

$$\underbrace{6+14}_{20} + \underbrace{8+5+10}_{23} = 43$$

## Solutions Assignment 4

Ex. 1: Who invented the t-distribution? What was his occupation?

William Gosset, (Head) Brewer (of Guinness)

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Ex. 2: Let X denote the sample of undergraduate students taking the midterm and Y the sample of graduates students taking the midterm. Assume that both samples are i.i.d. normal with means  $\mu_1$  and  $\mu_2$  respectively and the same (unknown) variance. Test the hypothesis  $H_0: \mu_1 = \mu_2$  against the alternative using the test statistic from the lectures. You are given the following values:  $m=90$ ,  $n=8$ ,  $\bar{X}=29.405$ ,  $\bar{Y}=31.38$ ,  $S_X^2=7875.41$ ,  $S_Y^2=671.88$ ,  $\alpha=0.25$ . What do you conclude? What is the p-value in this case?

We conduct a 2-sample t-test. We calculate

$$U_{90,8} = \frac{(90+8-2)^{1/2} (29.405 - 31.38)}{\sqrt{\left(\frac{1}{90} + \frac{1}{8}\right) (7875.41 + 671.88)}} = -0.558$$

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$$T_{90+8-2}^{-1}(1-0.125) = T_{96}^{-1}(0.875) = 1.157$$

$$\Rightarrow \text{Fail to reject } H_0 \Rightarrow \text{p-value } T_{96}(-0.558) + 1 - T_{96}(0.558) = 0.56 !!!$$

Ex. 3: Show that the two-sample test statistic U has t-distribution with  $m+n-2$  degrees of freedom, if  $\mu_1 = \mu_2$ .

See textbook Theorem 9.6.1 & proof

Not graded

9.1.2

*1 point for each sentence  
(if students directly conclude correctly, 2 pts. each) (4)*

- (a) We know that if  $0 < y < \theta$ , then  $\Pr(Y_n \leq y) = (y/\theta)^n$ . Also, if  $y \geq \theta$ , then  $\Pr(Y_n \leq y) = 1$ . Therefore, if  $\theta \leq 1.5$ , then  $\pi(\theta) = \Pr(Y_n \leq 1.5) = 1$ . If  $\theta > 1.5$ , then  $\pi(\theta) = \Pr(Y_n \leq 1.5) = (1.5/\theta)^n$ .

- (b) The size of the test is

$$\alpha = \sup_{\theta \geq 2} \pi(\theta) = \sup_{\theta \geq 2} \left( \frac{1.5}{\theta} \right)^n = \left( \frac{1.5}{2} \right)^n = \left( \frac{3}{4} \right)^n.$$

*not expected*

(2)

9.1.4

The null hypothesis  $H_0$  is simple. Therefore, the size  $\alpha$  of the test is  $\alpha = \Pr(\text{Rejecting } H_0 \mid \mu = \mu_0)$ . When  $\mu = \mu_0$ , the random variable  $Z = n^{1/2}(\bar{X}_n - \mu_0)$  will have the standard normal distribution. Hence, since  $n = 25$ ,

$$\alpha = \Pr(|\bar{X}_n - \mu_0| \geq c) = \Pr(|Z| \geq 5c) = 2[1 - \Phi(5c)].$$

Thus,  $\alpha = 0.05$  if and only if  $\Phi(5c) = 0.975$ . It is found from a table of the standard normal distribution that  $5c = 1.96$  and  $c = 0.392$ .

(1)

(1)

(8)

(1)

(1)

### 9.1.6

If  $H_0$  is true, then  $X$  will surely be smaller than 3.5. If  $H_1$  is true, then  $X$  will surely be greater than 3.5. Therefore, the test procedure which rejects  $H_0$  if and only if  $X > 3.5$  will have probability 0 of leading to a wrong decision, no matter what the true value of  $\theta$  is.

(2)

### 9.2.8

- (a) The p.d.f.'s  $f_0(x)$  and  $f_1(x)$  are as sketched in Fig. S.9.3. Under  $H_0$  it is impossible to obtain a value of  $X$  greater than 1, but such values are possible under  $H_1$ . Therefore, if a test procedure rejects  $H_0$  only if  $x > 1$ , then it is impossible to make an error of type 1, and  $\alpha(\delta) = 0$ . Also,

$$\beta(\delta) = \Pr(X < 1 | H_1) = \frac{1}{2}.$$

(Any value (1,2) instead of 1 is also fine.)

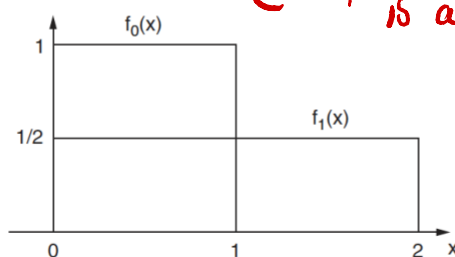


Figure S.9.3: Figure for Exercise 8a of Sec. 9.2.

- (b) To have  $\alpha(\delta) = 0$ , we can include in the critical region only a set of points having probability 0 under  $H_0$ . Therefore, only points  $x > 1$  can be considered. To minimize  $\beta(\delta)$  we should choose this set to have maximum probability under  $H_1$ . Therefore, all points  $x > 1$  should be used in the critical region.

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### 9.5.2 (a,b only)

When  $\mu_0 = 20$ , the statistic  $U$  given by Eq. (9.5.2) has a  $t$  distribution with 8 degrees of freedom. The value of  $U$  in this exercise is 2.

(a) We would reject  $H_0$  if  $U \geq 1.860$ . Therefore, we reject  $H_0$ .

(b) We would reject  $H_0$  if  $U \leq -2.306$  or  $U \geq 2.306$ . Therefore, we don't reject  $H_0$ .

$$(a) U = 9^{1/2} \frac{22-20}{\left(\frac{72}{8}\right)^{1/2}} = 2 \quad T_8^{-1}(0.95) = 1.860 \Rightarrow \text{Reject}$$

$$(b) T_8^{-1}(0.975) = 2.306 \Rightarrow \text{Fail to reject}$$

### 9.7.7

(a) Here,  $\bar{X}_m = 84/16 = 5.25$  and  $\bar{Y}_n = 18/10 = 1.8$ . Therefore,  $S_1^2 = \sum_{i=1}^{16} X_i^2 - 16(\bar{X}_m^2) = 122$  and

$$S_2^2 = \sum_{i=1}^{10} Y_i^2 - 10(\bar{Y}_n^2) = 39.6. \text{ It follows that}$$

$$\hat{\sigma}_1^2 = \frac{1}{16} S_1^2 = 7.625 \quad \text{and} \quad \hat{\sigma}_2^2 = \frac{1}{10} S_2^2 = 3.96.$$

(b) If  $\sigma_1^2 = \sigma_2^2$ , the following statistic  $V$  will have the  $F$  distribution with 15 and 9 degrees of freedom:

$$V = \frac{S_1^2/15}{S_2^2/9}.$$

(b) If the test is to be carried out at the level of significance 0.05, then  $H_0$  should be rejected if  $V > 3.01$ . It is found that  $V = 1.848$ . Therefore, we do not reject  $H_0$ .