

Stochastic optimization. Randomized methods

Gasnikov Alexander (MIPT, IITP RAS)

gasnikov@yandex.ru

Yandex

November 11, 2016.

Main books:

Polyak B.T., Juditsky A.B. Acceleration of stochastic approximation by averaging // SIAM J. Control Optim. – 1992. – V. 30. – P. 838–855.

Sridharan K. Learning from an optimization viewpoint. PhD Thesis, 2011.

Juditsky A., Nemirovski A. First order methods for nonsmooth convex large-scale optimization, I, II. // Optimization for Machine Learning. // Eds. S. Sra, S. Nowozin, S. Wright. – MIT Press, 2012.

Shapiro A., Dentcheva D., Ruszczyński A. Lecture on stochastic programming. Modeling and theory. – MPS-SIAM series on Optimization, 2014.

Guiges V., Juditsky A., Nemirovski A. Non-asymptotic confidence bounds for the optimal value of a stochastic program // e-print, 2016 [arXiv:1601.07592](https://arxiv.org/abs/1601.07592)

Gasnikov A. Searching equilibriums in large transport networks. Doctoral Thesis. MIPT, 2016.

<https://arxiv.org/ftp/arxiv/papers/1607/1607.03142.pdf>

<https://www.youtube.com/user/PreMoLab> (see course of A. Gasnikov)

Structure of the talk

- Auxiliary facts
- Stochastic Mirror Descent
 - Rate of convergence
 - Lower bounds
- Nesterov's problem about Mage and Experts (Parallelization)
 - Conditional Stochastic optimization
 - SAA *vs* SA
- Acceleration of Stochastic Approximation by proper Averaging
 - Randomized MD for huge QP
 - Randomized MD for Antagonistic matrix game

Auxiliary facts

Azuma–Hoeffding’s inequality: Let $\{\chi_t\}_t$ – a scalar random sequence is martingale-difference

$$\chi_t = Y_t - Y_{t-1}, \quad E\left[Y_t \mid F_{\sigma\text{-algebra}}(Y_1, \dots, Y_{t-1})\right] = Y_{t-1},$$

such that

$$E\left[\exp(\chi_t^2 / M^2) \mid \chi_1, \dots, \chi_{t-1}\right] \leq \exp(1) \text{ for all } t = 1, 2, \dots, N.$$

Then ($s > 0$)

$$P\left(\sum_{t=1}^N \gamma_t \chi_t \geq sM \sqrt{\sum_{t=1}^N \gamma_t^2}\right) \leq \exp(-s^2/3),$$

$$P\left(\sum_{t=1}^N \gamma_t \chi_t^2 \geq M^2 \sum_{t=1}^N \gamma_t + M^2 \max\left\{\sqrt{6.6s \sum_{t=1}^N \gamma_t^2}, 6.6s \frac{1}{N} \sum_{t=1}^N \gamma_t\right\}\right) \leq \exp(-s).$$

Heavy-tails, large deviations: Let scalar random sequence $\{\chi_t\}_t$ – i.i.d., $E[\chi_t] = 0$, $\text{Var}[\chi_t] \leq D$, $P(\chi_t > s) = V(s) = O(s^{-\alpha})$, $\alpha > 2$.

Then
$$P\left(\sum_{t=1}^N \chi_t \geq s\right)_{N \gg 1} \simeq 1 - \Phi\left(\frac{s}{\sqrt{DN}}\right) + N \cdot V(s), \quad \Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy,$$

$$P\left(\sum_{t=1}^N \chi_t \geq s\right)_{N \gg 1} \simeq 1 - \Phi\left(\frac{s}{\sqrt{DN}}\right), \quad s \leq \sqrt{(\alpha - 2)DN \ln N}, \quad (\text{CLT regime})$$

$$P\left(\sum_{t=1}^N \chi_t \geq s\right)_{N \gg 1} \simeq N \cdot V(s), \quad s > \sqrt{(\alpha - 2)DN \ln N}. \quad (\text{heavy-tails regime})$$

Note:
$$0.2e^{-2x^2/\pi} \leq 1 - \Phi(x) \leq e^{-x^2/2}, \quad x \gg 1.$$

These estimations can be generalized for the weighted sums of scalar martingale-differences and weighted sums of squares of martingale-differences.

Two coins comparison: Consider two coins: $p = 0.5$ and $p = 0.5 + \varepsilon$. How many observations $y = (y^1, y^2, \dots, y^N)$ we have to do to decide with probability $\geq 1 - \sigma$ what is a best coin? Let's introduce some decision rule $\varphi(y)$ that takes values $[0, 1]$ (we interpret $\varphi(y)$ as a probability to decide in favor of the second coin if we observe y). Then the probability of wrong decision is

$$\left| E[\varphi(y) | p = 0.5 + \varepsilon] - E[\varphi(y) | p = 0.5] \right| \geq 1 - 2\sigma.$$

Since for all measurable $0 \leq \varphi(y) \leq 1$ (Pinsker's inequality + chain rule)

$$\left| E_{P^N}[\varphi(y)] - E_{Q^N}[\varphi(y)] \right| \leq \|P^N - Q^N\|_1 \leq 2KL(P^N, Q^N) = 2N \cdot KL(P, Q),$$

$$KL(P, Q) = (0.5 + \varepsilon) \ln((0.5 + \varepsilon)/0.5) + (0.5 - \varepsilon) \ln((0.5 - \varepsilon)/0.5) \simeq 4\varepsilon^2,$$

we have that $N \geq C\varepsilon^{-2}$. One can show that indeed: $\boxed{N \geq C \ln(\sigma^{-1}) \varepsilon^{-2}}$. For that one should use Rao–Cramer's inequality for Bernoulli scheme (Lect. 2).

Stochastic Mirror Descent

Consider convex optimization problem

$$f(x) \rightarrow \min_{x \in Q},$$

with stochastic oracle, return such stochastic subgradient $\partial f(x, \xi)$ that:

$$E_{\xi} [\partial_x f(x, \xi)] = \partial f(x), \quad E_{\xi} [\|\partial_x f(x, \xi)\|_*^2] \leq M^2.$$

Method (the main tools for numerical stochastic programming!)

$$x^{k+1} = \text{Mirr}_{x^k} \left(h \partial_x f(x^k, \xi^k) \right), \quad \text{Mirr}_{x^k}(v) = \arg \min_{x \in Q} \left\{ \langle v, x - x^k \rangle + V(x, x^k) \right\}.$$

The main property of MD-step ($\{\xi^k\}$ – i.i.d.)

$$2V(x, x^{k+1}) \leq 2V(x, x^k) + 2h \langle \partial_x f(x^k, \xi^k), x - x^k \rangle + h^2 \|\partial_x f(x^k, \xi^k)\|_*^2.$$

$$f(x^k) - f(x) \leq \langle \partial f(x^k), x^k - x \rangle \leq \langle \partial f(x^k) - \partial_x f(x^k, \xi^k), x^k - x \rangle + \\ + \frac{1}{h} (V(x, x^k) - V(x, x^{k+1})) + \frac{h}{2} \left\| \partial_x f(x^k, \xi^k) \right\|_*^2 \Big| E[\cdot | \xi^1, \dots, \xi^{k-1}],$$

$$f(x^k) - f(x) \leq \langle \partial f(x^k), x^k - x \rangle \leq \\ \leq \frac{1}{h} (V(x, x^k) - E[V(x, x^{k+1}) | \xi^1, \dots, \xi^{k-1}]) + \frac{h}{2} E \left[\left\| \partial_x f(x^k, \xi^k) \right\|_*^2 \Big| \xi^1, \dots, \xi^{k-1} \right].$$

If we sum all these inequalities from $k = 0, \dots, N-1$ and take the total mathematical expectation from the both sides of the result with $x = x_*$, then due to the convexity of $f(x)$ we obtain (as in deterministic case)

$$E[f(\bar{x}^N)] - f_* \leq (hN)^{-1} V(x_*, x^0) + M^2 h/2 \leq \sqrt{2M^2 R^2 / N},$$

where

$$R^2 = V(x_*, x^0), \bar{x}^N = \frac{1}{N} \sum_{k=0}^{N-1} x^k, h = \frac{R}{M} \sqrt{\frac{2}{N}} = \frac{\varepsilon}{M^2}.$$

In other words, after $\boxed{N = 2M^2 R^2 / \varepsilon^2}$ oracle calls $\boxed{E[f(\bar{x}^N)] - f_* \leq \varepsilon}$.

Absolutely the same result (even constants) as it was in deterministic case!

From this result due to the Markov's inequality

$$P(f(\bar{x}^N) - f_* \geq 2\varepsilon) \leq \frac{E[f(\bar{x}^N)] - f_*}{2\varepsilon} \leq \frac{1}{2}.$$

So we can run in parallel $\sim \log_2(\sigma^{-1})$ MD-trajectories. Let's denote by \bar{x}_{\min}^N such \bar{x}^N from these trajectories that minimize $f(\bar{x}^N)$. Here we assume that we have an oracle for the value of function $f(x)$.

So after

$$N = \frac{8M^2 R^2}{\varepsilon^2} \log_2(\sigma^{-1})$$

oracle calls one can obtain

$$P\left(f\left(\bar{x}_{\min}^N\right) - f_* \geq 2\varepsilon\right) \leq \sigma.$$

But what we should do if there is no oracle for the value of the function?

Assume that $\|\partial_x f(x, \xi)\|_* \leq M$ a.s. for ξ , then

$$P\left(f\left(\bar{x}^N\right) - f_* \leq M \sqrt{\frac{2}{N}} \left(R + 2\tilde{R} \sqrt{\ln(2/\sigma)}\right)\right) \geq 1 - \sigma,$$

where $\tilde{R} = \sup_{x \in \tilde{Q}} \|x - x_*\|$, $\tilde{Q} = \left\{x \in Q : \|x - x_*\|^2 \leq 65R^2 \ln(4N/\sigma)\right\}$.

More generally, one can show (using Azuma–Hoeffding’s inequality) that

- if $\|\partial_x f(x, \xi)\|_* \leq M$, then

$$N \sim \frac{M^2 R^2 \ln(\sigma^{-1})}{\varepsilon^2};$$

- if $E\left(\exp\left(\|\partial_x f(x, \xi)\|_*^2 / M^2\right)\right) \leq \exp(1)$ and $\varepsilon \leq MR$ then

$$N \sim \frac{M^2 R^2 \ln(\sigma^{-1})}{\varepsilon^2}.$$

Using heavy-tails large deviations estimations one can obtain

- if $P\left(\left\|\partial_x f(x, \xi)\right\|_*^2 / M^2 \geq s\right) = O(s^{-\alpha})$, $\alpha > 2$ then

$$N \sim M^2 R^2 \max \left\{ \frac{\ln(\sigma^{-1})}{\varepsilon^2}, \left(\frac{1}{\sigma \varepsilon^\alpha} \right)^{\frac{2}{3\alpha-2}} \right\}.$$

All these bounds are optimal up to a multiplicative constants.

Using the restarts technique one can generalize all the results mentioned above to **μ -strongly convex functions** in norm $\|\cdot\|$. In all the estimations we leave non-euclidian prox-factor $\omega_n = O(\ln^\beta n)$ ($Q \subseteq \mathbb{R}^n$).

Juditsky A., Nesterov Yu. Deterministic and stochastic primal-dual subgradient algorithms for uniformly convex minimization // Stoch. System. – 2014. – V. 4. – no. 1. – P. 44–80.

- if $\|\partial_x f(x, \xi)\|_* \leq M$, then $N \sim \frac{M^2 \ln((\ln N)/\sigma)}{\mu \varepsilon}$;
- if $E\left(\frac{\|\partial_x f(x, \xi)\|_*^2}{M^2}\right) \leq \exp(1)$ and $\varepsilon \leq MR$ then $N \sim \frac{M^2 \ln((\ln N)/\sigma)}{\mu \varepsilon}$;
- if $P\left(\|\partial_x f(x, \xi)\|_*^2 / M^2 \geq s\right) = O(s^{-\alpha})$, $\alpha > 2$ then
$$N \sim \max \left\{ \frac{M^2 \ln((\ln N)/\sigma)}{\mu \varepsilon}, \left(\frac{M^2}{\mu \varepsilon} \right)^{\frac{\alpha}{3\alpha-2}} \left(\frac{\ln N}{\sigma} \right)^{\frac{2}{3\alpha-2}} \right\}.$$

All these bounds are optimal up to a $\ln N$ -factor of σ . We don't know at the moment is it possible to eliminate this factor and the ω_n -factor.

Is Markov's inequality always rough?

Consider sum-type convex optimization problem

$$f(x) = \frac{1}{m} \sum_{k=1}^m f_k(x) + h(x) \rightarrow \min_{x \in Q},$$

where $\|\nabla f_k(y) - \nabla f_k(x)\|_2 \leq L\|y - x\|_2$ and $h(x)$ is μ -strongly convex in $\|\cdot\|_2$. As we've seen later one can obtain $E[f(x^{N(\varepsilon)})] - f_* \leq \varepsilon$ after $N(\varepsilon) \sim \left(m + \min\{L/\mu, \sqrt{mL/\mu}\}\right) \ln(\Delta f / \varepsilon)$ iterations (calculations of $\nabla f_k(x)$ solely). Using rough Markov's inequality

$$P\left(f(x^{N(\varepsilon\sigma)}) - f_* \geq \varepsilon\sigma / \sigma\right) \leq \frac{E[f(x^{N(\varepsilon\sigma)})] - f_*}{\varepsilon\sigma / \sigma} \leq \sigma,$$

one can obtain unimprovable large deviations bound $\sim \ln(\sigma^{-1})$.

Simple lower bounds

Consider **non strongly convex case**

$$\varepsilon x \rightarrow \min_{x \in [-1,1]}.$$

Assume that the oracle return $\nabla f(x, \xi) = \varepsilon + \xi$, $\xi \in N(0,1)$. At each call ξ chooses independently. Assume we know in advance all the details except of ε sign – but we can observe $y^k = \varepsilon + \xi^k$. So we know in advanced that we should choose $x = \pm 1$. How many oracle's calls we need to determine with probability $\geq 1 - \sigma$ the right sign? Due to Neyman–Pirson's lemma the

best strategy is $\hat{x}_N = -\text{sign} \sum_{k=1}^N y^k$. $P(\hat{x}_N = 1 | \varepsilon > 0) = P\left(\sum_{k=1}^N y^k < 0\right) \simeq C e^{-\varepsilon^2 N}$,

when $\varepsilon > 0$, we have the following lower bound $\boxed{N \geq C \ln(\sigma^{-1}) / \varepsilon^2}$.

Consider **strongly convex case**. Probabilistic model:

$$y^k = x + \xi^k, \xi^k \in N(0,1) // \text{loglikelihood: } -(y-x)^2/2;$$

$$x_* = \arg \min_x (x - x_*)^2/2 = \arg \min_x E[(y-x)^2/2], y \in N(x_*,1). \quad (*)$$

One can consider (*) to be the stochastic programming problem with the oracle returns stochastic gradients $y^k - x$, $y^k \in N(x_*,1)$. Due to Rao–Cramer’s inequality we have $E\left[\left(\hat{x}_N(y^1, \dots, y^N) - x_*\right)^2\right] \geq N^{-1}$. Since normal distribution (with mathematical expectation as parameter) belongs to Exponential family, for MLE $\hat{x}_N = \arg \min_x \frac{1}{2} \sum_{k=1}^N (y^k - x)^2 = \frac{1}{N} \sum_{k=1}^N y^k$ we have equality in Rao–Cramer’s inequality. Since that we have a precise lower bound for that case $\boxed{N \simeq C \ln(\sigma^{-1})/\varepsilon}$. The other example – Bernoulli scheme (here one can also use lower bound for two coins comparison).

General lower bounds (A. Nemirovski)

Consider convex optimization problem

$$f(x) \rightarrow \min_{x \in B_p^n(R)}$$

with stochastic oracle, return such $\partial f(x, \xi)$ that:

$$E_\xi [\partial f(x, \xi)] = \partial f(x), \quad E_\xi [\|\partial f(x, \xi)\|_q^2] \leq M_p^2 \quad (1/p + 1/q = 1).$$

We'd like to obtain lower bound for the oracle calls N , that guarantee x^N

$$E[f(x^N)] - f_* \leq \varepsilon.$$

Nemirovski A. Efficient methods in convex programming. Technion, 1995.

http://www2.isye.gatech.edu/~nemirovs/Lec_EMCO.pdf

Lower bounds for the **Stochastic Oracle** are

- $N \geq c_p M_p^2 R^2 / \varepsilon^{\max(2,p)}$, under $N \ll n$, where $c_p = O(\ln n)$ (this estimation of c_p become precise when $p \rightarrow 1+0$);
- $N \geq c_p M_p^2 R^2 n^{1-2/\max(2,p)} / \varepsilon^2$, under $N \gg n$.

For the **Deterministic Oracle** (when oracle returns subgradient $\partial f(x)$ with the property $\|\partial f(x)\|_p \leq M_p$) we have lower bound

- $N \geq cn \ln(M_p R / \varepsilon)$, under $N \gg n$. // differs only in this regime

Agarwal A., Bartlett P.L., Ravikumar P., Wainwright M.J. Information-theoretic lower bounds on the oracle complexity of stochastic convex optimization // IEEE Trans. of Inform. – 2012. – V. 58. – № 5. – P. 3235–3249.

Nesterov's problem about Mage and Experts (Parallelization)

Assume that the optimal configuration determines by convex problem

$$f(x) \rightarrow \min_{x \in Q}.$$

But each day one can only observe independent stochastic subgradients

$$\partial_x f(x, \xi): E_\xi [\partial_x f(x, \xi)] = \partial f(x), \|\partial_x f(x, \xi)\|_* \leq M.$$

Mage can live $N \sim M^2 R^2 \ln(\sigma^{-1}) / \varepsilon^2$ iterations and Expert $N \sim M^2 R^2 / \varepsilon^2$.

What is better to ask a solution from Mage or from $K \sim \ln(\sigma^{-1})$ Experts?

Answer (A. Lagunovskaya, 2016): In both of the cases we obtain (up to constant factors) **the same** (ε, σ) -quality.

Indeed, as we've already known clever Mage (this Mage know MD algorithm) can give us (ε, σ) -solutions. That is return such a point that

$$P\left(f\left(\bar{x}^N\right)-f_* \leq \varepsilon\right) \geq 1-\sigma .$$

On the other hand clever Expert returns such $\bar{x}^{N,i}$ that $E\left[f\left(\bar{x}^{N,i}\right)\right]-f_* \leq \varepsilon$.

Therefore without loss of generality one can assume that (see above)

$$f\left(\bar{x}^{N,i}\right)-f_* \in N\left(\varepsilon, \varepsilon\right) .$$

Since we assume Experts to be independent and $f(x)$ is convex

$$f\left(\bar{x}^K\right)-f_* \leq \frac{1}{K} \sum_{i=1}^K\left(f\left(\bar{x}^{N,i}\right)-f_*\right) \in N\left(\varepsilon, \frac{\varepsilon}{K}\right), \quad \bar{x}^K=\frac{1}{K} \sum_{i=1}^K \bar{x}^{N,i}$$

Hence, $P\left(f\left(\bar{x}^K\right)-f_* \leq \varepsilon\right) \geq 1-\exp (-K) \simeq 1-\sigma$.

It'd be interesting to generalize this result for the other cases (see above).

Conditional Stochastic optimization

$$f(x) \rightarrow \min_{g(x) \leq 0; x \in Q},$$

where

$$E_{\xi} [\partial_x f(x, \xi)] = \partial f(x), \quad E_{\xi} [\partial_x g(x, \xi)] = \partial g(x),$$

$$E_{\xi} [\|\partial_x f(x, \xi)\|_*^2] \leq M_f^2, \quad E_{\xi} [\|\partial_x g(x, \xi)\|_*^2] \leq M_g^2.$$

Let's

$$h_g = \varepsilon_g / M_g^2, \quad h_f = \varepsilon_g / (M_f M_g),$$

$$\boxed{\begin{aligned} x^{k+1} &= \text{Mirr}_{x^k} \left(h_f \partial_x f(x^k, \xi^k) \right), \quad \text{if } g(x^k) \leq \varepsilon_g, \\ x^{k+1} &= \text{Mirr}_{x^k} \left(h_g \partial_x g(x^k, \xi^k) \right), \quad \text{if } g(x^k) > \varepsilon_g, \end{aligned}} \quad k = 1, \dots, N,$$

and the set I ($N_I = |I|$) of such indexes k , that $g(x^k) \leq \varepsilon_g$.

Then if $\boxed{N \geq 2M_g^2 R^2 / \varepsilon_g^2}$ then $E[N_I] \geq 1$ and

$$E\left[f\left(\bar{x}^N\right)\right] - f_* \leq \varepsilon_f = \frac{M_f}{M_g} \varepsilon_g, \quad g\left(\bar{x}^N\right) \leq \varepsilon_g, \quad \bar{x}^N = \frac{1}{N_I} \sum_{k \in I} x^k.$$

If additionally $\|\partial_x f(x, \xi)\|_* \leq M_f$, $\|\partial_x g(x, \xi)\|_* \leq M_g$, then for all

$$\boxed{N \geq \frac{9M_g^2 \tilde{R}^2}{\varepsilon_g^2} \ln(\sigma^{-1})}$$

up to a constant factor the same as it
was in unconditional case (see above)

with probability $\geq 1 - \sigma$ the following is true $N_I \geq 1$ and

$$f\left(\bar{x}^N\right) - f_* \leq \varepsilon_f, \quad g\left(\bar{x}^N\right) \leq \varepsilon_g,$$

where $\tilde{R} = \sup_{x \in \tilde{Q}} \|x - x_*\|$, $\tilde{Q} = \left\{x \in Q : \|x - x_*\|^2 \leq 65R^2 \ln(4N/\sigma)\right\}$.

A. Bayandina generalizes it to strongly convex case, using restarts technique.
Here we have still an open problem: to generalize on composite optimization.

SAA vs SA (Nemirovski–Juditsky–Lan–Shapiro, 2007)

Stochastic Average Approximation (Empirical Risk Minimization, Monte Carlo) approach proposes to change Stochastic convex optimization problem

$$E_{\xi} \left[f(x, \xi) \right] \rightarrow \min_{x \in Q}$$

by **non stochastic** sum-type **SAA-problem** ($\{\xi^k\}_{k=1}^m$ – i.i.d. realizations from ξ)

$$\frac{1}{m} \sum_{k=1}^m f(x, \xi^k) \rightarrow \min_{x \in Q}.$$

Unfortunately, for the absolutely accurate solution of SAA-problem to be (ε, σ) -solution of initial one, one should take at least $(\|\partial_x f(x, \xi)\|_* \leq M)$

$$m \geq C \cdot M^2 R^2 \left(n \ln(MR/\varepsilon) + \ln(\sigma^{-1}) \right) / \varepsilon^2 \text{ terms.}$$

Stochastic Approximation approach (Robbins–Monro, 1951) in our sense is nothing more than Mirror Descent. So we can find (ε, σ) -solution of initial stochastic programming problem for

$$N \sim M^2 R^2 \ln(\sigma^{-1}) / \varepsilon^2 \ll m // \text{SA is better SAA}$$

oracle calls (i.e. calculations of stochastic subgradients $\partial_x f(x, \xi)$). It seems too strange! But it should be mentioned that one can find (ε, σ) -solution of SAA-problem for

$$N \sim M^2 R^2 \ln(\sigma^{-1}) / \varepsilon^2$$

calculations of stochastic subgradients of the terms of the sum chose at random. Indeed, let's introduce

$$f(x,\eta)=\begin{cases} f(x,\xi^1), \text{ with probability } 1/m \\ \\ f(x,\xi^m), \text{ with probability } 1/m \end{cases}$$

Non stochastic sum-type SAA-problem can be considered as simple stochastic problem

$$E_{\eta} \left[f \left(x, \eta \right) \right] \rightarrow \min_{x \in Q},$$

with stochastic subgradient: $\partial_x f(x, \eta) = \partial f_\eta(x)$, $\eta \in R[1, \dots, m]$. One can generate η for $O(\log_2 m)$ arithmetic operations. Since $\|\partial_x f(x, \eta)\|_* \leq M$ one can easily obtain that $N \sim M^2 R^2 \ln(\sigma^{-1}) / \varepsilon^2$ QED. But sometimes SAA-approach isn't substantial at all instead of SA (K. Sridharan's example).

Acceleration of Stochastic Approximation by proper Averaging

Let $\mathbf{x}_k, k = 1, \dots, N$ – i.i.d. with density function $p_{\mathbf{x}}(\mathbf{x}|\theta)$ (supp. doesn't depend on θ), depends on unknown vector of parameters θ . Then for all statistics $\tilde{\theta}(\mathbf{x})$ ($E_{\mathbf{x}}[\tilde{\theta}(\mathbf{x})^2] < \infty$): $E_{\mathbf{x}}\left[\left(\tilde{\theta}(\mathbf{x}) - \theta\right)\left(\tilde{\theta}(\mathbf{x}) - \theta\right)^T\right] \succ [I_{p,N}]^{-1}$,

$$I_{p,N} = E_{\mathbf{x}}\left[\nabla_{\theta} \ln p_{\mathbf{x}}(\mathbf{x}|\theta)\left(\nabla_{\theta} \ln p_{\mathbf{x}}(\mathbf{x}|\theta)\right)^T\right] = NI_{p,1}.$$

In 1990 B. Polyak (see also Polyak–Juditsky, 1992) showed that for

$$\theta^{k+1} = \theta^k + \gamma_k \nabla_{\theta} \ln p_{\mathbf{x}}(\mathbf{x}_k|\theta^k), \quad \bar{\theta}^N = \frac{1}{N} \sum_{k=1}^N \theta^k, \quad \gamma_k = \gamma \cdot k^{-\beta}, \quad \beta \in (0,1),$$

$$\sqrt{N} \cdot (\bar{\theta}^N - \theta_*) \xrightarrow{d} N\left(0, [I_{p,1}]^{-1}\right), \quad E_{\mathbf{x}}\left[N \cdot (\bar{\theta}^N - \theta_*)(\bar{\theta}^N - \theta_*)^T\right] \rightarrow [I_{p,1}]^{-1}.$$

SAA approach leads to analogous result (Fisher's theorem).

Randomized MD for huge QP (Juditsky–Nemirovski randomization)

Let's consider QP problem ($n \times n$ matrix $A \succ 0$ is fully completed, $|A_{ij}| \leq M$)

$$\frac{1}{2} \langle x, Ax \rangle \rightarrow \min_{x \in S_n(1)}.$$

Using STM (see Lectures 3, 4), one can find ε -solution for

$O\left(n^2 \sqrt{M \ln n / \varepsilon}\right)$ arithmetic operations. // not good since $n \gg 1$ is huge

But if one use randomized MD with stochastic gradient $A^{\langle i[x] \rangle} - i[x]$ -column of matrix A and $P(i[x] = j) = x_j$, $j = 1, \dots, n$ (one can generate $i[x]$ for $O(n)$ arithmetic operations), than one can find (ε, σ) -solutions for

$O\left(n M^2 \ln n \cdot \ln(\sigma^{-1}) / \varepsilon^2\right)$ arithmetic operations.

Randomized MD for Antagonistic matrix game (Grigoriadis–Khachiyan)

Google problem can be reduced to the saddle-point problem (\tilde{A} is s -row and s -column sparse)

$$\min_{x \in S_n(1)} \max_{\omega \in S_{2n}(1)} \langle \omega, \tilde{A}x \rangle.$$

Assume that there are two players A and B. All the players know matrix $\tilde{A} = \|\tilde{a}_{ij}\|$, where $|\tilde{a}_{ij}| \leq 1$, \tilde{a}_{ij} – prize of A (loss of B) in case when A plays i and B plays j . We play for the player B. Assume that the game is repeated $N \gg 1$ times. Let's introduce loss-function at the step k

$$f_k(x) = \langle \omega^k, \tilde{A}x \rangle, \quad x \in S_n(1),$$

where $\omega^k \in S_{2n}(1)$ – such a vector with all zero components except one component, that component corresponds to the A's choice at the step k –

this components equals 1. This vector in principle could depends on all the history for that moment (but it can't depends on the realization of the randomized strategy of player B at the step k). Analogously, vector x^k has only one non zero component, corresponds to the choice of player B at the step k . One can introduce the price of the game ($C = 0$)

$$C = \max_{\omega \in S_{2n}(1)} \min_{x \in S_n(1)} \langle \omega, \tilde{A}x \rangle = \min_{x \in S_n(1)} \max_{\omega \in S_{2n}(1)} \langle \omega, \tilde{A}x \rangle. \text{ (von Neumann theorem)}$$

The solution of the saddle-point problem (ω, x) is Nash equilibrium. Since that (Hannan)

$$\min_{x \in S_n(1)} \frac{1}{N} \sum_{k=1}^N f_k(x) \leq C.$$

So if we (player B) will choose $\{x^k\}$ at random according to the following randomized MD-strategy (randomization under KL-projection!):

1. $p^1 = (n^{-1}, \dots, n^{-1})$;
2. Choose at random $j(k)$ such, that $P(j(k) = j) = p_j^k$;
3. Put $x_{j(k)}^k = 1$, $x_j^k = 0$, $j \neq j(k)$;
4. Recalculate

$$p_j^{k+1} \sim p_j^k \exp\left(-\sqrt{\frac{2 \ln n}{N}} \tilde{a}_{i(k)j}\right), \quad j = 1, \dots, n,$$

where $i(k)$ – the choice of A at the step k ;

then with probability $\geq 1 - \sigma$

$$\frac{1}{N} \sum_{k=1}^N f_k(x^k) - \min_{x \in S_n(1)} \frac{1}{N} \sum_{k=1}^N f_k(x) \leq \sqrt{\frac{2}{N}} \left(\sqrt{\ln n} + 2\sqrt{2 \ln(\sigma^{-1})} \right),$$

i.e. with probability $\geq 1 - \sigma$ our (B's player) loss can be bounded

$$\frac{1}{N} \sum_{k=1}^N f_k(x^k) \leq C + \sqrt{\frac{2}{N}} \left(\sqrt{\ln n} + 2\sqrt{2 \ln(\sigma^{-1})} \right).$$

The worst case – when A is also know this strategy and use it when choosing $\{\omega^k\}$ (it should be mentioned that A solve max-type problem). If A and B will use this strategy then they converges to Nash's equilibrium according to the following estimation.

With probability $\geq 1 - \sigma$

$$\begin{aligned}
0 \leq \|A\bar{x}^N\|_\infty &= \max_{\omega \in S_{2n}(1)} \langle \omega, \tilde{A}\bar{x}^N \rangle - \max_{\omega \in S_{2n}(1)} \min_{x \in S_n(1)} \langle \omega, \tilde{A}x \rangle \leq \\
&\leq \max_{\omega \in S_{2n}(1)} \langle \omega, \tilde{A}\bar{x}^N \rangle - \min_{x \in S_n(1)} \langle \bar{\omega}^N, \tilde{A}x \rangle \leq \\
&\leq \max_{\omega \in S_{2n}(1)} \langle \omega, \tilde{A}\bar{x}^N \rangle - \frac{1}{N} \sum_{k=1}^N \langle \omega^k, \tilde{A}x^k \rangle + \frac{1}{N} \sum_{k=1}^N \langle \omega^k, \tilde{A}x^k \rangle - \min_{x \in S_n(1)} \langle \bar{\omega}^N, \tilde{A}x \rangle \leq \\
&\leq \sqrt{\frac{2}{N}} \left(\sqrt{\ln(2n)} + 2\sqrt{2\ln(2/\sigma)} \right) + \sqrt{\frac{2}{N}} \left(\sqrt{\ln n} + 2\sqrt{2\ln(2/\sigma)} \right) \leq \\
&\leq 2\sqrt{\frac{2}{N}} \left(\sqrt{\ln(2n)} + 2\sqrt{2\ln(2/\sigma)} \right),
\end{aligned}$$

where

$$\bar{x}^N = \frac{1}{N} \sum_{k=1}^N x^k, \bar{\omega}^N = \frac{1}{N} \sum_{k=1}^N \omega^k.$$

So when

$$N = 16 \frac{\ln(2n) + 8 \ln(2/\sigma)}{\varepsilon^2},$$

then with probability $\geq 1 - \sigma$ one can guarantee $\|A\bar{x}^N\|_\infty \leq \varepsilon$. The total number of arithmetic operations can be estimated as follows

$$O\left(n + \frac{s \ln n \cdot \ln(n/\sigma)}{\varepsilon^2}\right).$$

To be continued...