

# Stationary Processes

A process X(t) is stationary when:

$$F_{X(t_1),...,X(t_k)}(x_1,...,x_k)$$

$$=F_{X(t_1+\tau),...,X(t_k+\tau)}(x_1,...,x_k)$$

$$\forall \tau, k, t_1,...,t_k$$

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# Stationary Processes

■ The "first-order cdf" of a stationary process is independent of time:

$$F_{X(t)}(x) = F_{X(t+\tau)}(x) \quad \forall t, \tau$$

$$F_{X(t)}(x) = F_X(x) \quad \forall t, \tau$$

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#### Stationary Processes

- All k<sup>th</sup> moments E[(X(t))<sup>k</sup>] of a stationary process are independent of time.
- In particular, the mean and variance are independent of time:

$$m_{x}(t) = E[X(t)] = m \quad \forall t$$

$$VAR[X(t)] = E[(X(t) - m)^{2}] = \sigma^{2} \qquad \forall t$$

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#### Wide Sense Stationary Proc.

- A process is Wide Sense Stationary (WSS) if:
  - (a) the mean is constant;
  - (b) the autocorrelation function depend only on the time difference  $(t_2-t_1)$ :

$$m_X(t) = E[X(t)] = m \quad \forall t$$

$$R_{X}(t_{1},t_{2}) = R_{X}(t_{2}-t_{1}) \qquad \forall t_{1},t_{2}$$

$$\Rightarrow C_{X}(t_{1},t_{2}) = C_{X}(t_{2}-t_{1}) \forall t_{1},t_{2}$$

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#### Autocorrelation of WSS Proc.

- $R_X(t_2-t_1) = R_X(\tau)$  where  $\tau = t_2-t_1$
- $\blacksquare$  The autocorrelation is a symmetric function of the time difference  $\tau$  :

$$R_{x}(\tau) = R_{x}(-\tau)$$

■ The autocorrelation function has a maximum magnitude at the origin  $\tau$  = 0:

$$|R_{\mathsf{X}}(\tau)| \leq R_{\mathsf{X}}(0)$$

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# Example: Sinusoidal Phase

■ Let  $\Theta$  be a uniform random variable over the interval  $(-\pi, \pi)$ , and let the random process  $X(t, \Theta)$ :

$$X(t) = \cos(t + \Theta)$$

Is X(t) a WSS process?

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# Example: Sinusoidal Phase

■ Remember that the mean:

$$m_{x}(t) = E[\cos(t + \Theta)]$$

$$m_{x}(t) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos(t + \theta) d\theta$$

$$m_{x}(t) = 0$$

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# Example: Sinusoidal Phase

And the autocorrelation function:

$$R_{x}(t_{1},t_{2}) = E[cos(t_{1} + \Theta)cos(t_{2} + \Theta)]$$

$$\boxed{\mathsf{R}_{\mathsf{x}}\left(\mathsf{t}_{\mathsf{1}},\mathsf{t}_{\mathsf{2}}\right) = \frac{1}{2}\cos\left(\mathsf{t}_{\mathsf{1}}-\mathsf{t}_{\mathsf{2}}\right) = \frac{1}{2}\cos\left(\mathsf{\tau}\right)}$$

Therefore, X(t) is a WSS process

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# Example: Sinusoidal Phase

- Note that  $R_X(\tau)$  is symmetric and has a maximum at  $\tau=0$
- In this example,  $R_X(\tau)$  is also periodic with a period of  $2\pi$
- This leads to the following general attribute:

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# Stationary & WSS Processes

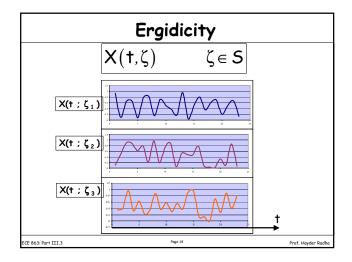
- A stationary process is a WSS process
- However, the inverse is not always true
- For a Gaussian process X(t):

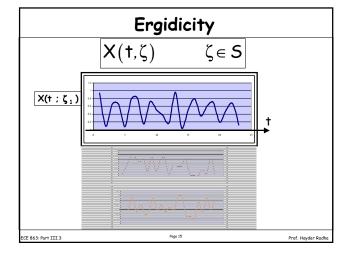
If X(t) is a Gaussian WSS process then X(t) is a stationary process

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- Remember that a random process  $X(t)=X(t,\zeta)$  is a collection of time functions, where each function  $X(t,\zeta_1)$  is generated from one random outcome  $\zeta_1$  of a random experiment
- Under certain conditions, we may be able to compute some statistics (e.g. the mean) of the random process X(t) by focusing on a <u>single</u> time function

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# Ergodicity

 For example, for a given outcome ζ of a random experiment, the <u>time average</u> over some time interval [-T, T]:

$$\langle X(t) \rangle_T = \frac{1}{2T} \int_{-T}^{T} X(t,\zeta) dt$$

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- As T gets large, and if the "time average"  $\langle X(t) \rangle_T$  converges to the true mean E[X(t)], then we say that X(t) is:
  - ergodic in the mean or "mean ergodic"

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#### **Ergodicity**

- It can be shown that, for WSS processes, the ergodic mean does converge to the mean E[X(t)] = m under a certain condition.
- This is formulated in a theorem
  - (The complete proof of the theorem is included as an Appendix.)

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# Ergodicity

Before stating the theorem, it is important to highlight three key expressions that result from the proof of the theorem:

$$\boxed{\mathsf{E}\!\left[\!\left\langle \mathsf{X}(\mathsf{t})\right\rangle_{\!\mathsf{T}}\right]\!=\mathsf{m}}$$

$$VAR\left[\left\langle X(t)\right\rangle_{T}\right] = E\left[\left(\left\langle X(t)\right\rangle_{T} - m\right)^{2}\right]$$

$$VAR\left[X\left\langle (t)\right\rangle_{T}\right] = \frac{1}{2T}\int_{-2T}^{2T} \left(1 - \frac{|\tau|}{2T}\right)C_{X}(\tau)d\tau$$

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#### "Mean Ergodic" Theorem

■ Let  $\langle X(t) \rangle_T$  be the time-average of a WSS process X(t) with mean E[X(t)] = m and covariance  $C_X(\tau)$ , then:

 $\langle X(t) \rangle_T$  converges to the mean E[X(t)] = m

if and only if VAR(  $\left\langle X(t)\right\rangle _{T}$  ) converges to zero

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# Example: Random Telegraph Process

- Let X(t) be a <u>random telegraph</u> process with parameter  $\alpha$ 
  - $\bullet$   $\alpha$  is the average number of occurrences per unit-of-time of a Poisson process

Evaluate the variance of the time-average of X(t). Does the time-average converges to the mean E[X(t)] in a mean-square-sense?

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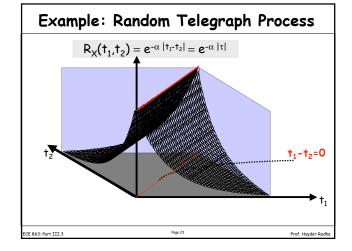
#### Example: Random Telegraph Process

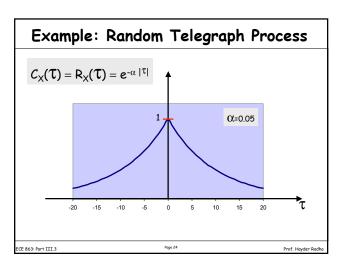
X(t) has a zero mean E[X(t)] = m = 0, and a covariance function:

$$C_{\mathsf{X}}(\tau) = \mathsf{R}_{\mathsf{X}}(\tau) = e^{-2\alpha|\tau|}$$

(See example 6.22 in the book)

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#### Example: Random Telegraph Process

Therefore, X(t) is a WSS process, and we can evaluate the variance of its timeaverage using the following:

$$VAR\Big[X\left\langle \left( t\right) \right\rangle _{T}\Big]=\frac{1}{2T}\int_{2T}^{2T}\!\!\left(1-\frac{\left| \tau \right|}{2T}\right)\!\!C_{X}\left( \tau \right)d\tau$$

$$VAR\left[X\left\langle \left(\dagger\right)\right\rangle _{T}\right]=\frac{1}{2T}\int_{-2T}^{2T}\left(1-\frac{\left|\tau\right|}{2T}\right)e^{-2\alpha\left|\tau\right|}d\tau$$

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#### Example: Random Telegraph Process

Since the integrand is symmetric then:

$$VAR\Big[X\left\langle \left(\uparrow\right)\right\rangle _{T}\Big]=\frac{1}{2T}\int_{-2T}^{2T}\left(1-\frac{\left|\tau\right|}{2T}\right)\!e^{-2\alpha\left|\tau\right|}d\tau$$

$$VAR\left[X\left\langle \left(\dagger\right)\right\rangle _{T}\right]=\frac{2}{2T}\int_{0}^{2T}\left(1-\frac{\tau}{2T}\right)e^{-2\alpha\tau}d\tau$$

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#### Example: Random Telegraph Process

■ Therefore,

$$\begin{split} VAR \Big[ X \big\langle (\tau) \big\rangle_T \, \Big] &= \frac{2}{2T} \int_0^{2T} \! \left( 1 - \frac{\tau}{2T} \right) e^{-2\alpha\tau} d\tau \\ &= \frac{1}{T} \Bigg\{ \int_0^{2T} e^{-2\alpha\tau} d\tau - \frac{1}{2T} \int_0^{2T} \tau \ e^{-2\alpha\tau} d\tau \Bigg\} \end{split}$$

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#### Example: Random Telegraph Process

■ This leads to the following:

$$VAR[X\langle(t)\rangle_{T}] = \frac{1 - e^{-4\alpha T}}{2\alpha T} - \frac{1}{8\alpha^{2}T^{2}} \left[1 - (1 + 4\alpha T)e^{-4\alpha T}\right]$$

$$VAR[X\langle(\dagger)\rangle_{T}] = \frac{4\alpha T + e^{-4\alpha T} - 1}{8\alpha^{2}T^{2}}$$

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#### Example: Random Telegraph Process

As T goes to infinity, the variance of the time-average goes to zero:

$$\overline{\lim_{T \to \infty} VAR \Big[ X \big\langle (t) \big\rangle_T \Big] = 0}$$

■ Therefore, for the "random telegraph process" the time-average converges to the mean (E[X(t)]=m=0) in a mean-square-sense:

$$\lim_{\mathsf{T}\to\infty} \ \mathsf{E}\bigg[ \Big( \big\langle \mathsf{X}(\mathsf{t}) \big\rangle_{\mathsf{T}} - \mathsf{m} \Big)^2 \bigg] = 0$$

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#### Example: Random Telegraph Process

Since the following condition is satisfied:

$$\lim_{\mathsf{T}\to\infty} \ \mathsf{E}\bigg[\big(\big\langle \mathsf{X}(\mathsf{t})\big\rangle_{\mathsf{T}} - \mathsf{m}\big)^2\bigg] = 0$$

then, we say that the random telegraph process X(t) is "mean ergodic" (or "ergodic in the mean").

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#### APPENDIX A

Proof of the

"Mean Ergodic" Theorem

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## "Mean Ergodic" Theorem

■ Let  $\langle X(t) \rangle_T$  be the time-average of a WSS process X(t) with mean E[X(t)]=m and covariance  $C_X(t)$ , then:

$$\lim_{\mathsf{T}\to\infty}\ \mathsf{E}\bigg[\Big(\big\langle \mathsf{X}(\mathsf{t})\big\rangle_{\mathsf{T}}-\mathsf{m}\Big)^2\,\bigg]=0$$

if and only if:

$$\boxed{\lim_{T\to\infty}\frac{1}{2T}\int_{-2T}^{2T}\!\left(1\!-\!\frac{\left|\tau\right|}{2T}\right)\!C_{X}\left(\tau\right)d\tau=0}$$

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- We need to show that as T gets large, then the time average <u>converges</u> to the mean E[X(t)]; i.e. we need to show that X(t) is: ergodic in the mean or "mean ergodic"
- A key point here is the type of convergence (section 5.5)
- Here, we focus on "mean-square convergence"

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#### **Ergodicity**

The time-average converges in the "mean-square" sense to some number (a) when the following is satisfied:

$$\lim_{T \to \infty} E \left[ \left( \left\langle X(t) \right\rangle_T - a \right)^2 \right] = 0$$

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# Ergodicity

- We are interested in the convergence of the timeaverage to the mean E[X(t)] (i.e. a=E[X(t)])
- Although the time-average  $\langle X(t) \rangle_T$  is a function of T,  $\langle X(t) \rangle_T$  is NOT a function of time t
- Consequently, we are interested in processes with E[X(t)] that is NOT a function of time: E[X(t)]=m

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## **Ergodicity**

- Therefore, we will consider Wide-Sense-Stationary (WSS) processes.
- We need to derive the condition that leads to the following "mean-square-convergence" of the time-average to the mean (m) of a WSS process:

$$\lim_{\mathsf{T}\to\infty} \ \mathsf{E}\bigg[ \Big( \big\langle \mathsf{X}(\mathsf{t}) \big\rangle_{\mathsf{T}} - \mathsf{m} \Big)^2 \bigg] = 0$$

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■ First, lets look at the expected value of the time average of a WSS process

$$E\left[\left\langle X(t)\right\rangle_{T}\right] = E\left[\frac{1}{2T}\int_{-T}^{T}X(t)dt\right]$$

$$E\left[\left\langle X(t)\right\rangle_{T}\right] = \frac{1}{2T}\int_{-T}^{T}E\left[X(t)\right]dt$$

$$E\left[\left\langle X(t)\right\rangle_{T}\right] = m$$

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#### **Ergodicity**

Since the expected value of the timeaverage is the mean (m), then

$$\begin{split} & E \bigg[ \Big( \big\langle X(t) \big\rangle_T - m \Big)^2 \bigg] = \\ & E \bigg[ \Big( \big\langle X(t) \big\rangle_T - E \Big[ \big\langle X(t) \big\rangle_T \Big] \Big)^2 \bigg] \\ & E \bigg[ \Big( \big\langle X(t) \big\rangle_T - m \Big)^2 \bigg] = VAR \Big[ \big\langle X(t) \big\rangle_T \bigg] \end{split}$$

# **Ergodicity**

Therefore, convergence of the time-average to the mean m (in a MS sense) is equivalent to the convergence of the "time-averagevariance" to zero:

$$\begin{bmatrix} \lim_{\mathsf{T}\to\infty} \ \mathsf{E}\bigg[\big(\big\langle \mathsf{X}(\mathsf{t})\big\rangle_{\mathsf{T}} - \mathsf{m}\big)^2 \hspace{0.1cm}\bigg] = 0 \\ \Leftrightarrow \\ \lim_{\mathsf{T}\to\infty} \ \mathsf{VAR}\bigg[\big\langle \mathsf{X}(\mathsf{t})\big\rangle_{\mathsf{T}} \hspace{0.1cm}\bigg] = 0 \\ \end{split}$$

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# Ergodicity

Now, let's look at the condition that leads to the desired convergence:

$$VAR\left[\left\langle X(t)\right\rangle_{T}\right] = E\left[\left(\left\langle X(t)\right\rangle_{T} - m\right)^{2}\right]$$

$$= E\left[\left\{\frac{1}{2T}\int_{-T}^{T}(X(t) - m)dt\right\}\left\{\frac{1}{2T}\int_{-T}^{T}(X(t') - m)dt'\right\}\right]$$

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$$VAR[\langle X(t) \rangle_{T}] = E\left[\left\{\frac{1}{2T}\int_{-T}^{T}(X(t) - m) dt\right\}\left\{\frac{1}{2T}\int_{-T}^{T}(X(t') - m) dt'\right\}\right]$$

$$= \frac{1}{4T^2} \int_{-T}^{T} \int_{-T}^{T} E[(X(t) - m)(X(t') - m)] dt dt'$$

$$VAR[\langle X(t) \rangle_{T}] = \frac{1}{4T^2} \int_{-T}^{T} \int_{-T}^{T} C_X(t, t') dt dt'$$

This is a general expression which is applicable not only for WSS processes

# **Ergodicity**

■ If X(t) is a WSS process:

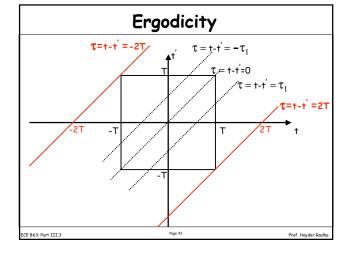
$$C_{\times}(\dagger,\dagger')=C_{\times}(\dagger-\dagger')$$

Then:  

$$VAR\left[\left\langle X(t)\right\rangle_{T}\right] = \frac{1}{4T^{2}} \int_{-T}^{T} \int_{-T}^{T} C_{X}(t,t') dt dt'$$

$$\bigcup$$

$$VAR\left[\left\langle X(t)\right\rangle_{T}\right] = \frac{1}{4T^{2}} \int_{-T}^{T} \int_{-T}^{T} C_{X}(t-t') dt dt'$$



# **Ergodicity**

- lacktriangle Since the covariance  $C_X$  is a function of the time difference  $\tau = t - t'$ , we can express the 2dimensional integral as a one-dimensional integral of
- The function  $C_X$  is constant across the line  $\tau = t t'$ in the (t,t) space
- It can be seen that  $\tau$  ranges from (-2T) to (+2T)

It can be shown that the 2-dimensional integral:

$$VAR\left[\left\langle X(t)\right\rangle _{T}\right]=\frac{1}{4T^{2}}\int_{-T}^{T}\int_{-T}^{T}\mathcal{C}_{X}\left(t-t^{\prime}\right)dt\ dt^{\prime}$$

is equivalent to the following integral:

$$VAR\left[X\left\langle \left(\dagger\right)\right\rangle _{T}\right]=\frac{1}{4T^{2}}\int_{-2T}^{2T}\left(2T-\left|\tau\right|\right)C_{X}\left(\tau\right)d\tau$$

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#### **Ergodicity**

■ Therefore, the variance of the time average  $\langle X(t) \rangle_T$  of a WSS process X(t) can be expressed as:

$$VAR\left[X\left\langle (\dagger)\right\rangle_{T}\right] = \frac{1}{2T} \int_{-2T}^{2T} \left(1 - \frac{|\tau|}{2T}\right) C_{X}(\tau) d\tau$$

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# **Ergodicity**

Consequently, if the following is satisfied:

$$\underset{T\rightarrow\infty}{lim}\frac{1}{2T}\int_{2T}^{2T}\!\!\left(1\!-\!\frac{\left|\tau\right|}{2T}\right)\!\!C_{X}\left(\tau\right)d\tau=0$$

then the variance of the time-average converges to zero which leads to the convergence of the time-average to the mean (m) in a "mean-square-sense".

END of the Proof

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#### APPENDIX B

A stationary process is

a Wide-Sense-Stationary
(WSS) Process

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# Stationary Processes

■ The "second-order cdf":

$$\mathsf{F}_{\mathsf{X}(\mathsf{t}_1),\mathsf{X}(\mathsf{t}_2)}(\mathsf{x}_{\!\scriptscriptstyle{1}},\!\mathsf{x}_{\!\scriptscriptstyle{2}})$$

of a stationary process is only dependent on the time difference  $(t_2-t_1)$ :

$$F_{X\left(t_{1}\right),X\left(t_{2}\right)}\left(x_{1},x_{2}\right)=F_{X\left(t_{1}+\tau\right),X\left(t_{2}+\tau\right)}\left(x_{1},x_{2}\right) \qquad \forall \ t_{1},t_{2},\tau$$

By setting  $\tau = -t_1$ :

$$\boxed{F_{X(t_1),X(t_2)}\big(x_1,x_2\big) = F_{X(0),X(t_2-t_1)}\big(x_1,x_2\big) \qquad \forall\, t_1,t_2}$$

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## Stationary Processes

■ Therefore, for a stationary process, the autocorrelation and autocovariance functions depend only on the time difference  $(t_2-t_1)$ :

$$R_X(t_1,t_2) = R_X(t_2-t_1) \quad \forall t_1,t_2$$

$$C_{X}(t_{1},t_{2}) = C_{X}(t_{2}-t_{1}) \quad \forall t_{1},t_{2}$$

■ Therefore, a stationary process is a WSS process

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