale mixture of Gaussian distributions with the mixing density depending on two hyperparameters to control the **global** shrinkage and the **local** deviations



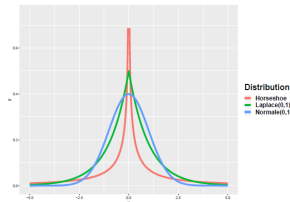
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The proposed approach

We propose to combine:

- A global-local prior, the horseshoe (HS) prior (Carvalho et al., 2010), for its **efficiency and flexibility** in terms of selection and estimation

↪ allows to shrink towards zero small coefficients while allowing large signals to escape from the overall shrinkage



- With a Gaussian Markov random field (GMRF) for its **appealing connection** with undirected graphs (Rue and Held, 2005)

↪ allows to impose the dependence structure between the parameters via the precision matrix of a conditionally Gaussian prior

↪ An extension of the approach by Faulkner and Minin (2018); Faulkner (2019) to the more general context of graph-structured variable selection.



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Bayesian hierarchical model

We assume that $\mathcal{G} = \bigcup_{i=1}^I \mathcal{G}_i = \bigcup_{i=1}^I (V_i, E_i)$ a disjoint union of I subgraphs and \mathcal{S} the set of indices associated to one representative of each of the I subgraphs.

HS-GMRF model

$$\mathbf{y} | \boldsymbol{\beta}, \sigma^2 \sim \mathcal{N}_n(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I}_n)$$

$$\beta_j - s_{jj'} \beta_{j'} | \tau_{jj'}^2, \lambda^2 \sim \mathcal{N}(0, \lambda^2 \tau_{jj'}^2) \text{ for } (j, j') \in \bigcup_{i=1}^I E_i$$

$$\beta_j | \tau_j^2, \lambda^2 \sim \mathcal{N}(0, \lambda^2 \tau_j^2) \text{ for } j \in \mathcal{S}$$

$$\tau_{jj'} \sim \mathcal{C}^+(0, 1) \text{ for } (j, j') \in \bigcup_{i=1}^I E_i; \tau_j \sim \mathcal{C}^+(0, 1) \text{ for } j \in \mathcal{S}$$

$$\lambda | \sigma \sim \mathcal{C}^+(0, \sigma); \sigma^2 \sim \mathcal{IG}(a_0, b_0)$$

with $s_{jj'} = \text{sign}\{\text{cor}(X_j, X_{j'})\}$ to encourage regression coefficients of negatively correlated variables to take opposite signs.



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Simulation study

Objectives

- To evaluate the performances of the proposed approach with and without incorporating the sign of the sample correlation (HS-GMRF and HS-GMRF-nosign),
- To compare the results with two other approaches: the HS and the spike-and-slab with Ising prior (SS-Ising) (Smith and Fahrmeir, 2007; Li and Zhang, 2010) and when the true graph is known and unknown.



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Simulation study

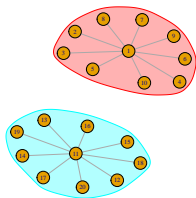
$$Y = \sum_{g=1}^G \mathbf{X}_g \beta_g + \varepsilon \text{ with } X_{i,g} = (X_{i,g1}, \dots, X_{i,gk})' \sim \mathcal{N}_k(0, \Sigma_g) \text{ and } \varepsilon \sim \mathcal{N}_n(0, \sigma^2 I_n)$$

12 simulated scenarios

- Two covariance structures
- Two levels of correlation ($\rho = 0.5, 0.9$)
- Three regression coefficients

Simulations

- $G = 14$ groups of $k = 10$ predictors with 5 groups with non-zero effects,
- $\sigma^2 = \sum_{g=1}^G \beta_g^2 / 5$
- Repetitions: 50



↗ Focus on the scenario where half of groups with Σ_g and

$$\beta_g = (5, -\frac{5}{\sqrt{10}}, -\frac{5}{\sqrt{10}}, \underbrace{\frac{5}{\sqrt{10}}, \dots, \frac{5}{\sqrt{10}}}_{k-3})$$



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Simulation study

Performance criteria

- Variable selection criteria:
 - ↪ For HS-based: variable selected if 95% HPD interval does not contain 0,
 - ↪ For SS-Ising: variable selected if marginal inclusion posterior probability greater than 0.5.
- Matthews correlation coefficient (MCC),
- Mean squared error (MSE) of the regression coefficients,
- Mean squared prediction error (MSPE).

MCMC settings:

- Iterations : 6,000,
- Burn-in: 1,000.



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Results using the true graph

Table 1: Average MCC, MSE and MSPE (with SE) over 50 simulated replications.

		MCC	MSE	MSPE
$\Sigma_{g, \text{half}}$ $\rho = 0.5$	HS-GMRF	0.708 (± 0.018)	0.513 (± 0.067)	94.871 (± 13.632)
	HS-GMRF-nosign	0.624 (± 0.034)	0.728 (± 0.155)	122.188 (± 21.609)
	HS	0.240 (± 0.041)	1.009 (± 0.200)	126.252 (± 19.657)
	SS-Ising	0.323 (± 0.054)	1.386 (± 0.204)	149.294 (± 27.384)
$\Sigma_{g, \text{half}}$ $\rho = 0.9$	HS-GMRF	0.668 (± 0.046)	0.541 (± 0.089)	84.954 (± 14.485)
	HS-GMRF-nosign	0.444 (± 0.117)	1.038 (± 0.259)	99.123 (± 17.694)
	HS	0.219 (± 0.038)	2.243 (± 0.551)	95.219 (± 19.279)
	SS-Ising	0.312 (± 0.048)	2.359 (± 0.437)	109.387 (± 23.713)

- HS-GMRF-based approaches lead to the best results in terms of MCCs, MSEs, and MSPEs,
- HS-GMRF outperforms HS-GMRF-nosign especially when $\rho = 0.9$



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Results using the true graph

Table 2: Average MCC and MSE for connected and non-connected covariates over 50 simulated replications.

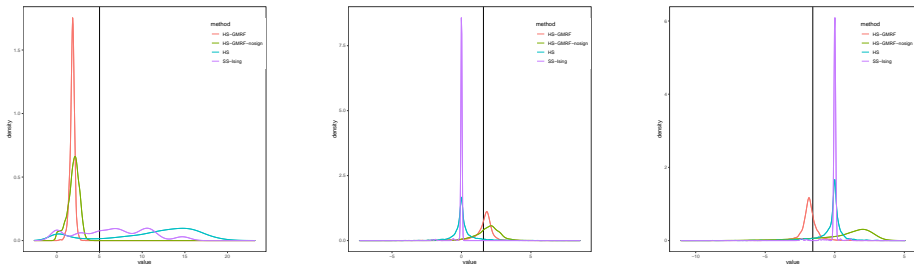
	MCC		MSE	
	Connected	Non-connected	Connected	Non-connected
	$\Sigma_{g, \text{half}} \rho = 0.9$			
HS-GMRF	0.883 (± 0.078)	0.278 (± 0.049)	0.611 (± 0.138)	0.470 (± 0.091)
HS-GMRF-nosign	0.526 (± 0.177)	0.265 (± 0.053)	1.582 (± 0.465)	0.495 (± 0.112)
HS	0.188 (± 0.043)	0.271 (± 0.046)	3.998 (± 1.105)	0.488 (± 0.103)
SS-Ising	0.310 (± 0.047)	0.304 (± 0.081)	4.055 (± 0.855)	0.662 (± 0.135)

- Performances for non-connected predictors are similar for HS and HS-GMRF-based approaches.
- For connected variables the integration of the dependence structure helps to select variables with small effects



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Results using the true graph

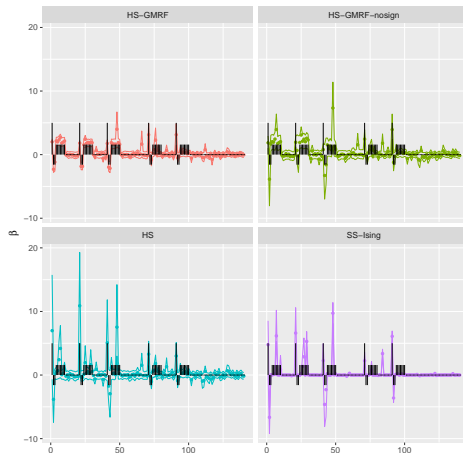


- HS leads to a bimodal posterior density or a distribution concentrated around 0 with large tails
- HS-GMRF-based approaches give narrower posterior densities away from 0
- For $\beta = -5/\sqrt{(10)}$ HS-GMRF-nosign spreads out posterior density around the average of β 's



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Results using the true graph



- HS-GMRF-based approaches give similar estimates for highly correlated covariates,
- HS-GMRF yields narrower HPD intervals with good coverage and fairly accurate estimates for regression coefficients with opposite signs.
- HS and SS-Ising tend to select one representative of a group of correlated variables,
- HS gives wide HPD intervals,

Estimated coefficients along with 80% HPD intervals in one simulated replication ($\Sigma_{g, \text{half}}$, $\rho = 0.9$).



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Results using an estimated graphs

Graph structure may not be known and may need to be estimated. Graphical Lasso approach used to estimate the graph (Friedman et al., 2008)

- HS-GMRF-based approaches outperform the other approaches,
- For moderate correlation: graph is underestimated \Rightarrow slightly poorer selection and estimation for the HS-GMRF-based approaches than with the true graph,
- For high correlation: graph is overestimated \Rightarrow improved selection for the HS-GMRF-based approaches compared to the true graph but with an over smoothing of the regressions coefficients



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Application

Objective

To identify gene expressions involved in the variability of riboflavin production using data on 71 samples

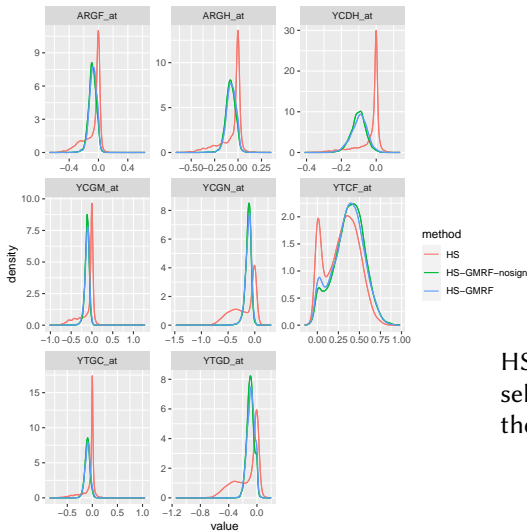
- A total of 142 gene expressions considered
- Estimation of an undirected graph with 157 edges
- 5-fold cross-validation procedure

Methods	CV-MSPE	Selected genes
HS-GMRF	0.29	4 (90% HPD) 8 (80% HPD)
HS-GMRF-nosign	0.31	4(90% HPD) 6 (80% HPD)
HS	0.33	0(90% HPD) 0 (80% HPD)
SS-Ising	0.37	21 ($PPI > 0.5$) 16 ($PPI > 0.8$)
Lasso	0.41	16

↪ HS-GMRF yields the smallest CV-MSPE



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For moderate non-zero effects:

- HS estimates densities concentrated around 0 with long tails or bimodal densities with one of the modes around 0,
- HS-GMRF-based methods estimate unimodal densities or bimodal densities with the mode around 0 less than with HS.

HS-GMRF-based approaches select groups of genes involved in the same biological pathway.



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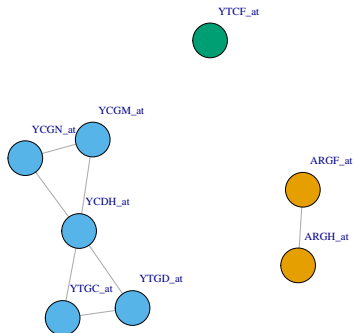


Figure 1: The estimated network



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Conclusion/Perspective

The proposed approaches allow to:

- consider a broad type of dependence structures,
- achieve flexibility in the estimation and the selection due to the local and global shrinkage hyperparameters,
- need to consider the sign of the sample correlation,
- give better predictive performances notably by selecting groups of connected variables,
- give good results even when true graph is unknown and needs to be estimated.

Limitation:

- tend to encourage similar values for connected variables, especially for highly correlated variables or overestimated graphs.

For future research:

- Extension to non-Gaussian distributions,
- Integration of prior knowledge on strengths of connections between variables.



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Thanks for your attention !

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MCMC implementation

MCMC implementation

A Gibbs sampling algorithm is straightforward to fit the hierarchical models:

- by using the parametrization of a half-Cauchy as a mixture of inverse-gamma distributions (Makalic and Schmidt, 2016),
- by introducing a q -dimensional vector $\phi = (\phi_1, \dots, \phi_q)' = C\beta$ (Martínez-Beneito and Botella-Rocamora, 2019) where $q = |E| + |\mathcal{S}|$ and C is a contrast matrix such that:

$$\phi \sim \mathcal{N}_q(0, \Sigma_\phi),$$

with $\Sigma_\phi = \text{diag}(\lambda^2 \tau^2)$.



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Results using the true graph

Table 3: Coverage probability (CP) and width of 95% HPD intervals averaged over the 50 simulated replications.

		CP of 95% HPD	Width of 95% HPD
$\Sigma_{g, \text{half}}$	$\rho = 0.5$	HS-GMRF	0.923 (± 0.026)
		HS-GMRF nosign	0.931 (± 0.027)
		HS	0.894 (± 0.037)
		SS-Ising	0.751 (± 0.026)
	$\rho = 0.9$	HS-GMRF	0.928 (± 0.019)
		HS-GMRF nosign	0.922 (± 0.031)
		HS	0.908 (± 0.05)
		SS-Ising	0.773 (± 0.029)

- CPs similar for HS-based approaches but wider intervals for HS



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