

$$35. \quad \eta \sim U[0, a] \\ x \sim U[\eta, a]$$

$$f(x|\eta) = \frac{f(x, \eta)}{f(\eta)} = \frac{1}{a-\eta}$$

$$\text{则 } E(x|\eta) = \frac{a+\eta}{2}$$

$$\therefore E(x|\eta) \sim U[\frac{a}{2}, a]$$

$$41. \quad x_1 \sim E(\lambda_1), \quad x_2 \sim E(\lambda_2)$$

$$f(x_1 < x_2) = \int_0^{+\infty} \int_0^y f_{x_1, x_2}(x, y) dx dy = 1 - \frac{\lambda_2}{\lambda_1 + \lambda_2}$$

$$E(x_1 | x_1 < x_2) = \int_0^{+\infty} x f(x_1 = x | x_1 < x_2) dx \\ = \frac{1}{\lambda_1 + \lambda_2}$$

47. (1) 由课堂结论: Gamma 分布为 iid 指数分布相加结果

直接可得: $Z = X + \gamma$ 的概率密度函数为

$$f_Z(z) = \begin{cases} \lambda^2 z e^{-\lambda z} & , z > 0 \\ 0 & , z \leq 0 \end{cases}$$

(2) 由对称性:

$$E(X | X + \gamma = z) = E(\gamma | X + \gamma = z)$$

$$\text{且 } E(X + \gamma | X + \gamma = z) = z$$

$$\text{则 } E(X | X + \gamma = z) = \frac{z}{2}$$

$$49. \quad \because x, \gamma \stackrel{\text{iid}}{\sim} E(1)$$

$$U = X + \gamma, \quad V = \frac{X}{\gamma}$$

$$\therefore \begin{cases} X = \frac{UV}{V+1} \\ \gamma = \frac{U}{V+1} \end{cases} \quad \left| \frac{\partial(x, \gamma)}{\partial(u, v)} \right| = \begin{vmatrix} \frac{V}{V+1} & \frac{U}{(V+1)^2} \\ \frac{1}{V+1} & \frac{-U}{(V+1)^2} \end{vmatrix} = \frac{U}{(V+1)^2}$$

$$\text{则 } f_{u, v}(u, v) = f_{x, \gamma}\left(\frac{UV}{V+1}, \frac{U}{V+1}\right) \frac{U}{(V+1)^2} e^{-u}$$

$$\text{类似地, 有: } f(u) = u \cdot e^{-u}$$

$$f(v) = \frac{1}{(v+1)^2}$$

$$\therefore f(u, v) = f(u) \cdot f(v) \quad \therefore \text{独立}$$

$$50. \quad U = \min(X, Y) \quad V = \max(X, Y)$$

在 $0 < u \leq v < 1$ 时

$$\begin{aligned} f(u, v) &= f(X=u, Y=v, X < Y) + f(X=v, Y=u, X > Y) \\ &= \frac{1^2}{7} (u+v)^2 \end{aligned}$$

$$\therefore f(u, v) = \begin{cases} \frac{1^2}{7} (u+v)^2 & 0 < u \leq v < 1 \\ 0 & \text{其他} \end{cases}$$

$$51. \quad X_1, X_2 \stackrel{iid}{\sim} U[0, 1]$$

$$\begin{cases} Y_1 = (-2 \ln X_1)^{\frac{1}{2}} \cos 2\pi X_2 \\ Y_2 = (-2 \ln X_1)^{\frac{1}{2}} \sin 2\pi X_2 \end{cases} \Rightarrow \begin{cases} X_1 = e^{-\frac{Y_1^2 + Y_2^2}{2}} \\ X_2 = \frac{1}{2\pi} \arctan \frac{Y_2}{Y_1} \end{cases}$$

$$\text{则 } \left| \frac{\partial (X_1, X_2)}{\partial (Y_1, Y_2)} \right| = \begin{vmatrix} -Y_1 e^{-\frac{Y_1^2 + Y_2^2}{2}} & -Y_2 e^{-\frac{Y_1^2 + Y_2^2}{2}} \\ \frac{1}{2\pi} \frac{-Y_2}{1 + (\frac{Y_2}{Y_1})^2} & \frac{1}{2\pi} \frac{Y_1}{1 + (\frac{Y_2}{Y_1})^2} \end{vmatrix} = \frac{1}{2\pi} e^{-\frac{Y_1^2 + Y_2^2}{2}}$$

$$\therefore f(y_1, y_2) = \frac{1}{2\pi} e^{-\frac{y_1^2 + y_2^2}{2}}$$

$$Y_1, Y_2 \sim N(0, 0, 1, 1)$$

可知 Y_1, Y_2 相互独立且服从 $N(0, 1)$

$$53. \quad G_n(x) = P(Z_n \leq x) = P(\max\{x_1, \dots, x_n\} \geq 1 - \frac{x}{n})$$

$$= 1 - \prod_{i=1}^n P(X_i < 1 - \frac{x}{n})$$

$$= 1 - (1 - \frac{x}{n})^{2n}$$

$$\lim_{n \rightarrow \infty} G_n(x) = 1 - e^{-2x} (x > 0)$$

$$G(x) = \int_0^x 2e^{-2x} dx = 1 - e^{-2x} (x > 0)$$

$$\therefore \text{证得 } \lim_{n \rightarrow \infty} G_n(x) = G(x)$$

P179.

$$1. \quad X \sim U(-a, a)$$

$$f(x) = \frac{1}{2a}$$

$$\varphi_x(\theta) = E(e^{i\theta x}) = \int_{-a}^a \frac{e^{i\theta x}}{2a} dx$$

$$\begin{aligned}
 &= \frac{1}{2a i \theta} (e^{i\theta a} - e^{-i\theta a}) \\
 &= \frac{e^{i\theta a} - e^{-i\theta a}}{2a \theta i} = \frac{\sin(\theta - a)}{\theta a}
 \end{aligned}$$

2. $X \sim \text{Cauchy}(m, a)$

$$f(x) = \frac{a}{\pi[(x-m)^2 + a^2]} = \frac{1}{a\pi[(\frac{x-m}{a})^2 + 1]}$$

$$\text{令 } y = \frac{x-m}{a}$$

$$\text{则 } f(y) = \frac{1}{a\pi(y^2 + 1)} \cdot a = \frac{1}{\pi(y^2 + 1)}$$

$$\text{则 } I(\theta) = \int_{-\infty}^{+\infty} \frac{e^{i\theta y}}{\pi(y^2 + 1)} dy = \frac{1}{\pi} \int_{-\infty}^{+\infty} \frac{e^{i\theta y}}{y^2 + 1} dy$$

经求导发现: $I'(\theta) = I(\theta)$

$$\therefore \varphi(\theta) = e^{\lambda^2 - 1} = e^{i\theta m - a|\theta|}$$

6. 若随机变量 Y 的特征函数为 $\varphi(\theta)$

$$\text{且 } X \sim U(-a, a), \text{ 则 } \varphi_X(\theta) = \frac{\sin(a\theta)}{a\theta}$$

$$\therefore X+Y \text{ 的特征函数为 } \frac{\sin(a\theta)}{a\theta} \varphi(\theta)$$

$$\therefore \varphi_{X+Y}(\theta) = \varphi_a(\theta)$$

$$\therefore \varphi_a(\theta) \text{ 为 } X+Y \text{ 的特征函数}$$

$$7. (a) f(x) = \begin{cases} e^{-x} & x > 0 \\ 0 & x < 0 \end{cases}$$

$$\therefore \varphi_X(\theta) = \int_{-\infty}^{+\infty} e^{i\theta x} f(x) dx = \frac{1}{1-i\theta}$$

(b) $\therefore X$ 与 Y 独立

$$\therefore \varphi_{X+Y}(\theta) = \varphi_X(\theta) \varphi_Y(\theta)$$

$$= \frac{1}{1+\theta^2}$$

(c) $Z = X - Y$

$$\text{则 } f(z) = \begin{cases} \frac{1}{2} e^z & , z < 0 \\ \frac{1}{2} e^{-z} & , z \geq 0 \end{cases}$$

$$\begin{aligned}
 \therefore \varphi_z(\theta) &= \int_{-\infty}^{+\infty} e^{i\theta z} f(z) dz \\
 &= \frac{1}{2} \frac{e^{(i\theta+1)z}}{i\theta+1} \Big|_{-\infty}^0 + \frac{1}{2} \frac{e^{(i\theta-1)z}}{i\theta-1} \Big|_0^{+\infty} \\
 &= \frac{1}{2} \frac{i\theta-1-i\theta-1}{(i\theta)^2-1} = \frac{1}{\theta^2+1}
 \end{aligned}$$

与前相同

$$\begin{aligned}
 \text{而 } \frac{1}{1+\theta^2} &= \int_{-\infty}^{+\infty} f(x) \cos \theta x dx \\
 &= \frac{1}{\pi} \int_{-\infty}^{+\infty} \frac{1}{1+\theta^2} \cos \theta x d\theta
 \end{aligned}$$

$$\therefore f(x) = \begin{cases} e^{-x} & x > 0 \\ 0 & x < 0 \end{cases}$$

10. 证明:

$$f_{x_i}(x) = \frac{a}{\pi[(x-m)^2+a^2]}$$

$$\varphi_{x_i}(\theta) = e^{i\theta m - a|\theta|}$$

$$\begin{aligned}
 \varphi_x(\theta) &= \prod_{i=1}^n \varphi_{x_i}\left(\frac{\theta}{n}\right) = \prod_{i=1}^n e^{i\frac{\theta}{n} \cdot m - a|\frac{\theta}{n}|} \\
 &= e^{i\theta m - a|\theta|} = \varphi_{x_i}(\theta)
 \end{aligned}$$