# Double exponential jump model (Kou)

#### DIA El Hadj Aly

Laboratoire Analyse et Mathématiques appliquées, Université de Marne-la-Vallée

March 1, 2012

# Premia 14

The double exponential jump model, initiated by Steven KOU (see [1]), is an exponential Levy model, which is a compromise between reality and tractability. It gives an explanation of the two empirical phenomena which received much attention in financial markets: the asymmetric leptokurtic feature and the volatility smile. It permits to obtain analytical solutions to the prices of many derivatives: European call and put options; interest rate derivatives, such as swaptions, caps, floors, and bond options; as well as path-dependant options, such as perpetual American options, barrier, and lookback options.

#### 1 The model

The behaviour of the asset price,  $S_t$ , under the risk neutral probability is modeled as followed:

$$\frac{dS_t}{S_{t^-}} = \mu dt + \sigma dW_t + d\left(\sum_{i=1}^{N_t} e^{Y_i} - 1\right)$$
 (1.1)

Where W is a standard brownian motion, N is a poisson process with rate  $\lambda$ , the constants  $\mu$  and  $\sigma > 0$  are drift and volatility of the diffusion part and the jump sizes  $\{Y_1, Y_2, \ldots\}$  are i.i.d random variables with a common asymmetric double exponential distribution, of density:

$$f_Y(y) = p\eta_1 e^{-\eta_1 y} \mathbb{1}_{\{y \ge 0\}} + q\eta_2 e^{\eta_2 y} \mathbb{1}_{\{y < 0\}}$$
(1.2)

where  $p,q \ge 0$  are constants, p + q = 1,  $\eta_1 > 1$  and  $\eta_2 > 0$ .

The random processes  $(W_t)_{t\geq 0}$ ,  $(N_t)_{t\geq 0}$ , and random variables  $\{Y_1,Y_2,\ldots\}$  are independent. Furthermore we have  $\mu=r-\lambda\xi$  with:

$$\xi = \frac{p\eta_1}{\eta_1 - 1} + \frac{q\eta_2}{\eta_2 + 1} - 1 \tag{1.3}$$

The condition on  $\mu$  hold in order to obtain  $(e^{-rt}S_t)_{t\geq 0}$  is a martingale. The caracteristic exponent G of  $\log{(S_t)}$  (i.e.  $\mathbb{E}\left[e^{\theta\log(S_t)}\right]=e^{G(\theta)t}$ ) is defined as:

$$G(x) = x \left( r - \frac{1}{2} - \lambda \xi \right) + \frac{1}{2} x^2 \sigma^2 + \lambda \left( \frac{p \eta_1}{\eta_1 - x} + \frac{q \eta_2}{\eta_2 + x} - 1 \right)$$

The equation  $G(x) = \alpha$  has exactly four roots (see [2]) :  $\beta_{1,\alpha}$ ,  $\beta_{2,\alpha}$ ,  $-\beta_{3,\alpha}$ ,  $-\beta_{4,\alpha}$ , where

$$0 < \beta_{1,\alpha} < \beta_{2,\alpha} < \infty, \quad 0 < \beta_{3,\alpha} < \beta_{4,\alpha} < \infty. \tag{1.4}$$

## 2 European call and put

Let us define some special functions (see pp. 1094 and 1099 in[1]):

$$Hh_{-1}(x) = e^{-\frac{x^2}{2}}$$

$$Hh_0(x) = \sqrt{2\pi}\Phi(-x)$$

$$Hh_n(x) = \int_x^{+\infty} Hh_{n-1}(y)dy = \frac{1}{n!} \int_x^{+\infty} (t-x)^n e^{-\frac{t^2}{2}} dt \quad \forall n \ge 0$$

$$I_n(c; \alpha, \beta, \gamma) = \int_c^{+\infty} e^{\alpha x} Hh_n(\beta c - \gamma) dx \quad \forall n \ge -1$$

where  $\Phi$  is the standard normal cumulative distribution. Then we have :

$$nHh_n(x) = Hh_{n-2}(x) - xHh_{n-1}(x) \quad \forall n > 1$$

And  $\forall n \geq -1$ :

$$I_{n}(c;\alpha,\beta,\gamma) = -\frac{e^{\alpha c}}{\alpha} \sum_{i=0}^{n} \left(\frac{\beta}{\alpha}\right)^{n-i} Hh_{i}(\beta c - \delta) + \left(\frac{\beta}{\alpha}\right)^{n+1} \frac{\sqrt{2\pi}}{\beta} e^{\frac{\alpha \delta}{\beta} + \frac{\alpha^{2}}{2\beta^{2}}} \Phi\left(-\beta c + \delta + \frac{\alpha}{\beta}\right) \quad \beta > 0, \alpha \neq 0$$

$$I_{n}(c;\alpha,\beta,\gamma) = -\frac{e^{\alpha c}}{\alpha} \sum_{i=0}^{n} \left(\frac{\beta}{\alpha}\right)^{n-i} Hh_{i}(\beta c - \delta) - \left(\frac{\beta}{\alpha}\right)^{n+1} \frac{\sqrt{2\pi}}{\beta} e^{\frac{\alpha \delta}{\beta} + \frac{\alpha^{2}}{2\beta^{2}}} \Phi\left(\beta c - \delta - \frac{\alpha}{\beta}\right) \quad \beta < 0, \alpha < 0$$

Introduce the following notation: For any given probability P, define:

$$\psi(\mu, \sigma, \lambda, p, \eta_1, \eta_2; a, T) = \mathbb{P}\left[Z_T \ge a\right] \tag{2.5}$$

where  $Z_T = \mu t + \sigma W_t + \sum_{i=1}^{N_t} Y_i$  and Y has a double exponential distribution with density as in (1.2), and N is a poisson process with rate  $\lambda$ . Theorem B.1. in [1] gives us:

$$\begin{split} \psi(\mu,\sigma,\lambda,p,\eta_1,\eta_2;a,T) &= & \frac{e^{(\sigma\eta_1)^2\frac{T}{2}}}{\sigma\sqrt{2\pi T}} \sum_{n=1}^{+\infty} \pi_n \sum_{k=1}^n P_{n,k} \left(\sigma\sqrt{T}\eta_1\right)^k I_{k-1} \left(a-\mu T;-\eta_1,-\frac{1}{\sigma\sqrt{T}},-\sigma\sqrt{T}\eta_1\right) \\ &+ \frac{e^{(\sigma\eta_2)^2\frac{T}{2}}}{\sigma\sqrt{2\pi T}} \sum_{n=1}^{+\infty} \pi_n \sum_{k=1}^n Q_{n,k} \left(\sigma\sqrt{T}\eta_2\right)^k I_{k-1} \left(a-\mu T;\eta_2,\frac{1}{\sigma\sqrt{T}},-\sigma\sqrt{T}\eta_2\right) \\ &+ \pi_0 Phi \left(-\frac{a-\mu T}{\sigma\sqrt{T}}\right) \end{split}$$

where

$$P_{n,i} := \sum_{j=i}^{n-1} p^{j} q^{n-j} \binom{n-i-1}{j-i} \left(\frac{\eta_{1}}{\eta_{1}+\eta_{2}}\right)^{j-i} \left(\frac{\eta_{2}}{\eta_{1}+\eta_{2}}\right)^{n-j}, \quad 1 \le i \le n-1$$

$$Q_{n,i} := \sum_{j=i}^{n-1} q^{j} p^{n-j} \binom{n-i-1}{j-i} \left(\frac{\eta_{2}}{\eta_{1}+\eta_{2}}\right)^{j-i} \left(\frac{\eta_{1}}{\eta_{1}+\eta_{2}}\right)^{n-j}, \quad 1 \le i \le n-1$$

$$P_{n,n} := p^{n}; \quad Q_{n,n} := q^{n}, \quad \pi_{n} = \frac{e^{-\lambda T} \lambda^{n}}{n!}$$

Using theorem 2, in [1], we know that the price of european call at inception and with maturity T is :

$$S_0\psi\left(r+\frac{1}{2}\sigma^2-\lambda\xi,\sigma,\widetilde{\lambda},\widetilde{p},\widetilde{\eta_1},\widetilde{\eta_2};\log\left(\frac{K}{S_0}\right),T\right)-Ke^{-rT}\psi\left(r-\frac{1}{2}\sigma^2-\lambda\xi,\sigma,\lambda,p,\eta_1,\eta_2;\log\left(\frac{K}{S_0}\right),T\right)$$

where

$$\widetilde{p} = \frac{p}{1+\xi} \frac{\eta_1}{\eta_1 - 1}, \quad \widetilde{\lambda} = \lambda(1+\xi), \quad \widetilde{\eta_1} = \eta_1 - 1, \quad \widetilde{\eta_2} = \eta_2 + 1$$

The put price can be obtain by using the call-put parity.

## 3 Finite time horizon american put option

Let EuP(v,t) be the price of a european put option with initial stock price v and maturity t,  $\mathbb{P}^v[S_t \leq K]$  the probability that the stock price at t is below K with initial stock price v,  $z = 1 - e^{-rT}$ ,  $\beta_3 \equiv \beta_{3,r/z}$ ,  $\beta_4 \equiv \beta_{5,r/z}$ ,  $C_\beta = \beta_3\beta_4(1 + \eta_2)$  (see (1.4)),  $D_\beta = \eta_2(1 + \beta_3)(1 + \beta_4)$ ,  $v_0 \equiv v_0(t) \in (0, K)$  the unique solution to the equation

$$C_{\beta}K - D_{\beta}(v_0 + EuP(v_0, t)) = (C_{\beta} - D_{\beta})Ke^{-rT}\mathbb{P}^{v_0}[S_t \le K]$$
 (3.6)

and

$$A = \frac{v_0^{\beta_3}}{\beta_4 - \beta_3} \left\{ \beta_4 K - (1 + \beta_4) \left[ v_0 + EuP(v_0, t) \right] + Ke^{-rT} \mathbb{P}^{v_0} \left[ S_t \le K \right] \right\} > 0,$$

$$B = \frac{v_0^{\beta_4}}{\beta_3 - \beta_4} \left\{ \beta_4 K - (1 + \beta_3) \left[ v_0 + EuP(v_0, t) \right] + Ke^{-rT} \mathbb{P}^{v_0} \left[ S_t \le K \right] \right\} > 0,$$

Then the price of a finite-horizon american put option with maturity t and strike K can be approximated by  $\psi(S_0, t)$  which is given by (see §3 in [3])

$$\psi(v,t) = \begin{cases} EuP(v,t) + Av^{-\beta_3} + Bv^{-\beta_4}, & if \ v \ge v_0 \\ K - v, & if \ v \le v_0 \end{cases}$$

## 4 Lookback option

The price of a lookback floating strike put option is given by :

$$LP(T) = \mathbb{E}\left[e^{-rT}\left(\max\left\{M, \max_{0 \le t \le T} S_t\right\} - S_T\right)\right]$$
$$= \mathbb{E}\left[e^{-rT}\left(\max\left\{M, \max_{0 \le t \le T} S_t\right\}\right)\right] - S_0$$

where  $M \geq S_0$  is a fixed constant representing the prefixed maximum at time 0. The Laplace transform of the lookback put, using notations in 1.4, is given by (see theorem 1 in [3])

$$\int_0^{+\infty} e^{-\alpha T} LP(T) dT = \frac{S_0 A_\alpha}{C_\alpha} \left(\frac{S_0}{M}\right)^{\beta_{1,\alpha+r}-1} + \frac{S_0 B_\alpha}{C_\alpha} \left(\frac{S_0}{M}\right)^{\beta_{2,\alpha+r}-1} + \frac{M}{\alpha+r} - \frac{S_0}{\alpha} \quad \forall \alpha > 0$$

where

$$A_{\alpha} = \frac{(\eta_1 - \beta_{1,\alpha+r})\beta_{2,\alpha+r}}{\beta_{1,\alpha+r} - 1}$$

$$B_{\alpha} = \frac{(\beta_{2,\alpha+r} - \eta_1)\beta_{1,\alpha+r}}{\beta_{2,\alpha+r} - 1}$$

$$C_{\alpha} = (\alpha + r)\eta_1(\beta_{2,\alpha+r} - \beta_{1,\alpha+r})$$

The put price is obtained by using an inversion of the Laplace transform. The call option price follows just by symmetry. For the lookback fixed strike, when we have  $M \geq max(S_0, K)$  for the put or  $m \leq min(S_0, K)$  for the call, we get similar results to those for floatings.

## 5 Barrier option

Since all eight types of barrier can be solved in similar way, we focus only on the price Up and In Call option defined as followed

$$UIC = \mathbb{E}\left[e^{-rT}\left(S_T - K\right)^{+} \mathbb{1}_{\left\{\max_{0 \le t \le T} S_t \ge H\right\}}\right]$$
(5.7)

where  $H > S_0$  is the barrier level. For any given probability P, define;

$$\Psi(\mu, \sigma, \lambda, p, \eta_1, \eta_2; a, b, T) = \mathbb{P}\left[Z_T \ge a, \max_{0 \le t \le T} Z_t \ge b\right]$$
 (5.8)

where  $Z_T = \mu t + \sigma W_t + \sum_{i=1}^{N_t} Y_i$  and Y has a double exponential distribution with density as in (1.2), and N is a poisson process with rate  $\lambda$ . Using formula (3.1) and the result before remark 3.1 in [2], we get

$$\int_{0}^{+\infty} e^{-\alpha T} \mathbb{P}\left[\max_{0 \le t \le T} Z_{t} \ge b\right] := \frac{1}{\alpha} \left(\frac{(\eta_{1} - \beta_{1,\alpha})\beta_{2,\alpha}}{\eta_{1}(\beta_{2,\alpha} - \beta_{1,\alpha})} e^{-b\beta_{1,\alpha}} + \frac{(\beta_{2,\alpha} - \eta_{1})\beta_{1,\alpha}}{\eta_{1}(\beta_{2,\alpha} - \beta_{1,\alpha})} e^{-b\beta_{2,\alpha}}\right)$$

By Inverting the Laplace transform we get  $\mathbb{P}[\max_{0 \le t \le T} Z_t \ge b]$ , which is useful for some types of barrier options. Let us now define some functions

$$\begin{split} H_{i}(a,b,c;n) &:= \frac{1}{\sqrt{2\pi}} \int_{0}^{\infty} e^{\left(\frac{1}{2}c^{2}-b\right)t} t^{n+\frac{i}{2}} H h_{i} \left(c\sqrt{t} + \frac{a}{\sqrt{t}}\right) dt \quad i \geq -1, n \geq 0 \\ A_{\alpha} &:= \mathbb{E}\left[e^{-\alpha\tau_{b}} \mathbb{1}_{X_{\tau_{b}} = b}\right] \\ &= \frac{\eta_{1} - \beta_{1,\alpha}}{\beta_{2,\alpha} - \beta_{1,\alpha}} e^{-b\beta_{1,\alpha}} + \frac{\beta_{2,\alpha} - \eta_{1}}{\beta_{2,\alpha} - \beta_{1,\alpha}} e^{-b\beta_{2,\alpha}} \\ B_{\alpha} &:= \mathbb{E}\left[e^{-\alpha\tau_{b}} \mathbb{1}_{X_{\tau_{b}} > b}\right] \\ &= \frac{(\eta_{1} - \beta_{1,\alpha})(\beta_{2,\alpha} - \eta_{1})}{\eta_{1}(\beta_{2,\alpha} - \beta_{1,\alpha})} \left[e^{-b\beta_{1,\alpha}} - e^{-b\beta_{2,\alpha}}\right] \end{split}$$

where  $\tau_b = \inf\{t \geq 0; \ X_t \geq b\}$ . Hh functions are defined in § 2, and  $\beta$  variables in (1.4). For  $i \geq 1$ , under assumption that b > 0 and  $c > -\sqrt{2b}$ , we have

$$H_i(a,b,c;n) = \frac{1}{i}H_{i-2}(a,b,c;n+1) - \frac{c}{i}H_{i-1}(a,b,c;n+1) - \frac{a}{i}H_{i-1}(a,b,c;n)$$

By knowing  $H_{-1}(a, b, c; n)$  and  $H_0(a, b, c; n)$ , this recursive formula allows us to determine all values of  $H_i$ . Lemmas A.1 and A.2 in [2] give us

$$H_{-1}(a,b,c;n) = e^{-ac - \sqrt{2a^2b}} \sqrt{\frac{1}{2b}} \left( \sqrt{\frac{a^2}{2b}} \right)^n \sum_{j=0}^n \frac{(-n)_j (n+1)_j}{j! \left( -2\sqrt{2a^2b} \right)^j}, \quad a \neq 0, n \geq 0$$

$$H_{-1}(a,b,c;n) = e^{-ac - \sqrt{2a^2b}} \sqrt{\frac{1}{2b}} \left( \sqrt{\frac{a^2}{2b}} \right)^n \sum_{j=0}^{n-n-1} \frac{(-n)_j (n+1)_j}{j! \left( -2\sqrt{2a^2b} \right)^j}, \quad a \neq 0, n \leq -1$$

$$H_{-1}(0,b,c;n) = \frac{(2n)!}{n! (4b)^n} \frac{1}{2b}, \quad n \geq 0$$

$$H_{0}(a,b,c;n) = \frac{c}{2(n+1)} H_{-1}(a,b,c;n+1) - \frac{a}{2(n+1)} H_{-1}(a,b,c;n), \quad b = \frac{1}{2}c^2, n \geq 0$$

And  $\forall n \geq 0$  et  $b \neq \frac{1}{2}c^2$ 

$$H_0(a,b,c;n) = \frac{n!}{\left(b - \frac{1}{2}c^2\right)^{n+1}} \sum_{i=0}^n \frac{\left(b - \frac{1}{2}c^2\right)^i}{i!} \left(\frac{a}{2}H_{-1}(a,b,c;i-1) - \frac{c}{2}H_{-1}(a,b,c;i)\right), \quad a > 0$$

$$H_0(a,b,c;n) = \frac{n!}{\left(b - \frac{1}{2}c^2\right)^{n+1}} \left(1 + \sum_{i=0}^n \frac{\left(b - \frac{1}{2}c^2\right)^i}{i!} \left(\frac{a}{2}H_{-1}(a,b,c;i-1) - \frac{c}{2}H_{-1}(a,b,c;i)\right)\right), \quad a < 0$$

$$H_0(a,b,c;n) = \frac{n!}{\left(b - \frac{1}{2}c^2\right)^{n+1}} \left(\frac{1}{2} + \sum_{i=0}^n \frac{\left(b - \frac{1}{2}c^2\right)^i}{i!} \frac{c}{2}H_{-1}(a,b,c;i)\right), \quad a = 0$$

where  $(n)_j = n(n+1) \dots (n+j-1)$ , with convention  $(n)_0 = 1$ . We can now determine the exact expression of the Laplace transform of  $\Psi$  when b > 0 and  $a \le b$  (see theorem 4.1 in [2])

$$\begin{split} \int_0^{+\infty} e^{-\alpha T} \mathbb{P} \left[ Z_T \geq a, \max_{0 \leq t \leq T} Z_t \geq b \right] dT &= A_\alpha \int_0^{+\infty} e^{-\alpha T} \mathbb{P} \left[ Z_T \geq a - b \right] dT \\ &+ B_\alpha \int_0^{+\infty} e^{-\alpha T} \mathbb{P} \left[ Z_T + \xi^+ \geq a - b \right] dT \\ &= \left( A_\alpha + B_\alpha \right) \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} H_0 \left( -h, \gamma_\alpha, -\frac{\mu}{\sigma}; n \right) \\ &+ e^{h\sigma \eta_1} \sum_{n=1}^{\infty} \sum_{j=1}^n \frac{\lambda^n}{n!} \left( A_\alpha P_{n,j} + B_\alpha \overline{P}_{n,j} \right) \sum_{i=0}^{j-1} (\sigma \eta_1)^i H_i \left( h, \gamma_\alpha, c_+; n \right) \\ &- e^{-h\sigma \eta_2} \sum_{n=1}^{\infty} \sum_{j=1}^n \frac{\lambda^n}{n!} \left( A_\alpha Q_{n,j} + B_\alpha \overline{Q}_{n,j} \right) \sum_{i=0}^{j-1} (\sigma \eta_1)^i H_i \left( -h, \gamma_\alpha, c_-; n \right) \\ &+ e^{h\sigma \eta_1} B_\alpha \sum_{n=1}^{\infty} \sum_{j=1}^n \frac{(\lambda)^n}{n!} (\sigma \eta_1)^i H_i \left( h, \gamma_\alpha, c_+; n \right) \\ &+ e^{h\sigma \eta_1} B_\alpha H_0 \left( h, \gamma_\alpha, c_+; 0 \right) \end{split}$$

where  $\xi^+$  has an exponential law with rate  $\eta_1$ , matrix P and Q are as defined in  $\S$  2, and

$$\overline{P}_{n,1} := \sum_{j=i}^{n-1} Q_{n_i} \left( \frac{\eta_2}{\eta_1 + \eta_2} \right)^i, \quad \overline{P}_{n,i} := P_{n,i-1}, \quad 2 \le i \le n+1$$

$$\overline{Q}_{n,i} := \sum_{j=i}^n \binom{n}{j} q^j p^{n-j} \binom{n-i}{j-i} \left( \frac{\eta_2}{\eta_1 + \eta_2} \right)^{j-i} \left( \frac{\eta_1}{\eta_1 + \eta_2} \right)^{n-j+1}, \quad 1 \le i \le n$$

$$c_+ := \sigma \eta_1 + \frac{\mu}{\sigma}, \quad c_- := \sigma \eta_2 - \frac{\mu}{\sigma}, \quad \gamma_\alpha := \alpha + \lambda + \frac{\mu^2}{2\sigma^2}, \quad h := \frac{b-a}{\sigma}$$

For to get numerically  $\mathbb{P}\left[Z_T \geq a, \max_{0 \leq t \leq T} Z_t \geq b\right]$  for a given T, i find that is better to inverse the right term in the first equality above, using some propreties of the Laplace inversion. Note that  $\mathbb{P}\left[Z_T \geq a - b\right]$  is given in § 2 and  $\mathbb{P}\left[Z_T + \xi^+ \geq a - b\right]$  is given in [2] (pp. 528, formula B.5):

$$\mathbb{P}\left[Z_{T} + \xi^{+} \geq a\right] = \frac{e^{(\sigma\eta_{1})^{2}\frac{T}{2}}}{\sigma\sqrt{2\pi T}} \sum_{n=1}^{+\infty} \pi_{n} \sum_{k=1}^{n+1} \overline{P}_{n,k} \left(\sigma\sqrt{T}\eta_{1}\right)^{k} I_{k-1} \left(a - \mu T; -\eta_{1}, -\frac{1}{\sigma\sqrt{T}}, -\sigma\sqrt{T}\eta_{1}\right) \\
+ \frac{e^{(\sigma\eta_{2})^{2}\frac{T}{2}}}{\sigma\sqrt{2\pi T}} \sum_{n=1}^{+\infty} \pi_{n} \sum_{k=1}^{n} \overline{Q}_{n,k} \left(\sigma\sqrt{T}\eta_{2}\right)^{k} I_{k-1} \left(a - \mu T; \eta_{2}, \frac{1}{\sigma\sqrt{T}}, -\sigma\sqrt{T}\eta_{2}\right) \\
+ \pi_{0}\eta_{1} \frac{e^{(\sigma\eta_{1})^{2}\frac{T}{2}}}{\sqrt{2\pi}} I_{0} \left(a - \mu T; -\eta_{1}, -\frac{1}{\sigma\sqrt{T}}, -\eta_{1}\sigma\sqrt{T}\right)$$

The price of the UIC option is obtained by, thanks to Kou and Wang (see theorem 2 in [3])

$$UIC = S_0 \Psi \left( r + \frac{1}{2} \sigma^2 - \lambda \xi, \sigma, \lambda, p, \eta_1, \eta_2; \log \left( \frac{K}{S_0} \right), \log \left( \frac{H}{S_0} \right), T \right)$$
$$-K e^{-rT} \Psi \left( r - \frac{1}{2} \sigma^2 - \lambda \xi, \sigma, \lambda, p, \eta_1, \eta_2; \log \left( \frac{K}{S_0} \right), \log \left( \frac{H}{S_0} \right), T \right)$$

where

$$\widetilde{p} = \frac{p}{1+\xi} \frac{\eta_1}{\eta_1 - 1}, \quad \widetilde{\lambda} = \lambda(1+\xi), \quad \widetilde{\eta_1} = \eta_1 - 1, \quad \widetilde{\eta_2} = \eta_2 + 1$$

#### References

- [1] KOU, S. G. (2002). A Jump-Diffusion Model for Option Pricing. Management Science Vol. 48, No. 8, August 2002, pp. 1086-1101. 1, 2, 3
- [2] KOU, S. G. AND WANG, H. (2003). First Passage Times Of A Jump Diffusion Process. Adv. Appl. Prob. 35, 504-531 (2003). 2, 4, 5, 6
- [3] KOU, S. G. AND WANG, H. (2003). Option Pricing Under a Double Exponential Jump Diffusion Model. Management Science Vol 50, No. 9, September 2004, pp. 1178-1192.

3, 4, 7