

SML ASSIGNMENT 2

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Exercice 1 Nearest Neighbors as a Kernel Method

1. Let explain clearly why and how $\hat{f}_{kNN}(x)$ does indeed compute the predicted label of x

$$\hat{f}_{kNN}(x) = \text{sign}\left(\sum_{i=1}^n y_i \alpha_i \mathcal{K}(x, x_i)\right)$$

$\mathcal{K}(x, x_i)$ is the weight attributed to the observation x_i .

It is in a such way that if $d(x, x_i) \leq d(x, x_j)$ then $\mathcal{K}(x, x_j) \leq \mathcal{K}(x, x_i)$. This mean that more close are the observation to x and more bigger is it weight.

$\sum_{i=1}^n y_i \alpha_i \mathcal{K}(x, x_i)$ is the summation of all the response of the observations that are in the neighborhood of x multiply by their weight.

With this principle x is classify in the class of obsevation that are more close of it.

2. Let explain succinctly what $\hat{\pi}(x)$ is estimating in this case

$$\hat{\pi}(x) = \sum_{i=1}^n \mathbb{1}(y_i = 1) \alpha_i \mathcal{K}(x, x_i)$$

$\hat{\pi}(x)$ compute the sum of the weight of all observations in the neighborhood of x that have response 1.

It is an estimation of $Pr(Y = 1|x)$

3. Let explain clearly the relationship between $\hat{f}_{kNN}(x)$ and

$$\hat{g}_{kNN}(x) = 2\mathbb{1}\left(\hat{\pi}(x) > \frac{1}{2}\right) - 1$$

- if $\hat{\pi}(x) > \frac{1}{2}$
 if $\hat{\pi}(x) > \frac{1}{2}$, we have $\hat{g}_{kNN}(x) = 2 \times 1 - 1 = 1$
 And at the same time by the way $\hat{\pi}(x)$ is defined, that's mean that ,

$$\sum_{i=1}^n \mathbb{1}(y_i = 1) \alpha_i \mathcal{K}(x, x_i) > \sum_{i=1}^n \mathbb{1}(y_i = -1) \alpha_i \mathcal{K}(x, x_i)$$

then, $\text{sign}(\sum_{i=1}^n y_i \alpha_i \mathcal{K}(x, x_i)) = 1$ i.e. $\hat{f}_{kNN}(x) = 1$

So if $\hat{\pi}(x) > \frac{1}{2}$ we have $\hat{g}_{kNN}(x) = \hat{f}_{kNN}(x)$

- if $\hat{\pi}(x) \leq \frac{1}{2}$
 if $\hat{\pi}(x) \leq \frac{1}{2}$, then $\hat{g}_{kNN}(x) = 2 \times 0 - 1 = -1$
 And at the same time by the way $\hat{\pi}(x)$ is defined, that's mean that ,

$$\sum_{i=1}^n \mathbb{1}(y_i = 1) \alpha_i \mathcal{K}(x, x_i) < \sum_{i=1}^n \mathbb{1}(y_i = -1) \alpha_i \mathcal{K}(x, x_i)$$

then, $\text{sign}(\sum_{i=1}^n y_i \alpha_i \mathcal{K}(x, x_i)) = -1$ i.e. $\hat{f}_{kNN}(x) = -1$

So if $\hat{\pi}(x) \leq \frac{1}{2}$ we have $\hat{g}_{kNN}(x) = \hat{f}_{kNN}(x)$

Conclusion : $\hat{g}_{kNN}(x) = \hat{f}_{kNN}(x)$

4. In this case we have $\mathcal{Y} = \{0, 1\}$ because $\mathbb{1}(\hat{\pi}(x) > \frac{1}{2}) \in \{0, 1\}$
5. We can state the logistic regression function space defined as follow:
 Let denote

$$\hat{\pi}(x) = \frac{1}{1 + e^{-\sum_i \beta_i x_i}}$$

We have,

$$\mathcal{H} := \left\{ f_{s.t}, \forall x \in \mathcal{X}, f(x) = \mathbb{1} \left(\hat{\pi}(x) > \frac{1}{2} \right) \right\}$$

Exercise 2: When can we compute the Bayes' Risk?

1. Let sketch the contour of the two classes in the plane.
 (SEE THE RMARKDOWN CODE)

2. Let find and write down the simplified expression of the Bayes' classifier for this task. We have,

$$\begin{aligned}
f^*(x) &= \mathbb{1} \left(\frac{\mathbb{P}[Y = 1|x]}{\mathbb{P}[Y = 0|x]} \geq 1 \right) \\
&= \mathbb{1} (\ln(\mathbb{P}[Y = 1|x]) - \ln(\mathbb{P}[Y = 0|x]) > 0) \\
&= \mathbb{1} \left(\ln\left(\frac{\mathbb{P}[x|Y = 1]\mathbb{P}[Y = 1]}{\mathbb{P}[x]}\right) - \ln\left(\frac{\mathbb{P}[x|Y = 0]\mathbb{P}[Y = 0]}{\mathbb{P}[x]}\right) > 0 \right) \\
&= \mathbb{1} (\ln(\mathbb{P}[x|Y = 1]) + \ln(\mathbb{P}[Y = 1]) - \ln(\mathbb{P}[x|Y = 0]) - \ln(\mathbb{P}[Y = 0]) > 0) \\
&= \mathbb{1} \left(\frac{-1}{2}(x - \mu_1)^T \Sigma^{-1}(x - \mu_1) + \frac{1}{2}(x - \mu_0)^T \Sigma^{-1}(x - \mu_0) + \ln(\psi) - \ln(1 - \psi) > 0 \right) \\
&= \mathbb{1} \left(\frac{1}{2}(-x^T \Sigma^{-1}x + x^T \Sigma^{-1}\mu_1 + \mu_1^T \Sigma^{-1}x - \mu_1^T \Sigma^{-1}\mu_1 + \right. \\
&\quad \left. x^T \Sigma^{-1}x - x^T \Sigma^{-1}\mu_0 - \mu_0^T \Sigma^{-1}x + \mu_0^T \Sigma^{-1}\mu_0) + \ln(\psi) - \ln(1 - \psi) > 0 \right) \\
&= \mathbb{1} \left(\frac{1}{2}(2x^T \Sigma^{-1}\mu_1 - \mu_1^T \Sigma^{-1}\mu_1 - 2x^T \Sigma^{-1}\mu_0 + \mu_0^T \Sigma^{-1}\mu_0) + \ln(\psi) - \ln(1 - \psi) > 0 \right) \\
&= \mathbb{1} \left((\mu_1 - \mu_0)^T \Sigma^{-1}x - \frac{1}{2}(\mu_1 + \mu_0)^T \Sigma^{-1}(\mu_1 - \mu_0) + \ln(\psi) - \ln(1 - \psi) > 0 \right)
\end{aligned}$$

We can then conclude that,

$$f^*(x) = \mathbb{1} (\omega^T x + b > 0)$$

where $\omega = \Sigma^{-1}(\mu_1 - \mu_0)$ and

$$b = -\frac{1}{2}(\mu_1 + \mu_0)^T \Sigma^{-1}(\mu_1 - \mu_0) + \ln(\psi) - \ln(1 - \psi)$$

3. Let denote DB the decision boundary. We have ,

$$DB = \{x \in \mathcal{X} \quad s.t \quad \omega^T x + b = 0\}$$

where $\omega = \Sigma^{-1}(\mu_1 - \mu_0)$ and

$$b = -\frac{1}{2}(\mu_1 + \mu_0)^T \Sigma^{-1}(\mu_1 - \mu_0) + \ln(\psi) - \ln(1 - \psi)$$

Let develop more. $\mu_0 = (-2 \ -1)^T$, $\mu_1 = (0 \ 1)^T$,

we have $\mu_1 - \mu_0 = (2 \ 2)$ and $\mu_1 + \mu_0 = (-2 \ 0)$

We have,

$$\Sigma = \begin{pmatrix} 1 & \frac{-3}{4} \\ \frac{-3}{4} & 2 \end{pmatrix} \quad \text{and} \quad \det(\Sigma) = \frac{23}{16} \quad \text{so,}$$

$$\Sigma^{-1} = \frac{16}{23} \begin{pmatrix} 2 & \frac{3}{4} \\ \frac{3}{4} & 1 \end{pmatrix}$$

Another side we have,

$$(\mu_1 - \mu_0)^T \Sigma^{-1} = (2 \quad 2) \frac{16}{23} \begin{pmatrix} 2 & \frac{3}{4} \\ \frac{3}{4} & 1 \end{pmatrix}$$

$$= \frac{8}{23} (11 \quad 7)$$

and also,

$$(\mu_1 + \mu_0)^T \Sigma^{-1} = (-2 \quad 0) \frac{16}{23} \begin{pmatrix} 2 & \frac{3}{4} \\ \frac{3}{4} & 1 \end{pmatrix}$$

$$= \frac{8}{23} (-8 \quad -3)$$

so,

$$\frac{-1}{2} (\mu_1 + \mu_0)^T \Sigma^{-1} (\mu_1 - \mu_0) = \frac{-1}{2} \frac{8}{23} (-8 \quad -3) \begin{pmatrix} 2 \\ 2 \end{pmatrix}$$

$$= \frac{88}{23}$$

We can then write,

$$\omega^T x + b = \frac{88}{23} x_1 + \frac{56}{23} x_2 + \frac{88}{23} + \ln(\psi) - \ln(1 - \psi)$$

Finally,

$$DB = \{x = (x_1, x_2) \in \mathcal{X} \quad \text{s.t} \quad \frac{88}{23} x_1 + \frac{56}{23} x_2 + \frac{88}{23} + \ln(\psi) - \ln(1 - \psi) = 0\}$$

4. Theoretical expression of the Bayes' risk. We have,

$$R^* = \min_f \mathbb{P}[Y \neq f(X)]$$

but because we know f^* , we have

$$R^* = \mathbb{P}[Y \neq \mathbf{1} (\omega^T x + b > 0)]$$

$$\begin{aligned} R^* &= \mathbb{P}[Y \neq \mathbf{1} (\omega^T x + b > 0)] \\ &= \mathbb{P}[Y = 1, \omega^T x + b \leq 0] + \mathbb{P}[Y = 0, \omega^T x + b > 0] \\ &= \mathbb{P}[\omega^T x + b \leq 0 | Y = 1] \mathbb{P}[Y = 1] + \mathbb{P}[\omega^T x + b > 0 | Y = 0] \mathbb{P}[Y = 0] \\ &= \psi \mathbb{P}[\omega^T x + b \leq 0 | Y = 1] + (1 - \psi) \mathbb{P}[\omega^T x + b > 0 | Y = 0] \\ &= \psi \mathbb{P}[\omega^T x \leq -b | Y = 1] + (1 - \psi) \mathbb{P}[\omega^T x > -b | Y = 0] \end{aligned}$$

But $X|Y$ is bivariate gaussian distribution for which we now the parameters, therefor $\omega^T X|Y$ is also a gaussian distribution . We are going to determine the parameters.

$$\begin{aligned} \mathbb{E}_0(\omega^T X | Y = 0) &= \omega^T \mathbb{E}_0(X | Y = 0) \\ &= \omega^T (\mu_0) \\ &= \frac{8}{23} \begin{pmatrix} 11 & 7 \end{pmatrix} \begin{pmatrix} -2 \\ -1 \end{pmatrix} \\ &= \frac{-232}{23} \end{aligned}$$

$$\begin{aligned} \mathbb{E}_1(\omega^T X | Y = 1) &= \omega^T \mathbb{E}_1(X | Y = 1) \\ &= \omega^T (\mu_1) \\ &= \frac{8}{23} \begin{pmatrix} 11 & 7 \end{pmatrix} \begin{pmatrix} 0 \\ 1 \end{pmatrix} \\ &= \frac{56}{23} \end{aligned}$$

$$\begin{aligned} \text{Var}(\omega^T X | Y = 1) &= \omega^T \Sigma \omega \\ &= \frac{8}{23} \begin{pmatrix} 11 & 7 \end{pmatrix} \begin{pmatrix} 1 & \frac{-3}{4} \\ \frac{-3}{4} & 2 \end{pmatrix} \begin{pmatrix} 11 \\ 7 \end{pmatrix} \frac{8}{23} \\ &= \frac{288}{23} \end{aligned}$$

With this we can deduce that,

$$Z_0 = \frac{(\omega^T X|Y=0) - (\frac{-232}{23})}{\sqrt{\frac{288}{23}}} \text{ follow a standard normal distribution}$$

$$\text{and } Z_1 = \frac{(\omega^T X|Y=1) - \frac{56}{23}}{\sqrt{\frac{288}{23}}} \text{ also .}$$

Let call Z a standard normal distribution and $\phi(x)$ its CDF
We have

$$R^* = \psi \phi\left(\frac{-b - \frac{56}{23}}{\sqrt{\frac{288}{23}}}\right) + (1 - \psi)\left(1 - \phi\left(\frac{-b + (\frac{232}{23})}{\sqrt{\frac{288}{23}}}\right)\right)$$

5. Let consider that $\psi = \frac{1}{2}$

With this will have $b = \frac{88}{23}$ and we have

$$\begin{aligned} R^* &= \frac{1}{2} \phi\left(\frac{-\frac{88}{23} - \frac{56}{23}}{\sqrt{\frac{288}{23}}}\right) + \frac{1}{2} \left(1 - \phi\left(\frac{-\frac{88}{23} + (\frac{232}{23})}{\sqrt{\frac{288}{23}}}\right)\right) \\ &= \frac{1}{2} \phi\left(\frac{-6\sqrt{46}}{23}\right) + \frac{1}{2} \left(1 - \phi\left(\frac{6\sqrt{46}}{23}\right)\right) \\ &= \frac{1}{2} \left(1 - \phi\left(\frac{6\sqrt{46}}{23}\right)\right) + \frac{1}{2} \left(1 - \phi\left(\frac{6\sqrt{46}}{23}\right)\right) \\ &= \left(1 - \phi\left(\frac{6\sqrt{46}}{23}\right)\right) \\ &= (1 - \phi(1.769)) \\ &= (1 - 0.9615784) \\ &= 0.03842162 \end{aligned}$$