# Linear regression model: classical approach

The observed data: (i). Any single observation is given by (Y, z)

Y is the observed variable

 $z = (z_0, \ldots, z_k)$  is the covariate

 $z_0 \in \{0,1,2\}$  is assumed to be a genotype

The distribution of Y is assumed to be a normal; any element  $z_i$  of the covariate z can be

nominal (categorical, factor)

ordinal

quantitative (numeric)

#### Covariates:

the covariate  $z_0 \in \{0,1,2\}$  (genotype) is assumed to be of special ordinal/nominal type (object "genotype")

type of each other covariate should be specified by user or selected by default

## By default

a binary categorical covariate (2 levels) is assumed to be a nominal, and should be converted to 0 and 1 levels:

a categorical covariate having 3-8 levels is assumed to be an ordinal; other covariates are assumed to be numeric.

(ii). The observed data is a sample (Y, z):

$$Y = (Y_1, \dots, Y_n)'$$
 is  $n \times 1$ -vector

$$\mathbf{z} = (z_1, \dots, z_n), \ z_i = (z_{i,0}, \dots, z_{i,k})$$
 are the covariates  $Y_1, \dots, Y_n$  are conditionally independent given  $\mathbf{z}$   $\mathbb{V}\mathbf{ar}(Y|\mathbf{z}) = \sigma^2$  for any  $\mathbf{z}$   $(\mathbb{V}\mathbf{ar}(Y_1) = \dots = \mathbb{V}\mathbf{ar}(Y_n))$ 

Linear regression model: (i). Regressor  $X: \mathbb{R}^k \to \mathbb{R}^m$ ; X(z) is defined for any observation

(ii). Linear regression model for single observation

$$\mathbb{E}_{\theta}(Y|\boldsymbol{z}) = \boldsymbol{X}(\boldsymbol{z})'\boldsymbol{\beta}$$

 $\boldsymbol{\beta} = (\beta_1, \dots, \beta_m)'$  is  $m \times 1$ -vector of parameters

 $Y \sim \mathcal{N}(\boldsymbol{X}(\boldsymbol{z})'\boldsymbol{\beta}, \sigma)$ , where  $\sigma$  is the nuisance variance parameter

 $\boldsymbol{\theta} = (\beta_1, \dots, \beta_m, \sigma)$  is the full parameter of the model

The linear regression model is determined by formula

$$Y \sim q(z_0) + f(z^*), \quad z^* = (z_1, \dots, z_k),$$

 $Y \sim g(z_0) + f(\boldsymbol{z}^*), \quad \boldsymbol{z}^* = (z_1, \dots, z_k),$  where  $g \in \{CD, D, R, A\}$  is the genetic model and f is the covariate term

$$f(z^*) = f_1(z^*) + \ldots + f_r(z^*),$$

 $f_i(\boldsymbol{z}^*) = z_1^{r_{1i}} * \dots * z_k^{r_{ki}}, r_{1i}, \dots, r_{ki} \in \{0,1\}$  is of multiple regression type  $f_i(\boldsymbol{z}^*) = z_{s_i}^{r_i}$ , where  $r_i \in \{2,3,\dots\}$  and  $s_i \in \{1,\dots,k\}$  is of polynomial type (not available for factors)

#### The parameter

The intercept parameter is included into the model by default

The covariate  $z_0$  generates

single parameter if  $g \in \{D, R, A\}$  (difference with the baseline level  $z_0 = 0$ ) vector of length two if g = CD (main effect with respect to the baseline level  $z_0 = 0$ )

Term  $f(z^*)$  should be transformed to  $\sum_{\kappa \in K} z_{\kappa}$ , where  $\kappa$  is a multi-index; K is a (ordered) set of subsets of  $\{1,\ldots,k\}$ 

the term  $f_i(\boldsymbol{z}^*) = z_1^{r_{1i}} * \ldots * z_k^{r_{ki}}$  generates  $K_i$  is the set of non empty subsets of  $\{r_{k,i} : r_{k,i} = 1\}$ then  $K = \bigcup_{i=1}^{r} K_i$ 

the elements of K should be sorted ( $\kappa$ 's of smaller length should be sorted before  $\kappa$ 's of larger length)

We associate a number  $d_{\kappa}$  to each index  $\kappa \in K$ 

 $d_i = 1$  if  $z_i$  is numeric or ordinal  $(\kappa = i)$ 

 $d_i = r \text{ for } z_i^r$ 

 $d_i + 1$  is a number of levels of the covariate  $z_i$ , if  $z_i$  is a factor  $(\kappa = i)$ 

in general case  $d_{\kappa} = \prod_{i \in \kappa} d_i$ .

The total length of the nuisance (part corresponding to f) parameter is  $d = \sum_{\kappa \in K} d_{\kappa}$ .

### The regressor

The regressor is a vector  $(x_0, x_K)$  such that

 $x_0 = (1, z_0) \text{ if } g \in \{D, R, A\}$ 

 $x_0 = (1, 1\!\!1_{\{z_0=1\}}, 1\!\!1_{\{z_0=2\}})$  if g = "CD" (Model for TEST)  $x_0 = (1, z_0)$  if g = "CD" (Model for QAS)

 $x_0 = 1$  (Model for base)

We use the ordered set K to create regressors for nuisance parameter

 $x_i = z_i$  if  $z_i$  is numeric or ordinal  $(\kappa = i)$ 

 $x_i = (z_i, z_i^2, \dots, z_i^r)$  for  $z_i^r$ 

 $x_i$  is a vector of length  $d_i$  of  $\{0,1\}$  with 1 on j-th position if  $z_i = a_{j+1}, a_j$  is the j-th level of  $z_i$  (for example (0,0,1,0,0)).

in general case  $x_{\kappa} = \text{as.vector}(\otimes_{i \in \kappa} x_i)$ , where  $\otimes$  is the outer product (vector should be sorted in a proper order).

the regressor  $x_K$ , which corresponds to the nuisance parameter is obtained by merging  $x_{\kappa}$  over  $\kappa \in K$  in the proper order.

(iii). Linear regression model for the observed data

$$\mathbb{E}(Y|\mathbf{z}) = \mathbf{X}'\boldsymbol{\beta}$$

 $\mathbf{X} = (\mathbf{X}(\mathbf{z}_1), \dots, \mathbf{X}(\mathbf{z}_n))$  is the  $m \times n$ -matrix of regressors

 $\boldsymbol{\beta}$  is the  $1 \times m$ -vector of parameters

assumption  $Y \sim \mathcal{N}(\mathbf{X}'\boldsymbol{\beta}, \sigma^2\mathbf{I}_n)$ ,  $\mathbf{I}_n$  is the identity matrix

the matrix  $\mathbf{X} = (\mathbf{X}'_q, \mathbf{X}^{*'})'$ , where  $\mathbf{X}_q$  contains the regressors related to g and  $\mathbf{X}^*$  are the last d rows of the matrix X

Nuisance parameter reduction

the parameter related to  $(z_1, \ldots, z_k)$  is not a subject of our interests

If the matrix  $\mathbf{X}^*$ , which contains the last d rows of the matrix  $\mathbf{X}$ , is not of the full rank d, we reduce the matrix  $\mathbf{X}^*$  to  $\mathbf{X}^\circ$  by choosing only basis vectors of the linear space generated by rows of the matrix  $X^*$ 

The reduced matrix X is determined by changing block  $X^*$  to  $X^\circ$  in the matrix X

QAS and P-value. Under XX' is of full rank (positively definite), the least square estimator (LSE) is given by

$$\widehat{\boldsymbol{\beta}} = (\mathbf{X}\mathbf{X}')^{-1}\mathbf{X}\,\mathbf{Y}$$

The sum of squares

$$S(\boldsymbol{\beta}) = \| \boldsymbol{Y} - \mathbf{X}' \boldsymbol{\beta} \|^2 = (\boldsymbol{Y} - \mathbf{X}' \boldsymbol{\beta})' (\boldsymbol{Y} - \mathbf{X}' \boldsymbol{\beta})$$

Notations ("Model for QAS" is equal to "Model for TEST" if g = A, D, R"):

 $\mathbf{X}/\mathbf{X}_q/\mathbf{X}_0$  is the regressors in the "Model for TEST"/"Model for QAS"/"Baseline model" respectively

 $\hat{\boldsymbol{\beta}}/\hat{\boldsymbol{\beta}}_q/\hat{\boldsymbol{\beta}}_0$  is the LSE from the "Model for TEST"/"Model for QAS"/"Baseline model" respectively

 $S(\widehat{\boldsymbol{\beta}})/S_q(\widehat{\boldsymbol{\beta}}_q)/S_0(\widehat{\boldsymbol{\beta}}_0)$  is the sums of squares from the "Model for TEST"/"Model for QAS"/ "Baseline model" respectively

#### The algorithm

The input data contain: "phenotype" (numeric vector), "genotype", "covariates" (matrix), "covariate types" (vector) and "formula"

- (i) Create regressors related to covariates  $X^*$
- (ii) Check that the rank of the matrix  $\mathbf{X}^*$  is full

if the rank of matrix  $\mathbf{X}^*$  is not equal to d, apply nuisance parameter reduction  $\mathbf{X}^* := \mathbf{X}^\circ$ .

# (iii) Loop by genotypes

Create regressors  $\mathbf{X}_g$  for "Model for TEST", for "Model for QAS" and for "Baseline model" Check that the rank of the matrix  $\mathbf{X}_g$  related to g is maximal (2 if g = A, D, R" or 3 if g = CD")

if the rank of the matrix  $\mathbf{X}_g$  in the "Model for TEST" is not maximal, report qas[i] = NA and pv[i] = NA otherwise, continue

Create matrix **X** by merging  $\mathbf{X}_g$  and  $\mathbf{X}^*$ 

If the rank of matrix  $\mathbf{X}$  is not equal to its number of rows use nuisance parameter reduction to whole  $\mathbf{X}$  ( $\mathbf{X}_q$  should not be changed)

Fit the "Model for QAS"

assign qas[i] the coefficient, related to  $z_0$  (second element of the vector  $\widehat{\beta}$ )

Fit the "Model for TESTS" and the "Baseline model"

$$SS_e = S(\widehat{\boldsymbol{\beta}}); \overline{SS}_e = SS_e/\mathrm{df}, \text{ where df} = n - \mathbf{rk}(\mathbf{X})$$
  
 $SS_h = S_0(\widehat{\boldsymbol{\beta}}_0) - SS_e; \overline{SS}_h = SS_H/\mathrm{df}_h, \text{ where df}_h = 2 \text{ if } g = CD \text{ and df}_h = 1 \text{ if } g = A,D,R$ 

 $\mathbb{F} = \overline{SS_h}/\overline{SS_e}$  is the F-statistic, which has under null hypothesis  $F_{\mathrm{df}_h,\mathrm{df}}$ -distribution assign  $pv[i] := 1 - F_{\mathrm{df}_h,\mathrm{df}}(\mathbb{F})$ , where  $F_{\mathrm{df}_h,\mathrm{df}}$  is the distribution function of the Fisher–Snedecor distribution  $F_{\mathrm{df}_h,\mathrm{df}}$ 

(iii) Return the vectors pv and qas.