

# Probability and Statistics

## \* Random Experiment

↳ An experiment

↳ More than one outcome.

↳ Not possible to predict exact outcome in advance.

## \* Set of all possible outcomes of a random experiment is called sample space.

Ex: Throwing a dice;  $S = \{1, 2, 3, 4, 5, 6\}$

## \* Event: subset of sample space.

Throwing a dice:-

Getting Even no.  $E = \{2, 4, 6\}$

Getting odd no.  $E = \{1, 3, 5\}$

## \* Calculate Probability:

If a random experiment has  $n$  possible outcomes out of which  $m$  outcomes are in favour of occurrence of an event  $E$ .

then  $P(E) = \frac{\text{No. of favourable cases for occurrence}}{\text{No. of total outcomes}}$

$$= \frac{n(E)}{n(S)}$$

Ex.  $S = \{1, 2, 3, 4, 5, 6\}$

$E_1 = \{2, 4, 6\}$

$E_2 = \{1, 3, 5\}$

$E_3 = \text{No. greater than } 5$

$E_4 = \text{No. less than } 7$

$$\Rightarrow P(E_1) = \frac{3}{6} \quad P(E_2) = \frac{3}{6} \quad P(E_3) = \frac{0}{6} \quad P(E_4) = \frac{6}{6}$$

NOTE :-  $0 \leq P(E) = \frac{n(E)}{n(S)} \leq 1$

$$0 \leq P(E) \leq 1$$

for any event E of a random experiment,

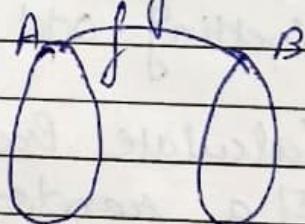
$$0 \leq P(E) \leq 1$$

$P(E)=0 \leftarrow$  Impossible event       $\rightarrow P(E)=1$  Sure event.

\* Function: A function f from a set A to B is rule that maps every element of set A to unique element of set B.

$$A = \{1, 2, 3\}$$

$$B = \{a, b, c\}$$



$$A \times B = \{(1, a), (1, b), (1, c), (2, a), (2, b), (2, c)\}$$

$$\text{Relation } R_1 = \{(1, a), (2, b), (3, c)\}$$

Also a function

$$\text{Relation } R_2 = \{(1, a), (1, b), (1, c)\}$$

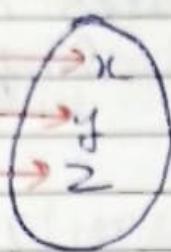
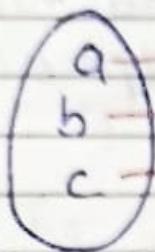
Not a function.

→ All functions are relations but not vice versa

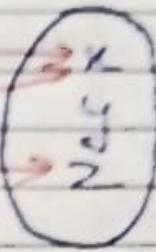
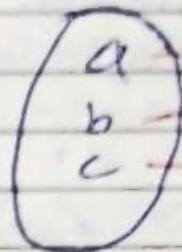
$$P(E^c) = 1 - P(E)$$

### Function :

1 to 1

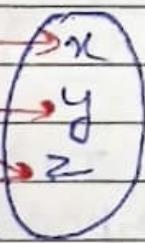
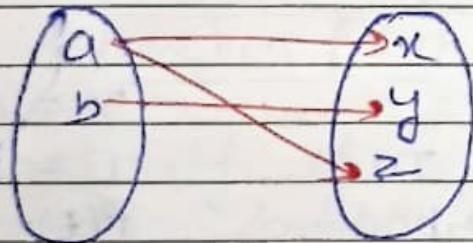


Many to 1

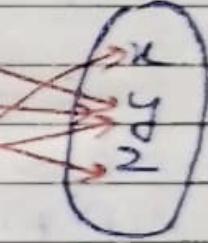
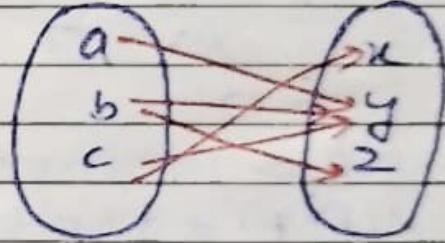


### Non Function:

1 to many



Many to Many



## Unit-1.

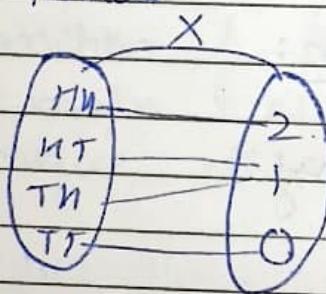
### Random Variable and probability distribution

- **Random Variable :** A random variable is a function that associates a real number with each element of sample space

Ex:- Two coins are tossed,

$$S = \{HH, HT, TH, TT\}$$

Let  $X$ : no. of Heads



$$X: 0 \quad 1 \quad 2$$

$$X=0$$

$$E = TT$$

$$P(E) = \frac{1}{4}$$

$$P(X=0) = \frac{1}{4}$$

$$X=1$$

$$E = HT, TH$$

$$P(E) = \frac{1}{2}$$

$$P(X=1) = \frac{1}{2}$$

$$X=2$$

$$E = HH$$

$$P(E) = \frac{1}{4}$$

$$P(X=2) = \frac{1}{4}$$

$X:$	$0$	$1$	$2$
$P(X=x):$	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$

$\Rightarrow$  Probability distribution

- **Discrete Sample Space:** If a sample space contains finite or countably infinite number of possibilities.

- Discrete Random Variable: A random variable  $X$  is called discrete random variable if its set of possible outcomes is finite or countably infinite.
- Probability function / Probability mass function / Probability distribution function:

The set of ordered pairs  $(x, P(x))$  is pmf or probability distribution function of a discrete random variable  $X$ , if for each outcome  $x$ :

$$(i) \quad P(x) \geq 0$$

$$(ii) \quad \sum_{x \in S} P(x) = 1$$

- Cumulative distribution (cdf) or distribution function  $F(x)$ :

The Cdf of a discrete random variable  $X$  with probability function  $(x, P(x))$  is defined as

$$F(x) = P(X \leq x)$$

$x =$	$x$	0	1	2	
$P(x) =$	$P(X=x)$	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$	$\Rightarrow \text{Pmf}$
Cdf	$F(x)$	$\frac{1}{4}$	$\frac{3}{4}$	1	

$$F(0) = P(X \leq 0) = P(X=0) \Rightarrow \frac{1}{4}$$

$$\begin{aligned} F(1) &= P(X \leq 1) \\ &= P(X=0) + P(X=1) \\ &= \frac{1}{4} + \frac{1}{2} = \frac{3}{4} \end{aligned}$$

$$\begin{aligned} F(2) &= P(X \leq 2) \\ &= P(X=0) + P(X=1) + P(X=2) \end{aligned}$$

$$\therefore \Rightarrow F(1) + P(X=2)$$

$$= \frac{3}{4} + \frac{1}{4} = 1$$

Numerical : A random variable  $X$  has following distribution :

$x$	0	1	2	3	4	5	6	7
$P(x)$	0	$k$	$2k$	$2k$	$3k$	$k^2$	$2k^2$	$7k^2+k$

- ① Evaluate  $k$ .
- ② Evaluate  $P(X < 6)$  &  $P(X \geq 6)$
- ③ Determine the Cdf of  $X$ .
- ④ If  $P(X \leq a) > \frac{1}{2}$

Find the minimum value of  $a$ .

Q) If  $F(4) = \frac{8}{10}$ ,  $F(5) = \frac{5}{10}$ , Find  $P(X=4)$

Sol: (1)  $\sum_n P(n) = 1$

$$0 + K + 2K + 2K + 3K + K^2 + 2K^2 + 3K^2 + K = 1$$

$$10K^2 + 9K - 1 = 0$$

$$K = -1, \frac{1}{10} \checkmark$$

$$\textcircled{2} \rightarrow P(X < 6) = P(X=0) + P(X=1) + P(X=2) + P(X=3) \\ + P(X=4) + P(X=5)$$

$$= 0 + \frac{1}{10} + \frac{2}{10} + \frac{2}{10} + \frac{3}{10} + \frac{1}{100}$$

$$\rightarrow P(X \geq 6) = 1 - P(X < 6)$$

$$\rightarrow P(0 < X < 5) \Rightarrow P(X=1) + P(X=2) + P(X=3) + P(X=4) \\ \Rightarrow \frac{1}{10} + \frac{2}{10} + \frac{2}{10} + \frac{3}{10} = \frac{8}{10}$$

(3)	$x$	0	1	2	3	4	5	6	7
pdf	$P(x)$	0	$\frac{1}{10}$	$\frac{2}{10}$	$\frac{2}{10}$	$\frac{3}{10}$	$\frac{1}{100}$	$\frac{1}{10}$	$\frac{17}{100}$
cdf	$CDF$	0	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{5}{10}$	$\frac{8}{10}$	$\frac{86}{100}$	$\frac{83}{100}$	$\frac{100}{100}$

Q)  $P(X \leq a) > \frac{1}{2}$

Ans.

Q)  $P(X=4) = \frac{8}{10} - \frac{5}{10} = \frac{3}{10}$

(Q)



Three electronic components.  
N - nondefective , D - defective.

$$S = \{ \text{NNN}, \text{NND}, \text{NON}, \text{DNN}, \text{OND}, \\ \{ \text{ODN}, \text{NDD}, \text{DDD} \} \}$$

$X$ : no. of defected items.

⇒

$X$	0	1	2	3	$\leftarrow \text{Pdf} \right.$
$P(X)$	$\frac{1}{8}$	$\frac{3}{8}$	$\frac{3}{8}$	$\frac{1}{8}$	
(df)	$\frac{1}{8}$	$\frac{4}{8}$	$\frac{2}{8}$	$\frac{8}{8}$	

\*

Continuous Random Variable : A random variable  $X$  is said to be continuous random variable if it takes all values in the given interval.

Continuous Sample Space : If it contains infinite number of possibilities equal to number of points on a line segment.

In case of continuous random variable, the probability is calculated for the interval.

$$P(a < x < b) = P(a \leq x \leq b) = P(a \leq x \leq b) = \\ = P(a \leq x \leq b) = \int_a^b f(x) dx$$

where  $f(x)$  is probability density function

\* Probability Density Function:- The function  $f(x)$  is a pdf for continuous random variable over the set of real numbers if

$$(i) f(x) \geq 0 \quad \forall x$$

$$(ii) \int_{-\infty}^{\infty} f(x) dx = 1$$

NOTE - When  $X$  is continuous random variable then,

$$P(X=x) = 0 \quad \text{for all values of } x.$$

Proof :-

$$P(a \leq x \leq b) = \int_a^b f(x) dx$$

now, both  $a$  &  $b$  are equal,  
so,  $P(X=x) = 0$

\* Cumulative distribution function : ( $F(x)$ ) (cdf)  
The cdf  $F(x)$  of a continuous random variable  $X$  with pdf  $f(x)$  is defined as

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt, \quad -\infty < x < \infty$$

$$\underline{\text{NOTE}} - P(a \leq x \leq b) = F(b) - F(a)$$

and  $\frac{dF(x)}{dx} = f(x)$ , if derivative exists.

Proof :- LHS.

$$\begin{aligned}
 & P(a \leq x \leq b) \\
 &= \int_a^b f(x) dx = \int_a^{\infty} f(x) dx + \int_{-\infty}^b f(x) dx \\
 &= \int_{-\infty}^b f(x) dx - \int_{-\infty}^a f(x) dx \\
 &= F(b) - F(a)
 \end{aligned}$$

Ex. Let  $X$  be a continuous random variable with pdf,

$$f(x) = \begin{cases} \frac{x^2}{3} & -1 < x < 2 \\ 0 & \text{elsewhere} \end{cases}$$

(1) Verify  $f(x)$  is pdf

(2) Find  $P(0 < x < 1)$

(3) Find the cdf.

(1)  $\Rightarrow$  (i)  $f(x) \geq 0 \quad \forall x$

$$\begin{aligned}
 & \text{(ii)} \quad \int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^{-1} f(x) dx + \int_{-1}^2 f(x) dx + \int_2^{\infty} f(x) dx \\
 &= \int_{-\infty}^{-1} 0 dx + \int_{-1}^2 \frac{x^2}{3} dx + \int_2^{\infty} 0 dx \\
 &= \left[ \frac{x^3}{9} \right]_{-1}^2 = \frac{8}{9} + \frac{1}{9} - 1 = 0
 \end{aligned}$$

$$\textcircled{2} \Rightarrow P(0 < n < 1)$$

$$\int_0^1 f(n) dn = \int_0^1 \frac{n^2}{3} dn = \left[ \frac{1}{3} \times \frac{n^3}{3} \right]_0^1 \\ = \frac{1}{9} n^3 \Big|_0^1 = \frac{1}{9}$$

$$\textcircled{3} \Rightarrow F(x) = P(X \leq n) = \int_{-\infty}^n f(t) dt$$

~~$$= \int_{-\infty}^{-1} f(t) dt + \int_{-1}^x f(t) dt$$~~

$$F(x) = \int_{-1}^x f(t) dt$$

$$= \int_1^x \frac{t^2}{3} dt = \frac{1}{9} \left[ t^3 \right]_1^x$$

$$F(x) = \frac{1}{9} (x^3 + 1)$$

$$F(x) = \begin{cases} 0 & x < -1 \\ \frac{x^3 + 1}{9} & -1 \leq x \leq 2 \\ 1 & x > 2 \end{cases}$$

$F(0) = 1$   
 $F(0) = F(0) + P(2 < n < 3)$   
 $= 1$

Eg. Let  $X$  be cont. R.V.,

$$f(x) = 6x(1-x) \quad 0 < x < 1$$

① Verify pdf  $\Rightarrow$

$$\text{(i)} = \int_0^1 f(x) dx$$

$$(i) f(x) \geq 0$$

$$= \int_0^1 6x(1-x) dx$$

$$= 6 \int_0^1 (x - x^2) dx$$

$$= 6 \left[ \frac{x^2}{2} \right]_0^1 - 6 \left[ \frac{x^3}{3} \right]_0^1$$

$$= 6 \left( \frac{1}{2} \right) - 6 \left( \frac{1}{3} \right) = 1$$

② Find cdf  $\Rightarrow$

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$$

$$= \int_{-\infty}^x 6t(1-t) dt$$

$$= 6 \left[ \int_{-\infty}^x (t - t^2) dt \right] = 6 \left[ \frac{t^2}{2} \right]_{-\infty}^x$$

## \* Joint Probability distribution

$f(x,y)$   
x & y

$$\textcircled{1} \quad f(x,y) = P(X=x, Y=y)$$

$$\textcircled{2} \quad f(x,y) \geq 0$$

$$\textcircled{3} \quad \sum_x \sum_y f(x,y) = 1$$

$f(x,y)$

$$\textcircled{1} \quad f(x,y) \geq 0$$

$$\textcircled{2} \quad \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dx dy = 1$$

$\Rightarrow$  Let  $X$  and  $Y$  be two discrete random variables  
then the function  $f(x,y)$  is said to be joint  
probability distribution if

$$\textcircled{1} \quad f(x,y) = P(X=x, Y=y)$$

$$\textcircled{2} \quad f(x,y) \geq 0 \quad \forall (x,y)$$

$$\textcircled{3} \quad \sum_x \sum_y f(x,y) = 1$$

$\Rightarrow$  If  $X$  and  $Y$  are continuous random variable  
then  $f(x,y)$  is Joint distribution density  
function if

$$\textcircled{1} \quad P[(x,y) \in A] = \iint_A f(x,y) dx dy$$

where  $A$  is any region in  $XY$  plane.

$$\textcircled{2} \quad f(x,y) \geq 0 \quad \forall (x,y)$$

$$\textcircled{3} \quad \iint_{-\infty}^{\infty} f(x,y) dx dy = 1$$

## \* Marginal distribution:

Let  $X$  and  $Y$  be discrete random variables; then marginal distribution of random variable  $X$  is given,

$$g(x) = \sum_y f(x,y)$$

and Marginal distribution of random variable  $Y$  is given by

$$h(y) = \sum_x f(x,y)$$

If  $X$  and  $Y$  are continuous random variable  
then marginal distribution of  
random variable  $X$  is given by

$$g(x) = \int_{-\infty}^{\infty} f(x,y) dy$$

and Marginal distribution of random variable  $Y$  is given by

$$h(y) = \int_{-\infty}^{\infty} f(x,y) dx$$

conditional Probability  $\Rightarrow$   
 $A, B$

NOTE :  $P(A|B) = \frac{P(A \cap B)}{P(B)}, P(B) > 0$

$$P(B|A) = \frac{P(A \cap B)}{P(A)}, P(A) > 0$$

### \* Conditional distribution:

Let  $X$  and  $Y$  be two discrete or continuous random variable then the conditional distribution of  $Y$  given  $X=x$ ,

$$f(y|x) = \frac{f(x,y)}{g(x)}, g(x) > 0$$

and conditional distribution of  $X$  given  $Y=y$  is

$$f(x|y) = \frac{f(x,y)}{h(y)}, h(y) > 0$$

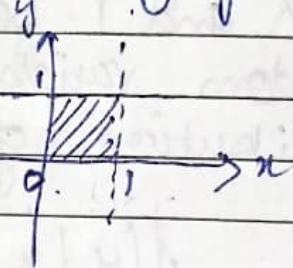
Example: Let  $X$  and  $Y$  be random variable with joint density function.

$$f(x,y) = \begin{cases} \frac{2}{5}(2x+3y), & 0 \leq x \leq 1 \\ 0, & \text{elsewhere} \end{cases}, \quad 0 \leq y \leq 1$$

① Verify  $f(x,y)$  is joint density function

Soln: Clearly,  $f(x,y) \geq 0$

$\forall (x,y)$



$$\text{(ii). } \int \int f(x,y) dx dy$$

$$= \int \int \frac{2}{5}(2x+3y) dx dy$$

$$= \frac{2}{5} \int \left[ x^2 + 3xy \right]_0^1 dy$$

$$= \frac{2}{5} \int (1+3y) dy$$

$$= \frac{2}{5} \left[ y + \frac{3y^2}{2} \right]_0^1 = \frac{2}{5} \times \frac{5}{2} = 1$$

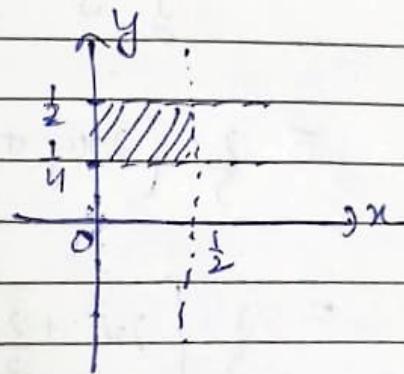
② Find  $P[(x,y) \in A]$  where

$$A = \{(x,y) : 0 < x < \frac{1}{2}$$

$$\frac{1}{4} < y < \frac{1}{2}$$

$\therefore P[(x,y) \in A]$

$$= \int_{\frac{1}{4}}^{\frac{1}{2}} \int_0^{\frac{1}{2}} \frac{2}{5} (2x+3y) dx dy$$



$$= \frac{2}{5} \int_{\frac{1}{4}}^{\frac{1}{2}} [x^2 + 3xy]_0^{\frac{1}{2}} dy$$

$$= \frac{2}{5} \int_{\frac{1}{4}}^{\frac{1}{2}} \left( \frac{1}{4} + \frac{3}{2}y \right) dy = \frac{2}{5} \left[ \frac{1}{4}y + \frac{3}{4}y^2 \right]_{\frac{1}{4}}^{\frac{1}{2}}$$

$$= \frac{2}{5} \times \frac{1}{4} \left[ \frac{1}{2} + \frac{3}{4} - \frac{1}{4} - \frac{3}{16} \right]$$

$$= \frac{1}{10} \left[ 1 - \frac{3}{16} \right] = \frac{13}{160}$$

③ Find the marginal distribution of  $X$ .

$$\begin{aligned} \text{Sol} \Rightarrow g(x) &= \int_0^x f(x,y) dy \\ &= \int_0^x \frac{2}{5} (2x+3y) dy \\ &= \frac{2}{5} \left( 2xy + \frac{3}{2} y^2 \right)_0^x \\ &= \frac{2}{5} \left[ 2x \cdot \frac{3}{2} \right] = \frac{2}{5} \left[ \frac{4x+3}{2} \right] \end{aligned}$$

$$\begin{aligned} g(y) &= \int_0^{\infty} f(x,y) dx \\ &= \int_0^{\infty} \frac{2}{5} (2x+3y) dx \\ &= \frac{2}{5} \left[ \frac{2x^2}{2} + 3yx \right]_0^{\infty} \end{aligned}$$

$$h(y) = \frac{2}{5} [1+3y] = \frac{2+6y}{5}$$

## \* mean / expectation of random variable :-

Let  $X$  be a random variable with probability distribution  $f(x)$ , then mean / expected value of random variable  $X$  is defined as:

$$\mu = E(X) = \sum_n x f(n)$$

if  $X$  is discrete random variable.

$$\text{and } \mu = E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

if  $X$  is continuous random variable.

Ex. Toss two coins :

$$S = \{HH, HT, TH, TT\}$$

$X$  = no. of heads.

$$E(X) = \sum_n n f(n)$$

$$= \sum_{n=0}^3 n f(n)$$

$$= 0 \times f(0) + 1 \times f(1) + 2 \times f(2)$$

$$= 0 + \frac{1}{2} + 2 \times \frac{1}{4} = 1$$

$X: 0 \quad 1 \quad 2$	$P(X=x): \frac{1}{4} \quad \frac{1}{2} \quad \frac{1}{4}$
$f(0) = P(X=0) = \frac{1}{4}$	

$f(1) = P(X=1) = \frac{1}{2}$	
$f(2) = P(X=2) = \frac{1}{4}$	

$f(2) = P(X=2) = \frac{1}{4}$	
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Ex. Let  $X$  be r.v. that denotes life of an electronic device in hours p.d.f:

$$f(x) = \begin{cases} \frac{20000}{x^3}, & x > 100 \\ 0, & \text{elsewhere} \end{cases}$$

Find the expected life of this device

$$\Rightarrow E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_{-\infty}^{\infty} x \times \frac{20000}{x^3} dx$$

$$= \int_{100}^{\infty} \frac{20000}{x^2} dx$$

$$E(X) = 20000 \int_{100}^{\infty} x^{-2} dx$$

$$= -20000 \left[ \frac{1}{x} \right]_{100}^{\infty}$$

$$= 200 \text{ hrs}$$

Expected life of device.

Def<sup>n</sup>: Let  $X$  be a random variable with probability distribution  $f(n)$  the expected value of random variable.

$$\mu_{g(x)} = E(g(x)) = \sum_n g(n) f(n)$$

if  $X$  is discrete random variable

and

$$\mu_{g(x)} = E(g(x)) = \int_{-\infty}^{\infty} g(x) f(x) dx$$

if  $X$  is continuous random variable

Ex

$$X: 0 \quad 1 \quad 2$$

~~$f(n)$~~

$$f(x) = P(X=x): \frac{1}{4} \quad \frac{1}{2} \quad \frac{1}{4}$$

Find the expected value of  $g(x) = 4x - 1$

$$\Rightarrow E[g(x)] = \sum_x g(x) f(n)$$

$$= \sum_{n=0}^2 (4n - 1) f(n)$$

$$= (4 \times 0 - 1) f(0) + (4 \times 1 - 1) f(1) + (4 \times 2 - 1) f(2)$$

$$= 3$$

Ques. Let  $X$  be a random variable with density function

$$f(x) = \begin{cases} \frac{x^2}{3} & -1 < x < 2 \\ 0 & \text{elsewhere} \end{cases}$$

Find expected value of  $g(X) = 4X + 3$

$$\text{Soln: } E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$$

$$= \int_{-1}^2 (4x+3) \frac{x^2}{3} dx$$

$$= 8.$$

### \* Properties of mean:-

i. If  $a$  and  $b$  are constants. then,

$$E[ax+b] = a E(x) + b$$

Note: (i). If  $a=0$  then,  $E(b) = b$

(ii) If  $b=0$  then,  $E(ax) = a E(x)$

2.  $E[g(x) \pm h(x)] = E[g(x)] \pm E[h(x)]$

Ex  $g(x) = 4x + 3$

$$\begin{aligned}E[g(x)] &= E[4x + 3] \\&= 4E(x) + 3\end{aligned}$$

Ex  $y = (x-1)^2$

$$\begin{aligned}E[y] &= E[(x-1)^2] \\&= E[x^2 + 1 - 2x] \\&= E(x^2) + E(1) - 2E(x)\end{aligned}$$

NOTE.

$$E[g(x, y)] = \sum_n \sum_y g(x, y) f(x, y)$$

$$E[g(x, y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

Ques. Find  $E(Y/X)$ , for the joint density function:

$$f(x, y) = \begin{cases} n(1+3y^2) & 0 < x < 2 \\ 0 & 0 < y < 1 \\ 0 & \text{elsewhere} \end{cases}$$

Sol. Here  $g(x, Y) = \frac{Y}{X}$

$$E[g(X, Y)] = E\left(\frac{Y}{X}\right)$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

$$= \int_0^1 \int_0^2 \frac{y}{x} \times n(1+3y^2) dx dy$$

$$E\left(\frac{Y}{X}\right) = \frac{1}{4} \int_0^1 \int_0^2 (y + 3y^3) dx dy$$

$$= \frac{1}{4} \int_0^1 \left[ yx + 3y^3 x \right]_0^2 dy$$

$$= \frac{1}{2} \left[ \frac{y^2}{2} + \frac{3}{4} y^4 \right]_0^1 = \frac{1}{2} \left[ \frac{1}{2} + \frac{3}{4} \right] = \frac{5}{8}$$

$$\text{Ex: } g(x, y) = x^2 + y$$

$$E[g(x, y)] = ?$$

Use  
Previous ex.)  
for poly

$$\Rightarrow g(x, y) = x^2 + y$$

$$E[g(x, y)] = \int_0^1 \int_0^2 (x^2 + y) (x + 3y^2) dx dy$$

$$= \int_0^1 \int_0^2 (x^2 + y) (x + 3y^2) dx dy$$

$$= \frac{1}{4} \int_0^1 \int_0^2 (x^3 + yx + 3y^2x^3 + 3y^3) dx dy$$

$$= \frac{1}{4} \int_0^1 \left[ \frac{x^4}{4} + yx^2 + 3y^2x^4 + 3y^3x^2 \right]_0^2 dy$$

$$= \frac{1}{4} \int_0^1 (4 + 2y + 12y^2 + 6y^3) dy$$

$$= \frac{1}{4} \left[ 4y + \frac{2y^2}{2} + \frac{12y^3}{3} + \frac{6y^4}{4} \right]_0^1$$

$$= \frac{1}{4} \left[ 4 + 1 + 4 + \frac{3}{2} \right] = \frac{1}{4} \left[ 9 + \frac{3}{2} \right]$$

$$= \frac{21}{8}$$

\* Variance of a Random Variable: Let  $X$  be a random variable with probability distribution  $f(x)$ , then variance of  $X$  is defined as (with mean  $\mu$ )

$$\sigma^2 = E[(X-\mu)^2]$$

If  $X$  is discrete r.v. then,

$$\sigma^2 = E[(X-\mu)^2] = \sum_w (x-w)^2 f(w) \\ (\text{as } E(X) = \sum w f(w))$$

If  $X$  is continuous R.V. then,

$$\sigma^2 = E[(X-\mu)^2] = \int_{-\infty}^{\infty} (x-\mu)^2 f(x) dx$$

NOTE- The variance of R.V.  $X$  is

$$\sigma^2 = E[X^2] - [E[X]]^2 \\ = E[X^2] - [\mu]^2$$

NOTE- The positive square root of variance is called standard deviation.

Ex. Let  $X$  be a R.V. with prob. distribution as

$X$	0	1	2	3
$f(x) = P(X=x)$	0.5	0.38	0.10	0.01

Calculate the variance of  $X$ .

Sol<sup>n</sup>

$$\begin{aligned} \mu = E(X) &= \sum_{x=0}^3 x f(x) \Rightarrow \\ &= 0 \times f(0) + 1 \times f(1) + 2 \times f(2) + 3 \times f(3) \\ &= 0 + 1 \times 0.38 + 2 \times 0.10 + 3 \times 0.01 \\ &= 0.61 \end{aligned}$$

Now,

$$\begin{aligned} E[X^2] &= \sum_{x=0}^3 x^2 f(x) \\ &= 0 \cdot 38 + 4 \times 0.10 + 9 \times 0.01 \\ &= 0.87 \end{aligned}$$

$$\begin{aligned} \sigma^2 &= E[X^2] - [\mu]^2 \\ &= 0.87 - 0.3721 = 0.4979 \end{aligned}$$

Ex Let  $X$  be a R.V. with density function

$$f(x) = \begin{cases} 2(x-1) & 1 < x < 2 \\ 0 & \text{elsewhere} \end{cases}$$

Find mean & variance.

Sol<sup>n</sup>:

$$\begin{aligned} \mu = E(X) &= \int_{-\infty}^{\infty} x f(x) dx \Rightarrow \int_{-\infty}^{\infty} 2x(x-1) dx \\ &= \int_1^2 (2x^2 - 2x) dx \\ &= \int_1^2 \left[ \frac{2x^3}{3} - x^2 \right] dx \Rightarrow \frac{2}{3} [8-1] - [4-1] \\ &= \frac{2}{3} (7) - 3 \Rightarrow \frac{14}{3} - 3 \end{aligned}$$

$$E[X^2] = \int_1^2 2x^2 (n-1) dx$$

$$\int_1^2 (2x^3 - 2x^2) dx \Rightarrow \int_1^2 \frac{2x^4}{4} - \frac{2x^3}{3}$$

$$\Rightarrow \frac{1}{2} [16-1] = \frac{2}{3} [?]$$

$$= \frac{15}{2} - \frac{14}{3} \Rightarrow \frac{17}{6}$$

$$\sigma^2 \rightarrow \frac{17}{6} - \frac{25}{9} \Rightarrow 153 - 150 \Rightarrow \frac{3}{54}$$

\*  $\sigma_{g(x)}^2 = E[(g(x) - \mu_{g(x)})^2]$

If  $X$  is ~~discrete~~ R.V., discrete R.V.

$$\sigma_{g(x)}^2 = \sum_n (g(n) - \mu_{g(x)})^2 f(x)$$

If  $X$  is continuous R.V.

$$\sigma_{g(x)}^2 = \int_{-\infty}^{\infty} [g(x) - \mu_{g(x)}]^2 f(x) dx$$

Ex. Calculate the variance of  $g(x) = 2x+3$   
where  $X$  is R.V. with prob. distribution

$X$	0	1	2	3
$f(x) = P(X=x)$	$1/4$	$1/8$	$1/2$	$1/8$

$$\Rightarrow \mu_{g(x)} = E[g(x)]$$

$$E(X) = \sum_{x=0}^3 x f(x)$$

$$\Rightarrow 0 \times \frac{1}{4} + 1 \times \frac{1}{8} + 2 \times \frac{1}{2} + 3 \times \frac{1}{8} = \frac{3}{2}$$

$$\mu_{g(x)} = E[2x+3] \Rightarrow 2E[X] + 3 = 1$$

$$E[X^2] = \sum_{n=0}^3 x^2 f(x)$$

$$= 0 \times \frac{1}{4} + 1 \times \frac{1}{8} + 4 \times \frac{1}{2} + 9 \times \frac{1}{8} +$$

$$= \frac{10}{8} + 2 = \frac{26}{8} = \frac{13}{4}$$

$$V(g(x)) = E[(g(x)) - \mu_{g(x)}]^2$$

$$= E[(2x+3-1)^2]$$

$$= E[(2x-3)^2] = E[4x^2+9-12x]$$

$$= E[4X^2]$$

$$= 4E[X^2] + 9 - 12E[X]$$

$$= 4 \times \frac{13}{4} + 9 - 12 \times \frac{3}{2} = 13 + 9 - 18 = 4$$

Ans  
~~143~~

$$Ex \quad f(x) = \begin{cases} 2(x-1) & 1 < x < 2 \\ 0 & \text{elsewhere} \end{cases}$$

$$g(u) = 2x+1 \quad E(x) = 2 \int_1^2 u(x-1) du$$

Soln:  $\sigma_{g(x)} \Rightarrow$

$$\mu_{g(x)} = 2E(x) + 1$$

$$= 2 \times \frac{5}{3} + 1 = \frac{10}{3} + 1 = \frac{13}{3}$$

$$= E[(2u+1 - \frac{13}{3})^2]$$

$$E(x^2) = 2 \int_1^2 u^2 du$$

$$= E[(2u - \frac{10}{3})^2]$$

$$= 2 \left( \frac{u^4}{4} - \frac{u^3}{3} \right)$$

$$= E[4u^2 + \frac{100}{9} - \frac{40}{3}u]$$

$$= \frac{17}{6}$$

$$= 4E(x)^2 + \frac{100}{9} - \frac{40}{3}E(x)$$

$$= 4 \times \frac{17}{6} + \frac{100}{9} - \frac{40}{3} \times \frac{5}{3}$$

$$= \frac{68}{6} + \frac{100}{9} - \frac{200}{9}$$

$$= \frac{68}{6} - \frac{100}{9}$$

$$= \frac{2}{9}$$

(Q) If  $X$  and  $Y$  are R.V. with joint prob. dist.  $f(x, y)$  then  $E[g(x, y)] \leq \sum \sum g(x, y) f(x, y)$  and

$$E[g(x, y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

$\Rightarrow$  If we have  $f(x, y)$  given,  
to find marginal dist.  
 $g(x) = \sum_y f(x, y)$  or  $\int_{-\infty}^{\infty} f(x, y) dy$

$$E(x) = \sum_x g(x) \quad \text{or} \quad \int_{-\infty}^{\infty} x g(x) dx$$

Also,  $h(y) = \sum_x f(x, y)$  or  $\int_{-\infty}^{\infty} f(x, y) dx$

$$E(y) = \sum_y y h(y) \quad \text{or} \quad \int_{-\infty}^{\infty} y h(y) dy.$$

## \* Covariance :-

NOTE: If  $X$  and  $Y$  are two random variables with joint probability distribution  $f(x,y)$  then,

$$E[X] = \sum_n n g(n)$$

and  $E[Y] = \sum_y y h(y)$

where  $g(n)$  and  $h(y)$  are marginal distribution of  $X$  and  $Y$ .

If  $X$  and  $Y$  are continuous random variable then,

$$E[X] = \int_{-\infty}^{\infty} x g(x) dx$$

and

$$E[Y] = \int_{-\infty}^{\infty} y h(y) dy$$

Covariance: Let  $X$  and  $Y$  be two random variables with joint probability dist.  $f(x,y)$  then covariance of  $X$  &  $Y$  is given by

$$\text{Cov}(X, Y) = \sigma_{XY} = E[(X - \mu_X)(Y - \mu_Y)]$$

If  $X$  and  $Y$  are discrete random variables  
then,

$$\sigma_{XY} = E[(X - \mu_X)(Y - \mu_Y)]$$

$$= \sum_x \sum_y (x - \mu_X)(y - \mu_Y) f(x, y)$$

If  $X$  and  $Y$  are continuous random variables.  
then,

$$\sigma_{XY} = \iint_{-\infty}^{\infty} (x - \mu_X)(y - \mu_Y) f(x, y) dx dy.$$

NOTE: When  $X$  &  $Y$  are independent then  $\sigma_{XY} = 0$   
but converse is not true

If  $X$  and  $Y$  are random variables with joint  
probability distribution  $f(x, y)$  then,

$$\sigma_{XY} = E[XY] - E[X]E[Y]$$

$$= E[XY] - \mu_X \mu_Y$$

Ex:- Two random variables  $X$  and  $Y$  with joint probability density function.

$$f(x, y) = \begin{cases} 2-x-y & 0 \leq x \leq 1 \\ 0 & 0 \leq y \leq 1 \\ 0 & \text{elsewhere} \end{cases}$$

Find covariance of  $X$  &  $Y$ .

Sol:  $\sigma_{XY} = E[XY] - E[X]E[Y]$

Marginal distribution of  $X$  is given by

$$g(x) = \int_0^1 f(x, y) dy$$

$$= \int_0^1 (2-x-y) dy = \left[ 2y - xy - \frac{y^2}{2} \right]_0^1$$

$$g(x) = \frac{3-2x}{2}$$

Also, Marginal distribution of  $Y$ ,

$$h(y) = \int_0^1 f(x, y) dx$$

$$= \int_0^1 (2-x-y) dx$$

$$h(y) = \frac{3-2y}{2}$$

$$E[X] = \int_0^1 x g(x) dx = \int_0^1 x \left(\frac{3-2x}{2}\right) dx$$

$$= \frac{1}{2} \left[ \frac{3x^2}{2} - \frac{2x^3}{3} \right]_0^1$$

$$= \frac{1}{2} \left[ \frac{3}{2} - \frac{2}{3} \right] = \frac{5}{12}$$

$$E[Y] = \int_0^1 y f(y) dy = \int_0^1 y \left(\frac{3-2y}{2}\right) dy = \frac{5}{12}$$

~~$E[g(XY)]$~~

$$E[XY] = \int_0^1 \int_0^1 xy f(x, y) dx dy$$

$$= \int_0^1 \int_0^1 xy (2-x-y) dx dy$$

$$= \int_0^1 \left[ \frac{2x^2y}{2} - \frac{x^3y}{3} - \frac{x^2y^2}{2} \right]_0^1 dy$$

$$= \int_0^1 \left( y - \frac{y^2}{3} - \frac{y^3}{6} \right) dy$$

$$= \left( \frac{y^2}{2} - \frac{y^3}{6} - \frac{y^4}{4} \right)_0^1 = \frac{1}{2} - \frac{1}{6} - \frac{1}{6} = \frac{1}{6}$$

$$\sigma_{XY} = E[XY] - E[X]E[Y]$$

$$= \frac{1}{6} - \frac{1}{12} \times \frac{5}{12} \quad \text{-ve sign}$$

$$= -\frac{1}{144}$$

{ means if  $X$  is taking large values  
 $Y$  will take small values  
 $\therefore X \propto \frac{1}{Y}$

## \* Chebyshov's Theorem:

Let  $X$  be a random variable with mean  $\mu$  and standard deviation  $\sigma$ , then for any positive number  $k$ , the probability that a value of  $X$  lie in interval  $[\mu - k\sigma, \mu + k\sigma]$  is at least  $1 - \frac{1}{k^2}$

that is,

$$P(\mu - k\sigma < X < \mu + k\sigma) \geq 1 - \frac{1}{k^2}$$

Ex. A random variable  $X$  has  $\mu = 8$  and  $\sigma^2 = 9$ . Then find ①  $P(-4 < x < 20)$

$\Rightarrow$

$$\begin{aligned} \mu - k\sigma &= -4 & \mu + k\sigma &= 20 \\ 8 - k(3) &= -4 & \text{1ve value} \\ & \text{---} & \end{aligned}$$

$$k = 4$$

$$\text{Ans: } 1 - \frac{1}{k^2} = 1 - \frac{1}{4^2} = \frac{15}{16}$$

(2) Find:  $P(|X-8| \geq 6)$

$$\begin{aligned} &= 1 - P(|X-8| < 6) \\ &= 1 - P(-6 < X-8 < 6) \\ &= 1 - P(8-6 < X < 8+6) \\ &= 1 - P(2 < X < 14) \end{aligned}$$

NOTE:

$$|x| < a$$

$$-a < x < a$$

# Formulas : Unit 1

Date: \_\_\_\_\_  
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①

pdf

$$P(x) \geq 0$$

$$cdf [ F(x) = P(X \leq x) ]$$

$$\sum_n P(x) = 1$$

②

$$P(a < x < b) = P(a \leq x < b) = P(a < x \leq b) = P(a \leq x \leq b) \\ = \int_a^b f(x) dx$$

③

$$\int_{-\infty}^{\infty} f(x) dx \geq 0$$

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

④

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$$

$$P(a < x \leq b) = F(b) - F(a)$$

⑤

$$f(x, y) = P(X=x, Y=y)$$

$$f(x, y) \geq 0 \quad \forall (x, y)$$

$$\sum_x \sum_y f(x, y) = 1$$

⑥

$$f(x, y) \geq 0 \quad \forall (x, y)$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

(7) marginal distribution

$$g(x) = \sum_y f(x,y) \quad g(x) = \int_{-\infty}^{\infty} f(x,y) dy$$

$$h(y) = \sum_x f(x,y) \quad h(y) = \int_{-\infty}^{\infty} f(x,y) dx$$

(8) Conditional distribution

$$f(y|x) = \frac{f(x,y)}{g(x)}, \quad g(x) > 0$$

$$f(x|y) = \frac{f(x,y)}{h(y)}, \quad h(y) > 0$$

(9) mean / expectation

$$\mu = E(x) = \sum_n x f(x)$$

$$\mu = E(x) = \int_{-\infty}^{\infty} x f(x) dx$$

(10)

$$\mu_{g(x)} = E(g(x)) = \sum_n g(x) f(x)$$

$$\mu_{g(x)} = E(g(x)) = \int_{-\infty}^{\infty} g(x) f(x) dx$$

(11)  $E[ax+b] = aE(x) + b$

where  $a+b \Rightarrow \text{const.}$

$$E[g(x) \pm h(x)] = E[g(x)] \pm E[h(x)]$$

(12)  $E[g(x, y)] = \sum_x \sum_y g(x, y) f(x, y)$

$$E[g(x, y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

(13) Variance

$$\sigma^2 = E[(x-\mu)^2]$$

$$\sigma^2 = E[(x-\mu)^2] = \sum_n (x-\mu)^2 f(n)$$

$$\sigma^2 = E[(x-\mu)^2] = \int_{-\infty}^{\infty} (x-\mu)^2 f(x) dx.$$

$$\begin{aligned} \sigma^2 &= E[x^2] - [E[x]]^2 \\ &= E[x^2] - [\mu]^2 \end{aligned}$$

$\sigma$   Standard deviation

$$(14) \quad \sigma_{g(x)}^2 = E \left[ (g(x) - \mu_{g(x)})^2 \right]$$

$$\sigma_{g(x)}^2 = \sum_n (g(n) - \mu_{g(x)})^2 f(n)$$

$$\sigma_{g(x)}^2 = \int_{-\infty}^{\infty} [g(x) - \mu_{g(x)}]^2 f(x) dx$$

$$(15) \quad E[X] = \sum_n x g(n) \quad E[X] = \int_{-\infty}^{\infty} x g(x) dx$$

$$E[Y] = \sum_y y h(y) \quad E[Y] = \int_{-\infty}^{\infty} y h(y) dy$$

(16) Covariance

$$\text{Cov}(X, Y) = \sigma_{XY} = E[(X - \mu_X)(Y - \mu_Y)]$$

$$\sigma_{XY} = \sum_x \sum_y (x - \mu_x)(y - \mu_y) f(x, y)$$

$$\sigma_{XY} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - \mu_x)(y - \mu_y) f(x, y) dx dy$$

$$(17) \quad \sigma_{XY} = E[XY] - E[X]E[Y]$$

$$= E[XY] - \mu_X \mu_Y$$

(18) Chebyshev's Theorem

$$P(\mu - k\sigma < X < \mu + k\sigma) \geq 1 - \frac{1}{k^2}$$