# Virtual whiteboard for Jan'22 math+econ+code

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#### Day 1 1

#### 1a. The diet problem

Consider  $N_{ij}$ =amount of nutrient i that is brought by one dollar's worth of food

Look for  $q_j =$ \$ invested in food jsuch that we minimize total cost  $\min \sum_{i} q_{i}$ 

subject to the constraint that the minimal intake in each food i is met  $\sum_{j} N_{ij} q_j \ge d_i$  for each nutrient i where  $d_i$  is the minimum quantity of nutrient i.

To summarize, we need to solve

$$\min_{q \ge 0} \sum_{j} q_j$$
  
s.t. 
$$\sum_{j} N_{ij} q_j \ge d_i$$

In matrix form this is

$$\min_{q \ge 0} c^{\top} q$$
  
s.t.  $Nq \ge d$ 

where 
$$c_j = 1$$
.

#### 1.1.1 Duality worked out by hand

$$\min_{q\geq 0} \sum_{j} c_{j}q_{j}$$
s.t. 
$$\sum_{j} N_{ij}q_{j} \geq d_{i}$$
transform into an unconstrained minimization problem

$$\min_{q\geq 0} \sum_{j} c_{j}q_{j} + \sum_{i} F\left(d_{i} - \sum_{j} N_{ij}q_{j}\right)$$
  
where  $F(z) = 0$  if  $z \leq 0$ , and  $F(z) = +\infty$  if  $z > 0$ .

How to represent F? How about

$$F(z) = \max_{\pi \ge 0} \{\pi z\}.$$

$$\min_{q\geq 0} \sum_{j} c_j q_j + \sum_{i} \max_{\pi_i \geq 0} \pi_i \left( d_i - \sum_{j} N_{ij} q_j \right)$$

$$\begin{aligned} & \min_{q \geq 0} \sum_{j} c_{j}q_{j} + \max_{\pi_{i} \geq 0} \sum_{i} \pi_{i} \left(d_{i} - \sum_{j} N_{ij}q_{j}\right) \\ & \min_{q_{j} \geq 0} \max_{\pi_{i} \geq 0} \sum_{j} c_{j}q_{j} + \sum_{i} \pi_{i} \left(d_{i} - \sum_{j} N_{ij}q_{j}\right) \\ & \text{Assuming min max} = \max \min, \text{ we can reformulate this as } \\ & \max_{\pi_{i} \geq 0} \min_{q_{j} \geq 0} \sum_{j} c_{j}q_{j} + \sum_{i} \pi_{i} \left(d_{i} - \sum_{j} N_{ij}q_{j}\right) \\ & \max_{\pi_{i} \geq 0} \min_{q_{j} \geq 0} \sum_{j} c_{j}q_{j} + \sum_{i} \pi_{i}d_{i} - \sum_{i,j} \pi_{i}N_{ij}q_{j} \\ & \text{going through the same steps in the opposite direction:} \\ & \max_{\pi_{i} \geq 0} \sum_{i} \pi_{i}d_{i} + \min_{q_{j} \geq 0} \sum_{j} c_{j}q_{j} - \sum_{i,j} \pi_{i}N_{ij}q_{j} \\ & \max_{\pi_{i} \geq 0} \sum_{i} \pi_{i}d_{i} + \min_{q_{j} \geq 0} \sum_{j} q_{j} \left(c_{j} - \sum_{i} \pi_{i}N_{ij}\right) \\ & \max_{\pi_{i} \geq 0} \sum_{i} \pi_{i}d_{i} + \sum_{j} \min_{q_{j} \geq 0} q_{j} \left(c_{j} - \sum_{i} \pi_{i}N_{ij}\right) \\ & \text{But } \left(\min_{q_{j} \geq 0} q_{j}w_{j}\right) = -\infty \text{ if } w_{j} < 0 \text{ and } = 0 \text{ if } w_{j} \geq 0 \end{aligned}$$

 $\max_{\pi_i \geq 0} \sum_i \pi_i d_i$ s.t.  $c_j \geq \sum_i \pi_i N_{ij}$ that is

 $\max_{\pi \geq 0} \pi^\top d$  s.t.  $N^\top \pi \leq c$ 

Theorem. Consider the "primal problem"

$$\min_{q\geq 0} \sum_{j} c_{j}q_{j}$$
s.t. 
$$\sum_{j} N_{ij}q_{j} \geq d_{i} \ [\pi_{i} \geq 0]$$
and the "dual problem"
$$\max_{\pi_{i}\geq 0} \sum_{i} \pi_{i}d_{i}$$
s.t. 
$$\sum_{i} \pi_{i}N_{ij} \leq c_{j} \ [q_{j} \geq 0]$$

Then if either of them is feasible [i.e. there is a variable that meets the constraints then

- (1) the value of the primal problem is equal to the value of the dual problem
- (2) complementary slackness:

Assume q is a solution to the primal problem and  $\pi$  is a solution to the dual

$$\begin{array}{l} \pi_i > 0 \text{ implies } d_i = \sum_j N_{ij} q_j \\ q_j > 0 \text{ implies } c_j = \sum_i \pi_i N_{ij} \end{array}$$

**Theorem.** If q is feasible for the primal and  $\pi$  is feasible for the dual and if complementary slackness holds, then q is optimal for the primal and  $\pi$  is optimal for the dual.

#### 1b. The optimal assignment problem

Joint surplus matrix

$$\Phi_{xy} = x^{\top} A y = \sum_{k,l} A_{kl} x_k y_l$$

Assume  $n_x$  men of type x and  $m_y$  women of type y.

Becker-Shapley-Shubik's model of matching.

Assume if man x matches with woman y, then:

x gets surplus  $\alpha_{xy}$ 

y gets surplus  $\gamma_{xy}$ 

Assume utility is transferable, ie if  $w_{xy}$  is the transfer from the woman to the man (either positive or negative),

x gets surplus  $\alpha_{xy} + w_{xy}$ 

y gets surplus  $\gamma_{xy}-w_{xy}$ 

This is the Transferable Utility assumption.

 $w_{xy}$  is determined at equilibrium.

Regardless of what  $w_{xy}$  is, the joint surplus

$$\Phi_{xy} = (\alpha_{xy} + w_{xy}) + (\gamma_{xy} - w_{xy}) = \alpha_{xy} + \gamma_{xy}$$
 is the same.

Roadmap:

- 1. Optimality
- 2. Equilibrium

#### 1.2.1Optimality

A matching is a  $\mu_{xy} \geq 0$  which is the number of men of type x matched with women of type y.

Constraint on 
$$\mu_{xy}$$
:  

$$\sum_{y} \mu_{xy} = n_{x}$$

$$\sum_{x} \mu_{xy} = m_{y}$$

Optimal matching consists of

$$\max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy}$$

$$\begin{array}{l} \sum_y \mu_{xy} = n_x \ [u_x] \\ \sum_x \mu_{xy} = m_y \ [v_y] \end{array}$$

Duality:

$$\max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy} + \sum_{x} \min_{u_x} u_x \left( n_x - \sum_{y} \mu_{xy} \right) + \sum_{y} \min_{v_y} v_y \left( m_y - \sum_{x} \mu_{xy} \right)$$

$$\min_{u_x, v_y} \sum_{x} n_x u_x + \sum_{y} m_y v_y + \max_{\mu \geq 0} \sum_{xy} \mu_{xy} \left( \Phi_{xy} - u_x - v_y \right)$$
that is, the dual problem is
$$\min_{u_x, v_y} \sum_{x} n_x u_x + \sum_{y} m_y v_y$$
s.t.  $u_x + v_y \geq \Phi_{xy} \left[ \mu_{xy} \geq 0 \right]$ 

By complementary slackness, we have  $\mu_{xy} > 0 \implies u_x + v_y = \Phi_{xy}$ .

#### Interpretation as a stable outcome.

We now consider the decentralized version of the problem.

Definition:  $(\mu, u, v)$  is a stable outcome if

(1)  $\mu$  is a matching:

$$\begin{array}{l} \sum_y \mu_{xy} = n_x \ [u_x] \\ \sum_x \mu_{xy} = m_y \ [v_y] \end{array}$$

(2) Stability: we have for all x and y that

$$u_x + v_y \ge \Phi_{xy}$$

[otherwise we would have  $u_x + v_y < \Phi_{xy}$  and xy would be a blocking pair, ie there is a way for x to have more than  $u_x$  and y to have more than  $v_y$  if x and y match together]

(3) Feasibility: if  $\mu_{xy} > 0$ , then  $u_x + v_y = \Phi_{xy}$ .

Note that (1) means that  $\mu$  is feasible for the primal problem

- (2) means that (u, v) is feasible for the dual problem
- (3) means complementary slackness.

Hence  $\mu$  is a solution to the primal problem

$$\max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy}$$

$$\sum_{y}^{y} \mu_{xy} = n_x [u_x]$$
$$\sum_{x}^{y} \mu_{xy} = m_y [v_y]$$

and (u, v) is a solution to the dual problem

$$\min_{u_x, v_y} \sum_x n_x u_x + \sum_y m_y v_y$$
  
s.t.  $u_x + v_y \ge \Phi_{xy} \left[ \mu_{xy} \ge 0 \right]$ 

**Recovering the transfers.** We had that if  $w_{xy}$  is the transfer from y to

x gets surplus  $\alpha_{xy} + w_{xy}$ 

y gets surplus 
$$\gamma_{mn} - w_{rn}$$

y gets surplus  $\gamma_{xy}-w_{xy}$ Hence the payoff of x at equilibrium is

$$u_x = \max_y \left\{ \alpha_{xy} + w_{xy} \right\}$$

$$v_y = \max_x \left\{ \gamma_{xy} - w_{xy} \right\}$$

 $u_x \ge \alpha_{xy} + w_{xy}$  for all x and y

$$v_y \ge \gamma_{xy} - w_{xy}$$
 for all  $x$  and  $y$   
This yields

$$\gamma_{xy} - v_y \le w_{xy} \le u_x - \alpha_{xy}$$

We have 
$$u_x - \alpha_{xy} \ge \gamma_{xy} - v_y$$
 – indeed  $u_x + v_y \ge \alpha_{xy} + \gamma_{xy}$ 

When 
$$\mu_{xy} > 0$$
, we have  $w_{xy} = \gamma_{xy} - v_y = u_x - \alpha_{xy}$ 

#### When partners can remain unmatched.

Assume that individuals get utility zero if they remain unmatched.

Then a feasible matching imposes

$$\sum_{y} \mu_{xy} \le n_x$$
$$\sum_{x} \mu_{xy} \le m_y$$

An optimal matching solves

$$\max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy}$$

$$\sum_{y} \mu_{xy} \le n_x \quad [u_x \ge 0]$$
$$\sum_{x} \mu_{xy} \le m_y \quad [v_y \ge 0]$$

#### Duality:

$$\max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy} + \sum_{x} \min_{u_x \geq 0} u_x \left( n_x - \sum_{y} \mu_{xy} \right) + \sum_{y} \min_{v_y \geq 0} v_y \left( m_y - \sum_{x} \mu_{xy} \right)$$
$$\min_{u_x \geq 0} \sum_{xy} n_x u_x + \sum_{y} m_y v_y + \max_{\mu \geq 0} \sum_{xy} \mu_{xy} \left( \Phi_{xy} - u_x - v_y \right)$$
that is, the dual problem is

$$\min_{u_x \geq v_y \geq 0} \sum_x n_x u_x + \sum_y m_y v_y$$
  
s.t.  $u_x + v_y \geq \Phi_{xy} \quad [\mu_{xy} \geq 0]$ 

By complementary slackness, we have  $\mu_{xy} > 0 \implies u_x + v_y = \Phi_{xy}$ .

$$\begin{aligned} & \min_{u_x \geq 0, v_y \geq 0} \sum_{x} n_x u_x + \sum_{y} m_y v_y \\ & \text{s.t. } u_x + v_y \geq \Phi_{xy} \ \left[ \mu_{xy} \geq 0 \right] \end{aligned}$$

Alternatively, the primal can be expresses as

$$\max_{\mu \ge 0} \sum_{xy} \mu_{xy} \Phi_{xy}$$

$$\begin{array}{l} \sum_y \mu_{xy} + \mu_{x0} = n_x \ [u_x \geq 0] \\ \sum_x \mu_{xy} + \mu_{0y} = m_y \ [v_y \geq 0] \end{array} \label{eq:local_equation}$$

where  $\mu_{x0}$  and  $\mu_{0y}$  act as slackness variable.

Stability interpretation: a stable outcome  $(\mu, u, v)$  in the problem with singles is such that

(1)  $\mu$  is a feasible partial matching:

$$\sum_{y} \mu_{xy} + \mu_{x0} = n_x$$
$$\sum_{x} \mu_{xy} + \mu_{0y} = m_y$$
(2) Stability holds

$$\sum_{x} \mu_{xy} + \mu_{0y} = m_y$$

$$(2)$$
 Stability holds

$$u_x + v_y \ge \Phi_{xy}$$

$$u_x \ge 0, v_y \ge 0$$

(3) Complementary slackness

$$\mu_{xy}>0 \implies u_x+v_y=\Phi_{xy}$$

$$\mu_{x0} > 0$$
 i.e  $\left(\sum_{y} \mu_{xy} < n_{x}\right) \implies u_{x} = 0$   
 $\mu_{0y} > 0$  i.e.  $\left(\sum_{x} \mu_{xy} < m_{y}\right) \implies v_{y} = 0$ 

$$\mu_{0y} > 0$$
 i.e.  $\left(\sum_{x} \mu_{xy} < m_y'\right) \implies v_y = 0$ 

#### 1.2.3 Indivisibilities (finite population)

When  $n_x = 1$  for each x and  $m_y = 1$  for each y, we should in principle impose an integrality constraint, that is

$$\mu_{xy} \in \{0,1\}$$
.

Then the problem is no longer a linear programming problem, but an integer programming problem.

However, in the bipartite case, one can abstract away from the integrality constraint.

#### 1.2.4 Computation

Consider the problem

 $\max_{\mu \ge 0} \sum_{xy} \mu_{xy} \Phi_{xy}$ s.t.  $\sum_{y} \mu_{xy} \le n_x \quad [u_x \ge 0]$ 

 $\begin{array}{l} \sum_y \mu_{xy} \leq n_x \ [u_x \geq 0] \\ \sum_x \mu_{xy} \leq m_y \ [v_y \geq 0] \\ \text{this is of the form} \end{array}$ 

 $\max_{\mu \ge 0} \mu^{\top} \Phi$  $M\mu \le \binom{n}{m}$ 

How to convert a matrix into a vector? Take a matrix M, we call vec(M) its vectorized version,

This can be done by:

\* stack the columns: Matlab, Julia, R, Fortran column-major ordering, or Fortran ordering.

\* stack the rows: done by NumPy by default, as well as C as well as some other languages: row-major ordering, or C ordering – primary convetion in this course.

Take the first set of constraints  $\sum_{y} \mu_{xy} \leq n_x$ . If  $\mu$  is understood as a matrix, this is

$$\mu 1_Y \leq n$$

by pre-multuplying by the identity this yields

$$I_X \mu 1_Y \le n$$

And we need to look at  $vec(I_X \mu 1_Y) = matrix.vec(\mu)$ Fundamental identity is (assuming row-major ordering)

$$vec(AXB^{\top}) = (A \otimes B) vec(X)$$
.

Here, our constraints become

$$vec\left(I_X\mu 1_Y\right) = \left(I_X\otimes 1_Y^\top\right)vec\left(\mu\right) \le n$$

Similarly, we can vectorize the other constraints  $\sum_x \mu_{xy} \leq m_y$  into  $1_X^\top \mu I_Y \leq m$ , that is

$$vec\left(1_X^{\top}\mu I_Y\right) = \left(1_X^{\top}\otimes I_Y\right)vec\left(\mu\right) \leq m$$

Hence our optimal assignment problem becomes in a vectorized fashion

$$\max_{vec(\mu) \ge 0} \quad vec(\mu)^{\top} vec(\Phi)$$

$$s.t. \quad \left(I_X \otimes 1_Y^{\top}\right) vec(\mu) \le n$$

$$\left(1_X^{\top} \otimes I_Y\right) vec(\mu) \le m$$

that is

$$\max_{v \ge 0} \qquad v^{\top} vec\left(\Phi\right)$$

$$s.t. \qquad Mv \le \binom{n}{m}$$

where

$$M = \begin{pmatrix} I_X \otimes 1_Y^\top \\ 1_X^\top \otimes I_Y \end{pmatrix}$$

is called the margining matrix.

## Day 2

#### About yesterday's exercises

$$\begin{pmatrix} 1/4 & 1/4 - 1/8 & 1/4 + 1/8 \\ & 2/8 + 1/8 & 1/8 - 1/8 \end{pmatrix}$$

networkx algorithms to detect loops

Birkhoff-von Neumann

$$\begin{aligned} & \min \sum_{xy} \mu_{xy} c_{xy} \\ & M\mu = \binom{n}{m} \\ & \text{If } c_{xy} \in \{0,1\}, \text{ then } \Gamma = \{xy : c_{xy} = 0\} \end{aligned}$$

$$0 = \min \sum_{xy} \mu_{xy} c_{xy}$$
$$M\mu = \binom{n}{m}$$

iff for every xy such that  $\mu_{xy} > 0$ , then  $xy \in \Gamma$ 

this means that there is a matching between n and m "compatible" with  $\Gamma$ .

The dual to 
$$\min \sum \mu_{mi} c$$
.

$$\min \sum_{xy} \mu_{xy} c_{xy}$$

$$M\mu = \binom{n}{m}$$
is

$$\max_{u,v} \sum_{x} u_x - \sum_{y} v_y$$
  
s.t.  $u_x - v_y \le c_{xy}$ 

s.t. 
$$u_x - v_y \le c_{xy}$$

$$\Gamma = egin{pmatrix} 1 & 1 & 1 & 1 \ & & & 1 \ & & 1 & \ & & & 1 \end{pmatrix}$$

## 2a. One-dimensional matching

We now assume that types of workers x and firms y belong in  $\mathbb{R}$ .

Assume that there is the same total mass of workers and firms.

The total mass of workers and firms is normalized to one.

n(x) is the density of probability associated with the distribution of workers m(y) is the density of probability associated with the distribution of firms

A matching is a distribution of probability  $\mu(x,y)$  over pairs x,y. It should

$$\int \mu(x, y) dy = n(x)$$
$$\int \mu(x, y) dx = m(y)$$

Assume that the economic value created by a CEO of type x with a firm of type y is

 $\Phi(x,y)$ .

The optimal assignment problem is

 $\max_{\mu} \int \Phi(x, y) \mu(x, y) dxdy$ 

s.t.

$$\int \mu(x, y) dy = n(x), x \in R$$
$$\int \mu(x, y) dx = m(y), y \in R$$

CEO application (Gabaix and Landier, Tervio):

x is the CEO's talent = extra % return on asset the CEO generates y is the firm's size (market cap)

In that case, we get that

$$\Phi\left(x,y\right) = xy$$

Agenda:

Goal = predict CEO compensation

Solve for the primal problem

From the solution to the primal problem, we will deduce the solution to the dual problem

The dual problem will give us predictions for CEO compensation

Intuition: If T(x) is the firm that CEO x is matched with, it makes sense to assume that

T(x) is increasing.

If that is the case, what is T(x)?

We know that if  $X \sim n$  is a random variable distributed according to n, then we  $T(X) \sim m$ .

We can introduce the cumulative distribution functions (CDF)  $F_n(x) = \int_{-\infty}^x n(z) dz$  and  $F_m(y) = \int_{-\infty}^y m(z) dz$  associated with n and m respectively, and we have

$$\Pr\left(T\left(X\right) \le y\right) = F_m\left(y\right)$$

take y = T(x) for a fixed value of x, we get

$$\Pr\left(T\left(X\right) \leq T\left(x\right)\right) = F_{m}\left(T\left(x\right)\right)$$

thus

$$Pr(X \le x) = F_m(T(x))$$

$$F_n(x) = F_m(T(x))$$

therefore

$$T\left(x\right) = F_{m}^{-1}\left(F_{n}\left(x\right)\right)$$

In particular, this means that the median CEO – ie the CEO x such  $F_n(x) =$ 1/2 is matched with the median firm indeed  $F_n(x) = 1/2 = F_m(T(x))$  – thus T(x) is the median firm.

Let's see how we can solve for the dual problem. The dual problem is  $\min_{u(x),v(y)} \int u(x) n(x) dx + \int v(y) m(y) dy$  s.t.  $u(x) + v(y) \ge \Phi(x,y) = xy$ 

Assume (u, v) is a \*\* feasible solution \*\* to the dual problem.

Then for every x and every y,

 $u(x) + v(y) \ge xy$ 

thus

 $v(y) \ge \max_{x} \{xy - u(x)\}$  holds for for every y.

Claim: for any \*\* optimal solution \*\* to the dual problem, we have

 $v(y) = \max_{x} \{xy - u(x)\}$  – interpreted as firm y's problem

Obviously, this is symmetric, so we have also

 $u(x) = \max_{y} \{xy - v(y)\}$  – interpreted as worker x's problem.

Let's write down optimality conditions in the firm y's problem y = u'(x).

But we are in the case where T(x) is known and  $T(x) = F_m^{-1}(F_n(x))$ .

We have therefore

$$u'(x) = F_m^{-1}(F_n(x))$$

as a result

$$u(x) = \int_{-\infty}^{x} F_{m}^{-1}(F_{n}(z)) dz + cte$$

 $u\left(x\right) = \int^{x} F_{m}^{-1}\left(F_{n}\left(z\right)\right) dz + cte.$  Deduce v (either by  $v\left(y\right) = \max_{x} \left\{xy - u\left(x\right)\right\}$ , or  $v\left(y\right) = \int^{y} F_{n}^{-1}\left(F_{m}\left(z\right)\right) dz + cte$ cte

**Theorem.** When  $\Phi$  is supermodular, i.e.  $\partial^2 \Phi(x,y)/\partial x \partial y \geq 0$ , then the positive assortative matching solution  $T\left(x\right) = F_m^{-1}\left(F_n\left(x\right)\right)$  is optimal. Until now we have assumed  $\Phi\left(x,y\right) = xy$ , so  $\partial^2\Phi\left(x,y\right)/\partial x\partial y = 1 \geq 0$ .

We have that if u and v are optimal dual solutions, then

 $v(y) = \max_{x} \{\Phi(x, y) - u(x)\}$  - interpreted as firm y's problem

Obviously, this is symmetric, so we have also

 $u\left(x\right) = \max_{y} \left\{\Phi\left(x,y\right) - v\left(y\right)\right\}$  – interpreted as worker x's problem.

By first order condition in the firm's T(x) problem, we have

$$u'(x) = \partial_x \Phi(x, T(x))$$

Now let's verify that if  $\Phi$  is supermodular, then T is increasing. Deriving the above wrt x, we have

$$u''\left(x\right) = \partial_{xx}^{2} \Phi\left(x, T\left(x\right)\right) + \partial_{xy}^{2} \Phi\left(x, T\left(x\right)\right) T'\left(x\right)$$
therefore

$$T'\left(x\right) = \frac{u''\left(x\right) - \partial_{xx}^{2}\Phi\left(x, T\left(x\right)\right)}{\partial_{xy}^{2}\Phi\left(x, T\left(x\right)\right)}$$

The denominator is positive  $\partial_{xy}^{2}\Phi\left(x,T\left(x\right)\right)>0$  by supermodularity We have  $u''\left(x\right)-\partial_{xx}^{2}\Phi\left(x,T\left(x\right)\right)\geq0$  by second order conditions. Indeed,  $\partial_{xx}^{2}\Phi\left(x,y\right)-u''\left(x\right)\leq0$ .

Hence

$$T'(x) \geq 0.$$

Exercise. Model of marriage with taxes.

If single individual get gross income w, then gets net amount  $N\left(w\right)$  where N is increasing and concave.

Assume x and y are the gross incomes of two marital partners. Then their combined gross income is x + y, and we create two fictious personnas which make  $\frac{x+y}{2}$  each. Thus their combined net income is  $2N\left(\frac{x+y}{2}\right)$ .

Thus we can consider a matching market where the matching surplus is  $\Phi(x,y) = 2N\left(\frac{x+y}{2}\right)$ 

- 1) Is there surplus to matching? i.e. do we have  $2N\left(\frac{x+y}{2}\right) \geq N\left(x\right) + N\left(y\right)$ ?
- 2) Derive the matching equilibrium on this market.

Solution.

- 1)  $N\left(\frac{x+y}{2}\right) \ge \frac{N(x)+N(y)}{2}$  because N is concave.
- 2) We have that the matching surplus  $\Phi\left(x,y\right)=2N\left(\frac{x+y}{2}\right)-N\left(x\right)-N\left(y\right)$ . The cross-derivative is

$$\frac{\partial^{2}\Phi\left(x,y\right)}{\partial x\partial y}=\frac{1}{2}N^{\prime\prime}\left(\frac{x+y}{2}\right)\leq0$$

by concavity.

How can we use what we saw before to tackle this? Define  $\tilde{x}=x$  and  $\tilde{y}=-y$  and the surplus becomes

$$\tilde{\Phi}(\tilde{x}, \tilde{y}) = \Phi(\tilde{x}, -\tilde{y})$$

We have x and y are matched if and only if

$$F_X(x) = 1 - F_Y(y).$$

#### 2.2 2b. Semi-discrete optimal transport

Inhabitant's problem

$$u\left(x\right) = \max_{j} \left\{x^{\top} y_{j} - v_{j}\right\}$$

I would like to get a formula for the market share of fountain j. Let us compute the aggregate indirect welfare of the consumer. This is

$$\int_{\mathcal{X}} u(x) n(x) dx$$

where n(x) is the density of inhabitants at x. Expressed as a function of the prices, this is

$$F(v) = \int_{\mathcal{X}} \max_{j} \left\{ x^{\top} y_{j} - v_{j} \right\} n(x) dx$$

Claim:

$$\frac{\partial F}{\partial v_i} = -D_j(v)$$

Thus the demand for fountain j is given

$$D_{j}\left(v\right) = -\frac{\partial F\left(v\right)}{\partial v_{j}}.$$

Recall that fountain j has fixed capacity  $q_j$ . Therefore the market-clearing prices  $v_j$  of the fountains are given by

$$D_i(v) = q_i,$$

that is

$$\frac{\partial F\left(v\right)}{\partial v_{j}} + q_{j} = 0$$

which we can rewrite as

$$\frac{\partial}{\partial v_j} \left\{ F(v) + \sum_k q_k v_k \right\} = 0$$

that is v is obtained by minimizing  $S\left(v\right):=F\left(v\right)+\sum_{k}q_{k}v_{k}$  over v. Thus, v is a solution to

$$\min_{v} \left\{ F\left(v\right) + \sum_{k} q_{k} v_{k} \right\}$$

that is

$$\min_{v} \left\{ \int_{\mathcal{X}} \max_{j} \left\{ x^{\top} y_{j} - v_{j} \right\} n\left(x\right) dx + \sum_{k} q_{k} v_{k} \right\}$$

but we can view this as

$$\min_{v} \left\{ \int_{\mathcal{X}} u(x) n(x) dx + \sum_{k} q_{k} v_{k} \right\}$$

$$s.t. \quad u(x) \ge \max_{j} \left\{ x^{\top} y_{j} - v_{j} \right\}$$

which reformates as

$$\begin{aligned} & \underset{v}{\min} & & \left\{ \int_{\mathcal{X}} u\left(x\right) n\left(x\right) dx + \sum_{k} q_{k} v_{k} \right\} \\ & s.t. & & u\left(x\right) + v_{j} \geq x^{\top} y_{j} \ \forall x, \forall j \end{aligned}$$

Tommaso' suggestion.

$$v_{j}^{t+1} = v_{j}^{t} + \varepsilon \left( D_{j} \left( v \right) - q_{j} \right).$$

We have  $S(v) = F(v) + \sum_{k} q_k v_k$ , and therefore

$$\frac{\partial S\left(v\right)}{\partial v_{j}} = -D_{j}\left(v\right) + q_{j}$$

thus this algorithm amounts to

$$v_{j}^{t+1} = v_{j}^{t} - \varepsilon \frac{\partial S(v)}{\partial v_{j}}.$$

This is gradient descent! ie

$$v^{t+1} = v^t - \varepsilon \nabla S(v).$$

Another possibility would be coordinate descent.

$$q_j = D_j \left( v_j^{t+1}, v_{-j}^t \right)$$
  
Parallel version: Jacobi

Sequential version: Gauss-Seidel

Note that 
$$q_j = D_j \left( v_j^{t+1}, v_{-j}^t \right) \text{ is equivalent to } v_j^{t+1} = \arg\min_{v_j} S\left( v_j, v_{-j}^t \right)$$

Final remark. Back to the central planner problem.

The assiment we found consists in mapping inhabitant x with fountain  $y_i$ such that  $j \in \arg\max_{j} \{x^{\top}y_{j} - v_{j}\}.$ 

Noting that  $u(x) = \max_{j} \{x^{\top}y_{j} - v_{j}\}$ , we have that  $\nabla u(x) = y_{j}$ . Thus  $T(x) = \nabla u(x)$  is the optimal assignment from inhabitants to fountains, in the sense that it solves the primal problem

$$\max_{\mu(x,y_j)} - \int |x - y_j|^2 \mu(x, y_j) dx$$

$$s.t. \qquad \int \mu(x, y_j) dx = q_j$$

$$\sum_j \mu(x, y_j) = n(x).$$

c Exercise. Implement a coordinate descent (Gauss-Seidel) version of the algorithm.

## 3 Day 3

#### 3.1 3a. Optimal transport with entropic regularization

Consider the problem

$$\begin{aligned} \max_{\mu \geq 0} & \sum_{xy} \mu_{xy} \Phi_{xy} - \sigma \sum_{xy} \mu_{xy} \ln \mu_{xy} \\ s.t. & \sum_{y} \mu_{xy} = n_x \ [u_x] \\ & \sum_{y} \mu_{xy} = m_y \ [v_y] \end{aligned}$$

Reformulate into

$$\max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy} - \sigma \sum_{xy} \mu_{xy} \ln \mu_{xy} + \min_{(u_x)} \sum_{x} u_x \left( n_x - \sum_{y} \mu_{xy} \right) + \min_{(v_y)} \sum_{y} v_y \left( m_y - \sum_{x} \mu_{xy} \right)$$

that is

$$\min_{u_x,v_y} \max_{\mu \geq 0} \sum_{xy} \mu_{xy} \Phi_{xy} - \sigma \sum_{xy} \mu_{xy} \ln \mu_{xy} + \sum_{x} u_x \left( n_x - \sum_{y} \mu_{xy} \right) + \sum_{y} v_y \left( m_y - \sum_{x} \mu_{xy} \right)$$

hence

$$\min_{u_x,v_y}\max_{\mu\geq 0}\sum_{xy}\mu_{xy}\Phi_{xy} - \sigma\sum_{xy}\mu_{xy}\ln\mu_{xy} + \sum_xu_xn_x - \sum_{xy}\mu_{xy}u_x + \sum_yv_ym_y - \sum_{xy}\mu_{xy}v_y$$

hence

$$\min_{u_x, v_y} \sum_{x} u_x n_x + \sum_{y} v_y m_y + \max_{\mu \ge 0} \sum_{xy} \mu_{xy} \left( \Phi_{xy} - u_x - v_y \right) - \sigma \sum_{xy} \mu_{xy} \ln \mu_{xy}$$

let's compute the soft penalization  $\max_{\mu \geq 0} \sum_{xy} \mu_{xy} (\Phi_{xy} - u_x - v_y) - \sigma \sum_{xy} \mu_{xy} \ln \mu_{xy}$ . This is

$$\sum_{xy} \max_{\mu_{xy} \ge 0} \left\{ \mu_{xy} \left( \Phi_{xy} - u_x - v_y \right) - \sigma \mu_{xy} \ln \mu_{xy} \right\}$$

Set  $a_{xy} = \Phi_{xy} - u_x - v_y$  and compute

$$\max_{\mu \ge 0} \left\{ \mu a - \sigma \mu \ln \mu \right\}$$

We have by first order conditions that

$$a = \sigma (1 + \ln \mu)$$

hence

$$\mu a = \sigma \mu + \sigma \mu \ln \mu$$

thus

$$\mu a - \sigma \mu \ln \mu = \sigma \mu$$

but  $\mu$  can be obtained by  $a = \sigma (1 + \ln \mu)$ 

$$\mu = \exp\left(\frac{a}{\sigma} - 1\right)$$

Therefore we have

$$\max_{\mu \ge 0} \left\{ \mu a - \sigma \mu \ln \mu \right\} = \sigma \exp \left( \frac{a}{\sigma} - 1 \right)$$

Hence the value of the problem is

$$\min_{u_x,v_y} \sum_x u_x n_x + \sum_y v_y m_y + \sum_{xy} \sigma \exp\left(\frac{\Phi_{xy} - u_x - v_y - \sigma}{\sigma}\right)$$

Let's replace  $v_y$  by  $v_y + \sigma$ , and this becomes

$$\min_{u_x, v_y} F\left(u, v\right) := \sum_{x} u_x n_x + \sum_{y} v_y m_y + \sum_{xy} \sigma \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right).$$

Let's look at the first order conditions 
$$n_x - \sum_y \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right) = 0$$
$$m_y - \sum_x \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right) = 0$$

How to recover the primal solution  $\mu_{xy}$  from the dual solution  $(u_x, v_y)$ ? Well, we have

$$\mu_{xy} = \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right).$$

How to solve for this problem?

\* gradient descent:

$$u_x^{t+1} = u_x^t + \epsilon \frac{\partial F}{\partial u_x} = u_x^t + \epsilon \left( n_x - \sum_y \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right) \right)$$
$$v_y^{t+1} = v_y^t + \epsilon \frac{\partial F}{\partial v_y} = v_y^t + \epsilon \left( m_y - \sum_x \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right) \right)$$

\* maximizing with respect to  $u_x$  only means setting the corresponding foc to zero, ie

$$n_x = \sum_{y} \exp\left(\frac{\Phi_{xy} - u_x - v_y}{\sigma}\right)$$

that is

$$n_x = \exp\left(-\frac{u_x}{\sigma}\right) \sum_y \exp\left(\frac{\Phi_{xy} - v_y}{\sigma}\right)$$

that is

$$\exp\left(\frac{-u_x}{\sigma}\right) = \frac{n_x}{\sum_y \exp\left(\frac{\Phi_{xy} - v_y}{\sigma}\right)}$$

ie

$$u_x = \sigma \ln \left( \frac{1}{n_x} \sum_{y} \exp \left( \frac{\Phi_{xy} - v_y}{\sigma} \right) \right)$$

This gives us the following coordinate algorithm

$$\begin{cases} u_x^{t+1} = \sigma \ln \left( \frac{1}{n_x} \sum_y \exp \left( \frac{\Phi_{xy} - v_y^t}{\sigma} \right) \right) \\ v_y^{t+1} = \sigma \ln \left( \frac{1}{m_y} \sum_x \exp \left( \frac{\Phi_{xy} - u_x^{t+1}}{\sigma} \right) \right) \end{cases}$$

or, setting  $A_x = \exp(-u_x/\sigma)$  and  $B_y = \exp(-v_y/\sigma)$ 

$$\begin{cases} A_x^{t+1} = \frac{n_x}{\sum_y \exp\left(\frac{\Phi_{xy}}{\sigma}\right) B_y^t} \\ B_y^{t+1} = \frac{m_y}{\sum_x \exp\left(\frac{\Phi_{xy}}{\sigma}\right) A_x^{t+1}} \end{cases}$$

This is called the Iterated Proportional Fitting Procedure (IPFP), or Sinkhorn's algorithm.

The log-sum-exp trick. We want to compute  $\log (e^a + e^b)$  and potentially a or b – or both – are very large; hence the exponentials will blow up.

Idea= for any constant c,  $\log(e^a + e^b) = c + \log(e^{a-c} + e^{b-c})$ . Indeed,

$$c + \log(e^{a-c} + e^{b-c}) = c + \log(e^{-c}(e^a + e^b)) = c - c + \log(e^a + e^b).$$

Take  $c = \max(a, b)$ , then  $a - c \le 0$  and  $b - c \le 0$  and at least one of these terms is zero. Thus

$$\log (e^a + e^b) = \max \{a, b\} + \log \left( e^{\min\{0, a - b\}} + e^{\min\{0, b - a\}} \right)$$

#### 3.1.1 Generalized linear models

Classical linear regression

$$Y = X\beta + \varepsilon$$
  
expresses as  
 $E[Y|X = x] = x^{\top}\beta$ 

We generalize this into

$$E[Y|X = x] = g^{-1}(X\beta)$$

In particular  $g = \ln$ , yields

$$E[Y|X=x] = \exp(X\beta)$$

How to estimate  $\beta$  then? Based on the observation of  $(x_i, y_i)$ .

We are going to build a method that will match the moments E[Xy] Idea is to compute

$$\min_{\beta} \left\{ \frac{1}{n} \sum_{i} e^{X_{i}\beta} - \frac{1}{n} \sum_{i} y_{i} X_{i}\beta \right\}$$

Population version is

$$\min_{\beta} \left\{ E\left[e^{X\beta}\right] - \hat{E}\left[YX\beta\right] \right\}$$

foc yield

$$E\left[Xe^{X\beta}\right] = \hat{E}\left[YX\right]$$

but by assumption,  $E\left[e^{X\beta}\right]=E\left[Y|X\right]$ , thus for the optimal value of  $\beta$ 

$$E[XY] = \hat{E}[YX]$$

#### 3.2 3b. Random utility models

Decision maker i

Decision maker needs to choose among alternatives  $y \in \mathcal{Y}$ .

Utility of decision maker i taking decision y is

$$U_y + \varepsilon_{iy}$$

where  $\varepsilon_{iy}$  is an individual-specific term which is distributed according to distribution P, and U is a vector called the vector of systematic utilities.

Decision maker i's problem is

$$\max_{y \in \mathcal{Y}} \left\{ U_y + \varepsilon_{iy} \right\}$$

Market share of y

$$Q_{y}(U) = P(U_{y} + \varepsilon_{iy} \ge U_{z} + \varepsilon_{iz} \forall z \in \mathcal{Y})$$

$$\begin{array}{lcl} \frac{\partial Q_y}{\partial U_y} & \geq & 0 \\ \frac{\partial Q_y}{\partial U_z} & \leq & 0, z \neq y \end{array}$$

Compute the mean indirect utilities of all the decision makers. That is

$$\frac{1}{N} \sum_{i} \max_{y \in \mathcal{Y}} \left\{ U_y + \varepsilon_{iy} \right\}$$

In a large population, by the law of large numbers, this becomes

$$G(U) = E_P \left[ \max_{y \in \mathcal{Y}} \left\{ U_y + \varepsilon_y \right\} \right]$$

We have

$$\frac{\partial G\left(U\right)}{\partial U_{y}} = P\left(U_{y} + \varepsilon_{iy} \ge U_{z} + \varepsilon_{iz} \forall z \in \mathcal{Y}\right)$$

indeed,

$$\frac{\partial G(U)}{\partial U_{y}} = E_{P} \left[ \frac{\partial}{\partial U_{y}} \max_{y \in \mathcal{Y}} \left\{ U_{y} + \varepsilon_{y} \right\} \right]$$
$$= E_{P} \left[ 1 \left\{ y \in \arg \max_{y \in \mathcal{Y}} \left\{ U_{y} + \varepsilon_{y} \right\} \right\} \right]$$

This means that

$$Q_{y}\left(U\right) = \frac{\partial G\left(U\right)}{\partial U_{y}}.$$

Assume that we observe  $(q_y)$  the market share of each y and we are after  $(U_y)$ . How do we solve the demand inversion problem, ie how do we recover  $(U_y)$  such that

$$Q_{u}\left(U\right) = q_{u}, \forall y \in \mathcal{Y}$$

We replace Q by its values as the gradient of G, and we have

$$\frac{\partial G\left(U\right)}{\partial U_{y}} = q_{y}, \forall y \in \mathcal{Y}$$

therefore the U's that we are looking for need to solve

$$\max_{(U_y)} \left\{ \sum_{y \in \mathcal{Y}} q_y U_y - G(U) \right\}.$$

This expression plays an important role, and we define

$$G^{*}\left(q\right) = \max_{\left(U_{y}\right)} \left\{ \sum_{y \in \mathcal{Y}} q_{y} U_{y} - G\left(U\right) \right\}$$

and we call  $G^*$  the generalized entropy of choice.

Example. The logit model.

Assume  $(\varepsilon_y)$  are iid random variables, and that the c.d.f of  $\varepsilon_y$  is  $\exp(-\exp(-(z+\gamma)))$  where  $\gamma = \text{Euler's constant}$ .

Then we have

$$G\left(U\right) = \log \sum_{y \in \mathcal{Y}} \exp\left(U_y\right)$$

(remember  $G(U) = E_P \left[ \max_y \left\{ U_y + \varepsilon_y \right\} \right]$ 

More generally, we have

$$E_P\left[\max_{y}\left\{U_y + \sigma \varepsilon_y\right\}\right] = \sigma \log \sum_{y \in \mathcal{Y}} \exp\left(\frac{U_y}{\sigma}\right).$$

Exercise for tomorrow. Show that in the case of the logit model,

$$G^*(q) = \sum_{y \in \mathcal{Y}} q_y \ln q_y \text{ if } \sum_{y \in \mathcal{Y}} q_y = 1 \text{ and } q_y > 0 \text{ for all } y$$
  
=  $+\infty$ 

## 4 Day 4

#### 4.1 4a. Random utility models (ctd)

## 4.2 Nonparametric inversion

Compute

$$G^{*}\left(q\right) = \max_{U_{y}} \left\{ \sum_{y} q_{y} U_{y} - \log \sum \exp U_{y} \right\}$$

Assume an interior solution. By FOC

$$q_y = \frac{e^{U_y}}{\sum_y \exp U_y}$$

thus  $\sum q_y = 1$  and  $q_y > 0$  necessarily. Taking logs,

$$\ln q_y = U_y - \log \sum_y \exp U_y$$

and multiplying by  $q_y$  and summing over y, we get

$$\sum_{y} q_y \ln q_y = \sum_{y} q_y U_y - \left(\sum_{y} q_y\right) \log \sum_{y} \exp U_y$$
$$= \sum_{y} q_y U_y - \log \sum \exp U_y = G^* (q).$$

Now, let's interpret  $G^*(q)$ . When  $q_y = \frac{e^{U_y}}{\sum_y \exp U_y}$ , we have that

$$G^{*}\left(q\right) = \sum_{y} q_{y} U_{y} - G\left(U\right)$$

thus

$$G(U) = \sum_{y} q_{y}U_{y} - G^{*}(q).$$

Recall

$$G(U) = E_P \left[ \max_{y} \left\{ U_y + \varepsilon_y \right\} \right]$$

Introduce  $Y^{\varepsilon} = \arg\max_{y} \{U_{y} + \varepsilon_{y}\}$  – this is the alternative chosen by an agent with  $\varepsilon$ .

$$G(U) = E_{P} \left[ \max_{y} \left\{ U_{y} + \varepsilon_{y} \right\} \right] = E_{P} \left[ U_{Y^{\varepsilon}} + \varepsilon_{Y^{\varepsilon}} \right]$$

$$= E_{P} \left[ U_{Y^{\varepsilon}} \right] + E_{P} \left[ \varepsilon_{Y^{\varepsilon}} \right]$$

$$= \sum_{y} U_{y} P\left( Y^{\varepsilon} = y \right) + E_{P} \left[ \varepsilon_{Y^{\varepsilon}} \right]$$

$$= \sum_{y} q_{y} U_{y} + E_{P} \left[ \varepsilon_{Y^{\varepsilon}} \right]$$

compare this with  $G\left(U\right)=\sum_{y}q_{y}U_{y}-G^{*}\left(q\right)$ , we get that  $-G^{*}\left(q\right)=E_{P}\left[\varepsilon_{Y^{\varepsilon}}\right].$ 

How can we deduce U from  $G^*$ ? Well recall that

$$G^*\left(q\right) = \max_{U_y} \left\{ \sum_{y} q_y U_y - \log \sum \exp U_y \right\}$$

Before we go on, we need to remark that a normalization is needed. Indeed, if  $U_u$  is a solution to

$$q_y = \frac{e^{U_y}}{\sum_y e^{U_y}}$$

then  $(U_y + c)_y$  is also a solution for any constant c.

Assume there is a default option 0 and we call  $Y_0 = Y \cup \{0\}$  the set of all alternatives including the default option, and normalize  $U_0 = 0$ . Then we are back to a unique solution. Consider incorporating the normalization to the problem. We have

$$G(U) = E\left[\max_{y \in Y} \left\{U_y + \varepsilon_y, 0 + \varepsilon_0\right\}\right]$$

In the logit model

$$G(U) = \log \sum_{y \in Y_0} \exp U_y = \log \left(1 + \sum_{y \in Y} \exp U_y\right)$$

now the market shares associated with the nondefault options  $y \in Y$  are given by

$$\frac{\partial G\left(U\right)}{\partial U_{y}} = \frac{\exp U_{y}}{1 + \sum_{y \in Y} \exp U_{y}}$$

and now the generalized entropy of choice can be expressed by

$$G^{*}\left(q\right) = \max_{U \in R^{Y_{0}}} \left\{ \sum_{y \in Y_{0}} q_{y}U_{y} - G\left(U\right) : U_{0} = 0 \right\}$$
$$= \max_{U \in R^{Y}} \left\{ \sum_{y \in Y} q_{y}U_{y} - G\left(U\right) \right\}$$

so I can view  $G^*$  as a function of  $(q_y)_{y\in Y}$  ie the market shares of nondefault alternatives.

In the logit model, we have

$$G^*\left(\left(q\right)_{y\in Y}\right) = \max_{U\in R^Y} \left\{ \sum_{y\in Y} q_y U_y - \log\left(1 + \sum_y \exp U_y\right) \right\}$$

We are getting

$$G^* \left( (q)_{y \in Y} \right) = \sum_{y \in Y} q_y \log q_y + q_0 \log q_0$$

where  $q_0 = 1 - \sum_{y \in Y} q_y$ , if  $q_y > 0$  for all  $y \in Y$  and  $1 > \sum_{y \in Y} q_y$ ;  $+\infty$  otherwise.

In other terms

$$G^*\left(\left(q\right)_{y\in Y}\right) = \sum_{y\in Y} q_y \log q_y + \left(1 - \sum_{y\in Y} q_y\right) \log \left(1 - \sum_{y\in Y} q_y\right).$$

We have that

$$\frac{\partial G^* \left( (q)_{y \in Y} \right)}{\partial q_y} = U_y$$

where U is such that  $Q_y(U) = q_y$ .

In the case of the entropy, we have

$$\begin{array}{rcl} U_y & = & \displaystyle \frac{\partial G^* \left( \left( q \right)_{y \in Y} \right)}{\partial q_y} \\ \\ & = & \displaystyle 1 + \log q_y + \frac{\partial}{\partial q_y} \left\{ q_0 \log q_0 \right\} \\ \\ & = & \displaystyle 1 + \log q_y + \frac{\partial}{\partial q_0} \left\{ q_0 \log q_0 \right\} \times \frac{\partial q_0}{\partial q_y} \\ \\ & = & \displaystyle 1 + \log q_y + \left( 1 + \log q_0 \right) \times -1 \\ \\ & = & \displaystyle \log q_y - \log q_0 \\ \\ & = & \displaystyle \log \frac{q_y}{q_0} \end{array}$$

this is a well-known formula called the log-odds ratio formula.

A bit more on  $G^*$ . Assume now that P is no longer the iid Gumbel distribution – we don't have an explicit way to recover U from q. Instead we will resort to simulated methods.

We have seen that U is identified from  $(q_y)$  by

$$\min_{U} \left\{ G\left(U\right) - \sum_{y} q_{y} U_{y} \right\} \tag{1}$$

What kind of problem is this? Recall the expression for G(U):

$$G\left(U\right)=E_{P}\left[\max_{y}\left\{ U_{y}+\varepsilon_{y}\right\} \right].$$

Take a sample of individuals  $i \in \mathcal{I}$  where  $|\mathcal{I}| = n$ . Each i draws a vector  $(\varepsilon_{iy})_{y \in Y}$  from the distribution P. The sample analog of G becomes

$$G_n(U) = \frac{1}{n} \sum_{i=1}^{n} \max_{y \in Y} \{U_y + \varepsilon_{iy}\}$$

Assume that we observe  $(q_y)_y$ . We would like to recover  $U_y$ . The sample analog of problem (1) is

$$\min_{U} \left\{ G_n\left(U\right) - \sum_{y} q_y U_y \right\}$$

that is

$$\min_{U} \left\{ \frac{1}{n} \sum_{i=1}^{n} \max_{y \in Y} \left\{ U_{y} + \varepsilon_{iy} \right\} - \sum_{y} q_{y} U_{y} \right\}$$

Introduce  $v_y = -U_y$ , and we get

$$\min_{v_y} \left\{ \frac{1}{n} \sum_{i=1}^n \max_{y \in Y} \left\{ -v_y + \varepsilon_{iy} \right\} + \sum_y q_y v_y \right\}$$

Claim: this rewrites as

$$\min_{u_i, v_y} \left\{ \sum_{i=1}^n \frac{1}{n} u_i + \sum_y q_y v_y \right\}$$
s.t. 
$$u_i + v_y \ge \varepsilon_{iy}$$

Indeed, the constraint in the latter problem implies  $u_i \ge \max_y \{-v_y + \varepsilon_{iy}\}$ , but optimality implies that this holds as an equality.

Hence: inverting a discrete choice model is an optimal transport problem. See G+Salanie – Cupids invisible hand.

# 4.3 Parametric estimation – characteristics approach / logit model

For now, let's assume we observe no characteristics about the decision makers. We do observe characteristics about alternatives.

 $U_y = (\Phi \beta)_y$  where for each alternative y,  $\Phi_{yk}$  is the k-th characteristics associated with y.

For example – y are car models, and  $\Phi_{y1}$  is fuel efficiency,  $\Phi_{y2}$  is number of seats,  $\Phi_{y3}$  is a safety rating,

We are assuming that  $U_y = \sum_k \Phi_{yk} \beta_k$ , where  $\beta_k$  are coefficients to be estimated

Assume that the random utilities are iid Gumbel.

Then the overall indirect utility is

$$G(\Phi\beta) = \log \sum_{y} \exp(\Phi\beta)_{y}$$

Let's derive  $G(\Phi\beta)$  with respect to  $\beta_k$ . We have

$$\frac{\partial G\left(\Phi\beta\right)}{\partial \beta_{k}} = \sum_{y} \frac{\partial G}{\partial U_{y}} \left(\Phi\beta\right) \frac{\partial U_{y}}{\partial \beta_{k}}$$

$$= \sum_{y} Q_{y} \left(\Phi\beta\right) \Phi_{yk}$$

$$= E_{\beta} \left[\Phi_{Yk}\right]$$

this is the moment of the k-th characteristics predicted by the model with parameter  $\beta$ .

I am observing the market shares  $\hat{q}_y$  of the models actually purchased, so my estimation strategy will consist of matching those moments, i.e. looking for the value of the parameter  $\beta$  such that the predicted moments match with the observed moments. That is, look for  $\beta$  such that

$$E_{\beta} \left[ \Phi_{Yk} \right] = \hat{E} \left[ \Phi_{Yk} \right], i.e.$$

$$\sum_{y} Q_{y} \left( \Phi \beta \right) \Phi_{yk} = \sum_{y} \hat{q}_{y} \Phi_{yk}$$

ie

$$\frac{\partial}{\partial \beta_k} G(\Phi \beta) = \sum_{y} \hat{q}_y \Phi_{yk}$$

which are moment matching conditions: predicted moments=observed moments.

We are going to view this as the first order conditions associated with a convex optimization problem and use that reformulation to estimate  $\beta$ .

This is the idea of generalized linear models (GLM). Let's see how it works. Introduce

$$\max_{\beta} \left\{ \sum_{k} \sum_{y} \beta_{k} \hat{q}_{y} \Phi_{yk} - G\left(\Phi\beta\right) \right\}$$

THe first order conditions for this problem are the moment-matching conditions above. The problem is a convex optimization problem, so gradient descent (or extensions) will lead to the estimator of  $\beta$ .

In the logit model, we have

$$\max_{\beta} \left\{ \sum_{k} \hat{q}_{y} \left( \Phi \beta \right)_{y} - \log \sum_{y} \exp \left( \left( \Phi \beta \right)_{y} \right) \right\}$$

which, when

Alternative approach via maximum likelihood. The probability that a decision maket picks up choice y is

$$l_{y}\left(\beta\right) = \frac{\partial G}{\partial U_{y}}\left(\Phi\beta\right)$$

Thus the likelihood of the sample is (up to rescaling)

$$\sum_{y} \hat{q}_{y} \ln l_{y} (\beta) = \sum_{y} \hat{q}_{y} \ln \frac{\partial G}{\partial U_{y}} (\Phi \beta)$$

and max-likelihood consists of

$$\max_{\beta} \sum_{y} \hat{q}_{y} \ln \frac{\partial G}{\partial U_{y}} \left( \Phi \beta \right).$$

In general (ie. for general structures of distribution of random utility), this is NOT a convex optimization problem.

HOWEVER, in the logit case, both approached coincide. Indeed, in that case,

$$l_{y}\left(\beta\right) = \frac{\partial G}{\partial U_{y}}\left(\Phi\beta\right) = \frac{e^{\left(\Phi\beta\right)_{y}}}{\sum_{y} e^{\left(\Phi\beta\right)_{y}}}$$

and thus

$$\sum_{y} \hat{q}_{y} \ln l_{y} (\beta) = \sum_{y} \hat{q}_{y} \left( (\Phi \beta)_{y} - \log \left( \sum_{y} e^{(\Phi \beta)_{y}} \right) \right)$$

so in this case (but in this case only), this is

$$\sum_{y} \hat{q}_{y} \ln l_{y} (\beta) = \sum_{y} \hat{q}_{y} \left( (\Phi \beta)_{y} - G (\Phi \beta) \right).$$

#### 4.4 With characteristics about decision-makers observed.

Let us now assume we observe some characteristics such as income about decision maker i.

The likelihood that individual i will make decision y is now

$$\frac{\exp\left(\left(\Phi\beta\right)_{iy}\right)}{\sum_{y}\exp\left(\left(\Phi\beta\right)_{iy}\right)}$$

where  $(\Phi \beta)_{iy} = \sum_k \Phi_{iy,k} \beta_k$ . We now observe which decision maker makes which decision, that is we observe

$$\hat{\mu}_{iy} = 1 \{ i \text{ chooses } y \}.$$

Let us compute the log-likelihood of the sample. This is

$$\sum_{iy} \hat{\mu}_{iy} \log \frac{\exp\left((\Phi\beta)_{iy}\right)}{\sum_{y} \exp\left((\Phi\beta)_{iy}\right)}$$

which is

$$\sum_{iy} \hat{\mu}_{iy} \left( (\Phi \beta)_{iy} - \log \sum_{y} \exp \left( (\Phi \beta)_{iy} \right) \right)$$

first order condition is

$$\sum_{iy} \hat{\mu}_{iy} \Phi_{iy,k} = \sum_{iy} \frac{\exp\left((\Phi\beta)_{iy}\right)}{\sum_{y} \exp\left((\Phi\beta)_{iy}\right)} \Phi_{iy,k}$$

How to reformulate as a Poisson regression?

#### 4.5 Reminders on Poisson regression

Assume that  $\mu_{iy}|\Phi$  follows a Poisson distribution of parameter  $\theta_{iy} = e^{(\Phi\beta)_{iy}}$ .

$$\Pr\left(\mu_{iy}|\Phi_{iy}\right) = \frac{\theta_{iy}^{\mu_{iy}}}{\mu_{iy}!}e^{-\theta_{iy}}$$

Let's compute the log-likelihood of z. We have

$$l\left(\hat{\mu}_{iy}|\Phi_{iy}\right) = \hat{\mu}_{iy}\log\theta_{iy} - \theta_{iy} - \log\left(\hat{\mu}_{iy}!\right)$$

that is

$$l_{iy}\left(\hat{\mu}_{iy}|\Phi_{iy}\right) = \hat{\mu}_{iy}\left(\Phi\beta\right)_{iy} - e^{(\Phi\beta)_{iy}} - \log\left(\hat{\mu}_{iy}!\right)$$

Hence the conditional log-likelihood of the sample is

$$\sum_{iy} \hat{\mu}_{iy} \left(\Phi \beta\right)_{iy} - \sum_{iy} e^{\left(\Phi \beta\right)_{iy}} - \sum_{iy} \log \left(\hat{\mu}_{iy}!\right)$$

therefore the Poisson regression consists of the MLE for Poisson distribution that is

$$\max_{\beta} \left\{ \sum_{iy} \hat{\mu}_{iy} \left( \Phi \beta \right)_{iy} - \sum_{iy} e^{(\Phi \beta)_{iy}} \right\}$$

Start from the Poisson regression and introduce an individual fixed effect.  $\sum_k \Phi_{iyk} \beta_k$  is now replaced by  $\sum_k \Phi_{iyk} \beta_k - u_i$ , and the parameter  $\beta$  is replaced by  $\theta = \left(\beta^\top, u^\top\right)^\top$ .

How do we write this as  $X\theta$ ? which matrix should we take for X?

$$X = \begin{pmatrix} \Phi_{iy,k} & -(I_I \otimes 1_Y) \end{pmatrix}$$

$$\begin{aligned} & \text{[[ indeed, recall that } \\ & \max vec\left(\mu\right)^{\top}vec\left(\Phi\right) \\ & \text{s.t.} \\ & \binom{M_X}{M_Y}vec\left(\mu\right) = \binom{n}{m} \\ & \text{where} \\ & M_X = I_X \otimes 1_Y^{\top} \\ & M_Y = 1_X^{\top} \otimes I_Y \end{aligned}$$

Dual 
$$\min \sum n_x u_x + \sum m_y v_y$$

$$u_x + v_y \ge \Phi_{xy}$$
can be written in a matrix form 
$$\max_{(u,v)} \binom{n}{m}^\top \binom{u}{v}$$
s.t.
$$\binom{M_X^\top M_Y^\top}{m} vec \left(\binom{u}{v}\right) \ge vec \left(\Phi\right)$$
this  $\binom{M_X^\top u}{m}_{xy} = u_x$ 
that is,  $((I_X \otimes 1_Y) u)_{xy} = u_x$ 
]

The Poisson regression now becomes

$$\max_{\beta_k, u_i} \left\{ \sum_{iy} \hat{\mu}_{iy} \left( \sum_k \Phi_{yk} \beta_k - u_i \right) - \sum_{iy} e^{(\Phi\beta)_{iy} - u_i} \right\}$$

Now we have to maximize over  $\beta_k$  the following expression

$$\max_{u_i} \left\{ \sum_{iy} \hat{\mu}_{iy} \left( \sum_k \Phi_{yk} \beta_k - u_i \right) - \sum_{iy} e^{(\Phi\beta)_y - u_i} \right\}$$

that is

$$\sum_{iyk} \hat{\mu}_{iy} \Phi_{yk} \beta_k + \max_{u_i} \left\{ \sum_{iy} \hat{\mu}_{iy} u_i - \sum_{iy} e^{(\Phi\beta)_{iy} - u_i} \right\}$$

by first order conditions, we have

$$\sum_{y} \hat{\mu}_{iy} = \sum_{y} e^{(\Phi\beta)_{iy} + u_i}$$

that is

$$1 = \sum_{y} e^{(\Phi\beta)_{iy} + u_i}$$

that is

$$e^{u_i} = \frac{1}{\sum_y e^{(\Phi\beta)_{iy}}}$$

hence  $u_i = -\log\left(\sum_y e^{(\Phi\beta)_{iy}}\right)$  and the expression becomes

$$\sum_{iyk} \hat{\mu}_{iy} \Phi_{yk} \beta_k + \max_{u_i} \left\{ \sum_{iy} \hat{\mu}_{iy} u_i - \sum_{iy} e^{(\Phi\beta)_y + u_i} \right\}$$

$$\sum_{i} \max_{\beta} \left\{ \sum_{k} \sum_{y} \beta_{k} \hat{q}_{y} \Phi_{yk} - \log \sum_{y} \exp \left( (\Phi \beta)_{y} \right) \right\}$$

Hence logit model = Poisson regression + consumer fixed effect.

## 5 Day 5

## 5.1 Poisson regression with 2-way fixed effects

Now introduce both i- and y- fixed-effects. We have

Start from the Poisson regression and introduce an individual fixed effect.

 $\sum_k \Phi_{iyk} \beta_k$  is now replaced by  $\sum_k \Phi_{iyk} \beta_k - u_i - v_y$ , and the parameter  $\beta$  is replaced by  $\theta = (\beta, u, v)$ .

Remember,  $\hat{\mu}_{iy}$  is indicator that individual i chose y.

Let's write down the Poisson regression.

$$\max_{\beta_k, u_i, v_y} \sum_{i, y} \hat{\mu}_{iy} \left( (\Phi \beta)_{iy} - u_i - v_y \right) - \sum_{i, y} \exp \left( (\Phi \beta)_{iy} - u_i - v_y \right)$$

Let's focus on the maximum over  $u_i$  and  $v_j$  for fixed  $\beta$ . We have

$$\max_{\beta} \left\{ \sum_{iy} \hat{\mu}_{iy} \left( \Phi \beta \right)_{iy} - W \left( \beta \right) \right\}$$

$$W(\beta) = -\max_{u_i, v_y} \left\{ \sum_{iy} \hat{\mu}_{iy} \left( -u_i - v_y \right) - \sum_{iy} \exp\left( (\Phi \beta)_{iy} - u_i - v_y \right) \right\}$$

and the latter problem becomes

$$W(\beta) = \min_{u_i, v_y} \left\{ \sum_{iy} \hat{\mu}_{iy} \left( u_i + v_y \right) + \sum_{iy} \exp \left( (\Phi \beta)_{iy} - u_i - v_y \right) \right\}$$

But  $\sum_{iy} \hat{\mu}_{iy} (u_i + v_y) = \sum_{iy} \hat{\mu}_{iy} u_i + \sum_{iy} \hat{\mu}_{iy} v_y = \sum_i u_i \sum_y \hat{\mu}_{iy} + \sum_y v_y \sum_i \hat{\mu}_{iy}$  but  $\sum_y \hat{\mu}_{iy} = 1$  and  $\sum_i \hat{\mu}_{iy} = q_y$  which is the total demand for good y. Rewrite the above problem as

$$W(\beta) = \min_{u_i, v_y} \left\{ \sum_i u_i + \sum_y q_y v_y + \sum_{iy} \exp\left((\Phi \beta)_{iy} - u_i - v_y\right) \right\}$$

## 5.2 5a. Estimation of matching models

Start with a matching model with individual men i and women j (one per index).

Assume that the observable characteristics of man i is  $x_i \in X$  finite set.

Assume that the observable characteristics of woman j is  $y_j \in Y$  finite set.

The analyst only observes the  $x_i$ 's and  $y_j$ 's.

Assume that the mass of men of type x is  $n_x$  and the mass of women of type y is  $m_y$ .

Utilities. Assume that if i and j match and decide on a transfer  $w_{ij}$  from j to i, then

 $i \text{ gets } \alpha_{ij} + w_{ij}$ 

 $j \text{ gets } \gamma_{ij} - w_{ij}$ 

and if they remain unmatched, they get respectively  $\alpha_{i0}$  and  $\gamma_{0j}$ .

Crucial assumption: separability. Assume that

 $\alpha_{ij} = \alpha_{x_i y_j} + \varepsilon_{i y_j}$ 

 $\gamma_{ij} = \gamma_{x_i y_j} + \eta_{x_i j}$ 

 $\alpha_{i0} = \varepsilon_{i0}$ 

 $\gamma_{0j} = \eta_{0j}$ 

G+Salanie (2022). Let's solve for this model. Let's write the dual optimal assignment problem.

$$\begin{split} \min_{u_i,v_j} & & \sum_{i \in I} u_i + \sum_{j \in J} v_j \\ s.t. & & u_i + v_j \geq \Phi_{ij} = \Phi_{x_iy_j} + \varepsilon_{iy_j} + \eta_{x_ij} \\ & & u_i \geq \varepsilon_{i0} \\ & & v_j \geq \eta_{0j} \end{split}$$

where

$$\Phi_{xy} = \alpha_{xy} + \gamma_{xy}$$

We can rewrite as

$$\begin{split} \min_{u_i,v_j} & & \sum_{i \in I} u_i + \sum_{j \in J} v_j \\ s.t. & & u_i - \varepsilon_{iy_j} + v_j - \eta_{x_ij} \geq \Phi_{x_iy_j} \\ & & u_i \geq \varepsilon_{i0} \\ & & v_j \geq \eta_{0j} \end{split}$$

Now  $u_i - \varepsilon_{iy_j} + v_j - \eta_{x_ij} \ge \Phi_{x_iy_j}$  means that

$$\underbrace{\min_{i:x_i=x} \left\{ u_i - \varepsilon_{iy} \right\}}_{U_{xy}} + \underbrace{\min_{j:y_j=y} \left\{ v_j - \eta_{x_ij} \right\}}_{V_{xy}} \ge \Phi_{xy}$$

and I claim that the problem reformulates as

$$\begin{aligned} \min_{u_i,v_j,U_{xy},V_{xy}} & & \sum_{i\in I} u_i + \sum_{j\in J} v_j \\ s.t. & & U_{xy} + V_{xy} \geq \Phi_{xy} \\ & & u_i \geq U_{xy} + \varepsilon_{iy} \\ & & u_i \geq \varepsilon_{i0} \\ & & v_j \geq V_{xy} + \eta_{xj} \\ & & v_j \geq \eta_{0j} \end{aligned}$$

 $U_{xy} = \min_{i:x_i=x} \{u_i - \varepsilon_{iy}\}$  implies that  $u_i \ge U_{xy} + \varepsilon_{iy}$ Finally, the problem rewrites as

$$\min_{U_{xy}, V_{xy}} \qquad \sum_{i \in I} \max \{U_{xy} + \varepsilon_{iy}, \varepsilon_{i0}\} + \sum_{j \in J} \max \{V_{xy} + \eta_{xj}, \eta_{0j}\}$$

$$s.t. \qquad U_{xy} + V_{xy} \ge \Phi_{xy}$$

actually, we may without changing the value problem saturate the constraint, that is

$$\min_{U_{xy}, V_{xy}} \qquad \sum_{i \in I} \max \{ U_{xy} + \varepsilon_{iy}, \varepsilon_{i0} \} + \sum_{j \in J} \max \{ V_{xy} + \eta_{xj}, \eta_{0j} \}$$

$$s.t. \qquad U_{xy} + V_{xy} = \Phi_{xy}$$

We have

$$\begin{split} \min_{U_{xy},V_{xy}} \quad & \sum_{x \in X} n_x \frac{1}{n_x} \sum_{i:x_i = x} \max_y \left\{ U_{xy} + \varepsilon_{iy}, \varepsilon_{i0} \right\} + \sum_{y \in Y} m_y \frac{1}{m_y} \sum_{j:y_j = y} \max_x \left\{ V_{xy} + \eta_{xj}, \eta_{0j} \right\} \\ s.t. \quad & U_{xy} + V_{xy} = \Phi_{xy} \end{split}$$

Now assume that there are many i per x and j per y, and assume that  $\varepsilon_{iy}$  is drawn from a probability distribution P. By the law of large numbers

$$\frac{1}{n_x} \sum_{i:x_i = x} \max_y \left\{ U_{xy} + \varepsilon_{iy}, \varepsilon_{i0} \right\}$$
 becomes 
$$G_x \left( U \right) = E_P \left[ \max_y \left\{ U_{xy} + \varepsilon_y, \varepsilon_0 \right\} \right]$$
 and similarly on the other side of the market, assuming  $\eta \sim Q$  
$$\frac{1}{m_y} \sum_{j:y_j = y} \max_x \left\{ V_{xy} + \eta_{xj}, \eta_{0j} \right\}$$
 becomes

 $G_y\left(V\right)=E_Q\left[\max_x\left\{V_{xy}+\eta_x,\eta_0\right\}\right],$  and therefore the matching problem becomes

$$\begin{aligned} \min_{U_{xy},V_{xy}} & & & \sum_{x \in X} n_x G_x \left( U \right) + \sum_{y \in Y} m_y G_y \left( V \right) \\ s.t. & & & U_{xy} + V_{xy} = \Phi_{xy} \ \left[ \mu_{xy} \right] \end{aligned}$$

Write this as

$$\min_{U_{xy},V_{xy}} \max_{\mu_{xy}} \sum_{x \in X} n_x G_x \left( U \right) + \sum_{y \in Y} m_y G_y \left( V \right) + \sum_{xy} \mu_{xy} \left( \Phi_{xy} - U_{xy} - V_{xy} \right)$$

thus

$$\max_{\mu_{xy}} \min_{U_{xy}, V_{xy}} \sum_{x \in X} n_x G_x\left(U\right) + \sum_{y \in Y} m_y G_y\left(V\right) + \sum_{xy} \mu_{xy} \left(\Phi_{xy} - U_{xy} - V_{xy}\right)$$

therefore by first order conditions wrt U and V in the inner problem, we have that

$$\min_{U_{xy},V_{xy}} \max_{\mu_{xy}} \sum_{x \in X} n_x G_x \left( U \right) + \sum_{y \in Y} m_y G_y \left( V \right) + \sum_{xy} \mu_{xy} \left( \Phi_{xy} - U_{xy} - V_{xy} \right)$$

$$\frac{\partial}{\partial U_{xy}} \left\{ \sum_{x \in X} n_x G_x \left( U \right) \right\} = \mu_{xy}$$

that is

$$\mu_{xy} = n_x \frac{\partial G_x \left( U \right)}{\partial U_{xy}}$$

and similarly,

$$\mu_{xy} = m_y \frac{\partial G_y \left( V \right)}{\partial V_{xy}}$$

hence we get that equilibrium is given by

$$n_{x}\frac{\partial G_{x}\left(U\right)}{\partial U_{xy}}=m_{y}\frac{\partial G_{y}\left(\Phi-U\right)}{\partial V_{xy}}.$$

that is

$$\max_{\mu_{xy}} \sum_{xy} \mu_{xy} \Phi_{xy} - \max_{U_{xy}, V_{xy}} \left\{ -\sum_{x \in X} n_x G_x (U) - \sum_{y \in Y} m_y G_y (V) + \sum_{xy} \mu_{xy} (U_{xy} + V_{xy}) \right\}$$

that

$$\max_{\mu_{xy}} \left( \begin{array}{c} \sum_{xy} \mu_{xy} \Phi_{xy} \\ -\max_{U_{xy}} \left\{ \sum_{xy} \mu_{xy} U_{xy} - \sum_{x \in X} n_x G_x(U) \right\} \\ -\max_{V_{xy}} \left\{ \sum_{xy} \mu_{xy} V_{xy} - \sum_{y \in Y} m_y G_y(V) \right\} \end{array} \right)$$

that is

$$\max_{\mu_{xy}} \left( \begin{array}{c} \sum_{xy} \mu_{xy} \Phi_{xy} \\ -\sum_{x} n_{x} \max_{U_{xy}} \left\{ \sum_{xy} \frac{\mu_{xy}}{n_{x}} U_{xy} - G_{x} \left( U \right) \right\} \\ -\sum_{y} m_{y} \max_{V_{xy}} \left\{ \sum_{xy} \frac{\mu_{xy}}{m_{y}} V_{xy} - G_{y} \left( V \right) \right\} \end{array} \right)$$

Notice that

$$G_{x}^{*}\left(\left(\frac{\mu_{xy}}{n_{x}}\right)_{y\in Y}\right) = \max_{U_{xy}}\left\{\sum_{xy}\left(\frac{\mu_{xy}}{n_{x}}\right)U_{xy} - G_{x}\left(U\right)\right\}$$

is called the entropy of choice of men of type x. Hence the problem reformulates as

$$\max_{\mu_{xy}} \left( \sum_{xy} \mu_{xy} \Phi_{xy} - \sum_{x} n_x G_x^* \left( \frac{\mu_x}{n_x} \right) - \sum_{y} m_y G_y^* \left( \frac{\mu_{,y}}{m_y} \right) \right)$$

which is the "dual of the dual", hence the primal aggregate problem. Let us work out the first order conditions in the primal problem. We have

$$\Phi_{xy} = \frac{\partial G_x^*}{\partial \mu_{y|x}} \left( \frac{\mu_{x.}}{n_x} \right) + \frac{\partial G_y^*}{\partial \mu_{x|y}} \left( \frac{\mu_{.y}}{m_y} \right)$$

To recap:

\* first order conditions in the dual problem gave us solution to the assignment problem: predict the optimal matching  $\mu_{xy}$  as a function of the utilities.

\* first order conditions in the primal problem gave us the solution to the inverse problem: get utilities  $\Phi$  based on the observation of matching patterns  $\mu_{xy}$ .

Now, assume a logit structure (Choo and Siow 2006). Recall that we had obtained by foc in the dual problem

$$\mu_{xy} = n_x \frac{\partial G_x \left( U \right)}{\partial U_{xy}}$$

and similarly,

$$\mu_{xy} = m_y \frac{\partial G_y\left(V\right)}{\partial V_{xy}}$$

In the case of the logit model, we have that the mass of xy matches is

$$\mu_{xy} = n_x \frac{e^{U_{xy}}}{e^{u_x}}$$

where  $e^{u_x}=1+\sum_{y\in Y}e^{U_{xy}}$  ie  $u_x=\log\left(1+\sum e^{U_{xy}}\right)$  and the mass of unmatched x's is

$$\mu_{x0} = n_x \frac{1}{e^{u_x}}$$

therefore

$$\mu_{xy} = \mu_{x0} e^{U_{xy}}.$$

Same thing on the side of women:

$$\mu_{xy} = \mu_{0y} e^{V_{xy}}.$$

by multiplying term-by-term the two inequalities, we get

$$\mu_{xy}^2 = \mu_{x0}\mu_{0y}e^{\Phi_{xy}}$$

which is Choo and Siow's formula.

Side remark. Note the inverse problem (ie recovering  $\Phi$  from  $\mu$ ) is easier than the direct problem (ie obtaining  $\mu$  from  $\Phi$ ). Indeed, we have

$$\Phi_{xy} = \log \frac{\mu_{xy}^2}{\mu_{x0}\mu_{0y}}.$$

What about the direct problem? Well,  $\mu_{x0}$  and  $\mu_{0y}$  are the solutions to

$$n_x = \mu_{x0} + \sum_y \sqrt{\mu_{x0}\mu_{0y}} e^{\frac{\Phi_{xy}}{2}}$$

$$m_y = \mu_{0y} + \sum_x \sqrt{\mu_{x0}\mu_{0y}} e^{\frac{\Phi_{xy}}{2}}$$

It turns out that this is the solution to

$$W\left(\Phi\right) = \min_{u,v} \left\{ \begin{array}{c} \sum_{x} n_{x} u_{x} + \sum_{y} m_{y} v_{y} \\ +2 \sum_{xy} \sqrt{n_{x} m_{y}} \exp\left(\frac{\Phi_{xy} - u_{x} - v_{y}}{2}\right) \\ + \sum_{x} n_{x} \exp\left(-u_{x}\right) + \sum_{y} m_{y} \exp\left(-v_{y}\right) \end{array} \right\}$$

[Recall we want to have  $\mu_{xy}=\sqrt{n_xm_y}\exp\left(\frac{\Phi_{xy}-u_x-v_y}{2}\right)$  and  $\mu_{x0}=n_x\exp\left(-u_x\right)$  and ...]

By FOC

$$n_x = n_x \exp(-u_x) + \sum_y \sqrt{n_x m_y} \exp\left(\frac{\Phi_{xy} - u_x - v_y}{2}\right)$$

$$m_y = m_y \exp(-v_y) + \sum_x \sqrt{n_x m_y} \exp\left(\frac{\Phi_{xy} - u_x - v_y}{2}\right)$$

QED.

#### 5.2.1 Estimation

Take a parameterization  $\Phi_{xy}^{\beta} = \sum_{k} \Phi_{xy,k} \beta_{k}$ . Denote

$$W(\beta) = \min_{u,v} \left\{ \begin{array}{c} \sum_{x} n_{x} u_{x} + \sum_{y} m_{y} v_{y} \\ +2 \sum_{xy} \sqrt{n_{x} m_{y}} \exp\left(\frac{\Phi_{xy}^{\beta} - u_{x} - v_{y}}{2}\right) \\ +\sum_{x} n_{x} \exp\left(-u_{x}\right) + \sum_{y} m_{y} \exp\left(-v_{y}\right) \end{array} \right\}.$$

Set 
$$\theta = (\beta, u, v)$$

Exercise. Compute  $\partial W(\beta)/\partial \beta_k$ . We have

$$\frac{\partial W(\beta)}{\partial \beta_k} = \sum_{xy} \sqrt{n_x m_y} \exp\left(\frac{\Phi_{xy}^{\beta} - u_x - v_y}{2}\right) \Phi_{xy,k}$$
$$= \sum_{xy} \mu_{xy}^{\theta} \Phi_{xy,k}$$

GLM idea – As always, we will look for  $\theta$  in order to match the predicted moments with the observed ones.

I.e.

$$\sum_{xy} \mu_{xy}^{\theta} \Phi_{xy,k} = \sum_{xy} \hat{\mu}_{xy} \Phi_{xy,k}$$

$$\sum_{y} \mu_{xy}^{\theta} = \sum_{y} \hat{\mu}_{xy} = n_{x}$$

$$\sum_{x} \mu_{xy}^{\theta} = \sum_{x} \hat{\mu}_{xy} = m_{y}$$

We are going to estimate  $\beta$  (and u and v) by

$$\max_{\beta} \left\{ \sum_{xyk} \hat{\mu}_{xy} \Phi_{xy,k} \beta_k - W(\beta) \right\}$$

that is

$$\max_{\beta,u,v} \left\{ \begin{array}{l} \sum_{xyk} \hat{\mu}_{xy} \Phi_{xy,k} \beta_k - \sum_x n_x u_x - \sum_y m_y v_y \\ -2 \sum_{xy} \sqrt{n_x m_y} \exp\left(\frac{\Phi_{xy}^{\beta} - u_x - v_y}{2}\right) \\ -\sum_x n_x \exp\left(-u_x\right) - \sum_y m_y \exp\left(-v_y\right) \end{array} \right\}$$

thus

$$\max_{\beta,u,v} \left\{ \begin{array}{c} \sum_{xyk} \hat{\mu}_{xy} \left( \Phi_{xy,k} \beta_k - u_x - v_y \right) \\ -2 \sum_{xy} \sqrt{n_x m_y} \exp\left( \frac{\Phi_{xy}^{\beta} - u_x - v_y}{2} \right) \\ -\sum_{x} n_x \exp\left( -u_x \right) - \sum_{y} m_y \exp\left( -v_y \right) \end{array} \right\}$$

## 6 Special Lecture 1, Jan 28

Class of problem = min-cost flow problems.

Silk road example.

Total set of nodes is Z

Set of producer cities  $X \subset Z$ : city x produces fixed quantity  $n_x$  of silk

Set of consumer cities  $Y \subset Z$ : city  $y \in Y$  consumes fixed quantity  $m_y$  of silk.

Transit nodes are those in Z that are not in X or Y

(Directed) network. A network is a set of nodes Z, and arcs A.  $xy \in A \subset Z \times Z$  is an arc with origin x and destination y.

A flow is the specification of the mass of commodity  $\mu_a \geq 0$  that flows across each arc a.

Hence a flow is an element of  $R_{+}^{A}$ .

Mass balance equation: for each node  $z \in Z$ 

everything that arrives - everything that leaves = everything consumed - everything produced  $\,$ 

That is for

$$\sum_{x:xz\in A}\mu_{xz}-\sum_{y:zy\in A}\mu_{zy}=q_z, \forall z\in Z$$

where  $q_z$  is everything consumed at z minus everything produced at z.

In particular, if  $z \in X$  is a producer node, then  $q_z = -n_z$ 

if  $z \in Y$  is a consumer node, then  $q_z = m_y$ 

if  $z \in \mathbb{Z} \setminus \{X \cup Y\}$  is a transit node, then  $q_z = 0$ 

We would like to write equations

$$\sum_{x:xz\in A}\mu_{xz}-\sum_{y:zy\in A}\mu_{zy}=q_z, \forall z\in Z$$

as

$$D.\mu = q$$

for some matrix D. What should D be?

- \* D should be a  $|Z| \times |A|$  matrix.
- \* We should have

 $D_{z,a} = 1$  if z is the endpoint of a; and -1 if z is the startpoint of a.

Note that  $D^{\top}$  is  $|A| \times |Z|$  matrix. And for  $p \in \mathbb{R}^{Z}$  we have

$$\left(\boldsymbol{D}^{\top}\boldsymbol{p}\right)_{xy} = \sum_{z} \left(\boldsymbol{D}^{\top}\right)_{xy,z} p_{z} = \sum_{z} D_{z,xy} p_{z} = p_{y} - p_{x}$$

so that we have

$$\left(D^{\top}p\right)_{xy} = p_y - p_x$$

which is why I will now on refer to  $D^{\top}$  as  $\nabla$ . It is called the node-arc incidence matrix. And we have that mass conservation is expressed as

$$\nabla^{\top} \mu = q.$$

Assume that now we need to dispatch the silk in a cost-efficient manner. The cost of using flow  $\mu$  is assumed to be

$$\sum_{a} \mu_a c_a = c^{\top} \mu$$

the cost-efficient dispatch is then

$$\begin{aligned} & \min_{\mu_a \geq 0} & & c^\top \mu \\ & s.t. & \nabla^\top \mu = & q \end{aligned}$$

(Just a remark, we need to have  $\sum_z q_z = 0$ ). 1. Recovering optimal transport. Consider the situation where there are no transit nodes, ie  $Z = X \cup Y$ .

and assume that there is an arc from every x to every Y and only between these, ie  $A = X \times Y$ .

In this case  $\mu_a$  identifies with  $\mu_{xy}$ , and mass conservation  $(\nabla^{\top}\mu) = q_z$ reads as:

$$\sum_{y:zy\in A}\mu_{zy} = n_z, z\in X$$
 
$$\sum_{x:xz\in A}\mu_{xz} = m_z, z\in Y$$

- 2. (Left as an exercise) What about matching problem with unassigned agents / partial assignements?
- 3. Shortest path problems: the problems solved by Google maps: What is the shortest way to go from s to t.

Take  $q_s = -1$  and  $q_t = 1$  and  $q_z = 0$  for  $z \neq s, t$ . and consider

$$\begin{aligned} & \min_{\mu_a \geq 0} & & c^\top \mu \\ & s.t. & \nabla^\top \mu = & q \end{aligned}$$

Duality and its interpretation. Recall the primal problem

$$\begin{aligned} & \min_{\mu_a \geq 0} & & c^\top \mu \\ & s.t. & \nabla^\top \mu = & q \end{aligned}$$

Dual problem Minimax formulation

$$\min_{\mu_a \geq 0} \max_{p_z} c^\top \mu + p^\top \left( q - \nabla^\top \mu \right) = \max_{p_z} p^\top q + \min_{\mu_a \geq 0} \left( c - \nabla p \right)^\top \mu$$

thus the dual is

$$\max_{p_z} \qquad p^{\top} q$$
 
$$s.t. \quad \nabla p \leq c \ [\mu \geq 0]$$

Interpret  $p_z$  as the price of the silk at z. How do we interpret  $\nabla p \leq c$ ?

$$(\nabla p)_{xy} \le c_{xy}$$

$$p_y - p_x \le c_{xy}$$

$$p_y \le p_x + c_{xy}.$$

Carry trade = buy at x, ship to y, sell at y. Profit =  $p_y - p_x - c_{xy} \le 0$  unless there is an arbitrage.

Thus the condition interprets as no arbitrage

Complementary slackness:  $\mu_{xy} > 0 \implies p_y - p_x = c_{xy}$  – this is interpreted as individual rationality.

Finally, how do we interpret  $p^{\top}q = \sum_{z} p_{z}q_{z}$ 

 $[\nabla^{\top}\nabla$  is the analog of a Laplacian]

## 7 Computing shortest path from s to t

At each step k, keep track of

$$C_{zt}^k$$

which is the optimal cost of going from any  $z \in Z$  to t in at most k steps or minus.

$$C_{zt}^0 = 0$$
 if  $z = t$ , and  $+\infty$  if  $z \neq t$ .

Suppose by induction I know  $C_{zt}^k$ . Then

$$C_{zt}^{k+1} = \min_{y} \left\{ c_{zy} + C_{yt}^{k}, C_{zt}^{k} \right\}$$

For a solution to the problem to exist, it is necessary that there is no loop with negative cost. If that is the case, any optimal path is of length at most |Z|.

Therefore the algorithm above (which is called Bellman-Ford) terminates in at most |Z| steps.