

BEFORE YOU START WRITING, check the page numbers at the bottom to confirm that all the pages are present. You have **ONE** hour to complete this exam. You must explain your work clearly to get credit for your answer. Use the available space judiciously.

Name: \_\_\_\_\_ Reg. No: \_\_\_\_\_ Tutorial section: \_\_\_\_\_

Question:	1	2	3	4	5	Total
Points:	6	3	3	2	6	20
Score:						

1. People arrive at a queue according to the following scheme: During each minute of time either 0 or 1 person arrives. In a minute, the probability that 1 person arrives is  $p$  and that no person arrives is  $q = 1 - p$ . Answer the following questions (No explanation required):

(2 marks)

- (a) Let  $C$  be the number of customers arriving in the first 10 minutes.

- $P(C = 2) = \binom{10}{2} p^2 q^8$
- $E[C] = 10p$

(2 marks)

- (b) Let  $W$  be the time (in minutes) until the first person arrives.

- $P(W = 5) = (1 - p)^4 p$
- $E[W] = 1/p$

(2 marks)

- (c) Let  $T$  be the time (in minutes) until 4 people arrive.

- $P(T = 10) = \binom{9}{3} p^4 q^6$
- $E[T] = 4/p$

- (3 marks) 2. Consider two independent random variables:  $X \sim Poi(\lambda)$  and  $Y \sim Poi(\mu)$  for  $\lambda, \mu > 0$ . Determine the probability mass function of  $Z = X + Y$ .

**Solution:**

$$\begin{aligned} P(Z = k) &= \sum_{r=0}^k P(X = r)P(Y = k - r) \\ &= \sum_{r=0}^k e^{-\lambda} \frac{\lambda^r}{r!} e^{-\mu} \frac{\mu^{k-r}}{(k-r)!} \\ &= \frac{e^{-(\lambda+\mu)}}{k!} \sum_{r=0}^k \frac{k!}{r!(k-r)!} \lambda^r \mu^{k-r} \\ &= \frac{e^{-(\lambda+\mu)} (\lambda + \mu)^k}{k!} \end{aligned}$$

This implies  $Z \sim Poi(\lambda + \mu)$ .

3. Let  $X \sim Unif([0, 1])$ .

- (2 marks) (a) Determine the probability density function of the random variable  $X^2$ .  
(1 mark) (b) Compute  $E[X^2]$ .

**Solution:** (a) We first determine the distribution of  $X^2$ . For  $0 \leq x \leq 1$ ,

$$P(X^2 \leq x) = P(X \leq \sqrt{x}) = \sqrt{x}$$

This implies,  $f_{X^2}(x) = \frac{1}{2\sqrt{x}}$  for  $x \in [0, 1]$  and 0 otherwise.

(b)  $E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx$ . So,

$$E[X^2] = \int_0^1 x^2 dx = 1/3$$

- (2 marks) 4. Let  $Y \sim \mathcal{N}(3, 9)$  and  $\phi(x) = P(Z \leq x)$ , where  $Z \sim \mathcal{N}(0, 1)$ . Compute  $P(Y > 3 | Y > 1)$  in terms of  $\phi(2/3)$ .

**Solution:**

$$\begin{aligned}
 P(Y > 3 | Y > 1) &= \frac{P(Y > 3, Y > 1)}{P(Y > 1)} \\
 &= \frac{P(Y > 3)}{P(Y > 1)} \\
 &= \frac{P\left(\frac{Y-3}{3} > 0\right)}{P\left(\frac{Y-3}{3} > -2/3\right)} \\
 &= \frac{P(Z > 0)}{P(Z > -2/3)} \\
 &= \frac{1}{2(1 - \phi(-2/3))} \\
 &= \frac{1}{2\phi(2/3)}
 \end{aligned}$$

5. Define  $C(X, Y) = E[(X - E[X])(Y - E[Y])]$ . Show that:

- (2 marks) (a)  $C(X, Y) = E[XY] - E[X]E[Y]$ .
- (2 marks) (b) For all  $a, b \in \mathbb{R}$ ,  $a^2 E[X^2] + 2abE[XY] + b^2 E[Y^2] \geq 0$ .
- (2 marks) (c)  $E[XY]^2 \leq E[X^2]E[Y^2]$ .  
 (Hint: For  $A, B, C \in \mathbb{R}$ ,  $Ar^2 + 2Br + C \geq 0 \forall r \in \mathbb{R}$  implies  $B^2 \leq AC$ )

**Solution:**

- (a)  $E[(X - E[X])(Y - E[Y])] = E[XY] - E[XE[Y]] - E[YE[X]] + E[X]E[Y] = E[XY] - E[X]E[Y]$ .
- (b) Note that  $E[(aX + bY)^2] \geq 0$ . Expand and use linearity of Expectation to conclude.
- (c) In part (b), take  $b = 1$  and consider the quadratic in variable  $a$ . Then,

$$a^2 E[X^2] + 2aE[XY] + E[Y^2] \geq 0$$

implies the result.