# Summary of Continuous Distributions for a random Variable X

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- 1 Uniform  $X \sim U(a, b)$
- 1.1 Probability Density Function f(x)

$$f(x) = \begin{cases} \frac{1}{b-a} & a \le x \le b\\ 0 & \text{Otherwise} \end{cases}$$

1.2 Cumulative Distribution Function F(x)

$$F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \le x \le b \\ 1 & x > b \end{cases}$$

Proof.

$$F(X) = \int_{a}^{b} f(u)du$$

$$= \left[\frac{u}{(b-a)}\right]_{a}^{x}$$

$$= \frac{x}{b-a} - \frac{a}{b-a}$$

$$= \frac{x-a}{b-a}$$

QED

1.3 Moment Generating Function  $M(t) = E(e^{tX})$ 

$$M(t) = \frac{e^{tb} - e^{ta}}{t(b-a)}$$

Proof.

$$M(t) = E(e^{tX})$$

$$= \int_{a}^{b} e^{tx} f(x) dx$$

$$= \int_{a}^{b} \frac{e^{tx}}{(b-a)}$$

$$= \left[\frac{e^{tx}}{t(b-a)}\right]_{a}^{b}$$

$$= \frac{e^{tb} - e^{ta}}{t(b-a)}$$

QED

1.4 E(X)

$$E(X) = \frac{a+b}{2}$$

*Proof.* Symmetry Argument makes this proof trivial

QED

Proof. Integration

$$E(X) = \int_{a}^{b} x f(x) dx$$

$$= \int_{a}^{b} \frac{x}{b-a} dx$$

$$= \left[ \frac{x^{2}}{2(b-a)} \right]_{a}^{b}$$

$$= \frac{1}{2(b-a)} [b^{2} - a^{2}]$$

$$= \frac{1}{2(b-a)} [(b-a)(b+a)]$$

$$= \frac{a+b}{2}$$

QED

1.5 Var(X)

$$Var(X) = \frac{(b-a)^2}{12}$$

*Proof.* Integration

$$E(X^{2}) = \int_{a}^{b} x^{2} f(x) dx$$

$$= \int_{a}^{b} \frac{x^{2}}{b - a} dx$$

$$= \left[ \frac{x^{3}}{3(b - a)} \right]_{a}^{b}$$

$$= \frac{1}{3(b - a)} [b^{3} - a^{3}]$$

$$= \frac{1}{3(b - a)} [(b - a)(b^{2} + ab + a^{2})]$$

$$= \frac{a^{2} + ab + b^{2}}{3}$$

Then we minus the two to obtain

$$Var(X) = \frac{a^2 + ab + b^2}{3} - \frac{(a+b)^2}{4}$$

$$= \frac{a^2 + ab + b^2}{3} - \frac{a^2 + 2ab + b^2}{4}$$

$$= \frac{4a^2 + 4ab + 4b^2}{12} - \frac{3a^2 + 6ab + 3b^2}{12}$$

$$= \frac{a^2 - 2ab + b^2}{12}$$

$$= \frac{(b-a)^2}{12}$$

- 2 Exponential  $X \sim Exp(\lambda)$
- **2.1** Probability Density Function f(x)

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0\\ 0 & \text{Otherwise} \end{cases}$$

**2.2** Cumulative Distribution Function F(x)

$$F(x) = 1 - e^{-\lambda x}$$
 where  $x > 0$ 

Proof.

$$F(x) = \int_0^x \lambda e^{-\lambda u} du$$
$$= \left[ -e^{-\lambda u} \right]_0^x$$
$$= (-e^{-\lambda x}) - (-1)$$
$$= 1 - e^{-\lambda x}$$

QED

2.3 Moment Generating Function  $M(t) = E(e^{tX})$ 

$$M(t) = \frac{\lambda}{\lambda - t}$$

Proof.

$$M(t) = E(e^{t})$$

$$= \int_{\mathbb{R}} e^{tx} f(x) dx$$

$$= \int_{0}^{\infty} e^{tx} \cdot \lambda e^{-\lambda x} dx$$

$$= \int_{0}^{\infty} \lambda e^{(t-\lambda)x} dx$$

$$= \left[ \frac{\lambda}{(t-\lambda)} e^{(t-\lambda)x} \right]_{0}^{\infty}$$

$$= \frac{\lambda}{\lambda - t}$$

## 2.4 E(X)

$$E(X) = \frac{1}{\lambda}$$

Proof.

$$E(X) = \int_0^\infty x f(x) dx$$

$$= \int_0^\infty x \lambda e^{-\lambda x} dx$$

$$= [-xe^{-\lambda x}]_0^\infty - \int_0^\infty -e^{-\lambda x} dx$$

$$= \left[ -\frac{1}{\lambda} e^{-\lambda x} \right]_0^\infty$$

$$= 0 - \left( -\frac{1}{\lambda} \right)$$

$$= \frac{1}{\lambda}$$

QED

## $2.5 \quad Var(X)$

$$Var(X) = \frac{1}{\lambda^2}$$

Proof.

$$E(X^{2}) = \int_{0}^{\infty} x^{2} f(x) dx$$

$$= \int_{0}^{\infty} x^{2} \lambda e^{-\lambda x} dx$$

$$= \left[ -x^{2} e^{-\lambda x} \right]_{0}^{\infty} - \int_{0}^{\infty} -2x e^{-\lambda x} dx$$

$$= 2 \int_{0}^{\infty} x e^{-\lambda x} dx$$

$$= 2 \left[ -\frac{1}{\lambda} x e^{-\lambda x} \right]_{0}^{\infty} - 2 \int_{0}^{\infty} -\frac{1}{\lambda} e^{-x} dx$$

$$= 2 \int_{0}^{\infty} \frac{1}{\lambda} e^{-\lambda x} dx$$

$$= 2 \left[ -\frac{1}{\lambda^{2}} e^{-\lambda x} \right]_{0}^{\infty}$$

$$= \frac{2}{\lambda^{2}}$$

$$Var(X) = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2$$
$$= \frac{1}{\lambda^2}$$

## **3** Gamma $X \sim \Gamma(\alpha, \lambda)$

#### **3.1** Probability Density Function f(x)

$$f(x) = \begin{cases} \frac{\lambda^{\alpha} x^{\alpha - 1} e^{-\lambda x}}{\Gamma(\alpha)} & x \ge 0\\ 0 & \text{Otherwise} \end{cases}$$

Where we have

$$\Gamma(\alpha) = \int_0^\infty x^{\alpha - 1} e^{-x} dx$$
$$= (\alpha - 1)!$$

## **3.2** Moment Generating Function $M(t) = E(e^{tX})$

$$M(t) = \frac{\lambda^{\alpha}}{(\lambda - t)^{\alpha}}$$

Proof.

$$\begin{split} M(t) &= E(e^{tX}) \\ &= \int_0^\infty e^{tx} f(x) dx \\ &= \int_0^\infty \frac{1}{\Gamma(\alpha)} x^{\alpha - 1} \lambda^\alpha e^{-(\lambda - t)x} dx \\ &= \frac{\lambda^\alpha}{(\lambda - t)^\alpha} \int_0^\infty \frac{1}{\Gamma(\alpha)} x^{\alpha - 1} (\lambda - t)^\alpha e^{-(\lambda - t)x} dx \end{split}$$

However, provided that  $t < \lambda$  so  $t - \lambda > 0$  the integrand here is the density of  $\Gamma(\alpha, \lambda - t)$  hence would equal 1. Hence can conclude that

$$M(t) = \frac{\lambda^{\alpha}}{(\lambda - t)^{\alpha}}$$

QED

#### 3.3 E(X)

$$E(X) = \frac{\alpha}{\lambda}$$

Proof. From MGF

$$M(t) = \lambda^{\alpha} (\lambda - t)^{\alpha}$$

$$M'(t) = \lambda^{\alpha} \alpha (\lambda - t)^{-(\alpha+1)}$$

$$M'(0) = \frac{\lambda^{\alpha} \alpha}{\lambda^{\alpha+1}}$$

$$= \frac{\alpha}{\lambda}$$

QED

#### $3.4 \quad Var(X)$

$$Var(X) = \frac{\alpha}{\lambda^2}$$

Proof.

$$M'(t) = \lambda^{\alpha} \alpha (\lambda - t)^{-(\alpha+1)}$$

$$M''(t) = \lambda^{\alpha} \alpha (\alpha + 1)(\lambda - t)^{-(\alpha+2)}$$

$$M''(0) = \frac{\lambda^{\alpha} \alpha (\alpha + 1)}{\lambda^{\alpha+2}}$$

$$= \frac{\alpha(\alpha + 1)}{\lambda^{2}}$$

Then we use  $Var(X) = E(X^2) - [E(X)]^2$  to get the result

$$Var(x) = \frac{\alpha(\alpha+1)}{\lambda^2} - \frac{\alpha^2}{\lambda^2}$$
$$= \frac{\alpha}{\lambda^2}$$

- 4 Beta  $X \sim B(\alpha, \beta)$
- 4.1 Probability Density Function f(x)

$$f(x) = \begin{cases} \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha,\beta)} & x \in (0,1) \\ 0 & \text{Otherwise} \end{cases}$$

Where we have

$$B(\alpha, \beta) = \int_0^1 x^{\alpha - 1} (1 - x)^{\beta - 1} dx$$

4.2 E(X)

$$E(X) = \frac{\alpha}{\alpha + \beta}$$

 $4.3 \quad Var(X)$ 

$$Var(X) = \frac{\alpha\beta}{(\alpha\beta)^2(\alpha+\beta+1)}$$

- 5 Normal  $X \sim N(\mu, \sigma^2)$
- **5.1** Probability Density Function f(x)

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right]$$

Where we have  $x \in \mathbb{R}$ 

**5.2** Moment Generating Function M(t)

$$M(t) = exp\left[\mu t - \frac{\sigma^2 t^2}{2}\right]$$

5.3 E(X)

$$E(X) = \mu$$

5.4 Var(X)

$$Var(X) = \sigma^2$$

Also for normal distributions we can have that X can become the standard normal distribution  $Z \sim N(0,1)$  by the calculation

$$\frac{X-\mu}{\sigma^2}$$