

# AMATH 561: Homework 4

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1. Not required.

2. (a) In a random walk,  $X$  becomes  $X+1$  with probability  $p$  and  $X-1$  with probability  $1-p$ .

In this process,  $X$  becomes  $X+2$  with probability  $p^2$ ,  $X-2$  with probability  $(1-p)^2$  and stays at  $X$  with probability  $2p(1-p)$ .

The transition matrix will look like

$$\begin{bmatrix} \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots \\ \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots \\ \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots \\ 0 & (1-p)^2 & 0 & 2p(1-p) & 0 & p^2 & 0 & 0 \\ 0 & 0 & (1-p)^2 & 0 & 2p(1-p) & 0 & p^2 & 0 \\ \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots \end{bmatrix}$$

(b) If the initial generation is 0, the future generations are 0 as well.

If the generating function for a single individual is  $G(s)$ , the generating function for an individual after two generations is  $G(G(s))$  or  $G_2(s)$

Thus, we can obtain the transition probabilities by differentiating it.

$$P(i, j) = \frac{1}{j!} \frac{d^j}{ds^j} (G_2(s))^i \Big|_{s=0}$$

The transition matrix will look like

$$\begin{bmatrix} 1 & 0 & 0 & \dots \\ p(1,0) & p(1,1) & p(1,2) & \dots \\ p(2,0) & p(2,1) & p(2,2) & \dots \\ \vdots & \vdots & \vdots & \dots \end{bmatrix}$$

3. not required

4.

$$P(\tau_j < \tau_i | X_0 = i) = P(\tau_i < \tau_j | X_0 = j) = p$$

$$\text{Probability that } j \text{ is visited once} = P(X_0 = i)P(\tau_j < \tau_i | X_0 = i)P(\tau_i < \tau_j | X_0 = j) = p^2 P(X_0 = i)$$

Probability that j is visited once given you start at i =  $p^2$

Probability that j is visited n times given you start at i =  
 $= p^2(\text{Probability that it transitions between j and j n-1 times}) = p^2 P(\tau_j < \tau_i | X_0 = j)^{n-1} = (1-p)^{n-1} p^2$

$$\text{Expected number of visits} = \sum_{n=1}^{\infty} n(1-p)^{n-1} p^2 = \sum_{n=0}^{\infty} (n+1)(1-p)^n p^2$$

This is an arithmetic geometric progression. So, the infinite sum is given by

$$\text{Expected number of visits} = p^2 \left( \frac{1}{1-(1-p)} + \frac{1-p}{(1-(1-p))^2} \right) = p^2 \left( \frac{1}{p} + \frac{1-p}{p^2} \right) = 1$$

5. not required

6. Written at the end

7. not required

8. Since  $\sum_i P(i, j) = 1$ , we can replace all the  $\mu$ s to get the following transition matrix.

$$\begin{bmatrix} 1-\lambda_0 & \lambda_0 & 0 & 0 & \dots \\ 1-\lambda_1 & 0 & \lambda_1 & 0 & \dots \\ 0 & 1-\lambda_2 & 0 & \lambda_2 & \dots \\ \vdots & \vdots & \vdots & \vdots & \lambda_{n-1} \\ 0 & 0 & 0 & 1-\lambda_n & \lambda_n \end{bmatrix}$$

For equilibrium, we need

$$\pi P = P$$

Let  $\pi$  be

$$[x_0, x_1, x_2, x_3, \dots]$$

$$(1-\lambda_0)x_0 + (1-\lambda_1)x_1 = x_0$$

$$x_1 = \frac{\lambda_0}{(1-\lambda_1)} x_0$$

$$(\lambda_0)x_0 + (1-\lambda_2)x_2 = x_1$$

$$(1-\lambda_1)x_1 + (1-\lambda_2)x_2 = x_1$$

$$(1-\lambda_2)x_2 = \lambda_1 x_1$$

$$x_2 = \frac{\lambda_1}{(1-\lambda_2)} x_1$$

Similarly, for j upto n-1,

$$x_j = \frac{\lambda_{j-1}}{(1-\lambda_j)} x_{j-1}$$

For n,

$$\lambda_{n-1} x_{n-1} = \lambda_n x_n$$

$$x_n = \frac{\lambda_{n-1}}{1 - \lambda_n} x_{n-1}$$

For reversibility, we need

$$\pi(i)p(i, j) = \pi(j)p(j, i)$$

Since this MC has finite space and the  $\pi$ s should sum to 1, we can safely say all the  $\pi$ s are finite. For diagonal elements,

$$\pi(i)p(i, i) = \pi(i)p(i, i)$$

We only need to worry about off-diagonal elements of the form  $j=i+1$ . (because all other probabilities are zero). For  $i=0$ ,

$$\pi(0)p(0, 1) = x_0\lambda_0 = \frac{\lambda_0}{(1 - \lambda_1)} x_0(1 - \lambda_1) = \pi(1)p(1, 0)$$

For  $i$  upto  $n-2$ ,

$$\pi(i)p(i, i+1) = x_i\lambda_i = \frac{\lambda_i}{(1 - \lambda_{i+1})} x_0(1 - \lambda_{i+1}) = \pi(i+1)p(i+1, i)$$

For  $i = n-1$ ,

$$\pi(n-1)p(n-1, n) = x_{n-1}\lambda_{n-1} = \frac{\lambda_{n-1}}{(1 - \lambda_n)} x_0(1 - \lambda_n) = \pi(n)p(n, n-1)$$

Thus, the system is reversible in equilibrium.