

## CDS: Machine Learning

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3.12

Options

$w_1, w_2$

$w_1, B_2$

$B_1, w_2$

~~$B_1 B_2$~~   $\leftarrow$  impossible

$$P(w_2 | w_1) = \frac{P(w_1, w_2)}{P(w_1)}$$

$$= \frac{1/3}{2/3}$$

$$= \frac{1}{2}$$

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Extra opgave 3.1

a)  $N = 2$  options  $S = \{hh, ht, th, tt\}$

$H_0: p_h = 0.5$        $H_1: p_h \neq 0.5$

assume equal probabilities

$$P(H_0) = 0.5 \quad P(H_1) = 0.5$$

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$$N_H = 0$$

$$S = tt$$

$$P(S=tt | H_0) = 0.5^2 = 0.25$$

$$P(S=tt | H_1) = 1^2 = 1$$

$$\begin{aligned} P(S=tt) &= P(S=tt | H_0)P(H_0) + P(S=tt | H_1)P(H_1) \\ &= 0.25 \cdot 0.5 + 1 \cdot 0.5 \\ &= 0.625 \end{aligned}$$

$$P(C|H_0 | s=tt) = \frac{P(s=tt | H_0) P(H_0)}{P(s=tt)}$$

$$= \frac{0,25 \cdot 0,5}{0,625}$$

$$= 0,2$$

$$P(C|H_1 | s=tt) = \frac{P(s=tt | H_1) P(H_1)}{P(s=tt)}$$

$$= \frac{1 \cdot 0,5}{0,625}$$

$$= 0,8$$

$$N_H = 2$$

~~s = hh~~ equal to  $N_H = 0$

$$P(C|H_0 | s=hh) = 0,2$$

$$P(C|H_1 | s=hh) = 0,8$$

$$N_H = 1$$

$$s = \{ht, th\}$$

$$P(s=\{ht, th\} | H_0) = 0,5^2 + 0,5^2$$

$$= 0,5$$

$$P(s=\{ht, th\} | H_1) < 0,5 \text{ for any } H_1: p_h \neq s$$

$$P(CS = \{\text{ht, th}\}) = P(CS = \{\text{ht, th}\} | H_0)P(H_0) + P(CS = \{\text{ht, th}\} | H_1)P(H_1)$$

$$= 0,5 \cdot 0,5$$

$$+ P(CS = \{\text{ht, th}\} | H_1) \cdot 0,5$$

$$< 0,25 + 0,25$$

$$< 0,5$$

$$P(CH_0 | CS = \{\text{ht, th}\}) = \frac{P(CS = \{\text{ht, th}\} | H_0)P(H_0)}{P(CS = \{\text{ht, th}\})}$$

$$> \frac{0,5 \cdot 0,5}{0,5}$$

$$> \frac{0,25}{0,5}$$

$$> 0,5$$

$$P(CH_1 | CS = \{\text{ht, th}\}) = \frac{P(CS = \{\text{ht, th}\} | H_1)P(H_1)}{P(CS = \{\text{ht, th}\})}$$

$$= \frac{P(CS = \{\text{ht, th}\} | H_1)P(H_1)}{(P(CS = \{\text{ht, th}\} | H_0)P(H_0) + P(CS = \{\text{ht, th}\} | H_1)P(H_1))}$$

$$(P(CS = \{\text{ht, th}\} | H_0)P(H_0) +$$

$$P(CS = \{\text{ht, th}\} | H_1)P(H_1))$$

$$\begin{aligned}
 P(H_1 | s=\{\text{ht, th}\}) &= \frac{P(s=\{\text{ht, th}\} | H_1) \cdot 0,5}{(0,5 \cdot 0,5 + P(s=\{\text{ht, th}\} | H_2) \cdot 0,5)} \\
 &= \frac{P(s=\{\text{ht, th}\} | H_1)}{0,5 + P(s=\{\text{ht, th}\} | H_2)} \\
 &< \frac{0,5}{0,5+0,5} \\
 &< 0,5
 \end{aligned}$$

b An unfair coin will prefer one result (side) over the other. Heads-Heads or Tails-Tails will more likely be unfair. A fair coin will prefer a split result, such as Heads-Tails or reverse. In the end, it is important to realise that  $N=2$ , and the coin should be tossed more to draw meaningful conclusions.

28.1

(mackay)

we have to calculate the evidence, which is denoted as:  $P(D|H)$

For  $H = H_0$ :

$$P(D|H_0) = \prod_{i=1}^5 p(x_i|H_0) \text{ since all } x_i's \text{ are i.i.d.}$$

$$= \prod_{i=1}^5 \frac{1}{2} = \left(\frac{1}{2}\right)^5 = \frac{1}{32} = 0.031$$

possible range  
1/16 of  $\pi$ .  
1/16 possible

For  $H = H_1$ :

$$P(D|H_1) = \int_{-1}^1 dm p(D|H_1, m) p(m|H_1) = \frac{1}{2}$$

$$= \int_{-1}^1 dm \cdot \frac{5}{\pi} \frac{1}{2} (1 + mx_i) - \frac{1}{2}$$

$$= \frac{1}{64\pi} \int_{-1}^1 dm (1 + m0.3)(1 + 0.5m)(1 + 0.7m)(1 + 0.8m)(1 + 0.9m)$$

"This integral will be done numerically"

↳ we used the file "28.1 numerical integration.py"

$$= 4.929/64 = 0.077\dots$$

## 1.3 Chapter 27

1(a)

$$p(x|r) \propto p(r|x) \cdot p(x)$$

$$p(r|x) \propto \exp(-\lambda) \frac{x^r}{r!} \cdot \frac{1}{x^{r-1}}$$

$$\propto \exp(-\lambda) \frac{x^r}{r!}$$

$$Q = -\ln p(\lambda|r)$$

$$= -(\ln(e^{-\lambda}) + \ln(\lambda^r) - \ln(r!))$$

$$\propto \lambda - (r-1)\log(\lambda)$$

$$\frac{\partial Q}{\partial \lambda} = 1 - \frac{(r-1)}{\lambda}$$

$$0 = 1 - \frac{r-1}{\lambda} = r-1 \Rightarrow \lambda_{\text{map}}$$

$$A = \left. \frac{\partial^2 Q}{\partial \lambda^2} \right|_{\lambda=\lambda_{\text{map}}} = \frac{r-1}{(r-1)^2} = \frac{1}{r-1}$$

$$p_r(\lambda) = \sqrt{\frac{A}{2\pi}} \exp \left\{ -\frac{A}{2} (\lambda - \lambda_{\text{map}})^2 \right\}$$

$$= \frac{1}{\sqrt{2\pi(r-1)}} \exp \left( -\frac{1}{2(r-1)} (\lambda - (r-1))^2 \right)$$

$$= N(\lambda; r-1, r-1)$$

Reckendat

# 1.3 Chapter 27

1. b)

$$P(y) = \frac{dy}{d\lambda}^{-1} \cdot p(\lambda)$$

$$p(y) = \frac{1}{\lambda}^{-1} \cdot \frac{1}{\lambda} = \lambda \cdot \frac{1}{\lambda} = 1$$

c)  $p(\log \lambda | r) \propto p(r | \lambda) \cdot 1$   
 $\propto \exp(-\lambda) \frac{\lambda^r}{r!}$

$$\begin{aligned} Q &= -\ln p(\log \lambda | r) \\ &= -(\ln(e^{-\lambda}) + \ln(\lambda^r) - \ln(r!)) \end{aligned}$$

$$\propto \lambda - r \log(\lambda) = e^y - ry$$

$$\frac{\partial Q}{\partial y} = e^y - r$$

$$e^y - r = 0$$

$$y_{\text{map}} = \log(r)$$

$$A = \frac{\partial^2 Q}{\partial y^2} \Big|_{y=y_{\text{map}}} = e^y \Big|_{y=y_{\text{map}}} = e^{\log(r)} = r$$

$$p_2(y) = \sqrt{\frac{A}{2\pi}} \exp \left\{ -\frac{A}{2} (y - y_{\text{map}})^2 \right\}$$

$$= \frac{1}{\sqrt{2\pi \cdot 1/r}} \exp \left\{ -\frac{1}{2 \cdot 1/r} (y - y_{\text{map}})^2 \right\}$$

fill in  $y = \log(\lambda)$   
 $y_{\text{map}} = \log(r)$

$$N(\log \lambda; \log r; 1/r)$$

# CDS: Machine Learning

2017 28.1

$$\begin{aligned}
 a.) P(D|w_0, w_1) &= \prod_{i=1}^N P(t_i | w_0 + w_1 x_i, \sigma^2) \\
 &= \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left(-\frac{1}{2\sigma^2} (t_i - w_0 - w_1 x_i)^2\right) \\
 &= \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_i (t_i^2 + w_0^2 + w_1^2 x_i^2 - 2t_i w_0 - 2w_0 w_1 x_i - 2w_1 t_i x_i)\right) \\
 &= \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left(-\frac{1}{2\sigma^2} (\sum_i t_i^2 + w_0^2 N + w_1^2 \sum_i x_i^2 - 2w_0 \sum_i t_i - 2w_1 \sum_i t_i x_i)\right)
 \end{aligned}$$

Since we can use that  $\sum_i x_i = 0$

$$\begin{aligned}
 &= \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_i t_i^2\right) \exp\left(-\frac{1}{2\sigma^2} (w_0^2 N - 2w_0 \sum_i t_i)\right) \\
 &\quad \exp\left(-\frac{1}{2\sigma^2} (w_1^2 \sum_i x_i^2 - 2w_1 \sum_i t_i x_i)\right)
 \end{aligned}$$

$$\begin{aligned}
 b.) \frac{P(D|H_2)}{P(D|H_1)} &= \frac{\int \int dw_0 dw_1 P(D|w_0, w_1, H_2) P(w_0, w_1 | H_2)}{\int \int dw_0 dw_1 P(D|w_0, w_1, H_1) P(w_0, w_1 | H_1)}
 \end{aligned}$$

\*). we know that  $P(D|w_0, w_1, H_i)$  is the same for both and described by a.)

$$*) P(w_0, w_1 | H_2) = \mathcal{U}(w_0 | 0, 1) \mathcal{U}(w_1 | 0, 1)$$

$$*) P(w_0, w_1 | H_1) = \mathcal{U}(w_0 | 0, 1) \quad |_{w_1=0}$$

Since within the expression of a.) the functions factorize for  $w_0$  and  $w_1$  and the dependence is the same for both models, all terms independent of both  $w_0$  and  $w_1$  and the terms depended on  $w_0$  drop out.

Further all terms with  $w_1$  in a.) have  $w_1$  in the exponent  $\rightarrow$  for  $H_1: w_1 = 0 \rightarrow$  all denominator stays 1. in the nominator we are left with the following Z.O.L.

$$P(\text{CDM}) = \frac{1}{\sqrt{2\pi}} \int dw_i \exp \left( -\frac{w_i^2}{2\sigma^2} \sum x_i^2 + \frac{w_i}{\sigma^2} \sum x_i t_i - \frac{w_i^2}{2} \right)$$

where we filled in  $\mathcal{N}(w_i | 0, 1)$

$$= \frac{1}{\sqrt{2\pi}} \int dw_i \exp \left( -\frac{w_i^2 N}{2\sigma^2} \left( \frac{\langle x^2 \rangle}{\sigma^2} + \frac{1}{N} \right) + \frac{w_i N}{\sigma^2} \langle xt \rangle \right)$$

which is ~~the~~<sup>a</sup> standard integral with:

$$a = N \left( \frac{\langle x^2 \rangle}{\sigma^2} + \frac{1}{N} \right) \quad \text{and} \quad J = \frac{N}{\sigma^2} \langle xt \rangle$$

Thus we have:

$$\begin{aligned} &= \frac{1}{\sqrt{2\pi}} \frac{\sqrt{N}}{\left( N \left( \frac{\langle x^2 \rangle}{\sigma^2} + \frac{1}{N} \right) \right)^{1/2}} \exp \left( \frac{N^2 \langle xt \rangle^2}{2(\sigma^2)^2 (N \left( \frac{\langle x^2 \rangle}{\sigma^2} + \frac{1}{N} \right))} \right) \\ &= \frac{1}{\left( N \left( \frac{\langle x^2 \rangle}{\sigma^2} + \frac{1}{N} \right) \right)^{1/2}} \exp \left( \frac{N \langle xt \rangle^2}{2(\langle x^2 \rangle \sigma^2 + \frac{\sigma^2}{N})^2} \right) \end{aligned}$$

$$\text{c.) } \sigma^2 = 1 \rightarrow \left( \frac{1}{N \left( \langle x^2 \rangle + \frac{1}{N} \right)} \right)^{1/2} \exp \left( \frac{N \langle xt \rangle^2}{2(\langle x^2 \rangle + \frac{1}{N})} \right)$$

If  $\langle xt \rangle^2 = \frac{\log(N)}{N}$  we find

$$\frac{\sqrt{N}}{\left( \langle x^2 \rangle + \frac{1}{N} \right)^{1/2}} \exp \left( \left( 2 \langle x^2 \rangle + \frac{1}{N} \right)^{-1} \right)$$

~~if~~  $N \rightarrow \infty$ : if  $N$  becomes large we can neglect  $\frac{1}{N}$

$$\rightarrow \frac{\sqrt{N}}{\sqrt{2x^2}} \exp \left( \frac{1}{2\langle x^2 \rangle} \right)$$

thus  $N \rightarrow \infty$  gives  $> 1$  and thus

we prefer model H<sub>2</sub>.

if  $\langle xt \rangle^2 < \frac{\log(N)}{N}$  we will find a negative

dependence of  $N$  and thus  $c_1$ , which shows

a preference for H<sub>1</sub>.

2.8.2

a.)

$$P(D|H_0) = \prod_{i=1}^n P(x_i|H_0) \cdot CF$$

where  $CF$  denotes the combinatorial factor we need since we disregarded the order!

$$P(x_i|H_0) = \frac{1}{k}, \text{ since we assume a fair dice.}$$

$$\rightarrow P(D|H_0) = \left(\frac{1}{k}\right)^n \cdot CF = \frac{n!}{n_1! \cdots n_k! (k!)^n}$$

20.2

$$b.) P(D|H_1) = \int d\vec{p} \, P(D|\vec{p}, H_1) P(\vec{p}|H_1)$$

$$P(D|\vec{p}, H_1) = C_F \prod_{i=1}^k (p_i)^{n_i} \quad (\text{Similar to a.)})$$

$P(\vec{p}|H_1)$ : uniform dirichlet distribution

we use 23.30 from Mackay with  $\alpha_i = 1$

and  $\vec{\alpha} \in \mathbb{R}^k$ .

$$\rightarrow P(\vec{p}|H_1) = \frac{1}{Z(\vec{\alpha})} \prod_{i=1}^k p_i^{n_i-1} \int \left( \sum_{i=1}^k p_i - 1 \right)$$

$$Z(\vec{\alpha}) = (k-1)! \quad (\text{see also special case dirichlet})$$

distribution from wikipedia)

$$\rightarrow P(\vec{p}|H_1) = (k-1)! \int \left( \sum_{i=1}^k p_i - 1 \right)$$

$$\Rightarrow P(D|H_1) = C_F \int d\vec{p} \, (k-1)! \prod_{i=1}^k (p_i)^{n_i-1} \int \left( \sum_{i=1}^k p_i - 1 \right)$$

which is similar to the given within  
the question.

$$\rightarrow I = C_F (k-1)! I,$$

$$\text{where } I_i = \int_{\text{simplex}} d\vec{p} \, \prod_{i=1}^k (p_i)^{n_i-1} \int \left( \sum_{i=1}^k p_i - 1 \right)$$

Now we notice that this is just the normalization  
constant of a different dirichlet distribution, with

$$\alpha_i - 1 = n_i \rightarrow \alpha_i = n_i + 1$$

$$\rightarrow m_i = \frac{n_i+1}{\alpha}, \text{ and we need } \sum_i m_i = 1$$

$$\rightarrow \alpha = \sum_i (n_i + 1) = \sum_i m_i + k$$

$$\rightarrow m_i = \frac{n_i+1}{\sum n_i + k}$$

$$\text{we know } I_i = \prod_i \Gamma(\alpha_i) / \Gamma(\alpha)$$

$$= \prod_i \frac{\Gamma(n_i + 1)}{\Gamma(\sum n_i + k)}$$

$$\Rightarrow P(D|H_1) = C_F (k-1)! \prod_i \frac{\Gamma(n_i + 1)}{\Gamma(\sum n_i + k)}$$

$$\Rightarrow C_F = \frac{n!}{n_1! \dots n_k!} (k-1)! \frac{n_1! \dots n_k!}{(n+k-1)!} = \frac{n! (k-1)!}{(n+k-1)!}$$

23.2

c.)

Posterior probability given equal priors:

$$P(H_i | D) = \frac{P(D|H_i) P(H_i)}{P(D)} = \frac{P(D|H_i) P(H_i)}{P(D|H_1) P(H_1) + P(D|H_0) P(H_0)}$$

$$P(H_1) = P(H_0) = P(H)$$

$$\rightarrow P(H_i | D) = \frac{P(D|H_i)}{P(D|H_0) + P(D|H_1)}$$

for numerical calculations see:

"~~ModelExtra-Bayes.py~~" .

"23.2.py"