

- Q.1.** Assume that the alcohol content (in percentage) in a cough syrup taken from any particular brand is normally distributed with the true standard deviation 0.75 .
- Compute a 95% CI for the true average alcohol content of a certain brand if the average content for 20 specimens from the brand was 4.85 .
  - Compute a 98% CI for the true average alcohol content of a another brand based on 16 specimens with a sample average alcohol content of 4.56 .
  - How large a sample size is necessary if the width of the 95% interval is to be 0.40?
  - What sample size is necessary to estimate the true average alcohol content to within 0.2 with 99% confidence?

**Sol.1.** Given that population is normally distributed with  $\sigma = 0.75$  .

- (a) A 95% CI for mean of a normal population is given by:

$$\bar{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

Given that  $1 - \alpha = 0.95$ ,  $\alpha = 0.05$ ,  $n = 20$ ,  $\bar{x} = 4.85$

Desired 95% CI is:  $4.85 \pm z_{0.025} \frac{0.75}{\sqrt{20}} = 4.85 \pm 0.33 = (4.52, 5.18)$

(b) Given that  $\alpha = 0.02$ ,  $n = 16$ ,  $\bar{x} = 4.56$

Desired 98% CI is:  $4.56 \pm z_{0.01} \frac{0.75}{\sqrt{16}} = (4.12, 5.00)$

- (c) Width of  $100(1-\alpha)\%$  interval is given by:

$$w = 2z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

Therefore,  $n = \left[ \frac{2(1.96)(0.75)}{0.40} \right]^2 = 54.02 \approx 55$

(d) Now, given that the true average alcohol content must lie within the interval  $(-0.2, 0.2)$ .

Width =  $2(0.2) = 0.4$

Therefore,  $n = \left[ \frac{2(2.58)(0.75)}{0.40} \right]^2 = 93.61 \approx 94$

- Q.2.** A random sample of  $n=15$  light bulbs of a certain type yielded the following observations on lifetime (in years):

2.0   1.3   6.0   1.9   5.1   0.4   1.0   5.3  
15.7   0.7   4.8   0.9   12.2   5.3   0.6

- Assume that the lifetime distribution is exponential. Obtain a 95% CI for expected (true average) lifetime.
- What is a 95% CI for the standard deviation of the lifetime distribution?

**Sol.2.** We know that  $h(X_1, X_2, \dots, X_n; \lambda) = 2\lambda \sum X_i$  has a chi-squared distribution with  $2n$  degrees of freedom.

The relevant number of df here is  $2(15)=30$ . The  $\nu = 30$  row of chi-squared distribution table shows

that 46.979 captures upper-tail area 0.025 and 16.791 captures lower-tail area 0.025 .

Thus for  $n=15$ :

$$P(16.791 < 2\lambda \sum X_i < 46.979) = 0.95$$

Desired 95% CI for  $\mu = \frac{1}{\lambda}$  is:

$$\left( \frac{2 \sum x_i}{46.979}, \frac{2 \sum x_i}{16.791} \right)$$

Here  $\sum x_i = 63.2$

Therefore, desired interval is (2.69, 7.53).

(b) We know theta when X has an exponential distribution, then  $V(X) = \frac{1}{\lambda^2}$ . So, the standard deviation is  $\frac{1}{\lambda}$  which is same as the mean.

Thus the interval calculated in part'a' is also a 95% CI for the standard deviation of the lifetime distribution.

**Q.3.** A company manufactures taps. Out of 143 manufactured taps, 10 are found to be defective. Calculate a lower confidence bound at the 95% confidence level for the true proportion of such taps that are defective. Also, interpret the 95% confidence level.

**Sol.3.** Here,  $\alpha=0.05$ ,  $n=143$

For One sided bound, we need  $z_\alpha = z_{0.05} = 1.645$

Sample proportion of defective taps =  $\frac{10}{143} = 0.07$

Desired 95% lower confidence bound for the true proportion of defective taps:

$$\begin{aligned} & \frac{\hat{p} + \frac{z_\alpha^2}{2n}}{1 + \frac{z_\alpha^2}{n}} - z_\alpha \frac{\sqrt{\frac{\hat{p}\hat{q}}{n} + \frac{z_\alpha^2}{4n^2}}}{1 + \frac{z_\alpha^2}{n}} \\ &= \frac{.07 + \frac{1.645^2}{2(143)}}{1 + \frac{1.645^2}{143}} - 1.645 \frac{\sqrt{\frac{(0.07)(.93)}{143} + \frac{1.645^2}{4(143)^2}}}{1 + \frac{1.645^2}{143}} \\ &= 0.078 - 0.036 = 0.042 \end{aligned}$$

We are 95% confident that more than 4.2% of all manufactured taps are defective.

**Q.4.** In a survey of 2003 adults, 25% said they believed in God.

(a) Calculate and interpret a confidence interval at the 99% confidence level for the proportion of all adults who believe in God.

(b) What sample size would be required for the width of a 99% CI to be almost .05 irrespective of the value of  $\hat{p}$ ?

**Sol.4.** With such a large sample size, we use simplified CI formula that is given by:

$$\hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}\hat{q}}{n}}$$

Using,  $\hat{p}=0.25$ ,  $n=2003$ , and  $z_{\alpha/2} = z_{0.005}=2.576$

Desired 99% CI for p is:

$$\begin{aligned} 0.25 \pm 2.576 \sqrt{\frac{(0.25)(0.75)}{2003}} \\ = 0.25 \pm 0.025 = \\ (0.225, 0.275) \end{aligned}$$

(b) Here,  $\hat{p} = \hat{q} = 0.5$  Therefore,

$$\begin{aligned} n &= \frac{4z^2\hat{p}\hat{q}}{w^2} \\ n &= \frac{4(2.576)^2(0.5)(0.5)}{0.05^2} = 2654.31 \approx 2655 \end{aligned}$$

So, a sample of atleast 2655 is required.

**Q.5.** Let  $X_1, X_2, \dots, X_n$  be a random sample from a continuous probability distribution having median  $\tilde{\mu}$ . Show that

(a)  $P[\min X_i < \tilde{\mu} < \max X_i] = 1 - \frac{1}{2^{n-1}}$ , so that  $(\min X_i, \max X_i)$  is a  $100(1 - \alpha)\%$  confidence interval for  $\tilde{\mu}$  with  $\alpha = 1/2^{n-1}$ .

(b) For each of six normal male infants, the amount of the amino acid alanine (mg/100 mL) was determined while the infants were on an isoleucine-free diet, resulting in the following data:

2.84, 3.54, 2.80, 1.44, 2.94, 2.70

Compute a  $100(1 - \frac{1}{2^5})\%$  CI for the true median amount of alanine for infants on such a diet.

(c). Let  $X_{(2)}$  denote the second smallest of the  $x_i$ s and  $X_{(n-1)}$  denote the second largest of the  $x_i$ s. What is the confidence coefficient of the interval  $(X_{(2)}, X_{(n-1)})$  for  $\tilde{\mu}$ ?

**Sol.5.** (a)  $\tilde{\mu}$  is the median, therefore

$$P[X_i \leq \tilde{\mu}] = P[X_i \geq \tilde{\mu}] = \frac{1}{2}, \text{ for all } i = 1..n.$$

Furthermore,

$$\begin{aligned} P[\min X_i \leq \tilde{\mu} \leq \max X_i] &= 1 - P[(\tilde{\mu} < \min X_i) \text{ or } (\tilde{\mu} > \max X_i)] \\ &= 1 - P[\tilde{\mu} < \min X_i] - P[\tilde{\mu} > \max X_i] \\ &= 1 - P[\tilde{\mu} < X_1] \dots P[\tilde{\mu} < X_n] - P[\tilde{\mu} > X_1] \dots P[\tilde{\mu} > X_n] \\ &= 1 - \frac{1}{2^n} - \frac{1}{2^n} = 1 - \frac{1}{2^{n-1}}. \end{aligned}$$

(b) We have,

$$\min x_i = 1.44, \max x_i = 3.54 \text{ and } n = 6.$$

Therefore (from (a) part) (1.44, 3.54) is a  $100(1 - \frac{1}{2^5})\%$  confidence interval for  $\tilde{\mu}$ .

(c) We have

$$P[X_{(2)} \leq \tilde{\mu} \leq X_{(n-1)}] = 1 - P[\tilde{\mu} < X_{(2)}] - P[\tilde{\mu} > X_{(n-1)}]$$

where,

$$\begin{aligned} P[\tilde{\mu} < X_{(2)}] &= P[\text{at most one } X_i \text{ is below } \tilde{\mu}] = \frac{1}{2^n} + \binom{n}{1} \frac{1}{2} \frac{1}{2^{n-1}} = \frac{1}{2^n}(1+n) \\ P[\tilde{\mu} > X_{(n-1)}] &= P[\text{at most one } X_i \text{ is above } \tilde{\mu}] = \frac{1}{2^n}(1+n) \end{aligned}$$

Hence,

$$P[X_{(2)} \leq \tilde{\mu} \leq X_{(n-1)}] = 1 - \frac{1}{2^{n-1}}(1+n).$$

Therefore,  $(X_{(2)}, X_{(n-1)})$  is a  $100(1 - \frac{1}{2^{n-1}}(1+n))\%$  confidence interval for  $\tilde{\mu}$ .

Q.6. Let  $X_1, X_2, \dots, X_n$  be a random sample from a uniform distribution on the interval  $[0, \theta]$ , and let  $U = \frac{1}{\theta} \max X_i$ .

(a) Show that the rv  $U$  has density function

$$f_U(x) = \begin{cases} nx^{n-1} & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}.$$

(b) Verify that

$$P\left[\left(\frac{\alpha}{2}\right)^{1/n} \leq U \leq \left(1 - \frac{\alpha}{2}\right)^{1/n}\right] = 1 - \alpha,$$

and

$$P[\alpha^{1/n} \leq U \leq 1] = 1 - \alpha.$$

Use above identities to obtain two  $100(1 - \alpha)\%$  CI for  $\theta$ .

(c) If the waiting time for a bus during morning hours is uniformly distributed and observed waiting times are  $x_1 = 4.2$ ,  $x_1 = 3.5$ ,  $x_3 = 1.7$ ,  $x_4 = 1.2$ , and  $x_5 = 2.4$ , derive a 95% CI for  $\theta$  by using both intervals.

Sol.6. (a) The CDF for  $U$  is given as

$$\begin{aligned} F_U(x) &= P[U \leq x] = P[\max X_i \leq x\theta] \\ &= P[X_1 \leq x\theta] \dots P[X_n \leq x\theta] \\ &= x^n \end{aligned}$$

Therefore,

$$f_U(x) = F'_U(x) = \begin{cases} nx^{n-1} & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}.$$

(b) Integrating  $f_U(x)$  between  $(\frac{\alpha}{2})^{1/n}$  and  $(1 - \frac{\alpha}{2})^{1/n}$ , one obtain

$$P\left[\left(\frac{\alpha}{2}\right)^{1/n} \leq U \leq \left(1 - \frac{\alpha}{2}\right)^{1/n}\right] = \int_{(\frac{\alpha}{2})^{1/n}}^{(1-\frac{\alpha}{2})^{1/n}} f_U(x) dx = \left(1 - \frac{\alpha}{2}\right) - \frac{\alpha}{2} = (1 - \alpha)$$

i.e.,

$$P\left[\left(\frac{\alpha}{2}\right)^{1/n} \leq \frac{1}{\theta} \max X_i \leq \left(1 - \frac{\alpha}{2}\right)^{1/n}\right] = 1 - \alpha$$

or

$$P\left[\frac{\max X_i}{(1 - \frac{\alpha}{2})^{1/n}} \leq \theta \leq \frac{\max X_i}{(\frac{\alpha}{2})^{1/n}}\right] = 1 - \alpha \quad (6(a))$$

Which implies that  $\left[\frac{\max X_i}{(1 - \frac{\alpha}{2})^{1/n}}, \frac{\max X_i}{(\frac{\alpha}{2})^{1/n}}\right]$  is an  $100(1 - \alpha)\%$  CI for  $\theta$ .

Again, integrating  $f_U(x)$  between  $\alpha^{1/n}$  and 1, we get

$$P[\alpha^{1/n} \leq U \leq 1] = 1 - \alpha$$

or, equivalently

$$P\left[\max X_i \leq \theta \leq \frac{1}{\alpha^{1/n}} \max X_i\right] = 1 - \alpha. \quad (6(b))$$

The last equation gives a second  $100(1 - \alpha)\%$  CI for  $\theta$ , as  $[\max X_i, \frac{1}{\alpha^{1/n}} \max X_i]$ .

(c) For the given data  $\min X_i = 1.2$ ,  $\max X_i = 4.2$  and  $n = 5$ . Therefore 95% (i.e.,  $\alpha = 0.05$ ) CIs for  $\theta$  from (6(a)) and (6(b)) are evaluated  $[4.24, 8.78]$  and  $[4.2, 7.646]$ , respectively.

**Q.7.** Find the probability that a random sample of 25 observations, from a normal population with variance  $\sigma^2 = 6$ , will have a sample variance  $S^2 > 9.1$ .

**Sol.7.** We want to evaluate  $P[S^2 > 9.1]$ . We know that the rv

$$X = \frac{(n-1)S^2}{\sigma^2}$$

has chi-squared distribution with df  $\nu = (n-1) = 24$ . Thus

$$P(S^2 > 9.1) = P\left(\frac{24}{6}S^2 > \frac{24}{6}9.1\right) = P\left(\frac{24}{6}S^2 > 36.4\right).$$

From table  $\chi_{0.05,24}^2 \approx 36.4$ , therefore  $P(S^2 > 9.1) = 0.05$ .