

EE 574 Detection and Estimation Theory

Lecture Presentation 4

Aykut HOCANIN

Dept. of Electrical and Electronic Engineering
Eastern Mediterranean University

👉 Chapter 2: Classical Detection and Estimation Theory

👉 Estimation Theory

A model of a general estimation problem is shown in Fig. 1. The model has the following components.

➤ **Parameter Space.**

The output of the source is a parameter in a parameter space. For the **single** parameter space, this corresponds to the line $-\infty < A < \infty$.

- The parameter is a random variable.
- The parameter is an unknown quantity but not a random variable.

➤ **Probabilistic Mapping from Parameter Space to Observation Space.**

This is the probability law that governs the effect of a on the observation.

➤ **Observation Space.**

In the classical problem, this is a finite-dimensional space. It is denoted by the vector \mathbf{R} .

➤ **Estimation Rule.**

After observing \mathbf{R} an **estimate** of the value of a is made and this estimate is denoted by $\hat{a}(\mathbf{R})$. This mapping of the observation space into an estimate is called the estimation rule.

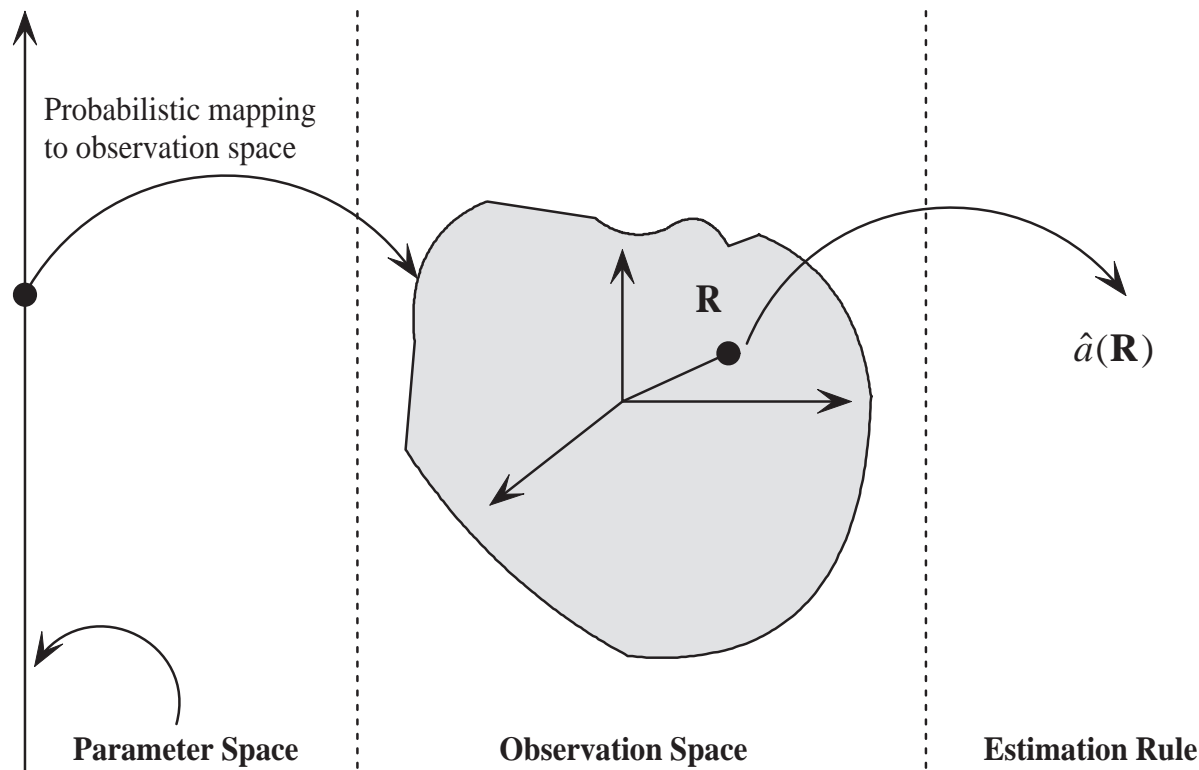


Figure 1: Estimation Model.

Random Parameters: Bayes Estimation

All pairs $[a, \hat{a}(\mathbf{R})]$ are assigned **costs** over the range of interest. This function of two variables is denoted as $C(a, \hat{a})$. Usually, the cost depends only on the **error** of the estimate which is given by

$$a_e(\mathbf{R}) = \hat{a}(\mathbf{R}) - a$$

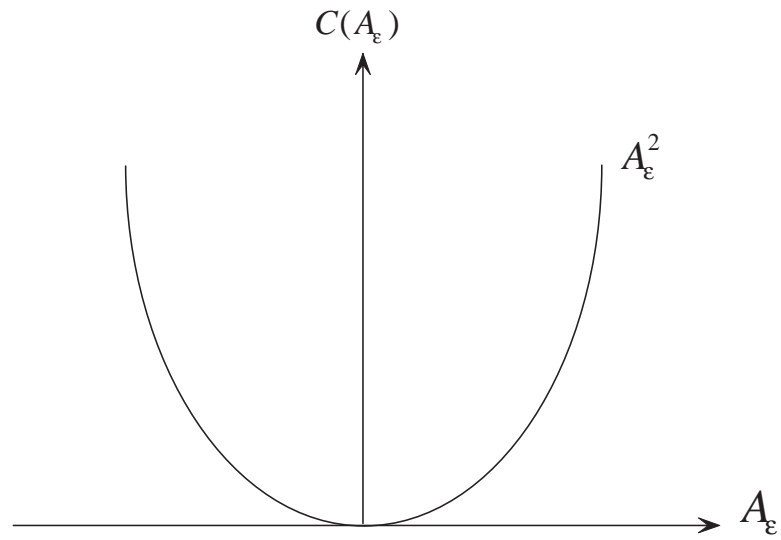


Figure 2: *Mean-square error cost function.*

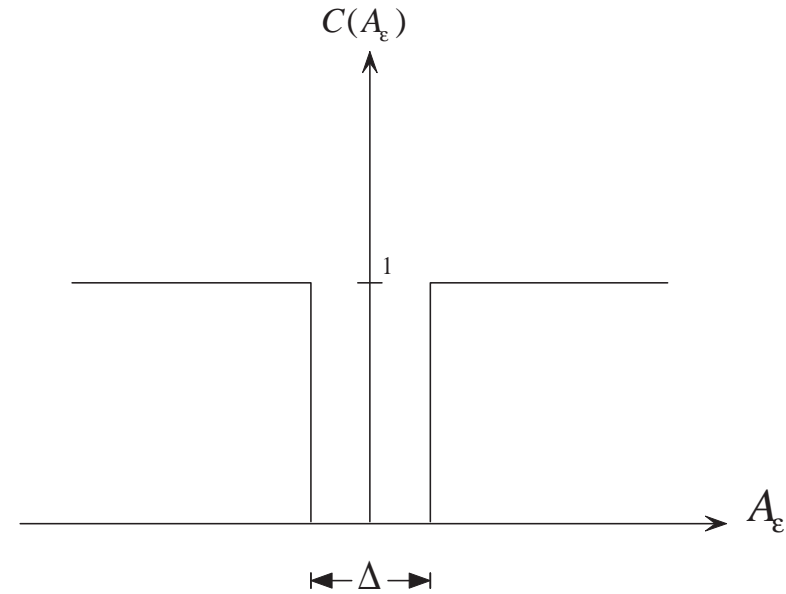
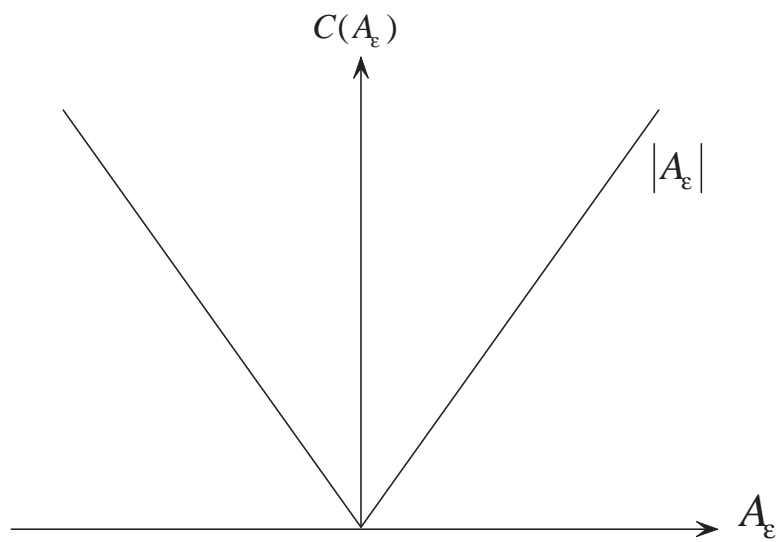


Figure 3: *Absolute error and uniform error cost functions*

The cost function $C(a_\epsilon)$ is a function of the error hence it is a function of a single variable. Some typical cost functions are shown Figs. 2 and 3.

➤ Squared error function.

$$C(a_\epsilon) = a_\epsilon^2$$

➤ Absolute value of the error function

$$C(a_\epsilon) = |a_\epsilon|$$

➤ Uniform error cost function

$$C(a_\epsilon) = \begin{cases} 0, & |a_\epsilon| \leq \frac{\Delta}{2} \\ 1, & |a_\epsilon| > \frac{\Delta}{2} \end{cases}$$

Just like the detection problem, it is assumed that the a priori probability density $p_a(A)$ in the random parameter estimation problem is known. If it is unknown, a procedure similar to the minimax test may be used.

The expression for the risk becomes:

$$\mathcal{R} = E\{C[A, \hat{a}(\mathbf{R})]\} = \int_{-\infty}^{\infty} C[a, \hat{a}(\mathbf{R})] p_{a,\mathbf{r}}(A, \mathbf{R}) dA d\mathbf{R}. \quad (1)$$

The expectation is over the random variable a and the observed variables \mathbf{r} . For costs which are defined on the error only Eq.(1) becomes

$$\mathcal{R} = \int_{-\infty}^{\infty} C[A - \hat{a}(\mathbf{R})] p_{a,\mathbf{r}}(A, \mathbf{R}) dA d\mathbf{R}. \quad (2)$$

Bayes estimate is the estimate that **minimizes the risk**. We write the joint density as:

$$p_{a,\mathbf{r}}(A, \mathbf{R}) = p_{\mathbf{r}}(\mathbf{R}) p_{a|\mathbf{r}}(A|\mathbf{R})$$

We then take the derivatives with respect to \hat{a} to obtain the estimates. The Bayes estimates for the cost functions defined above are as follows:

➤ **Bayes estimate for the square error cost function: (Mean Square Estimate)**

$$\hat{a}_{\text{ms}} = \int_{-\infty}^{\infty} A p_{a|\mathbf{r}}(A|\mathbf{R}) dA \quad (3)$$

The term on the right hand side of Eq. (3) is the **mean of the 'a posteriori density'** (conditional mean).

➤ **Bayes estimate for the absolute value criterion**

$$\int_{-\infty}^{\hat{a}_{\text{abs}}(\mathbf{R})} p_{a|\mathbf{r}}(A|\mathbf{R}) dA = \int_{\hat{a}_{\text{abs}}(\mathbf{R})}^{\infty} p_{a|\mathbf{r}}(A|\mathbf{R}) dA \quad (4)$$

This is given by the **median of the 'a posteriori' density**.

➤ **Bayes estimate for the uniform cost function:(Maximum a Posteriori (MAP) Estimate)**

This is obtained when $\lim \Delta \rightarrow 0$ for the uniform cost function. In order to find $\hat{a}_{\text{map}}(\mathbf{R})$ we must have the **location of the maximum of the 'a posteriori' density**.

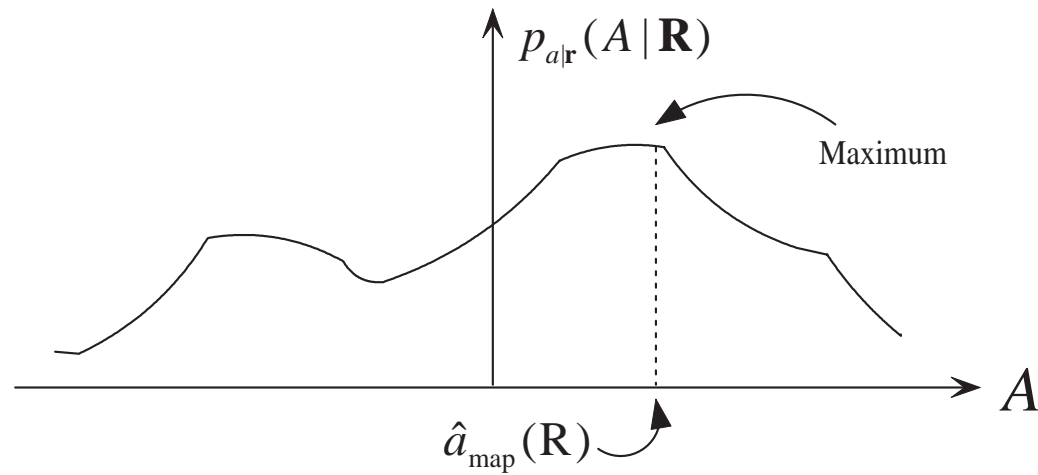


Figure 4: Maximum of the 'a posteriori' density.

Properties for the invariance of the cost functions for the estimators

1. We assume that the cost function $C(a_\epsilon)$ is a **symmetric**, **convex upward** function and that the a posteriori density $p_{a|R}(A|R)$ is **symmetric about its conditional mean**: that is,

$$C(a_\epsilon) = C(-a_\epsilon) \quad \text{Symmetry}$$

$$C(bx_1 + (1 - b)x_2) \leq bC(x_1) + (1 - b)C(x_2) \quad \text{Convexity}$$

For *non-convex functions*, we have property 2,

2. We assume that the cost function is a **symmetric**, **nondecreasing** function and that the a posteriori density $p_{a|\mathbf{r}}(A|\mathbf{R})$ is a symmetric function about its **conditional mean**. It is a **unimodal** function that satisfies the condition

$$\lim_{x \rightarrow \infty} C(x)p_{a|\mathbf{r}}(A|\mathbf{R}) = 0$$

The estimate \hat{a} that minimizes any cost function in this class (satisfying the properties) is identical to \hat{a}_{ms} .