HW2

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1 Problem 1

1.1 Subproblem 1

Proof.

Firstly we can write $\hat{\sigma}_a^2$ and $\hat{\sigma}_b^2$ as following:

$$\left[\begin{array}{c} \hat{\sigma}_a^2 \\ \hat{\sigma}_b^2 \end{array} \right] = \frac{1}{nq} \sum_{k=1}^n \left[\begin{array}{c} \sum_{j=1}^q (\varepsilon_{q(k-1)+j}^2) \\ (\sum_{j=1}^q \varepsilon_{q(k-1)+j})^2 \end{array} \right]$$

Since $\{\varepsilon_{qk}\}$ is an iid sequence, $\{\sum_{j=1}^{q}(\varepsilon_{q(k-1)+j}^2)\},\{(\sum_{j=1}^{q}\varepsilon_{q(k-1)+j})^2\}$ are also iid sequences. Which enables us to use central limit theorem to obtain the asymptotic distribution.

The expectation is easy to obtain:

$$\mathbb{E}\left[\begin{array}{c} \sum_{j=1}^{q} \left(\varepsilon_{q(k-1)+j}^{2}\right) \\ \left(\sum_{j=1}^{q} \varepsilon_{q(k-1)+j}\right)^{2} \end{array}\right] = \left[\begin{array}{c} q\sigma^{2} \\ q\sigma^{2} \end{array}\right]$$

The variance can be obtained as below:

$$Var(\sum_{j=1}^{q} (\varepsilon_{q(k-1)+j}^{2}))$$

$$= \sum_{j=1}^{q} (Var(\varepsilon_{q(k-1)+j}^{2}))$$

$$= q(A-1)\sigma^{4}$$

(Here I use A to represent $\mathbb{E}x^4$, x is subject to standard normal distribution, in fact A=3;)

$$Var((\sum_{j=1}^{q} \varepsilon_{q(k-1)+j})^{2})$$

$$=qA\sigma^{4} + 6 * C_{q}^{2}\sigma^{4} - q^{2} * \sigma^{4}$$

$$=(Aq + 2q^{2} - 3q)\sigma^{4}$$

and

$$Cov((\sum_{j=1}^{q} \varepsilon_{q(k-1)+j})^{2}, \sum_{j=1}^{q} (\varepsilon_{q(k-1)+j}^{2}))$$

$$= Var(\sum_{j=1}^{q} (\varepsilon_{q(k-1)+j}^{2})) + Cov(\sum_{i=1}^{q} \sum_{j\neq i,j=1}^{q} (\varepsilon_{q(k-1)+i}\varepsilon_{q(k-1)+j}), \sum_{j=1}^{q} (\varepsilon_{q(k-1)+j}^{2}))$$

$$= q(A-1)\sigma^{4}$$

Finally we get:

$$Var\left(\left[\begin{array}{c} \sum_{j=1}^q (\varepsilon_{q(k-1)+j}^2) \\ (\sum_{j=1}^q \varepsilon_{q(k-1)+j})^2 \end{array}\right]\right) = \left[\begin{array}{cc} 2q\sigma^4 & 2q\sigma^4 \\ 2q\sigma^4 & 2q^2\sigma^4 \end{array}\right]$$

Therefore, using CLT,

$$\sqrt{n} \left(\left[\begin{array}{c} \hat{\sigma}_a^2 \\ \hat{\sigma}_b^2 \end{array} \right] - \left[\begin{array}{c} \sigma^2 \\ \sigma^2 \end{array} \right] \right) \to_d \mathbb{N} \left(0, \left[\begin{array}{cc} \frac{2\sigma^4}{q} & \frac{2\sigma^4}{q} \\ \frac{2\sigma^4}{q} & 2\sigma^4 \end{array} \right] \right)$$

With the Delta method, representing $J_r = \frac{\hat{\sigma}_b^2}{\hat{\sigma}_a^2} - 1$ by $J_r = h(\hat{\sigma}_a^2, \hat{\sigma}_b^2)$, we have:

$$\sqrt{n}(h(\hat{\beta}) - h(\beta)) \to_d \mathbb{N}(0, \nabla h^T(\beta)D\nabla h(\beta))$$

$$\label{eq:where} \textit{where } \hat{\beta} = (\hat{\sigma}_a^2, \hat{\sigma}_b^2)^T, \; \beta = \vec{0}, \; \textit{and } D = \left[\begin{array}{cc} \frac{2\sigma^4}{q} & \frac{2\sigma^4}{q} \\ \frac{2\sigma^4}{q} & 2\sigma^4 \end{array} \right],$$

Namely, the distribution of J_r is,

$$\sqrt{n} (J_r - 0) \to_d \mathbb{N} \left(0, \begin{bmatrix} -\frac{1}{\sigma^2} & \frac{1}{\sigma^2} \end{bmatrix} \begin{bmatrix} \frac{A - \sigma^4}{2} & \frac{A - \sigma^4}{2} \\ \frac{A - \sigma^4}{2} & \frac{A + \sigma^4}{2} \end{bmatrix} \begin{bmatrix} -\frac{1}{\sigma^2} \\ \frac{1}{\sigma^2} \end{bmatrix} \right) = \mathbb{N}(0, 1)$$