

Convergence in real function. $f : \mathbb{R} \rightarrow [0, 1]$

① pointwise

$$f_n(x) \rightarrow f(x) \quad \forall x \in [0, 1]$$

② convergence in norm. $\|\cdot\|_p \quad V \rightarrow [0, \infty)$

Space of all fun. on $[0, 1]$

$$f_n \rightarrow f \text{ in } \|\cdot\|_p$$

$$\text{eq. } \|\cdot\|_p = \|\cdot\|_p$$

$$\text{if } \|f_n - f\|_p \rightarrow 0$$

$$\left(\int |f_n(x) - f(x)|^p dx \right)^{\frac{1}{p}} \rightarrow 0$$

$$\|\cdot\|_\infty$$

$$\sup_{x \in [0, 1]} |f_n(x) - f(x)| \rightarrow 0$$

$$\|\cdot\|_1$$

Lebesgue



$$\int |f_n(x) - f(x)| dx \rightarrow 0$$

③ convergence in measure.

$$\forall \varepsilon > 0 \quad \mu(\{x \in [0, 1] : |f_n(x) - f(x)| > \varepsilon\}) \rightarrow 0$$

$x_n, X : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R} \text{ Borel}, \cdot)$

④ a.s. convergence. $x_n \xrightarrow{\text{a.s.}} x$

relaxation of

$P(\{\omega \in \Omega : x_n(\omega) \rightarrow x(\omega)\}) = 1$ Pointwise convergence.



$P(\{\omega \in \Omega : x_n(\omega) \rightarrow x(\omega)\}) = 1$



If s.t. has prob = 0

we can omit it in

prob. theory

② Convergence in nth. mean. \Leftrightarrow Lr norm

$$x_n \xrightarrow{r} x$$

$$\int |X_n(\omega) - X(\omega)|^r d \, P(\omega) \rightarrow 0 \quad r > 1$$

\Updownarrow

$$\mathbb{E} |X_n - x|^r$$

③ Convergence in prob. $X_n \xrightarrow{P} X$

$$\lim_{n \rightarrow \infty} P(\{\omega : |X_n(\omega) - X(\omega)| > \epsilon\}) = 0$$

④. Convergence in distribution. $x_n \xrightarrow{D} x$ ($x_n \xrightarrow{d} x$)
 (weak convergence) $F_{x_n} \xrightarrow{} F_x$ $x_n \Rightarrow x$

real c.d.f. pointwise for those continuity points of $F_X(x)$

$F_{Xn}(x) \rightarrow F_X(x)$ for continuity point x of $F_X(x)$

四

$$x_n \rightharpoonup Fx$$

$$\text{eq. } x_n = \frac{1}{n} \quad F_{x_n}(x) = \begin{cases} 0 & \text{if } x < \frac{1}{n} \\ 1 & \text{if } x \geq \frac{1}{n} \end{cases}$$

$$Y_n = -\frac{1}{n} \quad F_{Y_n}(x) = \begin{cases} 0 & \text{if } x < -\frac{1}{n} \\ 1 & \text{if } x \geq -\frac{1}{n} \end{cases}$$

$$f_{xn}(x) \rightarrow \begin{cases} 0 & \text{if } x \leq 0 \\ 1 & \text{if } x > 0 \end{cases}$$

Not a d.f.

$$F_{\text{in}}(x) \rightarrow \begin{cases} 0 & x < 0 \\ 1 & x \geq 0 \end{cases} \rightarrow Y = 0$$

a d.f.

1) The first three convergence mode have to be defined on the same prob space

e.g. X_1 is die rolling X_2 is coin flipping X_3 X while convergence in distribution is ok for the above.

Thm.

$$\begin{aligned} \textcircled{1} \quad (X_n \xrightarrow{\text{a.s.}} X) &\Rightarrow (X_n \xrightarrow{P} X) \Rightarrow (X_n \xrightarrow{D} X) \\ (X_n \xrightarrow{r} X) &\Rightarrow \end{aligned}$$

$$\textcircled{2} \quad (X_n \xrightarrow{r} X) \Rightarrow (X_n \xrightarrow{s} X) \quad \text{if } r > s \geq 1$$

\textcircled{3} Converse statement are not true in general.

Proof of \textcircled{3}

$$\begin{aligned} \textcircled{1}^o \quad (X_n \xrightarrow{D} X) &\Rightarrow (X_n \xrightarrow{P} X) \quad P(|Y_n - Y| > \epsilon) \rightarrow 0 \\ F_{X_n}(x) \rightarrow F_X(x). \quad & \uparrow \\ & P(\{\omega \in \Omega : |X_{n(\omega)} - X(\omega)| > \epsilon\}) \rightarrow 0 \end{aligned}$$

$$\begin{aligned} \Omega &= \{H, T\} \\ X(\omega) \equiv X_{2n}(\omega) &= \begin{cases} 1 & \text{if } \omega = H \\ 0 & \text{if } \omega = T \end{cases} \quad \begin{array}{l} \text{coin is fair} \\ \text{Bernoulli r.v. with } p = \frac{1}{2}. \end{array} \\ X_{2n+1}(\omega) &= \begin{cases} 1 & \text{if } \omega = T \\ 0 & \text{if } \omega = H \end{cases} \end{aligned}$$

$$\forall \omega \quad P(|X_{2n+1}(\omega) - X_n(\omega)| = 1) = 1$$

2° $(x_n \xrightarrow{IP} x) \nRightarrow (x_n \xrightarrow{\text{wrt P}} x)$
 not wrt size of $|x_n - x|$

$$x_n = \begin{cases} n & \text{with prob. } \frac{1}{n} \\ 0 & \dots \quad 1 - \frac{1}{n} \end{cases} \quad x = 0$$

$$\mathbb{P}(|x_n - x| > \varepsilon) = \frac{1}{n} \rightarrow 0$$

$$\mathbb{E}|x_n - x|^s = n^s \cdot \frac{1}{n} + 0 \cdot (1 - \frac{1}{n}) = 1 \rightarrow 0$$

if. scr $(x_n \xrightarrow{s} x) \nRightarrow (x_n \xrightarrow{\delta} x)$

$$\mathbb{E}|x_n - x|^s = \int |y_n|^s d\mu_{y_n} = \int |y|^s f_{y_n}(y) dy \quad ①$$

$$\mathbb{E}|x_n - x|^r = \int |y|^r f_{y_n}(y) dy \quad ②$$

$f_{y_n}(y)$ is small enough for ① but not for ②

Choose

$$x_n = \begin{cases} n & \text{with prob. } n^{-\frac{n+s}{2}} \\ 0 & \text{with prob. } 1 - n^{-\frac{n+s}{2}} \end{cases} \quad x = 0$$

$$\mathbb{E}|x_n - x|^s = n^s n^{-\frac{n+s}{2}} + 0 \cdot (1 - n^{-\frac{n+s}{2}})$$

$$= n^{\frac{s-n}{2}} \rightarrow 0 \quad n \rightarrow \infty$$

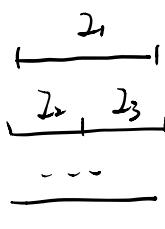
$$\mathbb{E}|x_n - x|^r = n^r n^{-\frac{n+s}{2}} + 0 \cdot (1 - n^{-\frac{n+s}{2}})$$

$$= n^{\frac{r-s}{2}} \rightarrow \infty \quad n \rightarrow \infty.$$

$(x_n \xrightarrow{IP} x) \nRightarrow (x_n \xrightarrow{a.s} x)$

$$\Omega = [0, 1]$$

$\omega \sim \text{Uniform}[0, 1]$



$$x_n(\omega) = \begin{cases} 1, & \omega \in I_n \\ 0, & \text{otherwise} \end{cases}$$

$$X \equiv 0 \quad \forall \omega \in [0, 1].$$

$$x_n \xrightarrow{\text{P}} X$$

$$\forall \varepsilon > 0 \quad P(|x_n - X| > \varepsilon) \rightarrow 0$$

$$x_n \xrightarrow{\text{a.s.}} X$$

$$P(x_n \rightarrow X) = 1$$

$I_{m+1} = I_m + m$. m -th. step.

Claim. $x_n \xrightarrow{\text{P}} X$ $x_n \not\xrightarrow{\text{a.s.}} X$

$$P\{ \{\omega : |x_n(\omega) - X(\omega)| > \varepsilon\} \} = P(I_n) \rightarrow 0 \quad n \rightarrow \infty.$$

Every step there is a I_{xx} st. $X_{xx}(\omega) = 0$
 \Rightarrow infinite many !

So. For any ω , $x_n \rightarrow X$ does not hold.

A The def of $\xrightarrow{\text{P}}$ only care about the size of bad event
but it doesn't care about the relation of bad event

A If ω fall into bad event occurs for only finite many times. Then. $x_n(\omega) \rightarrow X(\omega)$.

A In this eq. the bad event of x_n together covers.
The $[0, 1]$ again & again.

Pf of $(x_n \xrightarrow{P} x) \Rightarrow (x_n \xrightarrow{\text{P}} x)$

Aim. $\lim P(x_n < x) \rightarrow P(x < x)$ for all continuity point of $P(x \in x)$
 condition

$$P(|x_n - x| > \varepsilon) \rightarrow 0 \quad \forall \varepsilon > 0$$

$$\begin{aligned} P(x_n < x) &= P(x_n < x, |x_n - x| > \varepsilon) \\ &\quad + P(x_n < x, |x_n - x| \leq \varepsilon) \\ &\leq P(|x_n - x| > \varepsilon) + P(x_n < x, x - \varepsilon \leq x_n \leq x + \varepsilon) \\ &\leq P(|x_n - x| > \varepsilon) + P(x_n < x, x \in x_n + \varepsilon) \\ &= P(|x_n - x| > \varepsilon) + P(x \in x + \varepsilon) \quad \textcircled{1} \end{aligned}$$

Switch x_n and x

$$\underbrace{P(x < y)}_{x - \varepsilon} \leq P(|x_n - x| > \varepsilon) + P(x_n \leq \underbrace{y + \varepsilon}_{y + \varepsilon = x})$$

$$P(x < x - \varepsilon) = P(|x_n - x| > \varepsilon) \leq P(x_n < x) \quad \textcircled{2}$$

Discuss \liminf \limsup

$$P(x \leq x - \varepsilon) = 0 \leq \liminf_{n \rightarrow \infty} P(x_n < x) \leq \limsup_{n \rightarrow \infty} P(x_n \leq x) \leq P(x \leq x + \varepsilon) + 0$$

send $\varepsilon \downarrow 0$

Lemma. $(x_n \xrightarrow{r} x) \Rightarrow (x_n \xrightarrow{P} x)$
 $(x_n \xrightarrow{s} x) \Rightarrow (x_n \xrightarrow{r+s} x) \quad r \geq s \geq 1$

Markov inequality

$$\begin{aligned} & \Pr(|X_n - X| > \epsilon) \\ & \leq \frac{\mathbb{E}|X_n - X|}{\epsilon} \rightarrow 0 \end{aligned}$$

For any r.v. Z any $a > 0$

$$\Pr(|Z| > a) \leq \frac{\mathbb{E}|Z|^r}{a^r}$$

Hölder's inequality

$$(\mathbb{E}|Z|^s)^{\frac{1}{s}} \leq (\mathbb{E}|Z|^r)^{\frac{1}{r}} \quad r \geq s > 0$$

\uparrow

Jensen.

φ is convex Then. $\varphi(\mathbb{E}(x)) \leq \mathbb{E}(\varphi(x))$

$$x = |Z|^s \quad \varphi = x^{\frac{r}{s}}$$

$$\text{Then. } \varphi(x) = |Z|^r$$

$$\therefore (\mathbb{E}|Z|^s)^{\frac{r}{s}} \leq \mathbb{E}|Z|^r$$

$$\therefore (\mathbb{E}|Z|^s)^{\frac{1}{s}} \leq (\mathbb{E}|Z|^r)^{\frac{1}{r}}$$

\uparrow Hölder's

$$\mathbb{E}|XY| \leq (\mathbb{E}|X|^{\frac{r}{s}})^{\frac{s}{r}} (\mathbb{E}|Y|^{\frac{s}{r}})^{\frac{r}{s}}$$

$$\text{choose. } x = \frac{1}{s} \cdot Y = 1 \quad p = \frac{r}{s} \quad q = 1 - \frac{1}{r} = \frac{1}{\frac{r-s}{r}} = \frac{r}{r-s}$$

$$\therefore (\mathbb{E}|Z|^s)^{\frac{1}{s}} \leq (\mathbb{E}|X|^{\frac{r}{s}})^{\frac{r}{s}} \quad \text{上课讲的不对.}$$

$$\Pr(\{\omega : |X_n(\omega) - X(\omega)| > \epsilon\}) \rightarrow 0 \quad \forall \epsilon > 0$$

$\triangleq A_n(\epsilon)$ function of both. n and ϵ .
event of ω .

Lemma.:

Let $B_m(\epsilon) = \bigcup_{n=m}^{\infty} A_n(\epsilon)$ Take the union of the tail part of the event.

$$\textcircled{1}. \quad x_n \xrightarrow{\text{a.s.}} x \quad \text{iff} \quad \boxed{\lim_{m \rightarrow \infty} P(B_m(\varepsilon)) = 0 \quad \forall \varepsilon > 0} \quad \textcircled{1}$$

$$\textcircled{2}. \quad x_n \xrightarrow{\text{a.s.}} x \quad \text{if} \quad \boxed{\sum_{n=1}^{\infty} P(A_n(\varepsilon)) < \infty \quad \forall \varepsilon > 0} \quad \textcircled{2}$$

$$\textcircled{3}. \quad (x_n \xrightarrow{\text{a.s.}} x) \Rightarrow (x_n \xrightarrow{P} x)$$

$$B_m(\varepsilon) = \bigcup_{n=m}^{\infty} A_n(\varepsilon)$$

subadditivity

$$P(B_m(\varepsilon)) = P\left(\bigcup_{n=m}^{\infty} A_n(\varepsilon)\right) \leq \sum_{n=m}^{\infty} P(A_n(\varepsilon))$$

$$\text{If } \textcircled{2} \Rightarrow \textcircled{1} \Leftrightarrow x_n \xrightarrow{\text{a.s.}} x$$

So, we only need to prove \textcircled{1}

\textcircled{4} If $A_n(\varepsilon)$ overlaps a lot \textcircled{2} may fail
though $x_n \xrightarrow{\text{a.s.}} x$

\textcircled{5} Because $A_m(\varepsilon) \subset B_m(\varepsilon)$.

$$\text{If} \quad \lim_{m \rightarrow \infty} P(B_m(\varepsilon)) = 0 \quad \forall \varepsilon > 0$$

$$\text{of course.} \quad \lim_{m \rightarrow \infty} P(A_m(\varepsilon)) = 0 \quad \forall \varepsilon > 0 \quad \Leftrightarrow x_n \xrightarrow{P} x$$

Pf of \textcircled{1}

$$x_n \xrightarrow{\text{a.s.}} x$$

$$\Leftrightarrow P(\underbrace{\{\omega: x_n(\omega) \rightarrow x(\omega)\}}_{C}) = 1$$

$$C^c = \{\omega: x_n(\omega) \not\rightarrow x(\omega)\}$$

$$\Leftrightarrow \mathbb{P}(C^c) = 0$$

If. $w \in C$

$$x_n(w) \rightarrow x(w) \quad n \rightarrow \infty.$$

$\forall \varepsilon > 0$. $|x_n(w) - x(w)| > \varepsilon$ for only finite many n .

If $w \in C$.

$$\boxed{\exists \varepsilon > 0} \text{ st. } |x_n(w) - x(w)| > \varepsilon \text{ for infinite many } n.$$

$$\therefore C^c = \bigcup_{\varepsilon > 0} \{w : |x_n(w) - x(w)| > \varepsilon \text{ for infinite many } n\}$$

$$\beta_m(\varepsilon) \downarrow$$

$$\therefore \text{①} = \lim_{m \rightarrow \infty} \mathbb{P}(\beta_m(\varepsilon)) = \mathbb{P}\left(\bigcap_{m=1}^{\infty} \bigcup_{n=m}^{\infty} A_n(\varepsilon)\right)$$

$A_n(\varepsilon) \text{ i.o.} = \limsup_n A_n(\varepsilon).$

$$\{w \in \Omega : |x_n(w) - x(w)| > \varepsilon\} \text{ for infinite many } n.$$

$$\text{LHS} = \mathbb{P}(C^c) = 0$$

$$\text{RHS} = \mathbb{P}(\{w : |x_n(w) - x(w)| > \varepsilon \text{ for i.o. } n\}) = 0 \quad \forall \varepsilon.$$

LHS is stronger than RHS

uncountable - decreasing in. l.

$$\text{LHS} = \mathbb{P}\left(\bigcup_{\varepsilon > 0} \{A_n(\varepsilon) \text{ i.o.}\}\right) = 0. \quad \text{Just choose a sequence.}$$

$$\text{RHS} = \mathbb{P}(\underbrace{A_n(\varepsilon) \text{ i.o.}}_{\text{unions } A_n(\varepsilon)}) = 0. \quad \forall \varepsilon > 0. \quad \text{to do the union.}$$

$$A_n(\varepsilon) \downarrow \text{ if } \varepsilon \uparrow$$

$$\begin{aligned} & \bigcup_{\varepsilon > 0} \{A_n(\varepsilon) \text{ i.o.}\} \\ &= \bigcup_{k=1}^{\infty} \{A_n(\frac{1}{k}) \text{ i.o.}\} \end{aligned}$$

$$P\left(\bigcup_{n=0}^{\infty} \{A_n \text{ i.o.}\}\right)$$

$$= P\left(\bigcup_{k=1}^{\infty} \{A_k \text{ i.o.}\}\right) \leq \sum_{k=1}^{\infty} P(A_k \text{ i.o.})$$

$\underbrace{I_1}_{I_2}$

$$I_n = [0, \frac{1}{n}] \quad X_n(\omega) = \begin{cases} 1 & \omega \in I_n \\ 0 & \text{otherwise.} \end{cases} \quad X = 0$$

$$I_n \subset I_{n-1} \subset \dots \subset I_1$$

$$B_m(v) = \bigcup_{n=m}^{\infty} A_n(v) = \bigcup_{n=m}^{\infty} I_n = I_m.$$

\vdash

$$\lim_{m \rightarrow \infty} P(B_m(v)) = \lim_{m \rightarrow \infty} P(I_m) = \lim_{m \rightarrow \infty} \frac{1}{m} = 0. \quad \checkmark$$

We cannot use sufficient condition (2)

Because. $A_n(v)$ overlaps a lot.

From weak to strong

① If. $x_n \xrightarrow{P} c$ then. $x_n \xrightarrow{P} c$

In LN the limit is a number / no stochastic

② If. $x_n \xrightarrow{P} x$, $P(|x_n| \leq k) = 1$ for some k uniformly in n .
then. $x_n \xrightarrow{r} x \forall r \geq 1$

(Bounded Convergence Thm).

③ If $\sum_{n=1}^{\infty} P(|x_n - x| > \varepsilon) < \infty$. $\forall \varepsilon > 0$ then. $x_n \xrightarrow{a.s.} x$

$$\text{Pf: } P(|x_n - x| > \varepsilon) = P(x_n < x - \varepsilon) + P(x_n > x + \varepsilon)$$

$$= P(x_n < x - \varepsilon) + 1 - P(x_n \leq x + \varepsilon).$$

$$\xrightarrow{P} \underline{P(x < x - \varepsilon)} + 1 - \underline{P(x \leq x + \varepsilon)}.$$

$$= 0 + 1 - 1 = 0$$

$$\downarrow P(c \leq x + \varepsilon)$$

$$P(X_n \leq c - 2\sigma) \leq$$



$$\text{if } (c \in c - 2\sigma)$$

$$\leq P(X_n \leq c - \frac{\epsilon}{2})$$