### Week 1 – Thursday session

### **Statistical inference and Crude Analysis**

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### **Week 1: Discussion Topics**

#### 1. Statistical Inference

- Hypothesis testing
  - Null and alternative hypothesis
  - Form of the test statistic
  - P-values
  - Neyman-Pearson hypothesis tests
  - Limitations
- Confidence Intervals
  - Confidence, not probability
  - Relationship to Neyman-Pearson hypothesis testing
- Statistical versus biological/public health significance
- Sources of random variability

#### 2. Crude Analysis

- Person-time data testing, point estimates, confidence intervals
- Case-control data testing, point estimates, confidence intervals
- Binomial data (cumulative incidence; prevalence)
- 3. The EPI 202 Road map

## Statistical Significance Influence of the Size of the Numerator and Denominator

Recall the form of the statistic used for hypothesis testing:

$$Z^{2} = \frac{\left[X - E(X \mid H_{0})\right]^{2}}{Var(X \mid H_{0})} \sim \chi_{1}^{2}$$

The evidence against the null hypothesis increases as Z<sup>2</sup> gets larger as either the numerator gets larger or the denominator gets smaller

## **Statistical Significance Influence of Numerator Size**

- $Z^2$  increases as its numerator,  $[X E(X|H_0)]^2$ , increases
  - The value of the numerator is determined by the underlying phenomenon under investigation and cannot be altered by the researcher.
  - □ It is a function of the magnitude and direction of the exposure-disease association.
  - □ In a study of incidence rates in an exposed and unexposed group,  $X = \hat{R}D$ , the estimated incidence rate difference.  $E(X|H_0) = 0$ . The numerator of the test statistic,  $[X E(X|H_0)]^2$ , increases as the estimated rate difference becomes large (i.e. >> 0) or small (i.e. <<0).

## Statistical Significance Influence of Denominator Size

- Z<sup>2</sup> increases as the denominator, Var(X|H<sub>0</sub>), decreases N increases, more statistical information and greater precision, Var(X|H<sub>0</sub>) decreases; Z^2 increases, p gets smaller.
  - □ Var(X|H<sub>0</sub>) depends on both the value of X specified by H<sub>0</sub> and the sample size of the investigation.
  - All else being equal, the larger the sample size, the smaller the variance because there is more statistical information and the estimate of the exposure-outcome relationship can be estimated with greater precision.
  - Sample size is an arbitrary quantity; it is purely a function of study design and bears no relationship to any scientific quantity of interest.

### Z<sup>2</sup>, Sample Size and Statistical Significance

■ By making a study sufficiently large and thereby making the denominator of Z² sufficiently small, one can guarantee a "statistically significant" result no matter how small the difference between the groups are.

#### All else being equal, the value of the Z-squared test statistic gets larger when

$$Z^{2} = \frac{\left[X - E(X \mid H_{0})\right]^{2}}{Var(X \mid H_{0})} \sim \chi_{1}^{2}$$

The estimate of the association gets closer to the null

The estimate of the association gets further from the null

The variance gets larger

The variance gets smaller

Do not know

Total Results: 0

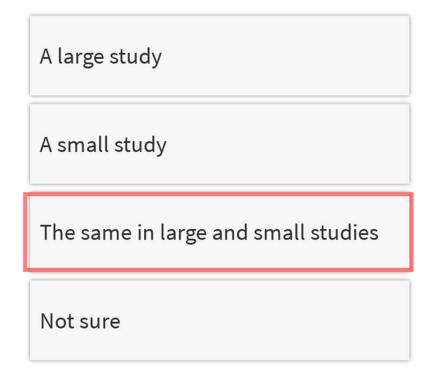
All else being equal, the Z^2 test statistic gets larger when either the numerator gets larger or the denominator gets smaller.

The nominator gets larger when the point estimate gets further from the null.

The denominator gets smaller when the variance under the null gets smaller. (More sample size)

# Assuming the null is true and there are no structural or other sources of bias, there will be a higher proportion of statistically significant findings at the alpha=0.05 level in





The probability of incorrectly rejecting the null when it is true is the same whether the study is large or small (because we have defined it as the alpha).

Unlike the p-value, alpha can be thought of as a long-term error risk.

Assuming the null is true and there are no structural or other sources of bias, the false positive findings (incorrectly rejecting the null), will on average be further from the null in





All else being equal, the variance of the estimator from a small study will be larger. Hence, the denominator of the test statistic will be larger in a small study.

In order for the test statistic to exceed the critical value, the numerator of the test statistic must be larger in a small study (on average be further from the null).

### Random Variability in Randomized Trials (1)

- The source of random variability in a randomized trial is well understood. Exposure (treatment) is assigned at random to each study participant.
- On average, both measured and unmeasured confounders are randomly distributed between the treatment and control groups. (Exchangeability is assumed)
- The source of random variability in this setting is identified as the random treatment assignment.

### Random Variability in Randomized Trials (2)

- The p-value has the interpretation that under the null, i.e. when there is no difference between the treated and untreated groups with respect to the outcome, the observed association is due to a random imbalance between the treated and untreated groups with respect to
  - □ unmeasured confounders or
  - □ confounders that were measured but not adjusted for statistically

### Random Variability in Observational Studies

- In an observational study, exposure is not randomly assigned.
  What, then, is the source of random variability to which the p-value refers in this setting?
- If exposure is not assigned by a chance mechanism, what is the meaning of the expression "These results could have been due to chance"?
- Because there is no clear physical source of randomness introduced into the design, the meaning of p-values and confidence intervals is in question.
- Additional assumptions are needed. These assumptions are largely not verifiable from observed data.

### Methods for Person-time Data Doll and Hill NCI Monograph 1966

 A classic prospective cohort study of cigarette smoking. The subjects are British male doctors. Here we investigate the relationship between cigarette smoking and coronary heart disease mortality (CHD).

<b>Notation</b>				Smoking			
	E	Ē	_		Yes	No	
Cases	а	b	M <sub>1</sub>	Deaths	630	101	731
Person-Time	N <sub>1</sub>	N <sub>o</sub>	T	Person-years	142,247	39,220	181,467

### Methods for Person-time Data Hypothesis Test – Choice of X

The test statistic is:

$$Z^{2} = \frac{\left[X - E(X \mid H_{0})\right]^{2}}{Var(X \mid H_{0})} \sim \chi_{1}^{2}$$

- X can be defined in several ways:
  - $\square$  X = number of exposed cases = a = 630
  - $\Box$  E(X|H<sub>o</sub>) = E(a|H<sub>o</sub>)
    - Total number of cases \* Pr(E)
    - $M_1(N_1/T)$
    - 731 \* (142,247/181,467)
    - 731 \* 0.784
    - 573.0

### Methods for Person-time Data Point estimate of the rate difference $(\widehat{IRD})$

$$I\widehat{RD} = \frac{a}{N_1} - \frac{b}{N_0} = \frac{630}{142,247 PY} - \frac{101}{39,220 PY}$$
$$= \frac{4.43}{1,000 PY} - \frac{2.58}{1,000 PY} = 1.85/1,000 PY$$

These data indicate that the rate of CHD deaths associated with smoking among British male doctors was 1.85 per 1,000 person-years higher among those who smoked compared with those who did not (assuming no confounding, selection bias, information bias or any other source of bias).

## Methods for Person-time Data Variance of the Incidence Rate Difference ( $\widehat{IRD}$ )

95% Confidence Interval  $X \pm 1.96\sqrt{Var(X)}$ 

$$X = I\hat{R}D = 0.00185$$

$$V\hat{a}r(X) = V\hat{a}r(I\hat{R}D) = \frac{\hat{I}_1}{N_1} + \frac{\hat{I}_0}{N_0} = \frac{a}{N_1^2} + \frac{b}{N_0^2}$$
$$= \frac{630}{142,247^2} + \frac{101}{39,220^2} = \frac{9.680}{10^8 PY^2}$$

Note: The variance of the IRD at it's observed value is not equal to the variance of the IRD under the null hypothesis when IRD=0.

### Methods for Person-time Data Point estimate of the incidence rate ratio $(\widehat{IRR})$

$$\widehat{IRR} = \frac{a}{N_1} / \frac{b}{N_0} = \frac{630}{142,247 \, PY} / \frac{101}{39,220 \, PY} = 1.72$$

 The CHD mortality rate was 72% higher among smokers compared with non-smokers (assuming no confounding or any other source of bias)

## Methods for Person-time Data Variance of the ln(Incidence Rate Ratio) ( $ln(\widehat{IRR})$ )

- 95% Confidence interval:  $X \pm 1.96\sqrt{\hat{Var}}(X)$
- Let  $X = \ln(\widehat{IRR}) = \ln(1.72) = 0.5422$

• 
$$\widehat{Var}(X) = \widehat{Var}(\ln(\widehat{IRR})) = \frac{1}{a} + \frac{1}{b}$$

$$=\frac{1}{630}+\frac{1}{101}=0.01149$$

## Methods for Person-time Data Confidence Interval for the Rate Ratio $(\widehat{IRR})$

• 95% Confidence interval for ln(IRR):

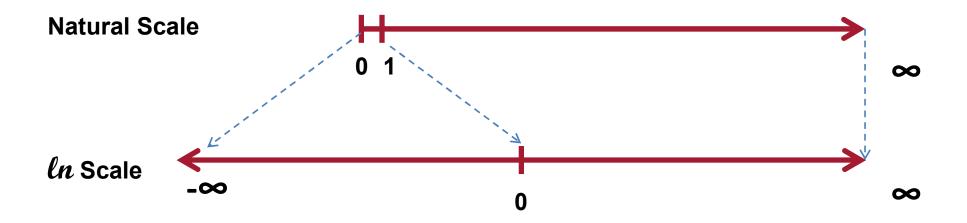
$$0.5422 \pm 1.96 \sqrt{0.01149}$$
$$= (0.3321, 0.7523)$$

95% Confidence interval for (IRR):

$$e^{(X\pm 1.96\sqrt{\widehat{Var}(X)})} = e^{(0.3321,0.7522)}$$
$$= (1.39, 2.12)$$

■ These data are consistent with rate ratios ranging from 1.4 to 2.1 with 95% confidence (assuming no confounding, selection bias or any other source of bias)

### Ratios on the *ln* Scale



	IRR	<i>ln</i> IRR
Bike helmet non-use vs. use	IRR=10	<i>ln</i> (10)=2.3
Bike helmet use vs. non-use	IRR=1/10=0.1	ln(1/10)=-2.3

### **Brief Review of Logs and Exponents**

■ For the purpose of this course and in nearly all work in epidemiology, biostatistics, and statistics, *log* refers to the natural log (ln)

```
\Box exp{X} = e<sup>X</sup>
\Box exp{0} = 1
\Box exp{1} = 2.718281828
\Box ln[exp(X)] = X
\Box ln[1] = 0
\Box ln[2.718281828] = 1
\Box ln[2.718281828^2] = ln[7.389056] = 2
\Box ln[0] = -\infty
\Box ln/\infty] = \infty
\Box exp[ln(X)] = X
\Box exp{A}*exp{B} = exp{A+B}
\Box exp{A}/exp{B} = exp{A-B}
\Box ln\{AB\} = ln\{A\} + ln\{B\}
\Box ln\{A/B\} = ln\{A\} - ln\{B\}
\Box ln[1/X]=-ln[X]
```

### **Test of No Exposure-Disease Association**

$$Z^{2} = \frac{[X - \hat{E}(X | H_{0})]^{2}}{V \hat{a} r(X | H_{0})} \sim X_{1}^{2}$$

Ho  $\hat{E}(X|H_0)$   $Var(\hat{X}|H_0)$ Methods for Count Data (Closed cohort and cross-sectional Studies)  $\frac{N_1 M_1}{T} \qquad \frac{M_1 M_0 N_1 N_0}{T^3}$ Unstratified  $E \overline{E}$ Non-cases N<sub>1</sub>-a N<sub>0</sub>-b M<sub>0</sub>  $\sum a_i \qquad \sum \frac{N_{1i}M_{1i}}{T} \qquad \sum \frac{M_{1i}M_{0i}N_{1i}N_{0i}}{T^3}$ Stratified Methods for Person-time Data (Open Cohort and Closed Cohort studies)  $\overline{\mathrm{E}}$ E Unstratified cases  $\sum a_i \qquad \sum \frac{N_{1i}M_{1i}}{T_i} \qquad \sum \frac{N_{1i}N_{0i}M_{1i}}{T_i^2}$ PT Stratified **Methods for Case-control Data** Ē  $\frac{N_1 M_1}{T} \qquad \frac{M_1 M_0 N_1 N_0}{T^2 (T-I)}$ E Unstratified cases OR=1  $I_1/I_0 = 1$  $d \mid M_0$ controls  $\sum a_i \qquad \sum \frac{N_{1i}M_{1i}}{T_i} \qquad \sum \frac{M_{1i}M_{0i}N_{1i}N_{0i}}{T_i^2(T_i-I)}$ Stratified

## **Confidence Intervals Count Data**

$$X \pm Z_{1-\alpha/2} \sqrt{V \widehat{a} r(X)}$$

X

 $w_i$ 

Var(X)

#### Methods for Count Data (closed cohort and cross-sectional studies)

Cumulative incidence difference

$$\frac{a}{N_1} - \frac{b}{N_0}$$

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$$\frac{ac}{N_1^3} + \frac{bd}{N_0^3}$$

Summary cumulative incidence difference

$$\frac{\sum w_i \left[ \frac{a_i}{N_{1i}} - \frac{b_i}{N_{0i}} \right]}{\sum w_i}$$

 $\frac{N_{1i}N_{0i}}{T_i}$ 

$$\frac{\sum \left(\frac{a_i c_i N_{0i}^2}{T_i^2 (N_{1i} - 1)} + \frac{b_i d_i N_{1i}^2}{T_i^2 (N_{0i} - 1)}\right)}{\left(\sum \frac{N_{1i} N_{0i}}{T_i}\right)^2}$$

Cumulative incidence ratio  $(\ell n)$ 

$$\ln \left\{ \frac{a}{N_1} / \frac{b}{N_0} \right\}$$

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$$\frac{c}{aN_1} + \frac{d}{bN_0}$$

Summary cumulative incidence ratio  $(\ell n)$ 

$$\ln \left\{ \frac{\sum w_i \left[ \frac{a_i}{N_{1i}} / \frac{b_i}{N_{0i}} \right]}{\sum w_i} \right\}$$

 $\frac{b_i N_{1i}}{T_i}$ 

$$\frac{\sum \! \left(\! M_{1i} N_{1i} N_{0i} - a_i b_i T_i \right) \! / T_i^2}{\left[ \sum \! \frac{a_i N_{0i}}{T_i} \right] \! \left[ \sum \! \frac{b_i N_{1i}}{T_i} \right]}$$

## **Confidence Intervals Person-time Data**

$$X \pm Z_{1-\alpha/2} \sqrt{V \widehat{a} r(X)}$$

 $W_i$ 

Var(X)

#### Methods for Person-Time Data (open cohort and closed cohort studies)

X

Rate difference  $\frac{a}{N_1} - \frac{b}{N_0}$ Summary rate  $\sum w_i \left[ \frac{a_i}{N_1} - \frac{b_i}{N_0} \right]$ 

 $\frac{\sum w_i \left[ \frac{a_i}{N_{1i}} - \frac{b_i}{N_{0i}} \right]}{\sum w_i}$ 

 $\frac{N_{1i}N_{0i}}{T_i}$ 

 $\frac{\sum \left(\frac{a_{i}N_{0i}^{2}+b_{i}N_{1i}^{2}}{T_{i}^{2}}\right)}{\left(\sum \frac{N_{1i}N_{0i}}{T_{i}}\right)^{2}}$ 

Rate ratio (ln)

difference

 $\ln \left\{ \frac{\mathsf{a}}{\mathsf{N}_1} \middle/ \frac{\mathsf{b}}{\mathsf{N}_0} \right\}$ 

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 $\frac{1}{a} + \frac{1}{b}$ 

Summary rate ratio ( $\ell n$ )

 $\ln \left\{ \frac{\sum w_i \left[ \frac{a_i}{N_{1i}} / \frac{b_i}{N_{0i}} \right]}{\sum w_i} \right\}$ 

 $\frac{b_i N_{1i}}{T_i}$ 

$$\frac{\sum \left(M_{1i}N_{1i}N_{0i}\right)\!\!/T_i^2}{\left[\sum \frac{a_iN_{0i}}{T_i}\right]\!\!\left[\sum \frac{b_iN_{1i}}{T_i}\right]}$$

## **Confidence Intervals**Case-control Data

$$X \pm Z_{1-\alpha/2} \sqrt{V \widehat{a} r(X)}$$

X

Wi

Var(X)

#### **Methods for case-control Data**

Odds ratio( $\ell n$ )

$$\ln \left\{ \frac{ad}{bc} \right\}$$

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$$\frac{1}{a} + \frac{1}{b} + \frac{1}{c} + \frac{1}{a}$$

Summarv odds ratio (*ln*)

$$\ln \left\{ \frac{\sum w_i \frac{a_i d_i}{b_i c_i}}{\sum w_i} \right\}$$

$$\frac{b_i c_i}{T_i}$$

**RGB** variance

**RGB** variance:

$$\frac{1}{2} \left[ \frac{\sum \left( \frac{a_i d_i}{T_i} \right) \left( \frac{a_i + d_i}{T_i} \right)}{\left( \sum \frac{a_i d_i}{T_i} \right)^2} + \frac{\sum \left( \frac{a_i d_i}{T_i} \right) \left( \frac{c_i + b_i}{T_i} \right) + \sum \left( \frac{b_i c_i}{T_i} \right) \left( \frac{a_i + d_i}{T_i} \right)}{\left( \sum \frac{a_i d_i}{T_i} \right) \left( \sum \frac{b_i c_i}{T_i} \right)} + \frac{\sum \left( \frac{b_i c_i}{T_i} \right) \left( \frac{c_i + b_i}{T_i} \right)}{\left( \sum \frac{b_i c_i}{T_i} \right)^2} \right]$$

### Computational formulas for point estimates

#### **Computational Form of the Mantel Haenszel Estimators**

#### Methods for Count Data (closed cohort and cross-sectional studies)

Summary cumulative incidence ratio

$$\hat{\text{CIR}}_{\text{MH}} = \frac{\sum\limits_{i=1}^{I} \frac{a_i N_{0i}}{T_i}}{\sum\limits_{i=1}^{I} \frac{b_i N_{1i}}{T_i}} \qquad \begin{array}{c} \text{Summary} \\ \text{cumulative} \\ \text{incidence} \\ \text{difference} \end{array} \qquad \frac{\sum \left(\frac{a_i N_{0i} - b_i N_{1i}}{T_i}\right)}{\sum \frac{N_{1i} N_{0i}}{T_i}}$$

$$\frac{\sum \left(\frac{a_i N_{0i} - b_i N_{1i}}{T_i}\right)}{\sum \frac{N_{1i} N_{0i}}{T_i}}$$

#### Methods for Person-Time Data (open cohort and closed cohort studies)

Summary rate ratio

$$\hat{IRR}_{\text{MH}} = \frac{\sum_{i=1}^{I} \frac{a_i N_{0i}}{T_i}}{\sum_{i=1}^{I} \frac{b_i N_{1i}}{T_i}} \qquad \text{Summary rate} \qquad \frac{\sum \left(\frac{a_i N_{0i} - b_i N_{1i}}{T_i}\right)}{\sum \frac{N_{1i} N_{0i}}{T_i}}$$

$$\frac{\sum \left(\frac{a_i N_{0i} - b_i N_{1i}}{T_i}\right)}{\sum \frac{N_{1i} N_{0i}}{T_i}}$$

#### **Methods for Case-control Data**

$$\hat{O}R_{MH} = \frac{\sum_{i=1}^{I} \frac{a_{i}d_{i}}{T_{i}}}{\sum_{i=1}^{I} \frac{b_{i}c_{i}}{T_{i}}}$$

### **Test of Homogeneity of Effect Measures**

$$H = \sum_{i=1}^{1} \frac{[\hat{X}_{i} - \hat{X}_{summary}]^{2}}{V \hat{a} r_{i} [\hat{X}_{i}]} \sim \chi_{i-1}^{2}$$

## **Confidence Intervals Matched Case-control Data**

$$\begin{array}{c|cccc} & \underline{Controls} \\ & E & \overline{E} \\ \underline{Case} & E & f_{11} & f_{10} \\ \overline{E} & f_{01} & f_{00} \\ \end{array}$$

$$OR_{MH} = f_{10} / f_{01}$$

Test statistic:

Ho:  $OR_{MH} = 1$ 

Ha:  $OR_{MH} \neq 1$ 

$$Z^{2} = \frac{(f_{10} - f_{01})^{2}}{f_{10} + f_{01}} \sim \chi_{1}^{2}$$

Variance formula for confidence interval:

$$Var(\ln (OR_{MH})) = \frac{1}{f_{10}} + \frac{1}{f_{01}}$$

### **HAVE A GOOD WEEKEND**